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CSE - S12 (HW 4)

i)

$$(a) P(x = red) = \frac{1}{2} \quad P(x = yellow) = \frac{1}{5}$$

$$P(x = blue) = \frac{1}{4} \quad P(x = black) = \frac{1}{20}$$

$$H(x) = - \sum_{x=x} P_x(x=x) \log_2 (P_x(x=x))$$

$$= - \left(\frac{1}{2} \log_2 \left(\frac{1}{2} \right) + \frac{1}{4} \log_2 \left(\frac{1}{4} \right) + \frac{1}{5} \log_2 \left(\frac{1}{5} \right) \right.$$

$$\left. + \frac{1}{20} \log_2 \left(\frac{1}{20} \right) \right)$$

$$= 1.68 \text{ bits}$$

b) $X = \text{color of a sock randomly picked}$

$Y = \text{which drawer (top or down)}$

$$P(\text{top drawer}) = 2 P(\text{bottom drawer}) \quad \textcircled{1}$$

$$P(\text{top drawer}) + P(\text{bottom drawer}) = 1 \quad \textcircled{2}$$

Solving the above 2 equations, we get

$$P(\text{top drawer}) = \frac{2}{3}$$

$$P(\text{bottom drawer}) = \frac{1}{3}$$

$$H(X|Y) = - \sum_{x=x, y=y} P(x=x, y=y) \log_2 (P(x=x | y=y))$$

calculating relevant Probabilities

$$P(x=x, y=y) = P(x|y) * P(y)$$

$$P(x=\text{red} | y=\text{top}) = 1$$

$$P(x=\text{red}, y=\text{top}) = 1 \times \frac{2}{3} = \frac{2}{3}$$

$$P(x=\text{red} | y=\text{bottom}) = 0$$

$$P(x=\text{blue} | y=\text{top}) = 0$$

$$P(x=\text{blue} | y=\text{below}) = \frac{1}{2}$$

$$P(x=\text{blue}, y=\text{below}) = \frac{1}{2} \times \frac{1}{3} = \frac{1}{6}$$

$$P_X(X = \text{yellow} | Y = \text{top}) = 0$$

$$P_X(X = \text{yellow} | Y = \text{below}) = \frac{2}{15}$$

$$P_X(X = \text{yellow}, Y = \text{below}) = \frac{2}{5} \times \frac{1}{3} = \frac{2}{15}$$

$$P_X(X = \text{black} | Y = \text{top}) = 0$$

$$P_X(X = \text{black} | Y = \text{below}) = \frac{1}{10}$$

$$P_X(X = \text{black}, Y = \text{below}) = \frac{1}{10} \times \frac{1}{3} = \frac{1}{30}$$

We will assume, $0 \log_2(0) \approx 0$

$$\begin{aligned} H(X|Y) &= -\left[\frac{2}{3} \log_2(1) + \frac{1}{5} \log_2\left(\frac{1}{2}\right) + \frac{2}{15} \log_2\left(\frac{2}{3}\right) \right. \\ &\quad \left. + \frac{1}{30} \log_2\left(\frac{1}{10}\right) \right] \\ &= 0.45 \text{ bits} \end{aligned}$$

c) Information gain:-

$$\begin{aligned} I(X:Y) &= H(X) - H(X|Y) \\ &= 1.68 - 0.45 \\ &= 1.23 \end{aligned}$$

2)

i) $P(\text{word} = \text{the}) = 6/14$

$$P(\text{word} = \text{rabbit}) = 3/14$$

$$P(\text{word} = \alpha) = 5/14$$

i) $P_A(\text{current word} = \text{rabbit} \mid \text{Previous word} = \text{the})$

$$\Rightarrow \frac{2}{6} = \frac{1}{3}$$

ii) $P_A(\text{current word} = \alpha \mid \text{Previous word} = \text{rabbit})$

$$= \frac{1}{3}$$

iii) $P_A(\text{current word} = \text{the} \mid \text{Previous word} = \text{rabbit})$

$$= \frac{2}{3}$$

iv) $P_A(\text{current word} = \text{the or } \alpha \mid \text{Previous word} = \text{rabbit})$

$$= 1$$

v) $P(A \cap B) = P(A) * P(B) \rightarrow \text{condition for Naive Bayes}$

$$P(\text{current word} = \text{the} \cap \text{Previous word} = \text{rabbit})$$

$$= \frac{2}{14}$$

$$P(\text{word} = \text{the}) = \frac{6}{141}$$

$$P(\text{word} = \text{rabbit}) = \frac{3}{141}$$

$$\Rightarrow \frac{2}{141} \neq \frac{6}{141} \times \frac{3}{141}$$

\therefore Naive Bayes assumption is not valid here

challenge

$$Q3) E[x] = \sum_{i=1}^n x_i p_i$$

$$E[f(x)] = \sum_{i=1}^n p_i f(x_i)$$

$$f(E[x]) = f(x_1 p_1 + x_2 p_2 + \dots + x_n p_n)$$

$$f(E[x]) = f(p_1 x_1 + (1-p_1) \left[\frac{p_2}{1-p_1} x_2 + \frac{p_3}{1-p_1} x_3 + \dots + \frac{p_n}{1-p_1} x_n \right])$$

$$\leq p_1 f(x_1) + (1-p_1) \left[\frac{p_2}{1-p_1} x_2 + \frac{p_3}{1-p_1} x_3 + \dots + \frac{p_n}{1-p_1} x_n \right]$$

(Assuming $\theta = p_1$)

$$f(E[x]) \leq p_1 f(x_1) + (1-p_1) f \left[\frac{p_2}{1-p_1} x_2 + \frac{p_3}{1-p_1} x_3 + \dots + \frac{p_n}{1-p_1} x_n \right] \quad - \textcircled{1}$$

Let us consider the term,

$$= f \left[\frac{p_2}{1-p_1} x_2 + \frac{p_3}{1-p_1} x_3 + \dots + \frac{p_n}{1-p_1} x_n \right]$$

$$= f \left[\frac{p_2 x_2}{1-p_1} + \frac{1-p_1-p_2}{1-p_1} \left[\frac{p_3 x_3}{1-p_1-p_2} + \frac{p_4 x_4}{1-p_1-p_2} + \dots + \frac{p_n x_n}{1-p_1-p_2} \right] \right]$$

$$\begin{aligned}
 & \leq \frac{P_2}{1-P_1} f(x_2) + \frac{1-P_1-P_2}{1-P_1} f \left[\frac{P_3 x_3}{1-P_1-P_2} + \frac{P_n x_n}{1-P_1-P_2} \dots \right] \\
 & \quad \uparrow \quad \uparrow \\
 & \theta = \frac{P_2}{1-P_1} \quad (-\theta) = 1 - \frac{P_2}{1-P_1} = \frac{1-P_1-P_2}{1-P_1}
 \end{aligned}$$

$$\Rightarrow (1-P_1) f \left(\frac{P_2 x_2}{1-P_1} + \frac{P_3 x_3}{1-P_1} \dots \frac{P_n x_n}{1-P_1} \right)$$

$$\begin{aligned}
 & \leq P_2 f(x_2) + (1-P_1-P_2) f \left(\frac{P_3 x_3}{1-P_1-P_2} + \dots + \frac{P_n x_n}{1-P_1-P_2} \right) \\
 & \quad - \textcircled{2}
 \end{aligned}$$

From ① & ②, we get,

$$\begin{aligned}
 f(E[x]) & \leq P_1 f(x_1) + P_2 f(x_2) \\
 & \quad + (1-P_1-P_2) f \left(\frac{P_3 x_3}{1-P_1-P_2} + \dots + \frac{P_n x_n}{1-P_1-P_2} \right)
 \end{aligned}$$

If we generalize the above term,

we get

$$\begin{aligned}
 f(E[x]) & \leq P_1 f(x_1) + P_2 f(x_2) + P_3 f(x_3) \dots \\
 & \quad + P_{n-1} f(x_{n-1}) + (1-P_1-P_2-\dots-P_{n-1}) f \left(\frac{P_n x_n}{1-P_1-P_2-\dots-P_{n-1}} \right)
 \end{aligned}$$

$$\Rightarrow p_1 + p_2 + p_3 + \dots + p_n = 1$$

$$\Rightarrow 1 - p_1 - p_2 - p_3 - \dots - p_{n-1} = p_n$$

$$f(E[x]) \leq p_1 f(x_1) + p_2 f(x_2) + \dots + p_{n-1} f(x_{n-1}) + p_n f(x_n)$$

$$\therefore f(E[x]) \leq E[f(x)]$$

Q4) $p_\lambda(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{else} \end{cases}$

(a) Likelihood (L),

$$L(\lambda) = \prod_{i=1}^n \lambda e^{-\lambda x_i}$$

$$= \lambda^n e^{-\lambda \sum_{i=1}^n x_i}$$

$$\log(L(\lambda)) = n \log \lambda - \lambda \sum_{i=1}^n x_i$$

For maximum, set derivative to 0.

$$\frac{\partial L(\lambda)}{\partial \lambda} = \frac{n}{\lambda} - \sum x_i = 0$$

$$\hat{\lambda}_{MLE} = \frac{n}{\sum x_i}$$

$$\hat{\lambda}_{MLE} = \frac{1}{\bar{\theta}} \quad (\bar{\theta} = \frac{\sum x_i}{n})$$

b) $E[\hat{\lambda}] = E\left[\frac{n}{\sum x_i}\right]$

$$= n \times E\left[\frac{1}{x_i}\right]$$

$$E[x_i] = \int_0^\infty \frac{1}{\lambda} \lambda e^{-\lambda x} dx = \frac{1}{n-1}$$

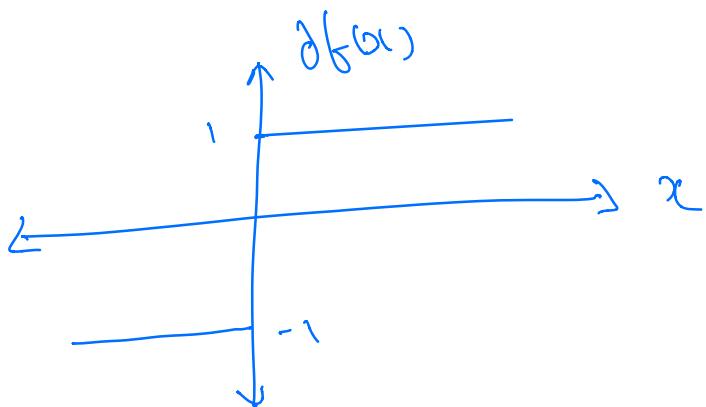
$$E(\hat{\lambda}) = \frac{n}{n-1} \lambda \neq \lambda$$

It is a
biased estimator

Q2)

(a) If f is convex and differentiable,
then gradient at x is a sub-gradient.

But a sub-gradient can exist
even when f is non-differentiable.



A function f is called sub-differentiable
at x if there exists at least one
sub-gradient at x .

Consider, $f(x) = |x|$

For, $x < 0 \rightarrow$ sub-gradient $\partial f(x) = -1$

For, $x > 0 \rightarrow$ sub-gradient $\partial f(x) = 1$

$$\text{At } x = 0$$

one sub-gradient is defined by the equality, $|z| \geq g z$ for z which is satisfied

$$\text{iff } g \in [-1, 1]$$

$$\therefore \partial f(0) = [-1, 1]$$

$$\therefore \partial f(x) = \begin{cases} g_1 g & \text{if } x > 0 \\ g_2 g & \text{if } x \leq 0 \\ [-1, 1] & x = 0 \end{cases}$$

b) A point x^* is a minimizer of a convex function iff f is sub-differentiable at x^* and

$$0 \in \partial f(x^*)$$

i.e. $g=0$ is a sub-gradient of f at x^* .

$$[\text{as } f(x) \geq f(x^*)]$$

$$\Rightarrow 0 \in \partial f(x^*) \text{ reduces to } 0 \in f'(x^*) = 0$$

if f is differentiable at x^* .