

Q1) MLE (MAXIMUM LIKELIHOOD ESTIMATOR) OF PARETO**(Total 5 points)**

Let X_1, X_2, \dots, X_n be i.i.d. as Pareto(θ), with the pdf of Pareto(θ) being $f(x) = \frac{\theta \cdot c^\theta}{x^{\theta+1}}$, where $c > 0$ is some given constant and assume that $\theta \geq 1$ and $x \geq c$. Find the MLE of θ .

$$L(\theta) = \prod_{i=1}^n f(X_i | \theta) = \prod_{i=1}^n \frac{\theta \cdot c^\theta}{X_i^{\theta+1}}$$

$$\begin{aligned} \log \text{ likelihood} \quad \ell(\theta) &= \log(L(\theta)) = \sum \log\left(\frac{\theta \cdot c^\theta}{X_i^{\theta+1}}\right) \end{aligned}$$

$$= \sum (\log \theta + \theta \log c - (\theta+1) \log(X_i))$$

$$= n \log \theta + n \theta \log c - (\theta+1) \sum \log(X_i)$$

$$\text{For MLE, } \frac{d\ell(\theta)}{d\theta} = 0$$

$$\Rightarrow \frac{d}{d\theta} (n \log \theta + n \theta \log c - (\theta+1) \sum \log(X_i)) = 0$$

$$\Rightarrow \frac{n}{\theta} + n \log c - \sum \log(X_i) = 0$$

$$\Rightarrow \frac{n}{\theta} = \sum \log(X_i) - n \log c$$

$$\Rightarrow \hat{\theta}_{MLE} = \frac{n}{(\sum \log(X_i)) - n \log c}$$

Q2) MME (METHOD OF MOMENTS ESTIMATOR) WITH DATA**(Total 3 points)**

Let $X = \begin{cases} 1 & \text{with prob } \theta \\ 3 & \text{otherwise} \end{cases}$, with θ being an unknown parameter. Let $D = \{1, 1, 1, 3\}$ be drawn i.i.d. from X . Derive $\hat{\theta}$, the MME of θ , using D as the sample data. Clearly show all your steps.

$K=1$ unknown parameter

$$\hat{\alpha}_1(\theta) = \frac{\sum X_i}{n} = \bar{X}$$

$$\alpha_1(\theta) = E[X] = 1 \cdot \theta + 3 \cdot (1 - \theta) = 3 - 2\theta$$

For MME, $\hat{\alpha}_1 = \alpha_1(\hat{\theta})$

$$\Rightarrow \bar{X} = 3 - 2\hat{\theta}$$

$$\Rightarrow \boxed{\hat{\theta} = \frac{3 - \bar{X}}{2}}$$

For $D = \{1, 1, 1, 3\}$

$$\bar{X} = \frac{1+1+1+3}{4} = \frac{3}{2}$$

$$\therefore \hat{\theta} = \frac{3 - \frac{3}{2}}{2} = \underline{\underline{\frac{3}{4}}}$$

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Q3) TOY EXAMPLE FOR PERMUTATION TEST**(Total 5 points)**

Let $X = \{1, 7\}$ and $Y = \{3\}$. The null hypothesis is that X & Y are from the same distribution. The threshold for Accept/Reject is $p=0.4$.

- (a) Use the permutation test to compute the p-value and decide the hypothesis. Please show all steps for each permutation clearly. (4 points)
- (b) Explain, in words, why the Permutation test works. Consider the case where X and Y are test scores of two classes with very different intellect levels. (1 point)

$$(a) T_{obs} = |\bar{X} - \bar{Y}| = \left| \frac{1+7}{2} - 3 \right| = |4 - 3| = 1 \quad \left. \vphantom{\frac{1+7}{2}} \right\} \frac{1}{2}$$

<u>Permutations</u>		<u>$T_i (\bar{X} - \bar{Y})$</u>	<u>$I(T_i > T_{obs})$</u>	
<u>X</u>	<u>Y</u>			
$\{1, 7\}$	$\{3\}$	1	0	} 3
$\{1, 3\}$	$\{7\}$	3 5	1	
$\{3, 1\}$	$\{7\}$	5	1	
$\{3, 7\}$	$\{1\}$	4	1	
$\{7, 1\}$	$\{3\}$	1	0	
$\{7, 3\}$	$\{1\}$	4	1	

$$p\text{-value} = \frac{\sum I(T_i > T_{obs})}{n} = \frac{4}{6} = \frac{2}{3} \quad \left. \vphantom{\frac{2}{3}} \right\} \frac{1}{2}$$

Since $p\text{-value} > 0.4$, we can't Reject, so accept null, ($X \equiv Y$)

(b) If X & Y are very different, $T_{obs} = |\bar{X} - \bar{Y}|$ is very large. So, $I(T_i > T_{obs})$ will likely be close to 0. Therefore, we will reject $\because p\text{-value} \rightarrow 0$.
so $X \neq Y$

Q4) POISSON-GAMMA CONJUGATE**(Total 5 points)**

Let X_1, X_2, \dots, X_n be i.i.d. as $\text{Poisson}(\lambda)$. Recall that the p.m.f. of $\text{Poisson}(\lambda)$ is $p(x) = \frac{e^{-\lambda} \lambda^x}{x!}$. Let the prior of λ be $\text{Gamma}(\alpha, \beta)$, where the p.d.f. of $\text{Gamma}(\alpha, \beta)$ is such that $f(x)$ is proportional to $x^{\alpha-1} e^{-x\beta}$. Show that the posterior is also a Gamma and find its parameters. Show all steps clearly. Feel free to ignore the constants and conclude that the posterior is a Gamma if its form resembles that of $f(x)$ above.

Bayesian

$$f(\lambda | \{X_1, \dots, X_n\}) \propto \underbrace{L(\lambda)}_{\text{likelihood}} \cdot \underbrace{f(\lambda)}_{\text{prior}}$$

$$= \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{X_i}}{(X_i!)} \times \lambda^{\alpha-1} \cdot e^{-\lambda\beta}$$

$$= \frac{e^{-n\lambda} \lambda^{\sum X_i}}{\prod (X_i!)} \cdot \lambda^{\alpha-1} \cdot e^{-\lambda\beta}$$

$$\propto e^{-\lambda(n+\beta)} \cdot \lambda^{\alpha-1+\sum X_i}$$

$$\sim \underline{\underline{\text{Gamma}(\alpha + \sum X_i, n + \beta)}}$$

Q5) WALD'S TEST

(Total 8 points)

You observe 36 successes in 100 flips. The null hypothesis is that the coin is unbiased.

- (a) Use the Wald's test with sample mean as estimator to Reject/Accept the null at the 95% confidence level (that is, $1-\alpha = 0.95$ and $z_{\alpha/2} = 1.96$). Use a calculator, if needed. Show all steps, including why the Wald's test is applicable here and the derivation of the standard error, with explanations. (4 points)
- (b) What is the p-value? You can leave your answer in terms of Φ , the cdf of the standard normal. Show the area represented by the p-value using a Normal diagram. (1) (2 points)
- (c) In the above test, what is the probability of a type I error? Show the area represented by this probability in the diagram from part (b). (1) (2 points)

(1) \because estimator is sample mean ($\frac{\sum X_i}{n}$), by CLT, it is Asym Normal. So, Wald's is applicable.

$$W = \left| \frac{\hat{\theta} - \theta_0}{se(\hat{\theta})} \right| \quad \hat{\theta} = \frac{\sum X_i}{n} = \frac{36}{100} \quad \theta_0 = 0.5$$

Low, iid $\left\{ \begin{array}{l} \text{iid} \\ \text{12} \end{array} \right\}$

$$se(\hat{\theta}) = \sqrt{Var(\hat{\theta})} = \sqrt{Var\left(\frac{\sum X_i}{n}\right)} \stackrel{(12)}{=} \sqrt{\frac{1}{n^2} \sum Var(X_i)}$$

$$\stackrel{\text{iid}}{=} \sqrt{\frac{1}{n^2} \times n Var(X_1)} = \sqrt{\frac{Var(X_1)}{n}} = \sqrt{\frac{p(1-p)}{n}}, \because X \sim \text{Bern.}$$

$$\hat{se}(\hat{\theta}) = \sqrt{\frac{p(1-p)}{n}} = \sqrt{\frac{\hat{\theta}(1-\hat{\theta})}{n}} = \sqrt{\frac{\frac{36}{100} \times \frac{64}{100}}{100}}$$

$$= 0.048$$

$$\therefore W = \left| \frac{0.36 - 0.5}{0.048} \right| = \left| \frac{0.14}{0.048} \right| = \frac{140}{48} = 2.9167$$

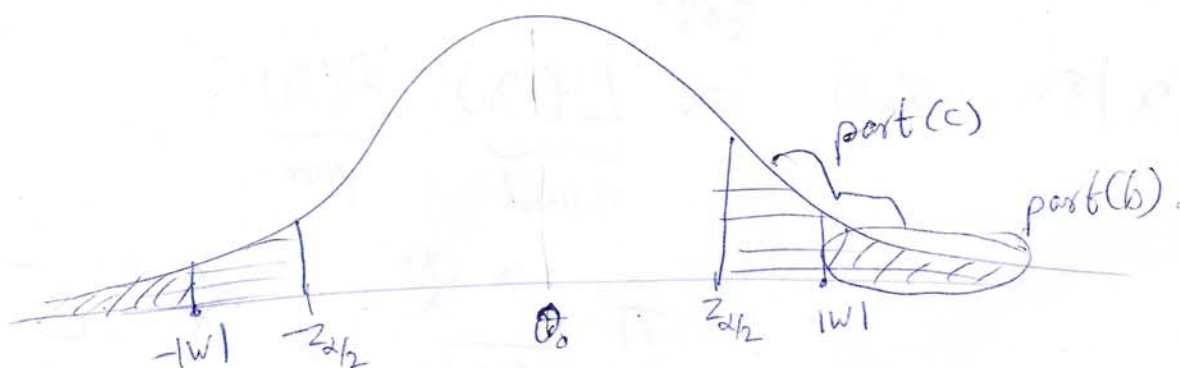
(1/2) $\because |W| > z_{\alpha/2} = 1.96$, we can't reject. ~~So, we accept.~~

$$(b) \text{ p-value} = \Pr(\text{more extreme than } |W|)$$

$$= \Pr(Z > |W|)$$

$$= 2(1 - \Pr(Z < |W|))$$

$$= 2(1 - \Phi(2.9167)).$$



(C) Type I error = incorrectly rejecting null

= Reject Null / Null is true

$$\Pr(\text{Type I}) = \Pr(\text{Reject Null} \mid \text{Null is true})$$

$$= \Pr(|W| > z_{\alpha/2})$$

$$\because \text{Null is true, } W \sim Z$$

$$= 2 \cdot \Pr(Z > z_{\alpha/2})$$

$$= 2(1 - \Phi(z_{\alpha/2}))$$

$$= 2(1 - (1 - \alpha/2)) = 2(\alpha/2) = \underline{\underline{\alpha}}$$

Q6) SIMPLE LINEAR REGRESSION**(Total 4 points)**

Given data $\{(Y_1, X_1); (Y_2, X_2); \dots; (Y_n, X_n)\}$, consider a different form of simple linear regression where the estimate of Y_i is $\hat{Y}_i = \hat{\beta} \cdot X_i$, with $\hat{\varepsilon}_i = Y_i - \hat{Y}_i$. Note that there is no $\hat{\beta}_0$ term here.

(a) Derive, from scratch, the OLS estimate of $\hat{\beta}$ that minimizes the SSE. (3 points)

(b) Find $\hat{\beta}$ for sample data $\{(3, 1); (7, 2); (9, 3)\}$, where the pairs are (Y_i, X_i) . (1 point)

$$\begin{aligned} \text{(a)} \quad SSE &= \sum \hat{\varepsilon}_i^2 = \sum (Y_i - \hat{Y}_i)^2 \\ &= \sum (Y_i - (\hat{\beta} X_i))^2 \end{aligned}$$

$$\frac{\partial SSE}{\partial \hat{\beta}} = 0 \Rightarrow \frac{\partial}{\partial \hat{\beta}} \left(\sum (Y_i - \hat{\beta} X_i)^2 \right) = 0$$

$$\Rightarrow \sum \frac{\partial}{\partial \hat{\beta}} (Y_i - \hat{\beta} X_i)^2 = 0$$

$$\Rightarrow \sum 2 \cdot (Y_i - \hat{\beta} X_i) \cdot (-X_i) = 0$$

$$\Rightarrow \left(\sum X_i Y_i \right) = \hat{\beta} \cdot \sum X_i^2 \Rightarrow \boxed{\hat{\beta}_{OLS} = \frac{\sum (X_i Y_i)}{\sum (X_i^2)}}$$

$$\text{(b)} \quad \sum (X_i Y_i) = 3 \cdot 1 + 7 \cdot 2 + 9 \cdot 3 = 3 + 14 + 27 = 44$$

$$\sum X_i^2 = 1^2 + 2^2 + 3^2 = 1 + 4 + 9 = 14$$

$$\therefore \hat{\beta} = \frac{44}{14} = \frac{22}{7}$$