

1. Model Appendix

This appendix provides a generic description of our model with relatively compact notation. I will then discuss below how this maps into the specific model of LMW, aggregate law's of motion, National Income and Product Accounting in the model, and then our equilibrium concept.

1.1. SpIM: The Spatial Incomplete Markets Model.

The state variables for a household can be divided into objects that an individual productivity state, endogenous state variables, aggregate state variables, and then iid preference shocks.

- **Individual productivity state.** Each household is subject to productivity shocks, s that is described by the Markov transition probabilities $\pi(s', s)$. The generic representation is that s could be anything, as long as it's Markov. So it could be a vector that spans the location space (as in the Roy model) or just a scalar value.
- **Endogenous state variables.** There are three endogenous (individual) state variables. The first is the household's asset holdings, a . The second is a variable that describes the household's location j . The third is whether or not the household is an inexperienced migrant, x , and, thus, whether or not it suffers disutility \bar{u} . As I describe below, this will evolve according to the transition probability function $\varphi(x', x, j)$ which may depend upon a households location.
- **Aggregate state variables.** There is an index i which describes the aggregate state.
- **Transitory moving shock.** Each household is subject to i.i.d. moving shocks ν for each location j . Here these moving shocks are independently and identically distributed across time and is distributed Type 1 extreme value distribution with scale parameter σ_ν .

Given this representation, we can express the household's problem as

$$v_j(a, s, x, i) = \max_{j'} \left\{ v_{j'j}(a, s, x, i) + \nu_{j'} \right\}, \quad (1)$$

and then the value functions associated with a move from j to j' is

$$v_{j'j}(a, s, x, i) = \max_{a' \in \mathcal{A}} \left\{ u(Ra + w_j(s, i) - a' - m_{j'j}, x) + \beta \mathbb{E}[v_{j'}(a', s', x', i')] \right\} \quad (2)$$

which says that the household chooses future asset holdings to maximize the expected present discounted value of utility. Utility depends upon consumption (the first argument) and experience status x (second argument). Given asset choices, a household's consumption equals the

gross return on current asset holdings, Ra , plus labor income $w_j(s, i)$, minus future asset holdings and minus any moving costs incurred $m_{j',j}$ for going from j to j' . The asset holdings must respect the borrowing constraint and, thus, must lie in the set \mathcal{A} . Next period's state variables are the new asset holdings, the individual productivity shock, the experience level, and the aggregate state.

Associated with this problem are the policy functions describing the optimal asset and migration choice. These will be denoted as: $g_{j',j}(a, s, x, i)$ and the indicator function $\iota_{j',j}(a, s, x, i, \nu)$. When integrating across the idiosyncratic preference shock in the migration choice gives rise to migration probabilities $\mu_{j',j}(a, s, x, i)$.

That's it. That's the household side of the model.

1.2. The LMW model.

The LMW model is a specific case of this representation.

Locations. One way to think about the LMW model is that there are essentially three locations: rural, seasonal migration, and urban. Then there is a technological restriction on the migration choices depending upon the location. So a rural household can: stay, transition to seasonal migration, or migrate to the urban area. Seasonal migrants can **only** move to the rural area. Urban migrants can: stay or migrate to the rural location. What this implies is that the migration cost matrix would take the form:

$$\begin{pmatrix} m_{\text{rural}', \text{rural}} = 0 & m_{\text{seas}', \text{rural}} = m_T & m_{\text{urban}', \text{rural}} = m_P \\ m_{\text{rural}', \text{seas}} = 0 & m_{\text{seas}', \text{seas}} = \infty & m_{\text{urban}', \text{seas}} = \infty \\ m_{\text{rural}', \text{urban}} = m_P & m_{\text{seas}', \text{urban}} = \infty & m_{\text{urban}', \text{urban}} = 0 \end{pmatrix}$$

For example, households seasonally migrating can not remain seasonal migrants in our model, thus $m_{\text{seas}', \text{seas}} = \infty$. And then the associated migration probabilities $\mu_{j',j}$ take the following form:

$$\begin{pmatrix} \mu_{\text{rural}', \text{rural}} & \mu_{\text{seas}', \text{rural}} & \mu_{\text{urban}', \text{rural}} \\ \mu_{\text{rural}', \text{seas}} & 0 & 0 \\ \mu_{\text{rural}', \text{urban}} & 0 & \mu_{\text{urban}', \text{urban}} \end{pmatrix}$$

Individual Productivity Shocks. One can always represent the state s as a pair (z, s) with $\pi(s', s)$ being non-zero for all $s' = (z, s')$ and zero for all $s' = (z', s')$ so that z is a permanent state. And thus the problem above describes the problem for a particular z . In the LMW setting,

we are then solving the problem many times for different z s. Below, we will explicitly carry around the dependence upon z to be consistent with the body of the paper.

Experience. In the LMW model, we posited a particular process for the evolution of x or experience that is indexed by the location and hence is influenced by the migration choices. To remind ourselves how this works, if a household in rural area does not have experience, it stays inexperienced; if the household is experienced, it stays experienced next period with probability π and inexperienced with probability $1 - \pi$. If a household is in the urban area and has experience it stays experienced; inexperienced urban households stay inexperienced in the next period with probability λ and become experienced with probability $1 - \lambda$. If a household is a seasonal migrant and does not have experience, then it can acquire experience with probability λ and become experienced with probability $1 - \lambda$.

The compact representation of this is as a location dependent Markov chain: $\varphi(x', x, j)$.

Aggregate State. In the LMW model, things are simple in that $i = (A_{r,i}, N_{r,i})$ with $A_{r,i}$ deterministically transitions from a good and bad season to mimic seasonal crop cycles. $N_{r,i}$ is an endogenous, aggregate state variable that depends upon location choices. It matters because of the decreasing returns to production in the rural area and, hence, it feeds back into the prices faced. In the stationary equilibrium, this will deterministically move with seasonal productivity and thus the index i is a sufficient statistic for the complete description of the aggregate states.

1.3. LMW Production

There is one homogeneous good produced in rural and urban locations by competitive producers. Seasonal migrants work in the urban location. Locations differ in the technologies they operate. The rural technology is

$$Y_{ri} = A_{ri}N_{ri}^\phi, \quad (3)$$

where N_{ri} are the effective labor units working in the rural area in season i . The parameter $0 < \phi < 1$, so that there are decreasing marginal product of labor in the rural area. And A_{ri} is rural productivity in season i . The urban technology is given by:

$$Y_{ui} = A_u N_{ui}, \quad (4)$$

where A_u captures urban productivity and N_{ui} is the effective labor units supplied by households working in the urban area in that season i .

Wages. In season i , with N_r effective labor units in the rural area, wages per efficiency unit are

$$\omega_{ri} = A_{ri}\phi N_{ri}^{\phi-1} \quad \text{and} \quad \omega_u = A_u. \quad (5)$$

Agents working in a particular location receive wages that are the product of (5) and the number of their efficiency units. Thus, the labor income that a household with transitory state s receives for working in location i is:

$$w_{ri}(s) = s\omega_{ri} \quad \text{and} \quad w_u(z, s) = zs^\gamma\omega_u, \quad (6)$$

Land Rents. There are rents that accrue to the owners of the fixed factor used in the rural production function. We assume that they are redistributed to “absentee landlords” in the economy. Below we discuss where this matters or not. For now, recognize that these rents are

$$\text{rents}_i = (1 - \phi)A_{ri}N_{ri}^\phi. \quad (7)$$

1.4. Laws of Motion and Accounting

First, it’s useful to define the migration probabilities for a given set of state variables as:

$$\mu_{j',j}(a, z, s, x, i) = \int_{\nu} \iota_{j',j}(a, z, s, x, i, \nu) d\nu, \quad (8)$$

where $\mu_{j',j}(a, z, s, x, i)$ is the probability mass of those moving from location j (rural, seasonal, and urban) to location j' . And these probabilities must respect the technological restrictions discussed above.

The distribution of households across states. Define the probability distribution of households across individual states in location j as $\lambda_j(a, z, s, x, i)$. This is the probability measure of households in location j with asset levels a , permanent shock z , individual shocks s , experience x , in season i . Furthermore, define the probability distribution of households in the next period as $\lambda_j(a, z, s, x, i')$ with the observation that the season deterministically changes with each time period.

The probability distribution $\lambda_j(a, z, s, x, i)$ evolves via the following law of motion:

$$\begin{aligned} \lambda_j(a', z, s', x', i') = & \underbrace{\int_x \int_s \int_{a': a' = g_{j,j}(a, z, s, x, i)} \mu_{j,j}(a, z, s, x, i) \lambda_j(a, z, s, x, i) \pi(s', s) \varphi(x', x, j) da ds dx}_{\text{stayers}} + \\ & \underbrace{\sum_{j' \neq j} \int_x \int_s \int_{a': a' = g_{j,j'}(a, z, s, x, i)} \mu_{j,j'}(a, z, s, x, i) \lambda_{j'}(a, z, s, x, i) \pi(s', s) \varphi(x', x, j') da ds dx}_{\text{movers}} \end{aligned} \quad (9)$$

The first bracket are those who stay in location j . This is the existing mass $\lambda_j(a, z, s, x, i)$ multiplied by the mass of households remaining in j or $\mu_{j,j}(a, z, s, x, i)$. This is then integrated over asset holdings of those households staying in j conditional on assets equalling a' tomorrow; the transition probability that s transits to s' ; that experience transits from x to x' which depends upon the location of the household.

The second bracket is the flow into location j from all other locations j' . That is the movers from locations j' into location j . Again, this is the existing mass $\lambda_{j'}(a, z, s, x, i)$ and multiplied by the mass of households moving into j or $\mu_{j,j'}(a, z, s, x, i)$. And then integration takes place over asset holdings, shock transitions, and experience transitions.

Population and Labor Supply. Define the population of location j in season i as

$$\lambda_j(i) = \int_x \int_s \int_z \int_a \lambda_j(a, z, s, x, i) da dz ds dx. \quad (10)$$

In words, the population is found integrating over the mass across asset holdings, shocks, experience levels, for a given season i . Another useful statistic is the measure of types of efficiency units in location and by season

$$\lambda_j(z, s, i) = \int_x \int_a \lambda_j(a, z, s, x, i) da dx. \quad (11)$$

Then the effective labor units in the urban area are

$$N_{ui} = \sum_{j=[\text{urban}, \text{seas}]} \int_s \int_z z s^\gamma \lambda_j(z, s, i) dz ds, \quad (12)$$

which includes the seasonal and permanent urban workforce. And in the rural area:

$$N_{ri} = \int_s \int_z s \lambda_{\text{rural}}(z, s, i) dz ds. \quad (13)$$

Asset Holdings. In each season, aggregate net-asset holdings are

$$\mathcal{A}'_i = \sum_j \sum_{j'} \int_x \int_s \int_z \int_a g_{j',j}(a, z, s, x, i) \mu_{j',j}(a, z, s, x, i) \lambda_j(a, z, s, x, i) da dz ds dx. \quad (14)$$

inside of the summation and integration this is the mass of households moving from j to j' (which is given by the migration rate $\mu_{j',j}(a, z, s, x, i)$) multiplied by the measure of households in that location $\lambda_j(a, z, s, x, i)$ multiplied by the asset policy function associated with that migration choice $g_{j',j}(a, z, s, x, i)$. This is then integrated across current period asset holdings,

individual shocks, experience levels and then summed across all location-destination pairs.

Connecting Consumption and National Income. Start from the production side of our economy. The value of aggregate production must equal aggregate payments to labor and payment to land owners who own the fixed factor that is used in rural area production.

$$Y_i = A_u N_{ui} + A_{ri} N_{ri}^\phi = \sum_j \int_z \int_s w_{ji}(z, s) \lambda_j(z, s, i) dz ds + \text{rents}_i \quad (15)$$

where the first term on the right side integrates over wage payments for each location, each shock state. The second term are the payments to land that is associated with rural location. Now by integrating over the consumers budget constraint and substituting (15) into the aggregated budget constraint, we arrive at the following:

$$Y_i - \text{rents}_i = C_i - R\mathcal{A}_i + \mathcal{A}'_i + \sum_j \sum_{j'} \int_x \int_s \int_a m_{j',j} \mu_{j',j}(a, s, x, i) \lambda_j(a, s, x, i) da ds dx \quad (16)$$

so aggregate labor income equals consumption minus (i) returns on assets (ii) new purchases of assets (iii) plus moving costs. Then defining moving costs each season as M_i and rearranging we have

$$Y_i = C_i + M_i + \left[-R\mathcal{A}_i + \mathcal{A}'_i + \text{rents}_i \right] \quad (17)$$

where C is like consumption in the standard GDP accounting identity, the moving cost is like investment, then the term in brackets represent net flows abroad or inward because (i) we are not clearing the asset market and (ii) who holds the rents are not explicitly accounted for.

1.5. The Decentralized Equilibrium

Here we define a **Stationary Decentralized Equilibrium** with the name signifying that this is the equilibrium which would arise in a “decentralized” market economy in contrast to the allocations that would be chosen by a centralized planner.

A Stationary Decentralized Equilibrium. A Stationary Decentralized Equilibrium are asset and moving policy functions $\{ g_{j',j}(a, z, s, x, i), \iota_{j',j}(a, z, s, x, i, \nu) \}$, a probability distribution $\lambda_j(a, z, s, x, i)$, and positive real numbers $N_{ri}, N_{ui}, w_{ri}(s), w_u(z, s)$ such that

- i The prices $(w_{ri}(s), w_u(z, s))$ satisfy (5);
- ii The policy functions solve the household’s optimization problem in (1, 2);
- iv The probability distribution $\lambda(a, j, z, s, x, i, \nu)$ induced by $\{ g_{j',j}(a, z, s, x, i), \iota_{j',j}(a, z, s, x, i, \nu), \pi(s', s), \varphi(x', x, j), \phi(\nu) \}$ satisfies (9) and is a stationary

distribution;

iv Effective labor units in the rural and urban areas satisfy (12, 13).

The final object we define is utilitarian social welfare function as

$$\mathcal{W}^{\text{DCE}} = \sum_j \int_a \int_x \int_s \int_z v_j(a, z, s, x, i) \lambda_j(a, z, s, x, i) ds dx da. \quad (18)$$

here the social welfare function is the average value of households across locations, shock states, experience, assets, and preference states. Below we provide a fuller discussion of this social welfare function in the context of our welfare analysis.

2. The Decentralized Equilibrium with Government Intervention

This section describes the model with a Government who is implementing a means-tested migration subsidy similar to Bryan et al. (2014) experiment and finances it through a tax on labor income. The households migration choice is the same as in (1), but the value function associated across different location options is now:

$$v_{j'j}(a, z, s, x, i) = \max_{a' \in \mathcal{A}} \left\{ u(Ra + (1 - \tau)w_j(z, s, i) - a' - [m_{j',j} + \tilde{m}_{j',j}(a)], x) + \beta \mathbb{E}[v_{j'}(a', z, s', x', i')] \right\} \quad (19)$$

where τ is the labor income tax rate and $\tilde{m}_{j',j}(a)$ is the migration subsidy that depends upon a households locatio-destination decision and asset holdings. In the context of Bryan et al. (2014), the value of migration subsidy is only positive for rural households seasonally migrating ($j = \text{rural}, j' = \text{seas}$) and when assets are below \bar{a} where \bar{a} corresponds with (approximately) the median level of assets in the rural village. For all other households the value of this subsidy is zero.

Define the Government's fiscal year budget constraint as:

$$G = \sum_i \sum_j \int_s \int_z \tau w_j(z, s, i) \lambda_j(z, s, i) - \sum_i \sum_j \sum_{j'} \int_x \int_s \int_z \int_a \tilde{m}_{j',j}(a) \mu_{j',j}(a, z, s, x, i) \lambda_j(a, z, s, x, i) da dz ds dx = 0 \quad (20)$$

where the first term are all labor income tax receipts and the second term are all the outlays of migration subsidies. We ask that the government balance it's budget, fiscal year by fiscal year. We are emphasizing fiscal year which is comprised of the lean and good season. Thus, we are allowing the Government to collect taxes and pay out subsidies that must balance across all seasons, but that may not balance within a season.

A Stationary Decentralized Equilibrium with Government Intervention A Stationary Decentralized Equilibrium with Government Intervention are asset and moving policy functions $\{ g_{j',j}(a, z, s, x, i), \iota_{j',j}(a, z, s, x, i, \nu) \}$, a probability distribution $\lambda_j(a, z, s, x, i)$, and positive real numbers $N_{ri}, N_{ui}, w_{ri}(s), w_u(z, s)$ and government policy $\{ \tilde{m}_{j',j}(a), \tau \}$ such that

- i The prices $(w_{ri}(s), w_u(s))$ satisfy (5);
- ii The policy functions solve the household's optimization problem in (1, 19);
- iv The probability distribution $\lambda(a, j, z, s, x, i, \nu)$ induced by $\{ g_{j',j}(a, z, s, x, i), \iota_{j',j}(a, z, s, x, i, \nu), \pi(s', s), \varphi(x', x, j), \phi(\nu) \}$ satisfies (9) and is a stationary distribution;
- iv Effective labor units in the rural and urban areas satisfy (12, 13).
- v The Government's fiscal year budget constraint (20) is satisfied.

The differences relative to the Decentralized Equilibrium are that the policy functions must satisfy the households problem with taxes and subsidies in (19) and that the Government's budget constraint is satisfied in (20)

3. Appendix The Centralized Equilibrium

3.1. The Social Welfare Function

We focus on a utilitarian planner placing equal weight on households. Define the social welfare function as

$$\mathcal{W}^{SP} = \sum_{t=0}^{\infty} \sum_j \int_z \int_s \int_x \int_{\nu} \beta^t u_{j',j}(c_{j',j}(z, s, x, t), x, \nu) \lambda_j(z, s, x, \nu, t) dz ds dx d\nu. \quad (21)$$

Here social welfare is the average value of households utility across locations j , productivity states z and s , experience x , and preference shocks ν . The average is computed with respect to the measure of households $\lambda_j(z, s, x, \nu, t)$ with those shock states, experience levels, and preference shocks at date all dates t . Utility depends directly upon the consumption allocation $c_{j',j}(z, s, x, t)$, but also directly on the location j through the \bar{u} , and the idiosyncratic preference shock across moving options j'

We cast the Planners Problem in terms of the planner choosing consumption allocations and migration probabilities. To cast the problem in terms of migration probabilities, we integrate out the preference shocks conditional on a set of migration probabilities for each household state.

These migration probabilities prescribe an assignment of those households with the largest relative preference shock to migrate or not. So given set of states j, z, s, x, t , utility is

$$u(c_{j',j}(z, s, x, t), x) + E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}]. \quad (22)$$

where $\mu_{j',j}(z, s, x, t)$ is the migration probability going from location j to location j' and then $E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}]$ is the expected value of the preference shock conditional on the migration probabilities. So, for example, if all people migrate location j to location j' , then this value is just the unconditional mean of a Type 1 extreme value shock. Now we can write the social welfare function as

$$\mathcal{W}^{SP} = \sum_{t=0}^{\infty} \sum_j \int_z \int_s \int_x \beta^t \left\{ u(c_{j',j}(z, s, x, t), x) + E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}] \right\} \lambda_j(z, s, x, t) dz ds dx. \quad (23)$$

3.2. The Law of Motion and Feasibility

The Planning Problem maximizes (23) subject to the law of motion describing how the population evolves across states and locations and then how many resources there are available, i.e., feasibility. We describe each of these aspects of the environment below.

Law of Motion. The law of motion describing how the measure of households evolves across states and locations is

$$\begin{aligned} \lambda_j(z, s', x', t+1) &= \int_s \int_x \mu_{j,j}(z, s, x, t) \pi(s', s) \varphi(x', x, j) \lambda_j(z, s, x, t) ds dx \\ &+ \sum_{j' \neq j} \int_s \int_x \mu_{j,j'}(z, s, x, t) \pi(s', s) \varphi(x', x, j') \lambda_{j'}(z, s, x, t) ds dx. \end{aligned} \quad (24)$$

This equation says, given the current distribution $\lambda_{j'}(z, s, x, t)$ in location j' , the measure of households $\lambda_j(z, s', x', t+1)$ reflects the migration probabilities of households in each location, how their productivity evolves over time ($\pi's$), and how their experience $\varphi(x', x, j)$ evolves.

Labor Supply, Aggregate Production, and the Resource Constraint. Given a distribution of

households, the effective labor units in the urban and rural area are

$$N_{u,t} = \sum_{j=[\text{urban,seas}]} \int_z \int_s \int_x z s^\gamma \lambda_j(z, s, x, t) dz ds dx,$$

$$N_{r,t} = \int_z \int_s \int_x s \lambda_{\text{rural}}(z, s, x, t) dz ds dx. \quad (25)$$

with the urban area includes the seasonal and permanent urban workforce. Aggregate production of the final good is

$$Y_t = A_u N_{u,t} + A_{r,t} (N_{r,t})^\phi. \quad (26)$$

Combining the amount of resources available in (26) with the consumption and moving decisions we have the following resource constraint:

$$Y_t \geq \sum_j \sum_{j'} \int_z \int_s \int_x c_{j',j}(z, s, x, t) \lambda_j(z, s, x, t) dz ds dx$$

$$+ \sum_j \sum_{j'} \int_z \int_s \int_x m_{j',j} \mu_{j',j}(z, s, x, t) \lambda_j(z, s, x, t) dz ds dx. \quad (27)$$

which says that production must be greater than or equal to consumption which is the first term on the righthand side of (27) and the moving costs associated with the migration of households across locations which is the second term on the righthand side. Here we compactly sum across all j' and j location pairs with the notion that the moving cost for staying in a location is zero, i.e., $m_{j,j} = 0$.

3.3. The Centralized Planners Problem

Given everything defined above, the **Centralized Planner's Problem** is the following:

$$\mathcal{W}^* = \max_{c_{j',j}(z,s,x,t), \mu_{j',j}(z,s,x,t)} \sum_{t=0}^{\infty} \sum_j \int_s \int_x \beta^t \left\{ u(c_{j',j}(z, s, x, t), x) + E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}] \right\} \lambda_j(z, s, x, t) dz ds dx$$

$$\text{subject to (26) (27) and (24) and an initial condition } \lambda_j(z, s, x, 0). \quad (28)$$

Then given this problem, define the following allocation:

A Stationary Centralized Planner Allocation. A Stationary Centralized Planner Allocation are time invariant policy functions $\{ c_j(z, s, x, i), \mu_{j',j}(z, s, x, i) \}$, a probability distribution

$\lambda_j(z, s, x, i)$, and positive real numbers $N_{j,i}$ for rural and urban areas and season i where:

- i The policy functions solve the Centralized Planner's Problem in (28);
- ii The probability distribution $\lambda_j(z, s, x, i)$ associated with $\{ \mu_{j',j}(z, s, x, i), \pi(s', s), \varphi(x', x, j), \phi(\nu) \}$ is a stationary distribution;
- iii Effective labor units in the rural and urban areas satisfy (25).

3.4. Solution to the Centralized Planner's Problem

The approach to solving the planning problem proceeds in the following way. First, we formulate the problem in (28) using Lagrangian methods under the operating assumption that these methods are usable. Second, we then derive the necessary conditions associated the planner's consumption allocation and migration probabilities.

We express the **Centralized Planner's Problem** as:

$$\begin{aligned}
\mathcal{L} = & \sum_{t=0}^{\infty} \sum_j \sum_{j'} \int_z \int_s \int_x \beta^t \left\{ u(c_{j',j}(z, s, x, t), x) + E[\nu \mid \{ \mu_{j',j}(z, s, x, t) \}_{j'}] \right\} \lambda_j(z, s, x, t) dz ds dx \quad (29) \\
& + \sum_{t=0}^{\infty} \chi(t) \left\{ Y_t - \sum_j \sum_{j'} \int_z \int_s \int_x [c_{j',j}(z, s, x, t) - m_{j',j} \mu_{j',j}(z, s, x, t)] \lambda_j(z, s, x, t) dz ds dx \right\} \\
& + \sum_{t=0}^{\infty} \sum_j \int_z \int_s \int_x \chi_{2j}(z, s, x, t) \left\{ 1 - \sum_{j'} \mu_{j',j}(z, s, x, t) \right\} \lambda_j(z, s, x, t) dz ds dx \\
& - \sum_{t=0}^{\infty} \sum_j \int_z \int_{s'} \int_{x'} \beta^{t+1} \chi_{3j}(z, s', x', t+1) \left\{ \lambda_j(z, s', x', t+1) - \right. \\
& \quad \left. \sum_{j'} \int_z \int_s \int_x \mu_{j,j'}(z, s, x, t) \pi(s', s) \varphi(x', x, j') \lambda_{j'}(z, s, x, t) dz ds dx \right\}
\end{aligned}$$

where the first term is the social welfare function and $\chi(t)$ is the lagrange multiplier on the resource constraint. The multipliers $\chi_{2j}(s, x, t)$ are on the constraint that the destination migration probabilities leaving source j must sum to one.

And then the multipliers $\chi_{3j}(s', x', t+1)$ are on the law of motion for the probability distribution, that is migration flows into j from all j' must equal the probability mass in j at date $t+1$. On this last multiplier, we do the following that will ease the algebra below: scale the multiplier

$\chi_{3j}(s', x', t + 1)$ by $-\beta^{t+1}$. This which does not change the problem, but only a scaling trick to see things quicker in the algebra below.

Consumption Allocations. Takeing the first order condition of (29) with respect to consumption gives:

$$\beta^t u'(c_{j',j}(z, s, x, t)) = \chi(t) \forall j, j', s, x, \text{ and } t. \quad (30)$$

which says set the marginal utility of consumption for every location, productivity state, and experience state equal its marginal cost at each date. In a time-invariant, stationary allocation, and implication of (30) is that for each season i :

$$u'(c_{j',j}(z, s, x, i)) = u'(c_{\tilde{j}',\tilde{j}}(z, s', x', i)) \forall j, s, x \quad (31)$$

so that marginal utility is equated within each season and **all** households bear the aggregate consequences of seasonality. There are several observations to make about this result. First, the destination does not matter in how much consumption a household recipes. The source does matter indirectly because of how a household's experience status and location affect its marginal utility of consumption. In other words, the planner will not equate the **level** of consumption across households. For example, compare the marginal utility of a rural household with a seasonal migrant who does not have experience

$$c_{\text{rural}}(s, x, i)^{-\alpha} = c_{\text{seas}}(s, x, i)^{-\alpha} \bar{u}. \quad (32)$$

But because the disutility of migration affects the marginal utility of consumption for non-experienced migrants, this implies that

$$c_{\text{rural}}(s, x, i) < c_{\text{seas}}(s, x, i) \quad (33)$$

so that non-experienced migrants must be compensated for migrating to equate the marginal utility of consumption across households.

Migration Probabilities. First, let's take stock of the challenge here. The key issue is that when deriving the first order conditions, the planner understands that he has control of the distribution λ next period via the migration probabilities chosen today. Thus, via the chain rule there are two effects. The direct effect arising from a small change in the migration rate and the indirect effect from how the distribution changes like this:

$$\left\{ \frac{\partial \text{stuff}}{\partial \mu_{j',j}(z, s, x, t)} \right\} \lambda_j(z, s, x, t) + \left\{ \text{other stuff} \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} = 0 \quad (34)$$

And the overall strategy is to find a solution that sets things on the inside of the brackets (for stuff and other stuff) equal to zero.

This is the first “stuff” component, holding fixed the distribution:

$$\left\{ \frac{\partial \text{stuff}}{\partial \mu_{j',j}(z, s, x, t)} \right\} \lambda_j(z, s, x, t) = \left\{ \beta^t \frac{\partial E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}]}{\partial \mu_{j',j}(z, s, x, t)} - \chi(t) m_{j',j} \right. \\ \left. - \chi_{2j}(z, s, x, t) + \underbrace{\beta^{t+1} \int_{s'} \int_{x'} \chi_{3j'}(z, s', x', t+1) \pi(s', s) \varphi(x', x, j)}_{\beta^{t+1} \mathbb{E}[\chi_{3j'}(t+1) \mid z, s, x]} \right\} \lambda_j(z, s, x, t) \quad (35)$$

Focusing on the right-hand side of the equation, the first term is how the expected value of the preference shocks change in that location. As the planner moves households more or less, the planner is affecting utility via the preference shock. The next the next term is the cost of moving evaluated at the marginal utility of consumption then minus a source multiplier $\chi_{2j}(s, t)$ and a discounted destination multiplier $\beta^{t+1} \mathbb{E}[\chi_{3j'}(t+1) \mid z, s, x]$. This last term integrates across all future shock and experience states for a particular destination j' , conditional on a households current state. As we show below, this represents the **expected** social value of the destination j' .

Let's set the term inside the brackets of (35) equal to zero by solving out for the migration rates as a function of the multipliers. The Type 1 extreme value assumption on the shocks implies that

$$E[\nu \mid \mu_{j',j}(z, s, x, t)] = -\sigma \sum_{j'} \mu_{j',j}(z, s, x, t) \log(\mu_{j',j}(z, s, x, t)) \quad (36)$$

and then it's derivative with respect to the migration rate is:

$$\frac{\partial E[\nu \mid \mu_{j',j}(z, s, x, t)]}{\partial \mu_{j',j}(s, x, t)} = -\sigma \log(\mu_{j',j}(z, s, x, t)) - \sigma \quad (37)$$

and with the proposed strategy to set things in the brackets equal to zero implies that

$$-\sigma\beta^t \log(\mu_{j'j}(z, s, x, t)) = \sigma + \chi(t)m_{j',j} + \chi_{2j}(z, s, x, t) - \beta^{t+1}\mathbb{E}[\chi_{3j'}(t+1)|z, s, x] \quad (38)$$

$$\mu_{j'j}(z, s, x, t) = \exp\left(\frac{\sigma + \chi(t)m_{j',j} + \chi_{2j}(z, s, x, t) - \beta^{t+1}\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{-\sigma\beta^t}\right) \quad (39)$$

$$= \exp\left(\frac{\sigma + \chi(t)m_{j',j} - \beta^{t+1}\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{-\sigma\beta^t}\right) \bigg/ \exp\left(\frac{\chi_{2j}(z, s, x, t)}{\sigma\beta^t}\right) \quad (40)$$

Where the last line follows from substituting out the source j , date t multiplier and properties of the exp function. The final step is to set $\chi_{2j}(z, s, x, t)$ source multiplier so that the migration probabilities sum to one. This implies that

$$\exp\left(\frac{\chi_{2j}(s, x, t)}{\beta^t\sigma}\right) = \sum_{j'} \exp\left(\frac{\sigma + \chi(t)m_{j',j} - \beta^{t+1}\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{-\sigma}\right) \quad (41)$$

$$\Rightarrow \chi_{2j}(s, t) = \sigma \log \left\{ \sum_{j'} \exp\left(\frac{\sigma + \chi(t)m_{j',j} - \beta^{t+1}\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{-\sigma\beta^t}\right) \right\}. \quad (42)$$

After canceling terms, we have that

$$\mu_{j'j}(z, s, x, t) = \exp\left(\frac{-u'(t)m_{j',j} + \beta\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{\sigma}\right) \bigg/ \sum_{j'} \exp\left(\frac{-u'(t)m_{j',j} + \beta\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{\sigma}\right). \quad (43)$$

The socially optimal migration rates take the familiar form associated with Type 1 extreme value shocks. And this can be separated into a source component (the denominator) and then destination j' components (the numerator). And the destination components depends on the costs of migration today relative to the discounted, expected value of the multiplier $\mathbb{E}[\chi_{3j'}(t+1)|z, s, x]$ in location j .

Recap. We have solved to the migration probabilities that set “stuff” in (34) equal to zero. We need the whole term to be set to zero, so the approach is to find the value of the multiplier(s)

$\chi_{3j'}(z, s', x', t + 1)$ so that the “other stuff” in (34) is also equal to zero. Step by step we have:

$$\left\{ \text{other stuff} \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} = \quad (44)$$

$$\sum_j \int_s \int_x \beta^{t+1} \left\{ \underbrace{u(c(z, s, x, t + 1), x) + E[\nu \mid \{\mu_{j',j}(z, s, x, t)\}_{j'}]}_{u_j(z, s, x, t+1)} \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx + \quad (45)$$

$$\sum_j \int_s \int_x \chi(t + 1) \left\{ \underbrace{\frac{\partial Y(t + 1)}{\partial \lambda_j(z, s, x, t + 1)} - c(z, s, x, t + 1) - \sum_{j'} m_{j',j} \mu_{j',j}(z, s, x, t + 1)}_{\kappa_j(z, s, x, t+1)} \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx \quad (46)$$

$$+ \underbrace{\sum_j \int_s \int_x \chi_{2j}(z, s, x, t + 1) \left\{ 1 - \sum_{j'} \mu_{j',j}(z, s, x, t + 1) \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx}_{=0} \quad (47)$$

$$- \beta^{t+1} \sum_j \int_s \int_x \chi_{3j}(z, s, x, t + 1) \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx \quad (48)$$

$$+ \beta^{t+2} \sum_j \int_s \int_x \left\{ \underbrace{\sum_{j'} \int_{s'} \int_{x'} \chi_{3j'}(z, s', x, t + 2) \mu_{j',j}(z, s, x, t + 1) \pi(s', s) \varphi(x', x, j)}_{\mathbb{E}[\chi_3(t+2)|z, s, x, j]} \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx. \quad (49)$$

If we put this all together than we have the following:

$$\sum_j \int_s \int_x \left\{ \beta^{t+1} u_j(z, s, x, t + 1) + \chi(t + 1) \kappa_j(z, s, x, t + 1) - \right. \quad (50)$$

$$\left. \beta^{t+1} \chi_{3j}(z, s, x, t + 1) + \beta^{t+2} \mathbb{E}[\chi_3(t + 2)|z, s, x] \right\} \frac{\partial \lambda_j(z, s, x, t + 1)}{\partial \mu_{j',j}(z, s, x, t)} ds dx = 0. \quad (51)$$

And the proposed solution is to set

$$\chi_{3j}(z, s, x, t + 1) = u_j(z, s, x, t + 1) + u'(t + 1) \kappa_j(z, s, x, t + 1) + \beta \mathbb{E}[\chi_3(t + 2)|z, s, x]. \quad (52)$$

The multiplier $\chi_{3j}(z, s, x, t + 1)$ takes on a recursive formulation where the multiplier equals utility, the κ term (more on this below), and then the expected, discounted multiplier tomorrow

taking into account all the different possible moving options (observe that our notion does not have a j in the expectation above, so this denotes given a realized states, integrate over shocks and different location possibilities).

A few details to arrive at Proposition 1. First, notice that $\frac{\partial Y(t+1)}{\partial \lambda_j(z, s, x, t+1)}$ is simply the marginal product of labor in location j multiplied by the relevant shock primitive, s for rural. Then combining the recursive expression for the multiplier in (52), the definition of $\kappa_{j'}(z, s, x, t+1)$, the migration probabilities, and consumption allocations we have Proposition 1:

Proposition 1 (Efficient Consumption and Migration) *The consumption and migration rates that solving the Planners Problem in (28) are: Consumption allocations equate the marginal utility of consumption in all locations, productivity and experience states for each date t :*

$$u'(t) = u'(c_{j',j}(z, s, x, t)) = u'(c_{\tilde{j}',\tilde{j}}(z, s', x', t)) \forall j, s, x. \quad (53)$$

Migration probabilities satisfy

$$\mu_{j',j}(z, s, x, t) = \exp\left(\frac{-u'(t) m_{j',j} + \beta \mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{\sigma}\right) \Bigg/ \sum_{j'} \exp\left(\frac{-u'(t) m_{j',j} + \beta \mathbb{E}[\chi_{3j'}(t+1)|z, s, x]}{\sigma}\right), \quad (54)$$

with the multiplier $\chi_{3j'}(z, s, x, t+1)$ satisfying the following recursive relationship

$$\chi_{3j'}(z, s, x, t+1) = u_{j'}(z, s, x, t+1) + u'(t+1)\kappa_{j'}(z, s, x, t+1) + \beta \mathbb{E}[\chi_3(t+2)|z, s, x], \quad (55)$$

where

$$\kappa_{j'}(z, s', x', t+1) = mp l_{j'}(z, s', t+1) - c(z, s', x', t+1) - \sum_{j''} m_{j'',j'} \mu_{j'',j'}(z, s', x', t+1). \quad (56)$$

3.5. Algorithm for Computation

The basic idea is to use a “value function iteration” like approach given the structure above to compute the migration probabilities. Then given the migration probabilities, we can compute where people are, production, aggregate consumption, then check if markets clear or not and then update. Also, in all of this we look for a stationary equilibrium, so the t state is dropped. Here are the details:

1. Guess consumption allocations, mpl 's, and χ 's. Given the solution above, the consumption allocations are those for experienced and non-experienced and monga and not monga. But because of the functional form on utility, non-experienced consumption is just a scalar

of experienced consumption, so we only need an initial guess of size 2×1 vector. And then the *mpls* are only equilibrium objects for the rural area in monga and not monga so an initial guess of 2×1 .

The χ 's then are just like a value function. A χ for each location, experience, shock state.

2. Given the guessed χ 's, compute the expected value of a transition to a new location, so $\mathbb{E}[\chi_{3j'}(t+1)|s, x]$. In words, this would be like what is the expected value of going to (staying) the rural area (j') next period, given your current state s and x .
3. Compute the migration probabilities using the expected values from 2. and marginal utility of consumption implied by 1.
4. Update the χ 's using (55) and (56) by computing all the objects on the right-hand side (given our guesses) and implying a new updated χ' .

A couple of details: remember that $u_{j'}(s, x, t+1)$ integrates out the preference shock. So this reflects expected utility given the migration probabilities for that location, it's the log sum thing. In words, it's like before I know where I'm going when in location j' , what is my expected utility. Second, $\mathbb{E}[\chi_3(t+2)|s, x]$ are the expected values for all realizations of the preference shock, so it is migration probabilities multiplied by $\mathbb{E}[\chi_{3j'}(t+1)|s, x]$. In words, this is saying, before I know where I'm going (my realization of the preference shock), what is my the expected value next period.

5. Given updated χ' , return to 2. and iterate until χ' and χ are close. This works well and different initial guesses go to the same place, etc.
6. Given the migration probability implied by the χ 's, compute aggregate production, consumption and *mpl*'s. Check if production is consistent with consumption and *mpl*'s, if not update guess of consumption and *mpl*'s and return to 1. The updating here was done with `fsolve` which works well.

4. Alternative Calibrations

Table 1: Targeted Moments in Data and Model with Different R values

Moments	Data	Model $R=0.93$	Model $R=0.95$	Model $R=0.97$
Control: Variance of log consumption growth in rural	0.19	0.19	0.19	0.19
Control: Percent of rural households with no liquid assets	47	60	48	38
Control: Seasonal migrants	36	38	36	35
Control: Consumption increase of migrants (OLS)	10	9	10	11
Treatment: Seasonal migration relative to control	22	20	21	21
Treatment: Seasonal migration relative to control in year 2	9	4	4	4
Treatment: Consumption of induced migrants (LATE)	30	30	29	27
Control: Probability of repeat migration	68	71	70	70
Urban-Rural wage gap	1.89	1.88	1.89	1.90
Percent in rural	62	59	60	60
Variance of log wages in urban	0.56	0.56	0.56	0.56

Note: The table reports the main moments of the paper for alternative values of R . The estimated model has $R = 0.95$. The model is not re-estimated in the cases of $R = 0.93$ and $R = 0.97$.

5. Alternative Calibrations

Table 2: Targeted Moments in Data and Model with Different β values

Moments	Data	Model $\beta=0.93$	Model $\beta=0.95$	Model $\beta=0.97$
Control: Variance of log consumption growth in rural	0.19	0.19	0.19	0.19
Control: Percent of rural households with no liquid assets	47	59	48	38
Control: Seasonal migrants	36	37	36	36
Control: Consumption increase of migrants (OLS)	10	8	10	12
Treatment: Seasonal migration relative to control	22	20	21	20
Treatment: Seasonal migration relative to control in year 2	9	4	4	4
Treatment: Consumption of induced migrants (LATE)	30	29	29	28
Control: Probability of repeat migration	68	71	70	71
Urban-Rural wage gap	1.89	1.90	1.89	1.87
Percent in rural	62	60	60	59
Variance of log wages in urban	0.56	0.56	0.56	0.56

Note: The table reports the main moments of the paper for alternative values of $\beta = 0.95$. The estimated model has $\beta = 0.95$. The model is not re-estimated in the cases of $\beta = 0.93$ and $\beta = 0.97$.

Table 3: Targeted Moments in Data and Models with no \bar{u} and $\rho = 0$

Moments	Data	Model Full	Model $\bar{u} = 1$	Model $\rho = 0$
Control: Variance of log consumption growth in rural	0.19	0.19	0.19	0.28
Control: Percent of rural households with no liquid assets	47	48	48	2
Control: Seasonal migrants	36	36	55	34
Control: Consumption increase of migrants (OLS)	10	10	-7	17
Treatment: Seasonal migration relative to control	22	21	10	22
Treatment: Seasonal migration relative to control in year 2	9	4	0	4
Treatment: Consumption of induced migrants (LATE)	30	29	23	21
Control: Probability of repeat migration	68	70	56	71
Urban-Rural wage gap	1.89	1.89	1.86	1.88
Percent in rural	62	60	73	57
Variance of log wages in urban	0.56	0.56	0.65	1.54

Note: The table reports the moments targeted using simulated method of moments and their values in the data and in the model.

Table 4: Welfare Under Alternative Models with no \bar{u} and $\rho = 0$

		Full Model		$\bar{u} = 1$		$\rho = 0$	
		Welfare	Migr. Rate	Welfare	Migr. Rate	Welfare	Migr. Rate
	Income Quintile						
	1	1.0	85	1.5	87	1.3	64
	2	0.4	62	0.7	74	0.8	61
	3	0.2	53	0.4	64	0.5	57
	4	0.1	43	0.3	54	0.3	52
	5	0.1	39	0.3	49	0.1	48
	Average	0.4	57	0.6	65	0.6	56

Note: The table reports the (lifetime) consumption-equivalent welfare gains from the conditional migration transfers relative to an unconditional transfer program costing the same total amount and to a rural workfare program that costs the same amount. The numbers in the table are the average percent increase in consumption each period that would make the households indifferent between the consumption increase and the transfers, and the seasonal migration rates, by quintile of the rural income distribution.

Table 5: Targeted Moments in Data and Model with Subsistence

Moments	Data	Model	
		Full Calibration	Full Calibration w/ Subsistence
Control: Variance of log consumption growth in rural	0.19	0.19	0.23
Control: Percent of rural households with no liquid assets	47	48	0
Control: Seasonal migrants	36	36	76
Control: Consumption increase of migrants (OLS)	10	10	5
Treatment: Seasonal migration relative to control	22	21	14
Treatment: Seasonal migration relative to control in year 2	9	4	4
Treatment: Cons of induced migrants relative to control (LATE)	30	29	46
Control: Probability of repeat migration	68	70	80
Urban-Rural wage gap	1.89	1.89	1.66
Percent in rural	62	60	56
Variance of log wages in urban	0.56	0.56	0.64

Note: The table reports the moments targeted using simulated method of moments and their values in the data and in the model. The final calibration reports the moments when a subsistence consumption constraint is added and set to equal 25 percent of average rural consumption in the lean season.

Table 6: Welfare Effects of Migration Subsidies with Subsistence

		Benchmark Model		Subsistence	
		Welfare	Migr. Rate	Welfare	Migr. Rate
Income Decile	1	1.5	93	2.0	92
	2	0.6	76	1.0	82
	3	0.5	68	0.8	77
	4	0.3	57	0.6	71
	5	0.3	54	0.5	67
	6	0.2	51	0.4	61
	7	0.2	46	0.3	56
	8	0.2	42	0.3	53
	9	0.1	41	0.2	52
	10	0.1	38	0.3	45
Average		0.4	57	0.6	65

Note: The table reports the (lifetime) consumption-equivalent welfare gains from the conditional migration transfers. The numbers in the table are the average percent increase in consumption each period that would make the households indifferent between the consumption increase and the transfers, and the seasonal migration rates, by decile of the rural income distribution. The first column is the benchmark model, and the second is a model with a subsistence constraint equal to 25 percent of average rural consumption in the lean season.

Table 7: Targeted Moments in Data and Model and Migration Costs

Moments	Data	Model	
		Full Cal $m_p = 2 * m_t$	Full Cal $m_p = m_t$
Control: Variance of log consumption growth in rural	0.19	0.19	0.18
Control: Percent of rural households with no liquid assets	47	48	45
Control: Seasonal migrants	36	36	32
Control: Consumption increase of migrants (OLS)	10	10	11
Treatment: Seasonal migration relative to control	22	21	23
Treatment: Seasonal migration relative to control in year 2	9	4	4
Treatment: Cons of induced migrants relative to control (LATE)	30	29	24
Control: Probability of repeat migration	68	70	60
Urban-Rural wage gap	1.89	1.89	1.79
Percent in rural	62	60	59
Variance of log wages in urban	0.56	0.56	0.59

Note: The table reports the moments targeted using simulated method of moments and their values in the data and in the model in the benchmark calibration and under alternative assumptions about migration costs.

Table 8: Welfare Under Alternative Assumptions About Migration Costs

		Benchmark Model		$m_p = m_t$	
		Welfare	Migr. Rate	Welfare	Migr. Rate
Income Quintile	1	1.05	85	0.91	81
	2	0.40	62	0.36	60
	3	0.26	54	0.22	50
	4	0.16	43	0.15	43
	5	0.11	40	0.10	39
Average		0.39	57	0.35	55

Note: The table reports the (lifetime) consumption-equivalent welfare gains from migration transfers by income quartile under alternative assumptions about migration costs.

6. Welfare Under Alternative Parameterizations

In this section we consider alternative parameterizations of the model that yield higher welfare gains from migration subsidies. The goal is to illustrate how our model allows for an interpretation of the experiments of Section 2 based on spatial misallocation, with credit constraints and migration risk driving migration outcomes. As we show below, such an interpretation would give rise to substantially larger welfare gains from migration subsidies than found in this paper, but at the cost of making counterfactual predictions about important aspects of the experimental data.

Table ?? summarizes the model’s welfare predictions under these alternatives. The first row reproduces the main experimental moments on which we will focus, and the second row reports the model’s predictions for the same moments plus the average welfare gain from the migration transfers. The third row raises γ from the estimated value of 0.57 up to 1.5, meaning that shocks are now relatively larger in the urban area. By itself this leads welfare gains to fall to 0.12 percent consumption equivalent, the OLS coefficient of consumption on migration to rise to a counterfactual 51 percent, and the treatment effect on migration to fall to a counterfactually low level of 10 percent. The fourth row sets $\bar{u} = 1$, which means there is no disutility from migration. Welfare gains raise substantially to 0.51 percent, but the LATE falls to a counterfactually low level of 9 percent, and the migration rate in the control group rises to 55 percent, well above the data. Clearly though, the lower value of \bar{u} is an important driver of the model’s welfare gains. The fifth column doubles A_u , the urban productivity, to a value of 3. The welfare gains now increase further to 1.29 percent, while other moments remain counterfactual, in particular the seasonal migration rate, which is now an implausible 84 percent.

Table 9: Alternative Estimation with Additive Migration Disutility

Moments	Data	Benchmark	Additive
Control: Variance of rural log consumption growth	0.19 (0.03)	0.19	0.19
Control: Percent of rural households with no liquid assets	47 (1.13)	48	50
Control: Seasonal migration rate	36 (2.64)	36	45
Control: Consumption increase of migrants (OLS)	10 (4.47)	10	5
Control: Probability of repeat migration	0.68 (0.46)	0.70	0.60
Treatment: Seasonal migration relative to control	22 (2.39)	21	30
Treatment: Seasonal migration relative to control in year 2	9 (2.44)	4	-1
Treatment: Consumption increase of induced migrants (LATE)	30 (9.67)	29	14
Urban-Rural wage gap	1.89 (0.18)	1.89	1.89
Percent in rural area	62 (1.36)	60	67
Variance of log urban wages	0.56 (0.06)	0.56	0.56

Note: The table reports the moments targeted using simulated method of moments and their values in the data, in the baseline model, in the model with additive disutility, and the standard errors of the empirical moments.

To lower the migration rate back to a level similar to the data, the last row increases the migration cost up to $m = 0.19$, which is the value that matches the 36 percent migration rate in the control group again. This change also raises the amount of the migration transfer, by construction, since the migration subsidies are intended to cover the migration cost and actually induce migration. Under this parameterization, the welfare gains from the transfers rise to 1.98 percent, or five times what they are in the benchmark calibration. The source of the welfare gains now become relaxing credit constraints, which keep risk-averse migrants from reaching a much more productive urban area, in the spirit of the model of Bryan et al. (2014). Yet the data do not support such an interpretation. As one example, the LATE effect of consumption on migration is counterfactually lower than the OLS coefficient, pointing to inaccurate sorting

Table 10: Welfare Gains Under Alternative Parameterizations

	Average Welfare Gains	LATE (Cons.)	OLS (Cons.)	Treatment Effect (Migration)	Seasonal Migration Control
Data	-	30	10	22	36
Benchmark calibration	0.39	29	10	21	36
+ Higher urban risk	0.12	27	51	10	16
+ No migration disutility	0.51	9	29	28	55
+ Higher urban TFP	1.29	33	51	15	84
+ Higher migration cost	1.98	16	34	62	36

Note: This table reports the average welfare gains implied by the model, the LATE and OLS effects of migration on consumption, seasonal migration in the control group, and the treatment effect on migration implied by the model for each specific calibration. Row 1 shows the data. Row 2 is the benchmark calibration that results from the simulated method of moments. Row 3 (“+ Higher urban risk”) changes the parameter shaping the urban relative shock by setting $\gamma = 1.5$. Row 4 (“+ No migration disutility”) further removes the disutility of migration by setting $\bar{u} = 1$. Row 5 (“+ Higher urban TFP”) further doubles the level of urban TFP of 3 (instead of $A_u = 1.5$). Row 6 (“+ Higher migration cost”) sets p_T to be 50 percent of rural consumption so that the model matches seasonal migrant rates in the control group.

patterns for migrants. As another example, the treatment effect of migration is far too large, pointing to the counterfactually large migration costs in this calibration of the model.

References

BRYAN, G., S. CHOWDHURY, AND A. M. MOBARAK (2014): “Underinvestment in a Profitable Technology: The Case of Seasonal Migration in Bangladesh,” *Econometrica*, 82, 1671–1748.