Econometrics using STATA

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1

INTRODUCTION

1.1 AN OVERVIEW OF STATA'S DISTICTIVE FEATURES

- You can easily learn STATA commands, even if you do not know the syntex
- You can use STATA's do-file Editor to save time developing you analysis
- A simple command performs all computations for all the desired obervations
- Looping over variables saves time & effort
- STATA's by-groups reduce the need for programming
- STATA has many statistical features that make it uniquely powerful
- You can avoid problems by keeping STATA up to date
- STATA is infinitely extensible
- STATA's user community provides a wealth of useful additions to STATA
- STATA is cross-platform compatible
- STATA can be fun

WORKING WITH ECONOMIC AND FINANCIAL DATA IN STATA

2.1 THE BASICS

2.1.1 The use Command

- . use census2c.(1980 Census data for NE and NC states)
- . list, sep(0)

| - | + | | | | | | |
|-----|---------------|---------|---------|---------|-------|-------|------|
| | state | | | | | | |
| _ | | | | | | | |
| | Connecticut | | 3107.6 | | | | |
| 2. | Illinois | N Cntrl | 11426.5 | 9518.0 | 29.90 | 109.8 | 51.0 |
| 3. | Indiana | N Cntrl | 5490.2 | 3525.3 | 29.20 | 57.9 | 40.0 |
| 4. | Iowa | N Cntrl | 2913.8 | 1708.2 | 30.00 | 27.5 | 11.9 |
| 5. | Kansas | N Cntrl | 2363.7 | 1575.9 | 30.10 | 24.8 | 13.4 |
| 6. | Maine | NE | 1124.7 | 534.1 | 30.40 | 12.0 | 6.2 |
| 7. | Massachusetts | NE | 5737.0 | 4808.3 | 31.20 | 46.3 | 17.9 |
| 8. | Michigan | N Cntrl | 9262.1 | 6551.6 | 28.80 | 86.9 | 45.0 |
| 9. | Minnesota | N Cntrl | 4076.0 | 2725.2 | 29.20 | 37.6 | 15.4 |
| 10. | Missouri | N Cntrl | 4916.7 | 3349.6 | 30.90 | 54.6 | 27.6 |
| 11. | Nebraska | N Cntrl | 1569.8 | 987.9 | 29.70 | 14.2 | 6.4 |
| 12. | New Hampshire | NE | 920.6 | 480.3 | 30.10 | 9.3 | 5.3 |
| 13. | New Jersey | NE | 7364.8 | 6557.4 | 32.20 | 55.8 | 27.8 |
| 14. | New York | NE | 17558.1 | 14858.1 | 31.90 | 144.5 | 62.0 |
| 15. | N. Dakota | N Cntrl | 652.7 | 318.3 | 28.30 | 6.1 | 2.1 |
| 16. | Ohio | N Cntrl | 10797.6 | 7918.3 | 29.90 | 99.8 | 58.8 |
| 17. | Pennsylvania | NE | 11863.9 | 8220.9 | 32.10 | 93.7 | 34.9 |
| 18. | Rhode Island | NE | 947.2 | 824.0 | 31.80 | 7.5 | 3.6 |
| 19. | S. Dakota | N Cntrl | 690.8 | 320.8 | 28.90 | 8.8 | 2.8 |
| 20. | Vermont | NE | 511.5 | 172.7 | 29.40 | 5.2 | 2.6 |
| 21. | Wisconsin | N Cntrl | 4705.8 | 3020.7 | 29.40 | 41.1 | 17.5 |
| _ | ' + | | | | | | |

2.1.2 Variable Types

- . use census2c.(1980 Census data for NE and NC states)
- . describe

Contains data from census2c.dta obs: 21

| vars: size: | 7 1,050 | | | 14 Jun 2006 08:48 | |
|--|--|--|--------------|--|--|
| variable name | type | display format | label | variable label | |
| state region pop popurb medage marr divr | str13 byte double double float double double | %-13s %-8.0g %8.1f %8.1f %9.2f %8.1f %8.1f | cenreg | State Census region 1980 Population, '000 1980 Urban population, '000 Median age, years Marriages, '000 Divorces, '000 | |
| Sorted by: | | | | | |
| 2.1.3 Genera | te & Rep | lace | | | |
| . use census2d | .(1980 Ce | ensus data f | for NE and N | NC states) | |
| . generate urb | anized = | popurb/pop | | | |
| . summarize ur | banized | | | | |
| Variable | (|)bs M | Mean Std | . Dev. Min Max | |
| • | | | | 00842 .3377319 .8903645 | |
| . use census2d | .(1980 Ce | ensus data f | for NE and N | NC states) | |
| . generate urb | anized = | popurb/pop | | | |
| . replace urba | | L00∗urbaniz∈ | ed | | |
| . summarize ur | banized | | | | |
| | | | | . Dev. Min Max | |
| • | | | | 00843 33.77319 89.03645 | |
| 2.1.4 sort & | gsort | | | | |
| . use census2d | | | | NC states) | |
| . list region | | | | | |
| region | state | | pop | | |

| 1. | NE | Connecticut | 3107.6 |
|-----|----------------------|---------------|---------------------|
| | N Cntrl N Cntrl | | 11426.5 5490.2 |
| | | | |
| | N Cntrl | Iowa | 2913.8 |
| 5. | N Cntrl | Kansas | 2363.7 |
| 6. | 1 | Maine | 1124.7 |
| 7. | NE | Massachusetts | 5737.0 |
| 8. | N Cntrl | Michigan | 9262.1 |
| 9. | • | Minnesota | 4076.0 I |
| 10. | N Cntrl | | 4916.7 |
| 11. | N Cntrl | | 1569.8 |
| 11. | | | 1309.0 |
| 12. | NE | New Hampshire | 920.6 |
| 13. | NE | New Jersey | 7364.8 |
| 14. | NE | New York | 17558.1 |
| 15. | N Cntrl | N. Dakota | 652.7 |
| 16. | N Cntrl | | 10797.6 |
| | | | |
| 17. | NE | Pennsylvania | 11863.9 |
| 18. | NE | Rhode Island | 947.2 |
| | | | |
| 19. | N Cntrl | | 690.8 |
| 20 | | | |
| 20. | 1 | Vermont | 511.5 |
| 21. | I | | 4705.8 |
| | + | | |
| | | | |

. sort region -pop

. list region state pop, sepby(region)

| | + | | + |
|-----|---------|---------------|-------------|
| | region | | pop |
| 1. | I NE | Vermont | 511.5 |
| 2. | I NE | New Hampshire | 920.6 |
| 3. | NE | Rhode Island | 947.2 |
| 4. | NE | Maine | 1124.7 |
| 5. | NE | Connecticut | 3107.6 |
| 6. | NE | Massachusetts | - |
| 7. | NE | New Jersey | 7364.8 |
| 8. | NE | Pennsylvania | 11863.9 |
| 9. | NE | New York | 17558.1 |
| | | | |
| 10. | N Cntrl | N. Dakota | 652.7 |
| 11. | N Cntrl | S. Dakota | 690.8 |
| 12. | N Cntrl | Nebraska | 1569.8 |
| 13. | N Cntrl | Kansas | 2363.7 |
| 14. | N Cntrl | Iowa | 2913.8 |
| 15. | N Cntrl | Minnesota | 4076.0 |
| 16. | N Cntrl | Wisconsin | 4705.8 |

| 17. | 1 | N | Cntrl | Missouri | 4916.7 | - |
|-----|---|---|-------|----------|---------|-----|
| 18. | 1 | N | Cntrl | Indiana | 5490.2 | 1 |
| 19. | 1 | N | Cntrl | Michigan | 9262.1 | 1 |
| 20. | 1 | N | Cntrl | Ohio | 10797.6 | ١ |
| 21. | 1 | N | Cntrl | Illinois | 11426.5 | - |
| | + | | | | | . + |

.

. gsort region -pop

. list region state pop, sepby(region)

| | + | | + |
|-----|---------|---------------|---------|
| | region | state | pop |
| 1. | NE | New York | 17558.1 |
| 2. | NE | Pennsylvania | 11863.9 |
| 3. | NE | New Jersey | 7364.8 |
| 4. | NE | Massachusetts | 5737.0 |
| 5. | NE | Connecticut | 3107.6 |
| 6. | NE | Maine | 1124.7 |
| 7. | NE | Rhode Island | 947.2 |
| 8. | NE | New Hampshire | 920.6 |
| 9. | NE | Vermont | 511.5 |
| | | | |
| 10. | N Cntrl | Illinois | 11426.5 |
| 11. | N Cntrl | Ohio | 10797.6 |
| 12. | N Cntrl | Michigan | 9262.1 |
| 13. | N Cntrl | Indiana | 5490.2 |
| 14. | N Cntrl | Missouri | 4916.7 |
| 15. | N Cntrl | Wisconsin | 4705.8 |
| 16. | N Cntrl | Minnesota | 4076.0 |
| 17. | N Cntrl | Iowa | 2913.8 |
| 18. | N Cntrl | Kansas | 2363.7 |
| 19. | N Cntrl | Nebraska | 1569.8 |
| 20. | N Cntrl | S. Dakota | 690.8 |
| 21. | N Cntrl | N. Dakota | 652.7 |
| | + | | + |

ORGANIZING & HANDLING ECONOMIC DATA

- 3.1 CROSS-SECTIONAL DATA & IDENTIFIER VARIABLES
- 3.2 TIME SERIES DATA
- 3.3 POOLED CROSS-SECTIONAL TIME SERIES DATA
- 3.4 PANEL DATA
- 3.5 TOOLS FOR MANIPULATING PANEL DATA
- 3.6 COMBINING CROSS-SECTIONAL & TIME SERIES DATASETS
- 3.7 CREATING LONG-FORMAT DATASETS WITH APPEND
- 3.8 THE RESHAPE COMMAND
- 3.9 USING STATA FOR REPRODUCIBLE RESEARCH
- 3.9.1 Using do-files
- 3.9.2 Data Validation: Assert & Duplicates

LINEAR REGRESSION

4.1 LINEAR REGRESSION

- . use hprice2a.(Housing price data for Boston-area communities)
- . summarize price lprice lnox ldist rooms stratio, sep(0)

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|----------|-----|----------|-----------|----------|----------|
| price | 506 | 22511.51 | 9208.856 | 5000 | 50001 |
| lprice | 506 | 9.941057 | . 409255 | 8.517193 | 10.8198 |
| lnox | 506 | 1.693091 | .2014102 | 1.348073 | 2.164472 |
| ldist | 506 | 1.188233 | .539501 | .1222176 | 2.495682 |
| rooms | 506 | 6.284051 | .7025938 | 3.56 | 8.78 |
| stratio | 506 | 18.45929 | 2.16582 | 12.6 | 22 |

- . use hprice2a.(Housing price data for Boston-area communities)
- . regress lprice lnox ldist rooms stratio

| Source | SS | df | MS | Number of obs | = | 506 |
|----------|------------|-----|------------|---------------|---|--------|
| + | | | | F(4, 501) | = | 175.86 |
| Model | 49.3987735 | 4 | 12.3496934 | Prob > F | = | 0.0000 |
| Residual | 35.1834974 | 501 | .070226542 | R-squared | = | 0.5840 |
| +- | | | | Adj R-squared | = | 0.5807 |
| Total | 84.5822709 | 505 | .167489645 | Root MSE | = | .265 |
| | | | | | | |
| | | | | | | |

| lprice | | | | • • | [95% Conf. | - |
|---|---|--|----------------------------------|----------------------------------|---|---|
| lnox ldist rooms stratio | 95354 1343401 .2545271 0524512 | .1167418 .0431032 .0185303 .0058971 | -8.17 -3.12 13.74 -8.89 | 0.000 0.002 0.000 0.000 | -1.182904 2190255 .2181203 0640373 | 7241762 0496548 .2909338 0408651 |
| _cons | 11.08387 | .3181115 | 34.84 | 0.000 | 10.45887 | 11.70886 |
| | | | | | | |

4.2 INFORMATION CRITERIA

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . estat ic

| Model | 0bs | ll(null) | ll(model) | df | AIC | BIC |
|-------|-----|-----------|-----------|----|-----|---------|
| · | 506 | -265.4135 | -43.49514 | 5 | | 118.123 |

Note: N=Obs used in calculating BIC; see [R] BIC note.

4.3 THE COEFFICIENT ESTIMATES & BETA COEFFICIENTS

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . regress, beta

| Source | SS | df | MS | Number of obs | = | 506 175.86 |
|----------|------------|----------|-------------|----------------------------|---|------------------|
| Model | 49.3987735 | 4 | 12.3496934 | , , | = | 0.0000 |
| Residual | 35.1834974 | | .070226542 | R-squared Adj R-squared | = | 0.5840 0.5807 |
| Total | 84.5822709 | | . 167489645 | | = | . 265 |
| lprice | Coef. | | | P> t | | Beta |
| lnox | 95354 | .1167418 | -8.17 | 0.000 | - | .4692738 |
| ldist | 1343401 | .0431032 | -3.12 | 0.002 | - | .1770941 |
| rooms | .2545271 | .0185303 | 13.74 | 0.000 | | .4369626 |
| stratio | 0524512 | .0058971 | -8.89 | 0.000 | - | .2775771 |
| _cons | 11.08387 | .3181115 | 34.84 | 0.000 | | • |

4.4 RECOVERING ESTIMATION RESULTS

- . use hprice2a.(Housing price data for Boston-area communities) $\,$
- . quietly regress lprice lnox ldist rooms stratio
- . ereturn list

scalars:

e(N) = 506 $e(df_m) = 4$ $e(df_r) = 501$ e(F) = 175.8550695227946 e(r2) = .5840322442976398 e(rmse) = .2650029089298266e(mss) = 49.39877352102587

e(rss) = 35.18349741237627 $e(r2_a) = .5807111444517128$ e(ll) = -43.4951392092929 $e(ll_0) = -265.4134648194153$ e(rank) = 5e(cmdline) : "regress lprice lnox ldist rooms stratio" e(title) : "Linear regression" e(marginsok) : "XB default" e(vce) : "ols" e(depvar) : "lprice" e(cmd) : "regress" e(properties) : "b V" e(predict) : "regres_p" e(model) : "ols" e(estat_cmd) : "regress_estat" $e(b) : 1 \times 5$ $e(V) : 5 \times 5$ e(sample) . use hprice2a.(Housing price data for Boston-area communities) . quietly regress lprice lnox ldist rooms stratio . matrix list e(b) lnox ldist rooms stratio _cons y1 -.95354002 -.13434015 .25452706 -.05245119 11.083865 . use hprice2a.(Housing price data for Boston-area communities) . quietly regress lprice lnox ldist rooms stratio

Covariance matrix of coefficients of regress model

macros:

matrices:

functions:

e(b)[1,5]

. estat vce

```
e(V) | lnox ldist rooms stratio _cons
------
     lnox | .01362865
    ldist | .00426247 .00185789
    rooms | .00035279 .00003043 .00034337
   stratio | 9.740e-07 .00002182 .00003374 .00003478
    _cons | -.03037429 -.01001835 -.00341397 -.00088151 .10119496
```

4.5 DETECTING COLLINEARLITY IN REGRESSION

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . estat vif

| Variable | VIF | 1/VIF |
|----------|------|----------|
| lnox | 3.98 | 0.251533 |
| ldist | 3.89 | 0.257162 |
| rooms | 1.22 | 0.820417 |
| stratio | 1.17 | 0.852488 |
| Mean VIF | 2.56 | |

4.6 PRESENTING REGRESSION ESTIMATES

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate rooms2 = rooms^2
- . quietly regress lprice rooms
- . estimates store model1
- . quietly regress lprice rooms rooms2 ldist
- . estimates store model2
- . quietly regress lprice ldist stratio lnox
- . estimates store model3
- . quietly regress lprice lnox ldist rooms stratio
- . estimates store model4
- . estimates table model1 model2 model3 model4, stat(r2_a, rmse) b(%7.3g) se(%6.
- > 3q) p(%4.3f)

| | | | |
|----------|-------|-------|-----------|
| Variable | | | |
| · | .369 | | . 255 |
| 1 | .0201 | . 183 | .0185 |

| | 0.000 | 0.000 | | 0.000 |
|---------|-------|-------|-------|------------|
| rooms2 | | .0889 | | |
| | | .014 | | |
| | | 0.000 | | |
| ldist | | . 237 | 157 | 134 |
| | | .0255 | .0505 | .0431 |
| | | 0.000 | 0.002 | 0.002 |
| stratio | | | 0775 | 0525 |
| | | | .0066 | .0059 |
| | | | 0.000 | 0.000 |
| lnox | | | -1.22 | 954 |
| | | | . 135 | .117 |
| | | | 0.000 | 0.000 |
| _cons | 7.62 | 11.3 | 13.6 | 11.1 |
| | . 127 | .584 | .304 | .318 |
| | 0.000 | 0.000 | 0.000 | 0.000 |
| r2_a | .399 | .5 | .424 | .581 |
| rmse | .317 | . 289 | .311 | . 265 |
| | | | lege | nd: b/se/p |

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate rooms2 = rooms^2
- . quietly regress lprice rooms
- . estimates store model1
- . quietly regress lprice rooms rooms2 ldist
- . estimates store model2
- . quietly regress lprice ldist stratio lnox
- . estimates store model3
- . quietly regress lprice lnox ldist rooms stratio
- . estimates store model4
- . estimates table model1 model2 model3 model4, stat($r2_a$ rmse ll) b(\$7.3g) star > title("Models of Median Housing Price")

Models of Median Housing Price

| Variable | model1 | model2 | model3 | model4 |
|----------|---------|--------|--------|---------|
| + | | | | |
| rooms | .369*** | 821*** | | .255*** |

| rooms2 ldist stratio lnox _cons | 7.62*** | .0889*** .237*** 11.3*** | 157** 0775*** -1.22*** 13.6*** | 134** 0525*** 954*** 11. 1*** |
|---|---------|--------------------------------|---|--|
| r2_a | .399 | .5 | .424 | .581 |
| rmse | .317 | .289 | .311 | .265 |
| ll | -136 | -88.6 | -124 | -43.5 |

legend: * p<0.05; ** p<0.01; *** p<0.001

4.7 HYPOTHESIS TESTS, LINEAR RESTRICTIONS, & CONSTRAINED LEAST SQUARES

4.7.1 Wald Tests with Test

- . use hprice2a.(Housing price data for Boston-area communities)
- . regress lprice lnox ldist rooms stratio

| Source | SS | df | MS | Number of obs | = | 506 175.86 |
|----------|----------------|-----|------------|----------------------------|---|------------------|
| Model | 49.3987735 | 4 | 12.3496934 | F(4, 501) Prob > F | = | 0.0000 |
| Residual | 35.1834974 | | .070226542 | R-squared Adj R-squared | = | 0.5840 0.5807 |
| Total | 84.5822709 | 505 | .167489645 | Root MSE | = | . 265 |
| | | | | | | |

| lprice | | Std. Err. | | | - | Interval] |
|---------|----------|-----------|-------|-------|-----------|-----------|
| lnox | 95354 | .1167418 | -8.17 | 0.000 | -1.182904 | 7241762 |
| ldist | 1343401 | .0431032 | -3.12 | 0.002 | 2190255 | 0496548 |
| rooms | .2545271 | .0185303 | 13.74 | 0.000 | .2181203 | . 2909338 |
| stratio | 0524512 | .0058971 | -8.89 | 0.000 | 0640373 | 0408651 |
| _cons | 11.08387 | .3181115 | 34.84 | 0.000 | 10.45887 | 11.70886 |
| | | | | | | |

. test rooms

(1) rooms = 0

$$F(1, 501) = 188.67$$

 $Prob > F = 0.0000$

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . test rooms = 0.33
- (1) rooms = .33

F(1, 501) = 16.59Prob > F = 0.0001

4.7.2 Wald Tests involving Linear Combinations of Parameters

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . lincom rooms + ldist + stratio
- (1) ldist + rooms + stratio = 0

lprice | Coef. Std. Err. t P>|t| [95% Conf. Interval]
(1) | .0677357 .0490714 1.38 0.168 -.0286753 .1641468

- . test ldist = stratio
- (1) ldist stratio = 0

$$F(1, 501) = 3.63$$

 $Prob > F = 0.0574$

- . test lnox = 10∗stratio
- (1) lnox 10*stratio = 0

$$F(1, 501) = 10.77$$

 $Prob > F = 0.0011$

- . use hprice2a.(Housing price data for Boston-area communities)
- . constraint def 1 rooms + ldist + stratio = 0
- . cnsreg lprice lnox ldist rooms stratio, constraint(1)

Constrained linear regression F(3, 502) = 233.42 Prob > F = 0.0000 Root MSE = 0.2652

(1) ldist + rooms + stratio = 0

| | | | | | [95% Conf. | - |
|---------|----------|----------|--------|-------|------------|----------|
| lnox | | .0691935 | | | -1.219337 | |
| ldist | 1880712 | .0185284 | -10.15 | 0.000 | 2244739 | 1516684 |
| rooms | .2430633 | .01658 | 14.66 | 0.000 | .2104886 | .2756381 |
| stratio | 0549922 | .0056075 | -9.81 | 0.000 | 0660092 | 0439752 |

_cons | 11.48651 .1270377 90.42 0.000 11.23691 11.7361

4.7.3 Joint Hypothesis Tests

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . test lnox ldist
- (1) lnox = 0
- (2) ldist = 0

$$F(2, 501) = 58.95$$

 $Prob > F = 0.0000$

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . test (lnox = 10*stratio) (ldist = stratio)
- (1) lnox 10*stratio = 0
- (2) ldist stratio = 0

$$F(2, 501) = 5.94$$

 $Prob > F = 0.0028$

- 4.7.4 Testing Nonlinear Restrictions & Forming Nonlinear Combinations
- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . testnl $_b[lnox] * _b[stratio] = 0.06$
 - (1) $_b[lnox] * _b[stratio] = 0.06$

$$chi2(1) = 1.44$$

Prob > chi2 = 0.2300

- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . testnl (b[lnox] * b[stratio] = 0.06) (b[rooms] * b[ldist] = 3 * b[lnox])

```
(1) _b[lnox] * _b[stratio] = 0.06
  (2) _b[rooms] * _b[ldist] = 3 * _b[lnox]
              chi2(2) = 184.94
          Prob > chi2 =
                            0.0000
4.7.5 Testing Competing (Non-Nested) Models
. use hprice2a.(Housing price data for Boston-area communities)
. nnest lprice lnox ldist rooms stratio (crime proptax ldist rooms stratio)
command nnest is unrecognized
r(199);
end of do-file
r(199);
4.8 COMPUTING RESIDUALS & PREDICTED VALUES
. use hprice2a.(Housing price data for Boston-area communities)
. quietly regress lprice lnox ldist rooms stratio
. predict double lpricehat, xb
. label var lpricehat "Predicted log price"
. twoway (scatter lpricehat lprice, msize(small) mcolor(black) msize(tiny)) ||
> (line lprice lprice if lprice <., clwidth(thin)), ytitle("Predicted log media
> n housing price") xtitle("Actual log median housing price") aspectratio(1) le
> gend(off)
4.8.1 Computing Interval Predictions
. use hprice2a.(Housing price data for Boston-area communities)
. quietly regress lprice lnox if _n <= 100
. predict double xb if e(sample)
(option xb assumed; fitted values)
(406 missing values generated)
. predict double stdpred if e(sample), stdp
(406 missing values generated)
. scalar tval = invttail(e(df_r), 0.975)
```

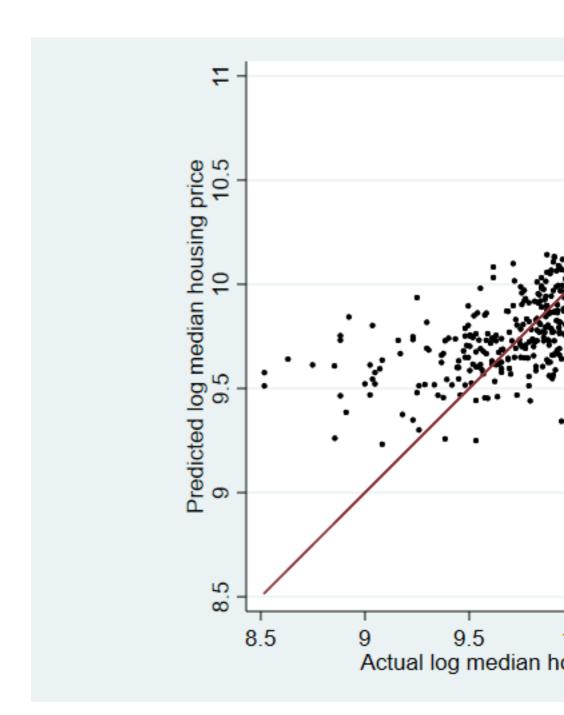


Figure 1: Actual versus predicted values from regression model

- . generate double uplim = xb + tval * stdpred
 (406 missing values generated)
- . generate double lowlim = xb tval * stdpred (406 missing values generated)
- . summarize lnox if e(sample), meanonly
- . generate lnoxbar = r(mean)
- . label var xb "Pred"
- . label var uplim "95% prediction interval"
- . label var lowlim "95% prediction interval"
- . twoway (scatter lprice lnox if e(sample), sort ms(0h) xline('lnoxbar')) (conn
- > ected xb lnox if e(sample), sort msize(small)) (rline uplim lowlim lnox if e(
- > sample), sort), ytitle(Actual and predicted log price) legend(cols(3))

4.9 COMPUTING MARGINAL EFFECTS

- . use hprice2a.(Housing price data for Boston-area communities)
- . regress price nox dist rooms stratio proptax

| Source | SS | dŤ | MS | Number of obs | = | 506 |
|----------|------------|-----|------------|---------------|---|--------|
| +- | | | | F(5, 500) | = | 165.85 |
| Model | 2.6717e+10 | 5 | 5.3434e+09 | Prob > F | = | 0.0000 |
| Residual | 1.6109e+10 | 500 | 32217368.7 | R-squared | = | 0.6239 |
| +- | | | | Adj R-squared | = | 0.6201 |
| Total | 4.2826e+10 | 505 | 84803032 | Root MSE | = | 5676 |

| price | Coef. | Std. Err. | t | P> t | - | Interval] |
|---------|-----------|-----------|-------|-------|-----------|-----------|
| nox | -2570.162 | 407.371 | -6.31 | 0.000 | -3370.532 | -1769.793 |
| dist | -955.7175 | 190.7124 | -5.01 | 0.000 | -1330.414 | -581.021 |
| rooms | 6828.264 | 399.7034 | 17.08 | 0.000 | 6042.959 | 7613.569 |
| stratio | -1127.534 | 140.7653 | -8.01 | 0.000 | -1404.099 | -850.9699 |
| proptax | -52.24272 | 22.53714 | -2.32 | 0.021 | -96.52188 | -7.963555 |
| _cons | 20440.08 | 5290.616 | 3.86 | 0.000 | 10045.5 | 30834.66 |

. mfx, eyex

Elasticities after regress
 y = Fitted values (predict)

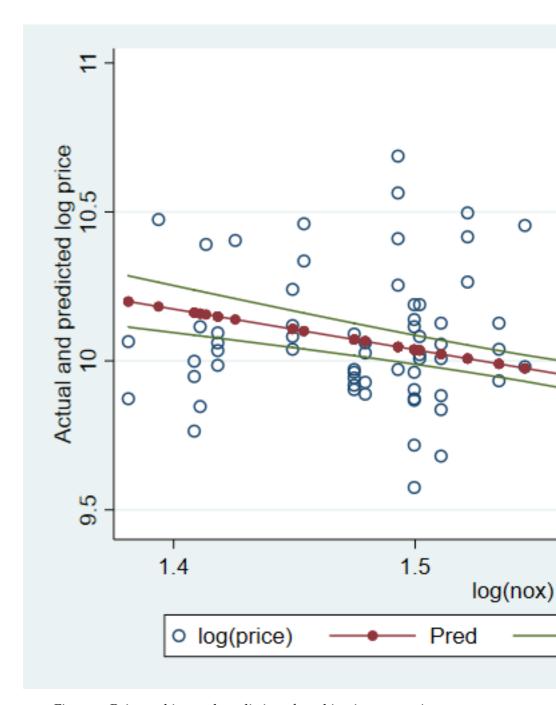


Figure 2: Point and interval predictions from bivariate regression

= 22511.51

| variable | ey/ex | Std. Err. | z | | - | C.I.] | Х |
|--|--|-------------------------------------|----------------------------------|----------------------------------|------------------|--|---|
| nox dist rooms stratio proptax | 6336244 1611472 1.906099 9245706 0947401 | .10068 .03221 .1136 .11589 | -6.29 -5.00 16.78 -7.98 | 0.000 0.000 0.000 0.000 | 830954 224273 | 436295 098022 2 .12876 697429 | 5.54978 3.79575 6.28405 18.4593 40.8237 |
| hiohrax | 094/401 | .04088 | -2.32 | 0.020 | 1/48/1 | 014009 | 40.8237 |

5.1 INTRODUCTION

5.2 SPECIFICATION ERROR

- 5.2.1 Omitting relevant Variables from the Model
- 5.2.2 Graphically Analyzing Regression data
- . use hprice2a.(Housing price data for Boston-area communities)
- . graph matrix lprice lnox ldist rooms stratio, ms(Oh) msize(tiny)

5.2.3 Added-Variable Plots

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate rooms2 = rooms^2
- . regress lprice lnox ldist rooms rooms2 stratio lproptax

| Source | SS | df | MS | Number of obs | = | 506 |
|----------|------------|-----|-------------|---------------|---|--------|
| | | | | F(6, 499) | = | 138.41 |
| Model | 52.8357813 | 6 | 8.80596356 | Prob > F | = | 0.0000 |
| Residual | 31.7464896 | 499 | .06362022 | R-squared | = | 0.6247 |
| | | | | Adj R-squared | = | 0.6202 |
| Total | 84.5822709 | 505 | . 167489645 | Root MSE | = | .25223 |
| | | | | | | |

lprice | Coef. Std. Err. t P>|t| [95% Conf. Interval]

lnox | -.6615694 .1201606 -5.51 0.000 -.8976524 -.4254864
ldist | -.095087 .0421435 -2.26 0.024 -.1778875 -.0122864
rooms | -.5625662 .1610315 -3.49 0.001 -.8789496 -.2461829
rooms2 | .0634347 .0124621 5.09 0.000 .0389501 .0879193
stratio | -.0362928 .0060699 -5.98 0.000 -.0482185 -.0243671
lproptax | -.2211125 .0410202 -5.39 0.000 -.301706 -.1405189
_cons | 14.15454 .5693846 24.86 0.000 13.03585 15.27323

. avplots, ms(Oh) msize(small) col(2)

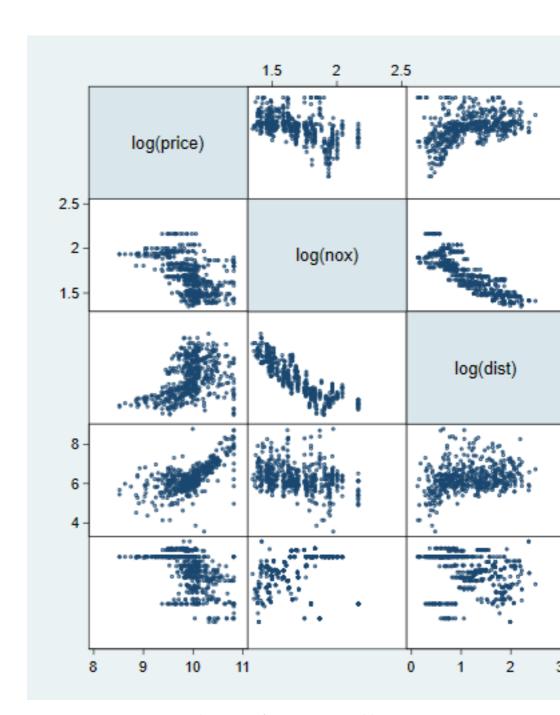


Figure 3: graph matrix of regression variables

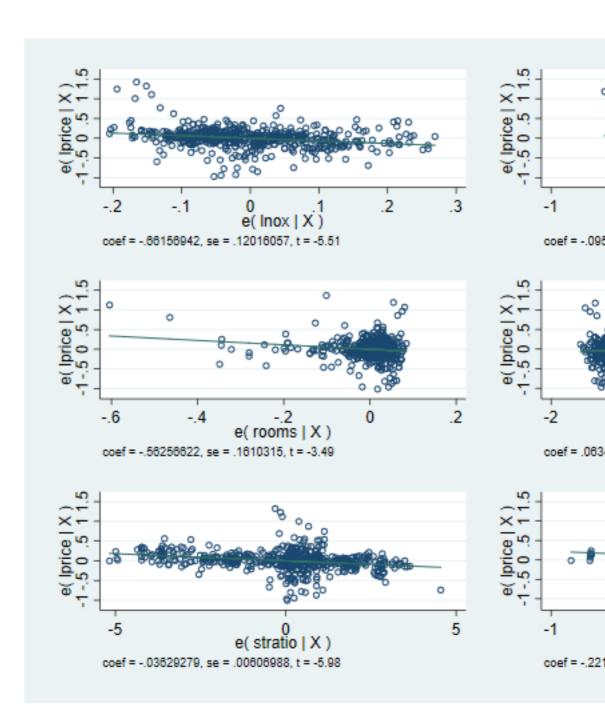


Figure 4: Added-variable plots

- 5.2.4 Including Irrelevant Variables in the Model
- 5.2.5 The Asymmetry of Specification Error
- 5.2.6 Misspecification of the Functional Form
- 5.2.7 Ramsey's RESET
- . use hprice2a.(Housing price data for Boston-area communities)
- . quietly regress lprice lnox ldist rooms stratio
- . estat ovtest

Ramsey RESET test using powers of the fitted values of lprice Ho: model has no omitted variables

> F(3, 498) = 9.69Prob > F = 0.0000

. estat ovtest, rhs

Ramsey RESET test using powers of the independent variables

Ho: model has no omitted variables

F(12, 489) = 11.79Prob > F = 0.0000

- . use hprice2a.(Housing price data for Boston-area communities)
- . $generate rooms2 = rooms^2$
- . regress lprice lnox ldist rooms rooms2 stratio lproptax

| Source | SS | df | MS | Number of obs | = | 506 |
|----------|------------|-----|------------|---------------|---|---------|
| +- | | | | F(6, 499) | = | 138.41 |
| Model | 52.8357813 | 6 | 8.80596356 | Prob > F | = | 0.0000 |
| Residual | 31.7464896 | 499 | .06362022 | R-squared | = | 0.6247 |
| +- | | | | Adj R-squared | = | 0.6202 |
| Total | 84.5822709 | 505 | .167489645 | Root MSE | = | . 25223 |

| lprice | Coef. | Std. Err. | t | P> t | [95% Conf. | Interval] |
|----------|----------|-----------|-------|-------|------------|-----------|
| lnox | 6615694 | .1201606 | -5.51 | 0.000 | 8976524 | 4254864 |
| ldist | 095087 | .0421435 | -2.26 | 0.024 | 1778875 | 0122864 |
| rooms | 5625662 | .1610315 | -3.49 | 0.001 | 8789496 | 2461829 |
| rooms2 | .0634347 | .0124621 | 5.09 | 0.000 | .0389501 | .0879193 |
| stratio | 0362928 | .0060699 | -5.98 | 0.000 | 0482185 | 0243671 |
| lproptax | 2211125 | .0410202 | -5.39 | 0.000 | 301706 | 1405189 |
| _cons | 14.15454 | .5693846 | 24.86 | 0.000 | 13.03585 | 15.27323 |
| | | | | | | |

Ramsey RESET test using powers of the fitted values of lprice

Ho: model has no omitted variables

F(3, 496) = 1.64Prob > F = 0.1798

5.2.8 Specification Plots

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate rooms2 = rooms^2
- . quietly regress lprice lnox ldist rooms rooms2 stratio lproptax
- . rvpplot ldist, ms(0h) yline(0)

5.2.9 Specification & Interaction Terms

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate taxschl = lproptax * stratio
- . regress lprice lnox ldist lproptax stratio taxschl

| Source | SS | df | MS | Number of obs | = | 506 |
|----------|------------|-----|-------------|---------------|---|--------|
| +- | | | | F(5, 500) | = | 84.47 |
| Model | 38.7301562 | 5 | 7.74603123 | Prob > F | = | 0.0000 |
| Residual | 45.8521148 | 500 | .09170423 | R-squared | = | 0.4579 |
| +- | | | | Adj R-squared | = | 0.4525 |
| Total | 84.5822709 | 505 | . 167489645 | Root MSE | = | .30283 |

| lprice | Coef. | Std. Err. | t | P> t | = | . Interval] |
|---|---|---|--|---|--|--|
| lnox ldist lproptax stratio taxschl | 9041103 1430541 -1.48103 4388722 .0641648 | .1441253 .0501831 .5163117 .1538321 .026406 | -6.27 -2.85 -2.87 -2.85 2.43 | 0.000 0.005 0.004 0.005 0.015 | -1.187276 2416499 -2.495438 7411093 .0122843 | 6209444 0444583 4666219 1366351 .1160452 |
| _cons | 21.47905 | 2.952307 | 7.28 | 0.000 | 15.6786 | 27.27951 |

5.2.10 Outlier Statistics & Measures of Leverage

- . use hprice2a.(Housing price data for Boston-area communities)
- . $generate rooms2 = rooms^2$

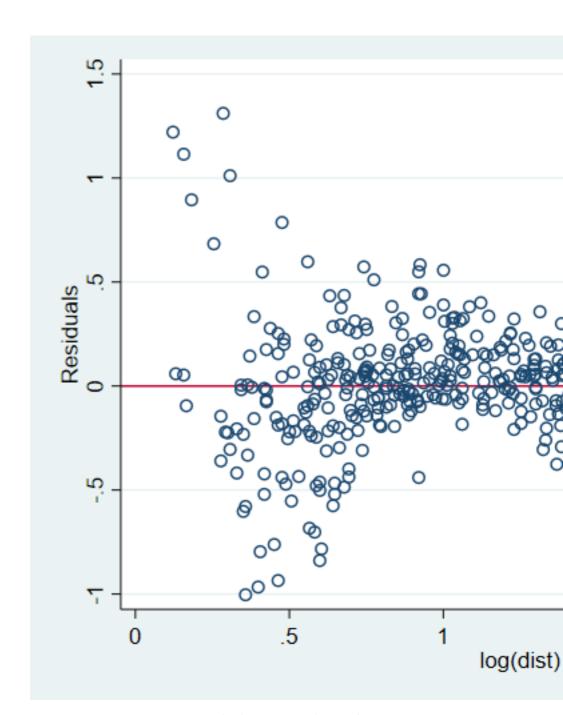


Figure 5: Residual-versus-predictor plot

- . quietly regress lprice lnox ldist rooms rooms2 stratio lproptax
- . $generate town = _n$
- . predict double lev if e(sample), leverage
- . predict double eps if e(sample), res
- . $generate eps2 = eps^2$
- . summarize price lprice

| Variable | 0bs | Mean | Std. Dev | . Min | Max |
|----------|-----|----------|----------|----------|---------|
| + | | | | | |
| price | 506 | 22511.51 | 9208.856 | 5000 | 50001 |
| lprice | 506 | 9.941057 | .409255 | 8.517193 | 10.8198 |

- . gsort -lev
- . list town price lprice lev eps2 in 1/5

| | т. | | | | | |
|----------------------------|------|---------------------------------|---|---|--|--|
| | | town | price | lprice | lev | eps2 |
| 1. 2. 3. 4. 5. | | 366 368 365 258 226 | 27499 23100 21900 50001 50001 | 10.2219 10.04759 9.994242 10.8198 10.8198 | .17039262 .11272637 .10947853 .08036068 .0799096 | .6181372 .3002205 .3308896 .0604706 .0338277 |
| | т. | | | | | |

- . gsort -eps2
- . list town price lprice lev eps2 in 1/5

| | + - | | | | | + |
|----------------------------|--------|---------------------------------|--|--|--|---|
| | | town | price | lprice | lev | eps2 |
| 1. 2. 3. 4. 5. | i I | 369 373 372 370 406 | 50001 50001 50001 50001 5000 | 10.8198 10.8198 10.8198 10.8198 8.517193 | .02250047 .01609848 .02056901 .0172083 .00854955 | 1.71812 1.489409 1.242105 1.022456 1.006366 |
| | + | | | | | + |

- . predict double dfits if e(sample), dfits
- . gsort -dfits
- . quietly generate cutoff = abs(dfits) > 2 * sqrt((e(df_m) +1)/e(N)) & e(sampl > e)
- . list town price lprice dfits if cutoff

```
1. | 366 27499 10.2219 1.5679033 |
 2. | 368 23100 10.04759 .82559867 |
 3. | 369 50001 10.8198 .8196735 |
 4. | 372 50001 10.8198 .65967704 |
 5. | 373 50001 10.8198 .63873964 |
   |-----|
 6. | 371 50001 10.8198 .55639311 |
 7. | 370 50001 10.8198 .54354054 |
 8. | 361 24999 10.12659 .32184327 |
9. | 359 22700 10.03012 .31516743 |
10. | 408 27901 10.23642 .31281326 |
   |-----|
11. | 367
         21900 9.994242 .31060611 |
12. | 360 22600 10.02571 .28892457 |
13. | 363 20800 9.942708 .27393758 |
14. | 358 21700 9.985067
                        .24312885 |
490. | 386 7200 8.881836 -.23838749 |
   |-----|
491. | 388 7400 8.909235 -.25909393 |
492. | 491 8100 8.999619 -.26584795 |
493. | 400 6300 8.748305 -.28782824 |
494. | 416 7200 8.881836 -.29288953 |
495. | 402 7200 8.881836 -.29595696 |
   |-----|
496. | 381 10400 9.249561 -.29668364 |
497. | 258 50001
               10.8198 -.30053391 |
498. | 385 8800 9.082507 -.302916 |
499. | 420
         8400 9.035987
                       -.30843965 |
500. | 490 7000 8.853665 -.3142718 |
   |-----|
501. | 401 5600 8.630522 -.33273658 |
502. | 417 7500 8.922658 -.34950136 |
503. | 399 5000 8.517193 -.36618139 |
504. | 406 5000 8.517193 -.37661853 |
|-----|
506. | 365 21900 9.994242
                      -.85150064 |
   +-----+
```

. dfbeta lnox

_dfbeta_1: dfbeta(lnox)

- . quietly generate $dfcut = abs(_dfbeta_1) > 2 * sqrt(e(N)) & e(sample)$
- . sort _dfbeta_1
- . summarize lnox

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|----------|-----|----------|-----------|----------|----------|
| | | | | | |
| lnox | 506 | 1.693091 | .2014102 | 1.348073 | 2.164472 |

. list town price lprice lnox _dfbeta_1

+----+

price lprice lnox | town _dfbeta_1 | |-----| 50001 10.8198 1.842136 1. | 369 -.4316933 | 2. | 372 50001 10.8198 1.842136 -.4257791 l 3. | 373 50001 10.8198 1.899118 -.3631822 | 4. | 371 50001 10.8198 1.842136 -.2938702 | 10.8198 -.2841335 | 5. | 370 50001 1.842136 |-----| 6. | 365 9.994242 21900 1.971299 -.2107066 | 7. | 408 27901 10.23642 1.885553 -.1728729 | 8. | 368 23100 10.04759 1.842136 -.1309522 | 9. | 11 15000 9.615806 1.656321 -.1172723 | 10. | 410 27499 10.2219 1.786747 -.1117743 | |-----| 11. | 413 17900 9.792556 1.786747 -.0959273 | 12. | 437 9600 9.169518 2.00148 -.0955826 | 13. | 146 13800 9.532424 2.164472 -.0914387 | -.0856147 | 14. | 438 8700 9.071078 2.00148 15. | 420 8400 9.035987 1.971299 -.085223 | |-----| 16. | 145 11800 9.375854 2.164472 -.0816827 | 17. | 439 2.00148 8400 9.035987 -.070508 | -.0668001 | 18. | 182 36199 10.49679 1.585145 19. | 423 9.942708 20800 1.814825 -.064928 | 20. | 157 9.480368 -.0622912 | 13100 2.164472 |------21. | 258 50001 10.8198 1.867176 -.0570776 | 22. | 158 41299 10.62859 1.800058 -.0563827 | 23. | 409 17200 9.752665 1.786747 -.0549674 | 24. | 343 16500 9.711116 1.644805 -.0542086 | 25. | 414 16300 9.69892 1.786747 -.0522002 | |-----26. | 143 13400 9.50301 2.164472 -.0499795 | 27. | 446 11800 9.375854 2.00148 -.0498875 | 28. | 31 12700 9.449357 1.682688 -.0480591 | 29. | 156 9.655026 -.0474292 | 15600 2.164472 30. | 493 20100 9.908475 1.806648 -.0472557 | |-----| 9.971147 1.814825 31. | 480 21400 -.0469688 | 32. | 481 23001 10.04329 1.671473 -.0465269 33. | 451 13400 9.50301 1.964311 -.0461007 | 34. I 9 16500 9.711116 1.656321 -.0449441 | 35. | 13200 9.487972 1.682688 33 -.0427313 | |-----| 36. | 28 14800 9.602383 1.682688 -.0383155 | 37. 32 14500 9.581903 1.682688 -.0374429 | 38. | 188 32000 10.37349 1.474763 -.0370926 | 39. | 435 11700 9.367344 1.964311 -.0370922 | 40. | 454 17794 9.786616 1.964311 -.0369095 | |-----| 41. | 9.50301 -.0368249 | 436 13400 2.00148 42. | 35 13500 9.510445 1.682688 -.036522 | 43. | 161 27000 10.20359 1.800058 -.035759 | 44. | 441 9.25913 2.00148 -.0354014 | 10500 45. | 26 13900 9.539644 1.682688 -.0336016 | |------46. | 445 10800 9.287301 2.00148 -.0330708 |

47. | 10 18900 9.846917 1.656321 -.0328998 | 48. | 101 27499 10.2219 1.648659 -.0317264 | 49. | 189 29801 10.3023 1.474763 -.0315825 | 50. I 12 18900 9.846917 1.656321 -.0311876 | -----| |----1.671473 -.0303079 | 51. | 472 19600 9.883285 34899 10.46021 1.393766 -.0302973 | 52. | 200 53. | 34 13100 9.480368 1.682688 -.0300799 | 54. | 470 20100 9.908475 1.757858 -.0299316 | 55. | 482 23699 10.07319 1.671473 -.0295599 | |-----| 50001 10.8198 1.800058 -.0294572 | 56. | 162 57. | 263 48801 10.79551 1.867176 -.0288593 | 58. | 23 15200 9.62905 1.682688 -.0284555 | 59. | 468 19100 9.857444 1.764731 -.0278265 | 60. | 448 12600 9.441452 2.00148 -.027274 | |------61. | 450 13000 9.472705 1.964311 -.0268602 | 62. | 24 14500 9.581903 1.682688 -.0267329 | 63. | 506 11900 9.384294 1.745715 -.0259339 | 64. | 75 24101 10.09001 1.474763 -.0255723 | 65. | 159 24299 10.09819 1.800058 -.0253549 | |-----| 37001 10.5187 66. | 191 1.474763 -.0246281 | 19500 9.87817 1.648659 -.0242416 | 67. | 106 68. | 292 37298 10.5267 1.413423 -.0242084 | 9.87817 69. | 107 19500 1.648659 -.0240593 | 70. | 484 21800 9.989665 1.671473 -.0238902 | |-----| 9.517825 71. | 21 13600 1.682688 -.0234366 | 72. | 230 1.617406 -.0231037 | 31499 10.35771 73. | 102 26500 10.1849 1.648659 -.0219822 | 74. | 62 16000 9.680344 1.510722 -.0219783 | 75. I 489 15200 9.62905 1.806648 -.0218215 | |-----| 76. | 25 15600 9.655026 1.682688 -.0215176 | 77. | 127 15700 9.661416 1.83098 -.0211289 | 78. | 430 9500 9.159047 1.915451 -.0207799 | 79. | 455 9.609117 1.964311 -.0206806 | 14900 80. | 108 20400 9.92329 1.648659 -.020656 | |-----| 81. | 173 23100 10.04759 1.629241 -.020537 | 82. | 215 23699 10.07319 1.587192 -.0205038 | 83. | 275 32400 10.38591 1.497388 -.0203821 | 84. | 184 10.389 32500 1.585145 -.019082 | 85. | 185 26399 10.18108 1.585145 -.0189528 | -----| |----86. | 190 34899 10.46021 1.474763 -.0188717 | 87. | 456 14100 9.55393 1.964311 -.0187349 | 88. | 483 24999 10.12659 1.671473 -.0186684 | 89. | 473 23200 10.05191 1.757858 -.0183966 | 90. | 224 30101 10.31231 1.623341 -.0176759 | |-----| 91. | 233 41702 10.6383 1.623341 -.0173393 | 92. | 180 37201 10.52409 1.585145 -.0173209 | 93. | 417 7500 8.922658 1.915451 -.0171022 |

94. | 375

13800

9.532424

1.899118 -.0166074 |

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| 95. | 316 | 16200 | 9.692766 | 1.693779 | 0165045 |
|------|---------|----------------|---------------------|----------------------|----------------------|
| 96. | 279 | 29100 | 10.27849 | 1.497388 | 0162037 |
| | 1119 | 20400 | 9.92329 | 1.699279 | 0161968 |
| | 176 | 29401 | 10.28878 | 1.629241 | 0161278 |
| | | | | | |
| | | 18400 | 9.820106 | 1.682688 | 0157718 |
| 100. | 374 | 13800 | 9.532424 | 1.899118 | 0145677 |
| 101. | 183 | 37900 | 10.54271 | 1.585145 | 0144407 |
| 102. | 40 | 30801 | 10.3353 | 1.453953 | 014429 |
| 103. | 495 | 24499 | 10.10639 | 1.766442 | 0138828 |
| 104. | 264 | 30999 | 10.34171 | 1.867176 | 013838 |
| 105. | 105 | 20100 | 9.908475 | 1.648659 | 013284 |
| | | | | | |
| 106. | 457 | 12700 | 9.449357 | 1.964311 | 0129107 |
| 107. | 341 | 18700 | 9.836279 | 1.638997 | 0129061 |
| | 193 | 36399 | 10.5023 | 1.474763 | 0128875 |
| 109. | 96 | 28399 | 10.25411 | 1.492904 | 0127983 |
| 110. | 201 | 32899 | 10.4012 | 1.393766 | 0126864 |
| 111 | | 20501 | 10 22551 | 1 474762 | 0125961 |
| | 192 | 30501 43998 | 10.32551 | 1.474763 | 0125861 |
| 112. | 257 | 43998 32000 | 10.6919 10.37349 | 1.371181 1.497388 | 0124312 0120773 |
| | 276 | | | | |
| 114. | 186 | 29599 | 10.2955 | 1.585145 | 0116235 |
| 115. | 338 | 18500 | 9.825526 | 1.638997 | 0116081 |
| 116. | 494 | 21800 | 9.989665 | 1.766442 | 0114905 |
| | 166 | 24999 | 10.12659 | 1.800058 | 0114317 |
| | 99 | 43800 | 10.68739 | 1.492904 | 0108364 |
| | 41 | 34899 | 10.46021 | 1.453953 | 0108243 |
| 120. | 471 | 19900 | 9.898475 | 1.757858 | 0104401 |
| | | | | | · |
| 121. | 278 | 33100 | 10.40729 | 1.497388 | 0101313 |
| 122. | 94 | 24999 | 10.12659 | 1.534714 | 0101088 |
| 123. | 111 | 21700 | 9.985067 | 1.699279 | 01002 |
| 124. | 317 | 17800 | 9.786954 | 1.693779 | 0099336 |
| 125. | 434 | 14300 | 9.568015 | 1.964311 | 0098488 |
| | | | | | |
| | | 19000 | | | 0097685 |
| 127. | 117 | 21200 | 9.961757 | 1.699279 | 0093754 |
| 128. | I 346 | 17500 | 9.769957 | 1.48614 | 0091785 |
| | 469 | 1/052 | 9.744022 | 1.757858 | 0091705 |
| | 452 | 15200 | | | 0091184 |
| | • | | | | 000026 |
| | - | | | | 0090826 |
| | - | | | | 008913 0087004 |
| | 179 | | | | |
| | | | | | 0086304 |
| 135. | 68 | | | | 0085998 |
| 136. | 347 | | | | 0085918 |
| | 53 | 22620 | 10.02659 | 1.479329 | 0084467 l |
| | 1 429 | 11000 | 9.305651 | 1.915451 | 0083504 |
| | 150 | 15400 | 9.642123 | 2.164472 | 0080456 |
| | 349 | 24499 | | 1.470176 | |
| | | | | | ·i |
| 141. | 213 | 22400 | 10.01682 | 1.587192 | 0074679 |

| 142. | 163 | 50001 | 10.8198 | 1.800058 | 0072357 |
|------|-------|-------|----------|----------|----------------------|
| 143. | 203 | 42302 | 10.65259 | 1.423108 | 0072147 |
| 144. | 214 | 28099 | 10.24349 | 1.587192 | 0071618 |
| 145. | 488 | 20600 | 9.933046 | 1.763017 | 0069738 |
| | | | | | |
| 146. | 169 | 23799 | 10.0774 | 1.800058 | 0067039 |
| 147. | 168 | 23799 | 10.0774 | 1.800058 | 0066772 |
| 148. | 426 | 8300 | 9.024011 | 1.915451 | 0066314 |
| 149. | I 393 | 9700 | 9.179881 | 1.94591 | 0066296 |
| 150. | 81 | 28001 | 10.24 | 1.449269 | 0064669 I |
| 1301 | | | | | |
| 151. | 140 | 17800 | 9.786954 | 1.83098 | 0063847 |
| 152. | 83 | 24800 | 10.1186 | 1.449269 | 0063639 |
| 153. | 52 | 20500 | 9.92818 | 1.479329 | 0061827 |
| 154. | 306 | 28399 | 10.25411 | 1.551809 | 0060002 |
| 155. | 104 | 19300 | 9.867861 | 1.648659 | 0059841 |
| | | | | | |
| 156. | 51 | 19700 | 9.888374 | 1.479329 | 0058348 |
| 157. | 236 | 24000 | 10.08581 | 1.623341 | 0057875 |
| 158. | 499 | 21200 | 9.961757 | 1.766442 | 0056033 |
| 159. | 1 | 24000 | 10.08581 | 1.682688 | 0055244 |
| 160. | l 254 | 42800 | 10.66429 | 1.460938 | 0054843 |
| | | | | | |
| 161. | 165 | 22700 | 10.03012 | 1.800058 | 0053979 |
| 162. | 216 | 24999 | 10.12659 | 1.587192 | 0053433 |
| 163. | 416 | 7200 | 8.881836 | 1.915451 | 0053274 |
| 164. | 228 | 31600 | 10.36091 | 1.617406 | 0051245 |
| | l 293 | 27901 | 10.23642 | 1.413423 | 0049782 |
| | | | | | |
| 166. | 27 | 16600 | 9.717158 | 1.682688 | 0049226 |
| 167. | 49 | 14400 | 9.574984 | 1.499623 | 0049053 |
| 168. | 229 | 46700 | 10.7515 | 1.617406 | 0047393 |
| 169. | 147 | 15600 | 9.655026 | 2.164472 | 0047173 |
| 170. | 175 | 22600 | 10.02571 | 1.629241 | 0047105 |
| | | | | | |
| 171. | 128 | 16200 | 9.692766 | 1.83098 | 0045839 |
| 172. | 312 | 22099 | 10.00329 | 1.693779 | 0044402 |
| 173. | 122 | 20300 | 9.918376 | 1.759581 | 0044247 |
| 174. | 235 | 29001 | 10.27509 | 1.623341 | 0042688 |
| | • | | | 1.534714 | • |
| | • | | | | |
| 176. | 73 | 22800 | 10.03452 | 1.418277 | 0041709 |
| 177. | 206 | 22600 | 10.02571 | 1.587192 | 004164 |
| 178. | 125 | 18800 | 9.841612 | 1.759581 | 0037207 |
| | | | | 1.545433 | |
| | | | | | 0034284 |
| | | | | | |
| 181. | | 19400 | 9.873029 | 1.499623 | 0033576 |
| | 39 | 24701 | 10.1146 | 1.607436 | 0031628 0030147 |
| 183. | 178 | 24600 | 10.1105 | 1.629241 | 0030147 |
| 184. | 110 | 19400 | | 1.648659 | 0029739 |
| 185. | 61 | 18700 | | 1.510/22 | 0029386 |
| 100 | | | | | 0020010 |
| | | | | | 0028818 |
| | | | | | 0027495 |
| | 120 | | | 1.759581 | |
| 189. | 124 | 17300 | 9.758462 | 1.759581 | 0025984 |

| 190. | 148 | 14600 | 0 500777 | | |
|------|-----------|-------|-----------|----------|----------------------------|
| | | 14000 | 9.588/// | 2.164472 | 0025087 |
| 191. | 123 | 20500 | 9,92818 | 1.759581 | 0022385. |
| | | | | 1.460938 | |
| | 167 | 50001 | | | 0022249 |
| | | | 10.0190 | 1.800058 | |
| | 500 | | 9.769957 | | 0021589 |
| 195. | 428 | | 9.296518 | 1.915451 | 0021474 |
| 196. | 196 | 50001 | 10.8198 | | 0021369 |
| 197. | 248 | 20500 | 9.92818 | 1.460938 | 002122 |
| | | | | 1.682688 | |
| | • | | | 1.499623 | |
| | - | | | 1.48614 | |
| 200. | 342 | | | | 1.0010905 |
| 201. | 48 | 16600 | 9.717158 | 1.499623 | 0018894 |
| 202. | 60 | | | 1.510722 | 0018844 |
| 203. | 226 | | 10.8198 | | 001832 |
| | 155 | | | 2.164472 | 0018229 |
| | 353 | | 9.830916 | | 0016996 |
| 206 | 137 | 17400 | 0 764226 | 1 02000 | 0016502 |
| | | 17400 | | 1.83098 | |
| | 394 | | | 1.93586 | |
| | | | | 1.477049 | • |
| | | | | 1.587192 | |
| 210. | 223 | 27499 | 10.2219 | 1.623341 | 0015933 |
| 211. | 330 | 22600 | 10.02571 | 1.526056 | 0015749- |
| | | | | 1.704748 | |
| | | 11500 | | 1.94591 | |
| | 272 | 25200 | | 1.534714 | 0014252 |
| | 209 | 24399 | 10.1023 | 1.587192 | 0014232 |
| | | | | | |
| | 395 | | 9.449357 | | 0013078 |
| | • | | | 1.745715 | |
| | | | | 1.4884 | |
| 219. | 63 | 22199 | 10.0078 | 1.510722 | 0011577 |
| 220. | 324 | 18500 | 9.825526 | 1.595339 | 0011329 |
| 221 | | 22100 | 10 04750 | 1 456207 | 0011200 |
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| | 74 | | | | 0010706 |
| | 142 | | | 1.83098 | |
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| | • | 28601 | | 1.474763 | 0009321 |
| | | | | | 0008757- |
| | 246 | | | | 0008337 |
| | | 22800 | 10 03/152 | 1 603770 | 0007587 |
| | | | | 1.693779 | |
| | | | | | |
| | | 22800 | | | 000556 |
| | | | | | 0005297 |
| | | | | | 0004579 |
| | | | | 1.474763 | |
| | | | | 1.587192 | |
| 234. | | | | | |
| | 126 | 21400 | 9.9/114/ | 1./39361 | 0002049 |
| 235. | | | | | 0002649 0001415- |

237. | 397 12500 9.433484 1.93586 -.0001163 | 238. | 477 16700 9.723164 1.814825 -.000081 |

| 238. | 477 | 16700 | 9.723164 | 1.814825 | 000081 |
|--------|---------|-----------|----------|----------|----------|
| 239. | 86 | 26601 | 10.1887 | 1.501853 | 0000522 |
| 240. | 232 | 31701 | 10.3641 | 1.617406 | 0000151 |
| | | | | | |
| 241. | 310 | 20300 | 9.918376 | 1.693779 | .0001054 |
| 242. | 219 | 21500 | 9.975808 | 1.704748 | .0001567 |
| 243. | | 46000 | 10.7364 | 1.4884 | .0001949 |
| 244. | | 14100 | 9.55393 | 1.964311 | .0002028 |
| 245. | 277 | 33200 | 10.41031 | 1.497388 | .0002292 |
| 245. | 2// | 33200 | 10.41031 | 1.497300 | |
| 246. | 44 | 24701 | 10.1146 | 1.499623 | .0002693 |
| 247. | 305 | 36098 | 10.49399 | 1.551809 | .000271 |
| 248. | 344 | 23899 | 10.43333 | 1.576915 | .000271 |
| | | | 10.00139 | | |
| 249. | 297 | 27100 | | 1.474763 | .0003851 |
| 250. | 84 | 22900 | 10.03889 | 1.449269 | .0005833 |
| 251. | 212 | 19300 | 9.867861 | 1.587192 | .0005905 |
| 252. | 43 | 25301 | 10.1386 | 1.499623 | .0006433 |
| 253. | | | 10.13058 | | .0006712 |
| | 237 | 25099 | | 1.623341 | |
| 254. | 37 | 20000 | 9.903487 | 1.607436 | .0006828 |
| 255. | 170 | 22299 | 10.0123 | 1.800058 | .0007145 |
| 256. l | 130 | 14300 | 9.568015 | 1.83098 | .0007622 |
| 257. | 70 | 20900 | 9.947504 | 1.408545 | .0007022 |
| 258. | ! | | 10.1023 | 1.587192 | |
| | 207 | 24399 | | | .000875 |
| 259. | 251 | 24399 | 10.1023 | 1.460938 | .00088 |
| 260. | 314 | 21600 | 9.980449 | 1.693779 | .0009465 |
| 261. | 501 | 16800 | 9.729135 | 1.766442 | .0010252 |
| 262. | 238 | 31499 | 10.35771 | 1.623341 | .0010367 |
| 263. | | 24200 | 10.09411 | 1.418277 | .0010956 |
| | | | | | |
| 264. | 57 | 24701 | 10.1146 | 1.410987 | .0011089 |
| 265. | 138 | 17100 | 9.746834 | 1.83098 | .0011234 |
| 266. | 252 | 24800 | 10.1186 | 1.460938 | .0011321 |
| 267. | 503 | 20600 | 9.933046 | 1.745715 | |
| | | | | 1.595339 | |
| 268. | 323 | 20400 | 9.92329 | | .0011816 |
| 269. | 227 | 37602 | 10.53481 | 1.617406 | .0013204 |
| 270. | 30 | 21000 | 9.952278 | 1.682688 | .0013535 |
| 271. | 56 | 35/01 | | 1.393766 | .0014627 |
| | 440 | 12800 | 9.4572 | 2.00148 | .0014027 |
| 273. | | 15600 | 9.655026 | 2.164472 | .0016168 |
| | | | | | • |
| | 313 | 19400 | 9.873029 | 1.693779 | .0017158 |
| 275. | 281 | 45401 | 10.72329 | 1.4884 | .0017327 |
| 276. | 217 | 23300 | 10.05621 | 1.704748 | .0018928 |
| | 136 | 18100 | 9.803667 | 1.83098 | .0019621 |
| | 150 | | 9.809176 | 1.682688 | |
| : | | | | | |
| | 271 | | 9.957028 | 1.534714 | .0020471 |
| 280. | 172 | 19100 | 9.857444 | 1.800058 | .0023016 |
| 281. | 267 | 30699 | 10.33199 | 1.867176 | .0023292 |
| : | 113 | 18800 | 9.841612 | 1.699279 | .0023232 |
| 283. | | 22501 | 10.02131 | 1.501853 | .0023302 |
| | | | | | |
| 284. | 475 | 13800 | 9.532424 | 1.764731 | .0023/62 |
| | | | | | |

| 285. | 352 | 24101 | 10.09001 | 1.413423 | .002526 |
|------|-----------|-------|--------------|----------|----------|
| 286. | 244 | 23699 | 10 07310 | 1.453953 | .0025868 |
| | | | | | |
| | 85 | 23899 | | 1.501853 | |
| | 72 | 21700 | 9.90500/ | 1.418277 | .0027141 |
| | 225 | 44802 | | 1.617406 | .0027266 |
| 290. | 304 | 33100 | 10.40729 | 1.551809 | .0027396 |
| 291. | 90 | 28701 | 10.26469 | 1.587192 | .0029018 |
| 292. | 403 | 12100 | 9.400961 | 1.93586 | .0029982 |
| | 135 | | | 1.83098 | |
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| | 384 | | | 1.94591 | .0031938 |
| | | | | | |
| | 38 | | 9.522813 | | .0032047 |
| 297. | 266 | 22800 | 10.03452 | 1.867176 | .0032496 |
| | 187 | 50001 | 10.8198 | 1.585145 | .0033213 |
| | 291 | | 10.25769 | | .0033408 |
| | 422 | 14200 | 9.560997 | | .0033722 |
| 301. | 351 | 22900 | 10.03889 | 1.456287 | .0033961 |
| | | | | 1.83098 | .003347 |
| | | | | 1.595339 | |
| | - | | | 1.460938 | .003478 |
| | • | | | 1.400938 | .0034818 |
| 303. | 404 | 40499 | 10./093 | 1.423313 | .0033203 |
| | 164 | 50001 | 10.8198 | 1.800058 | .0035734 |
| 307. | | | 9.985067 | | .0036516 |
| 308. | | 29801 | 10.3023 | 1.814825 | .0039417 |
| 309. | | | | 1.460938 | .0039509 |
| | 181 | 39799 | | 1.585145 | .0041167 |
| 311. | 79 | 21200 | 9.961757 | 1.474763 | .0041254 |
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| 314. | | | 10.78521 | | .0043115 |
| 515. | | | | | |
| | 308 | | | 1.551809 | |
| | 354 | | | 1.410987 | |
| 318. | 321 | 23799 | 10.0774 | 1.595339 | .0046011 |
| 319. | | | | 1.94591 | .0047599 |
| 320. | 389 | 10200 | 9.230143 | 1.94591 | .0047679 |
| | • | | | 1.388791 | |
| 322 | 260 | 30101 | 10.343 | 1.867176 | .0047700 |
| | | | | 1.474763 | |
| | • | | | | |
| | | | | 1.623341 | |
| | - | | | 1.759581 | |
| | • | | | 1.453953 | |
| | • | | | 1.800058 | |
| | | | | 1.693779 | |
| | 259 | 36001 | 10.4913 | | .0056653 |
| | | 24399 | 10.1023 | | .0056988 |
| | | | | | |
| 331. | 331 | 19800 | 9.893437 | 1.526056 | .0057538 |
| | | | | | |

| 332. | 109 | 19800 | 9.893437 | 1.648659 | .005793 |
|--------------------------------------|-------------------------------------|----------------------------------|---|---------------------------------|----------------------------------|
| 333. | 404 | 8300 | 9.024011 | 1.93586 | .0060482 |
| 334. | 319 | 23100 | 10.04759 | 1.693779 | .006068 |
| 335. | 444 | 15400 | 9.642123 | 2.00148 | .0062065 |
| | | | | | |
| 336. | 114 | 18700 | 9.836279 | 1.699279 | .0064821 |
| 337. | 47 | 20000 | 9.903487 | 1.499623 | .0065764 |
| | - | | | 1.456287 | .0065786 |
| | • | | 10.4545 | | .006829 |
| | - | | 10.27849 | | .0068445 |
| | , | | | | |
| 341. | 411 | 15000 | 9.615806 | 1.786747 | .0069358 |
| | 247 | 24299 | 10.09819 | 1.460938 | .0069601 |
| | • | 19700 | 9.888374 | 1.766442 | .0070464 |
| | 355 | 18200 | 9.809176 | 1.418277 | .0070865 |
| | 320 | | 9.952278 | | .0070909 |
| | | | | | |
| 346. | • | | 10.41631 | | .007096 |
| | • | | 9.588777 | | |
| | • | | 9.846917 | | |
| | | | 10.42819 | | .0072586 |
| 350. | 1116 | | 9.814656 | | .0072618 |
| | | | | | |
| 351. | 20 | 18200 | 9.809176 | 1.682688 | .0075003 |
| | 285 | 32199 | | 1.386294 | .007564 |
| | 318 | 19800 | 9.893437 | | .0076158 |
| | 329 | 19300 | 9.867861 | 1.526056 | .0076533 |
| 355. | • | 24999 | 10.12659 | 1.510722 | .0076632 |
| | | | | | |
| 356. | 46 | 19300 | 9.867861 | 1.499623 | .0077691 |
| | • | 7012 | 8.855378 | 1.93586 | .007824 |
| 358. | 419 | 8800 | 9.082507 | 1.915451 | .007874 |
| 359. | 458 | 13500 | 9.510445 | 1.964311 | .0080279 |
| 360. | 205 | 50001 | 10.8198 | 1.425515 | .0081019 |
| | | | | | |
| 361. | 447 | 14900 | 9.609117 | 2.00148 | .0081669 |
| | 298 | | 9.918376 | | .0085907 |
| 363. | 76 | 21400 | 9.971147 | 1.474763 | .0091335 |
| 364. | 82 | 23899 | 10.08159 | 1.449269 | .0092037 |
| 365. | | 20100 | 9.908475 | 1.348073 | .0093339 |
| | | | | | |
| 366. | 270 | 20700 | 9.937889 | 1.534714 | .0094668 |
| | 391 | | | 1.94591 | |
| | 387 | | 9.25913 | 1.94591 | |
| 369. | 332 | 17100 | 9.746834 | 1.477049 | .009686 |
| | | | | 1.964311 | |
| | | | | | |
| 371. | 103 | 18600 | 9.830916 | 1.648659 | .0097767 |
| | 7 | | 10.03889 | | .0099829 |
| | | | 9.994242 | 1.366092 | .0101005 |
| 373. | 255 | | 10 07210 | 1.366092 1.453953 | .0102395 |
| | • | 23699 | 10.07319 | | |
| 374. | • | | 9.883285 | 1.83098 | .0102551 |
| 374. | 239 132 | 19600 | 9.883285 | | .0102551 |
| 374. 375. | 239 132 | 19600 | 9.883285 | 1.83098 | |
| 374. 375. | 239 132 288 | 19600 23200 | 9.883285 | 1.83098 1.398717 | .0103194 |
| 374. 375. 376. 377. | 239 132 288 284 | 19600 23200 50001 | 9.883285 10.05191 10.8198 | 1.83098 1.398717 | .0103194 .0103747 |
| 374. 375. 376. 377. 378. | 239 132 288 284 | 19600 23200 50001 19100 | 9.883285 10.05191 10.8198 9.857444 | 1.83098 1.398717 1.388791 | .0103194 .0103747 .0104986 |

| 380. | 325 | 24999 | 10.12659 | 1.595339 | .0105819 |
|------|------------|-------|--------------|----------|----------|
| 381. | 433 | 16100 | 9.686575 | 1 764731 | .0107606 |
| | • | | | | |
| | 256 | | | | .0108386 |
| 383. | • | 24800 | 10.1186 | 1.398717 | .0114381 |
| 384. | 174 | 23600 | 10.069 | 1.629241 | .0116044 |
| 385. | 496 | 23100 | 10.04759 | 1.766442 | .0117644 |
| 386. | 65 | 33001 | 10.40429 | 1.425515 | .0121224 |
| | 78 | 20800 | 9.942708 | 1.474763 | .0121406 |
| | 265 | | | | .0122671 |
| | • | | 9.723164 | | .0123169 |
| | | | | | |
| 390. | 218 | | 10.26469 | 1.704748 | .012537 |
| 391. | 1 | | 10.06049 | 1.479329 | .0129076 |
| 392. | 55 | 18900 | 9.846917 | 1.410987 | .0129513 |
| 393. | 221 | | | 1.623341 | .0129747 |
| | 327 | 23001 | 10.04329 | 1.595339 | .0132197 |
| | 100 | 33200 | 10.41031 | 1.492904 | .01351 |
| 555. | ±00 | | | 1.752304 | .01331 |
| | 58 | 32562 | 10.3909 | 1.413423 | .0140252 |
| 397. | 202 | | | 1.423108 | .0143737 |
| 398. | 240 | 23300 | 10.05621 | 1.453953 | .0146344 |
| 399. | 93 | | | 1.534714 | .0146731 |
| 400. | 88 | 22199 | 10.0078 | 1.501853 | .0147033 |
| 401. | 220 | 22100 | 10 0079 | 1.595339 | .0147417 |
| | • | | | | |
| | | | | 1.474763 | |
| 403. | 91 | 22600 | 10.02571 | 1.587192 | .0149329 |
| 404. | 418 | 10400 | 9.249561 | 1.915451 | .0149713 |
| 405. | 459 | 14900 | 9.609117 | 1.964311 | .0150557 |
| 406. | 22 | 19600 | 9.883285 | 1.682688 | .0152086 |
| | 398 | | 9.047821 | 1.93586 | .0153095 |
| | | | 10.68049 | 1.7492 | .0162008 |
| | • | | 9.392662 | | |
| | | | | | .0163773 |
| 410. | 300 | 29001 | 10.27509 | 1.386294 | .0168745 |
| 411. | 139 | 13300 | 9.49552 | 1.83098 | .0171233 |
| 412. | 89 | 23600 | 10.069 | 1.587192 | .0173629 |
| | | | | 1.453953 | .0176759 |
| | | | 10.0123 | 1.398717 | .0176951 |
| 415. | 356 | 20600 | 9.933046 | 1.418277 | .0178045 |
| 41.5 | | | | | |
| 416. | • | | | 1.460938 | |
| | 6 | 28701 | 10.26469 | 1.521699 | .0180267 |
| | | 8500 | | 1.93586 | |
| 419. | 199 | 34600 | 10.45161 | 1.396245 | .0184836 |
| | 92 | 22000 | | 1.587192 | |
| | • | 24600 | | 1 505220 | |
| 421. | | | | 1.595339 | |
| 400 | 134 | | | 1.83098 | |
| | 241 | | 9.998797 | | .0201607 |
| 423. | | | 0 635600 | 2 164472 | .0202187 |
| 423. | | | | 2.164472 | |
| 423. | 153 4 | | 10.41631 | 1.521699 | .0206138 |

| 427. | 141 | | | 1.83098 | |
|--|--|---|---|---|--|
| 428. | | | 10.05621 | 1.510722 | .0209189 |
| 429. | 385 | 8800 | 9.082507 | 1.94591 | .0216851 |
| 430. | 485 | 20600 | 9.933046 | | .0229157 |
| 431. | 402 | 13600 | 0 517025 | 1 0066/10 | 0221727 |
| | 492 | | 9.517825 | | .0231727 |
| | 115 19 | 18500 | 9.825526 | 1.699279 | .0234778 |
| | | 20200 | 9.913438 9.957028 | 1.682688 | .0239359 |
| | | | | | .0240607 |
| | • | 23501 | 10.0648 | 1.381282 | .0240661 |
| | | | | 1.786747 | |
| 437. | 133 | 23001 | 10.04329 | 1.83098 | .0251679 |
| | • | | | 1.964311 | |
| | - | | | 1.396245 | |
| | | | 10.0078 | | |
| | | | | | |
| 141. | 388 | /400 | 8.909235 | | .0263166 |
| 142. | 69 | 7400 17400 36199 | 9.764226 | | .0263849 |
| 143. | 5 | 36199 | 10.49679 | 1.521699 | .0266658 |
| 144. | | 20600 | 9.933046 | 1.534714 | .0274035 |
| 45. | 407 | 11900 | 9.384294 | 1.885553 | .0274155 |
| 146. | 402 | 7200 | 8.881836 | 1.93586 | .0280494 |
| | • | | | 1.386294 | |
| | - | | | 1.492904 | |
| | | | 9.898475 | | |
| 50. | 345 | | | 1.576915 | .0318101 |
| | | | | | |
| ¥51. | 262 | 43101 | 10.6713 | 1.867176 | .032273 |
| 52. | 379 | 13100 | 9.480368 | 1.903599 | .0328231 |
| 153. | 198 | 30300 | 10.3189 | 1.396245 | .0330362 |
| 54. | 486 | 21200 | 9.961757 | 1.763017 | .0337355 |
| | 302 | 22000 | 9.998797 | 1.465567 | .0341089 |
| | • | 14500 | | 1 76/1721 | 0242425 |
| | • | | 9.581903 | | .0342425 |
| | • | | 10.1186 | | .0343599 |
| 58. | 77 | | | | .0356156 |
| | 97 | | 9.971147 | | |
| 160. | 476 | 13300 | 9.49552 | 1.764731 | .0361187 |
| 461. | 442 | 17100 | 9.746834 | 2.00148 | .0364184 |
| 462. | 14 | 20400 | 9.92329 | 1.682688 | .0369498 |
| 460 | 1 277 | 13900 | 9.539644 | 1.903599 | .0369729 |
| 463. | 3// | 13300 | | | |
| | 401 | 5600 | 8.630522 | 1.93586 | .0383786 |
| 464. | 401 367 | 5600 21900 | 8.630522 9.994242 | 1.971299 | .0430825 |
| 464. 465. | 401 367 | 5600 21900 | 8.630522 9.994242 | 1.971299 | .0430825 |
| 164. 165. 166. | 401 367 | 5600 21900 27100 | 8.630522 9.994242 | 1.971299 1.656321 | .0430825 .043922 |
| 164. 165. 166. 167. | 401 367 8 382 | 5600 21900 27100 10900 | 8.630522 9.994242 | 1.971299 1.656321 1.903599 | .0430825 .043922 .0451832 |
| 64. 65. 66. 67. 68. | 401 367 8 382 378 | 5600 21900 27100 10900 13300 | 8.630522 9.994242 10.20729 9.296518 9.49552 | 1.971299 1.656321 1.903599 1.903599 | .0430825 .043922 .0451832 .0464509 |
| 164. 165. 166. 167. 168. | 401 367 8 382 378 467 | 5600 21900 27100 10900 13300 19000 | 8.630522 9.994242 10.20729 9.296518 9.49552 9.852194 | 1.971299 1.656321 1.903599 1.903599 1.879465 | .0430825 .043922 .0451832 .0464509 .0484201 |
| 164. 165. 166. 167. 168. | 401 367 8 382 378 467 | 5600 21900 27100 10900 13300 19000 | 8.630522 9.994242 10.20729 9.296518 9.49552 9.852194 | 1.971299 1.656321 1.903599 1.903599 | .0430825 .043922 .0451832 .0464509 .0484201 .048621 |
| 464. 465. 466. 467. 468. 469. 470. | 401 367 | 5600 21900 27100 10900 13300 19000 | 8.630522 9.994242 10.20729 9.296518 9.49552 9.852194 | 1.971299 1.656321 1.903599 1.903599 1.879465 1.94591 | .0430825 .043922 .0451832 .0464509 .0484201 .048621 |
| 464. 465. 466. 467. 468. 469. 470. | 401 367 8 382 378 467 386 | 5600 21900 27100 10900 13300 19000 7200 15000 19400 | 8.630522 9.994242 10.20729 9.296518 9.49552 9.852194 8.881836 | 1.971299 1.656321 1.903599 1.903599 1.879465 1.94591 | .0430825 .043922 .0451832 .0464509 .0484201 .048621 |
| 465. 466. 467. 468. 469. 470. | 401 367 8 382 378 467 386 376 | 27100 10900 13300 19000 7200 | 8.630522 9.994242 10.20729 9.296518 9.49552 9.852194 8.881836 | 1.971299 1.656321 1.903599 1.903599 1.879465 1.94591 | .0430825 .043922 .0451832 .0464509 .0484201 .048621 |

| 475. | 443 | 18400 | 9.820106 | 2.00148 | .0566888 |
|------|--------------|---------------|----------------------|----------|------------------------|
| 476. | 286 | 22000 | 9.998797 | 1.358409 | .0573509 |
| 477. | 17 | 23100 | 10.04759 | 1.682688 | .0589057 |
| 478. | 400 | 6300 | 8.748305 | 1.93586 | .0589789 |
| 479. | 432 | 14100 | 9.55393 | 1.764731 | .0600489 |
| 480. | 465 | 21400 | 9.971147 | 1.879465 | .0677176 |
| | | | | | |
| 481. | 380 | 10200 | 9.230143 | 1.903599 | .0691514 |
| 482. | 364 | 16800 | 9.729135 | 2.04122 | .0709101 |
| 483. | 392 | 23200 | 10.05191 | 1.94591 | .0715526 |
| 484. | 381 | 10400 | 9.249561 | 1.903599 | .0725951 |
| 485. | 149 | 17794 | 9.786616 | 2.164472 | .0743867 |
| 486. | · 466 | 19900 | 9.898475 | 1.879465 | .0744588 |
| 487. | 399 | 5000 | 8.517193 | 1.93586 | .0797595 |
| 488. | l 366 | 27499 | 10.2219 | 1.971299 | .0797842 |
| 489. | 357 | 17794 | 9.786616 | 2.04122 | .0857706 |
| 490. | 154 | 19400 | 9.873029 | 2.164472 | .0910494 |
| | | | | | |
| 491. | 463 | 19500 | 9.87817 | 1.964311 | .0941472 |
| 492. | 464 | 20200 | 9.913438 | 1.964311 | .0974507 |
| 493. | 427 | 10200 | 9.230143 | 1.764731 | .1007114 |
| 494. | 406 | 5000 | 8.517193 | 1.93586 | .1024767 |
| 495. | 151 | 21500 | 9.975808 | 2.164472 | .1047597 |
| 406 | | 10600 | 0 002205 | 2 164472 | 1120427 |
| 496. | 152 | 19600 | 9.883285 | 2.164472 | .1120427 |
| | 460 | 20000 | 9.903487 | 1.964311 | .1142668 |
| 498. | 160 | 23300 | 10.05621 | 2.164472 | .1165014 |
| 499. | 491 362 | 8100 19900 | 8.999619 9.898475 | 1.806648 | .1222368 .1376445 |
| 500. | 302 | 19900 | 9.0904/3 | 2.04122 | .1370445 |
| 501. | 363 | 20800 | 9.942708 | 2.04122 | .1707894 |
| 502. | 490 | 7000 | 8.853665 | 1.806648 | .1791869 |
| 503. | 358 | 21700 | 9.985067 | 2.04122 | .1827834 |
| 504. | 360 | 22600 | 10.02571 | 2.04122 | .2209745 |
| 505. | 361 | 24999 | 10.12659 | 2.04122 | .2422512 |
| | | | | | |
| 506. | 359 | 22700 | 10.03012 | 2.04122 | .2483543 |

5.3 ENDOGENEITY & MEASUREMENT ERROR

+----+

REGRESSION WITH NON-IID ERRORS

6.1 THE GENERALIZED LINEAR REGRESSION MODEL

- 6.1.1 Types of Deviations from i.i.d. Errors
- 6.1.2 The Robust Estimator of VCE

. use fertil2.. describe

Contains data from fertil2.dta

obs: 4,361 vars: 30 size: 484,071

2 Dec 2004 00:16

| vaniahla nama | | display format | variable ' | lahal | |
|----------------------|-------|-------------------|------------|----------------|-------|
| variable name | type | TOTIIIat | variable | | |
| mnthborn | float | %9.0g | | | |
| yearborn | float | %9.0g | | | |
| age | float | %9.0g | | | |
| electric | float | %9.0g | | | |
| radio | float | %9.0g | | | |
| tv | float | %9.0g | | | |
| bicycle | float | %9.0g | | | |
| educ | float | %9.0g | | | |
| ceb | float | %9.0g | | | |
| agefbrth | float | %9.0g | | | |
| children | float | %9.0g | | | |
| knowmeth | float | %9.0g | | | |
| usemeth | float | %9.0g | | | |
| monthfm | float | %9.0g | | | |
| yearfm | float | %9.0g | | | |
| agefm | float | %9.0g | | | |
| idlnchld | float | %9.0g | | | |
| heduc | float | %9.0g | | | |
| agesq | float | %9.0g | | | |
| urban | float | %9.0g | | | |
| urbeduc | float | %9.0g | | | |
| spirit | float | %9.0g | | | |
| protest | float | %9.0g | | | |
| catholic | float | %9.0g | | | |
| frsthalf | float | %9.0g | | | |
| educ0 | float | %9.0g | | | |
| evermarr | float | %9.0g | | | |
| _est_OLS | byte | %8.0g | | from estimates | |
| $_{ m est_}$ robust | byte | %8.0g | • | from estimates | |
| _est_cluster | byte | %8.0g | esample() | from estimates | store |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|--------------|-------|----------|-----------|-----|------|
| mnthborn | 4,361 | 6.331346 | 3.323333 | 1 | 12 |
| yearborn | 4,361 | 60.43362 | 8.682723 | 38 | 73 |
| age | 4,361 | 27.40518 | 8.685233 | 15 | 49 |
| electric | 4,358 | .1402019 | .3472363 | 0 | 1 |
| radio | 4,359 | .7017665 | . 457535 | 0 | 1 |
| tv | 4,359 | .0929112 | .2903413 | 0 | 1 |
| bicycle | 4,358 | .2758146 | .4469751 | Θ | 1 |
| educ | 4,361 | 5.855996 | 3.927075 | 0 | 20 |
| ceb | 4,361 | 2.441642 | 2.406861 | 0 | 13 |
| agefbrth | 3,273 | 19.0113 | 3.092333 | 10 | 38 |
| children | 4,361 | 2.267828 | 2.222032 | 0 | 13 |
| knowmeth | 4,354 | .9632522 | .1881636 | 0 | 1 |
| usemeth | 4,290 | .5776224 | .4939956 | 0 | 1 |
| monthfm | 2,079 | 6.270322 | 3.619943 | 1 | 12 |
| yearfm | 2,079 | 76.91246 | 7.760183 | 50 | 88 |
| agefm | 2,079 | 20.68639 | 5.002383 | 10 | 46 |
| idlnchld | 4,241 | 4.615892 | 2.219303 | 0 | 20 |
| heduc | 1,956 | 5.144683 | 4.803028 | 0 | 20 |
| agesq | 4,361 | 826.46 | 526.9232 | 225 | 2401 |
| urban | 4,361 | .5166246 | . 4997808 | 0 | 1 |
| urbeduc | 4,361 | 3.469158 | 4.294228 | 0 | 20 |
| spirit | 4,361 | .4221509 | . 493959 | 0 | 1 |
| protest | 4,361 | .2277001 | .4193961 | 0 | 1 |
| catholic | 4,361 | .1024994 | .3033387 | Θ | 1 |
| frsthalf | 4,361 | .5404724 | . 4984164 | 0 | 1 |
| educ0 | 4,361 | .2077505 | .4057437 | 0 | 1 |
| evermarr | 4,361 | .4767255 | . 4995153 | Θ | 1 |
| _est_OLS | 4,361 | .7367576 | . 4404433 | Θ | 1 |
| _est_robust | 4,361 | .7367576 | . 4404433 | 0 | 1 |
| _est_cluster | 4,361 | .7367576 | .4404433 | 0 | 1 |

. use fertil2.. regress ceb age agefbrth usemeth

| Source | SS | df | MS | Number of obs | = | 3,213 |
|----------|------------|-----------|------------|---------------|------|-----------|
| + | | | | F(3, 3209) | = | 1433.16 |
| Model | 9202.53439 | 3 | 3067.51146 | Prob > F | = | 0.0000 |
| Residual | 6868.49331 | 3,209 | 2.14038433 | R-squared | = | 0.5726 |
| + | | | | Adj R-squared | = | 0.5722 |
| Total | 16071.0277 | 3,212 | 5.00343328 | Root MSE | = | 1.463 |
| | | | | | | |
| | | | | | | |
| ceb | Coef. | Std. Err. | t P | P> t [95% Co | onf. | Interval] |
| | | | | | | |

age | .2237368 .003448 64.89 0.000 .2169763 .2304974 agefbrth | -.2606634 .0087954 -29.64 0.000 -.2779085 -.2434184

| usemeth | . 1873702 | .0554298 | 3.38 | 0.001 | .0786888 | .2960516 |
|---------|-----------|----------|------|-------|----------|----------|
| _cons | 1.358134 | .1737828 | 7.82 | 0.000 | 1.017397 | 1.69887 |

- . estimates store nonRobust
- . summarize ceb age agefbrth usemeth children if e(sample)

| 0bs | Mean | Std. Dev. | Min | Max |
|-------|----------------------------------|--|--|--|
| 3,213 | 3.230003 | 2.236836 | 1 | 13 |
| 3,213 | 29.93931 | 7.920432 | 15 | 49 |
| 3,213 | 19.00498 | 3.098121 | 10 | 38 |
| 3,213 | .6791161 | .4668889 | Θ | 1 |
| 3,213 | 2.999378 | 2.055579 | Θ | 13 |
| | 3,213 3,213 3,213 3,213 | 3,213 3.230003 3,213 29.93931 3,213 19.00498 3,213 .6791161 | 3,213 3.230003 2.236836 3,213 29.93931 7.920432 3,213 19.00498 3.098121 3,213 .6791161 .4668889 | 3,213 3.230003 2.236836 1 3,213 29.93931 7.920432 15 3,213 19.00498 3.098121 10 3,213 .6791161 .4668889 0 |

. regress ceb age agefbrth usemeth, robust

| Linear regression | Number of obs | = | 3,213 |
|-------------------|---------------|---|--------|
| | F(3, 3209) | = | 874.06 |
| | Prob > F | = | 0.0000 |
| | R-squared | = | 0.5726 |
| | Root MSE | = | 1.463 |

| ceb | Coef. | Robust Std. Err. | t | P> t | [95% Conf. | - |
|-----------|----------|---------------------|--------|-------|------------|----------|
| age | .2237368 | .0046619 | 47.99 | 0.000 | .2145962 | .2328775 |
| agefbrth | 2606634 | .0095616 | -27.26 | 0.000 | 2794109 | 2419159 |
| usemeth | .1873702 | .0606446 | 3.09 | 0.002 | .0684642 | .3062762 |
| _cons | 1.358134 | .1675624 | 8.11 | 0.000 | 1.029593 | 1.686674 |

. estimates store Robust

- . estimates table nonRobust Robust, b(%9.4f) se(%5.3f) t(%5.2f) p(%4.3f) title(
- > Estimates of CEB with OLS and Robust Standard Errors)

Estimates of CEB with OLS and Robust Standard Errors

| Variable | nonRobust | Robust |
|----------|-----------|---------|
| age | 0.2237 | 0.2237 |
| I | 0.003 | 0.005 |
| I | 64.89 | 47.99 |
| I | 0.000 | 0.000 |
| agefbrth | -0.2607 | -0.2607 |
| I | 0.009 | 0.010 |
| I | -29.64 | -27.26 |
| I | 0.000 | 0.000 |
| usemeth | 0.1874 | 0.1874 |
| I | 0.055 | 0.061 |
| I | 3.38 | 3.09 |
| I | 0.001 | 0.002 |
| _cons | 1.3581 | 1.3581 |

| | | | | | | | | 1 | | | 0 | | 17 | 4 | | | 0 | . 1 | 68 | 3 |
|--|---|------|-------|---|---|---|---|---|---|---|------|------|-----|----|------|-------|------|-----|----|---|
| | | | | | | | | 1 | | | | 7 | . 8 | 32 | | | 8 | 3. | 1: | 1 |
| | | | | | | | | 1 | | | 0 | ١. ١ | 00 | 0 | | | 0 | . 0 | 0(| 9 |
| | _ | | _ | _ | _ | _ | _ | | _ | _ | | | _ | | | _ | | | | _ |

legend: b/se/t/p

6.1.3 The Cluster Estimator of VCE

. use fertil2.. regress ceb age agefbrth usemeth, cluster(children)

| Linear regression | Number of obs | = | 3,213 |
|-------------------|---------------|---|--------|
| | F(3, 13) | = | 20.91 |
| | Prob > F | = | 0.0000 |
| | R-squared | = | 0.5726 |
| | Root MSE | = | 1.463 |

(Std. Err. adjusted for 14 clusters in children)

| ceb | Coef. | Robust Std. Err. | t | P> t | - | Interval] |
|-----------|----------|---------------------|-------|-------|----------|-----------|
| age | .2237368 | .0315086 | 7.10 | 0.000 | .1556665 | .2918071 |
| agefbrth | 2606634 | .0354296 | -7.36 | 0.000 | 3372045 | 1841224 |
| usemeth | .1873702 | .0943553 | 1.99 | 0.069 | 016472 | .3912125 |
| _cons | 1.358134 | .4248589 | 3.20 | 0.007 | .4402818 | 2.275985 |

6.1.4 The Newey-West Estimator of VCE

. use ukrates.. describe

Contains data from ukrates.dta

obs: 526 vars: 3 size: 6,312

2 Dec 2004 10:43

| variable name | 3 | display format | variable label | |
|---------------|----------------|-------------------|----------------|--|
| rs month | float float | - 3 | | |

month float %9.0g
r20 float %9.0g

Sorted by: month

. summarize

| Variable | | | Std. Dev. | | |
|----------|-----|----------|----------------------|----------|-------|
| | 526 | 7.651513 | 3.553109 151.9874 | 1.561667 | 16.18 |

r20 | 526 8.863726 3.224372 3.35 17.18

- . quietly regress D.rs LD.r20
- . estimates store nonHAC
- . newey D.rs LD.r20, lag(5)

Regression with Newey-West standard errors Number of obs = 524 maximum lag: 5 F(1, 522) = 36.00 Prob > F = 0.0000

| D.rs | | Newey-West Std. Err. | | | [95% Conf. | Interval] |
|--------------|-----------|-------------------------|------|-------|------------|-----------|
| r20 LD. | . 4882883 | .0813867 | 6.00 | 0.000 | .3284026 | . 648174 |
| _cons | .0040183 | .0254102 | 0.16 | 0.874 | 0459004 | .0539371 |

- . estimates store NeweyWest
- . estimates table nonHAC NeweyWest, b(%9.4f) se(%5.3f) t(%5.2f) p(%4.3f) title(
- > Estimates of D.rs with OLS and Newey-West Standard Errors)

Estimates of D.rs with OLS and Newey-West Standard Errors

| Variable | nonHAC | NeweyWest |
|----------|--------|-----------|
| r20 | | |
| LD. | 0.4883 | 0.4883 |
| 1 | 0.067 | 0.081 |
| | 7.27 | 6.00 |
| | 0.000 | 0.000 |
| | | |
| _cons | 0.0040 | 0.0040 |
| | 0.022 | 0.025 |
| | 0.18 | 0.16 |
| 1 | 0.858 | 0.874 |
| | | |

legend: b/se/t/p

6.1.5 The Generalized Least Squares Estimator

6.2 HETEROSKEDASTICITY IN THE ERROR DISTRIBUTION

- 6.2.1 Heteroskedasticity Related to Scale
- 6.2.1.1 Testing for Heteroskedasticity Related to Scale

- . use hprice2a.(Housing price data for Boston-area communities)
- . regress lprice rooms crime ldist

| Source | SS | df | MS | Number of obs | = | 506 |
|----------|------------|-----|------------|---------------|---|--------|
| +- | | | | F(3, 502) | = | 219.03 |
| Model | 47.9496883 | 3 | 15.9832294 | Prob > F | = | 0.0000 |
| Residual | 36.6325827 | 502 | .072973272 | R-squared | = | 0.5669 |
| +- | | | | Adj R-squared | = | 0.5643 |
| Total | 84.5822709 | 505 | .167489645 | Root MSE | = | .27014 |
| | | | | | | |

| lprice | | | | [95% Conf. | _ |
|--------|----------|--------|-------|------------|----------|
| rooms | .3072343 | 17.24 | 0.000 | .2722172 | .3422514 |
| crime | 0174486 | -10.97 | 0.000 | 0205744 | 0143228 |
| ldist | .074858 | 2.93 | 0.004 | .0246115 | .1251045 |

. estat hettest, iid

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: fitted values of lprice

chi2(1) = 44.67Prob > chi2 = 0.0000

. estat hettest rooms crime ldist, iid

 ${\tt Breusch-Pagan / Cook-Weisberg \ test \ for \ heterosked a sticity}$

Ho: Constant variance

Variables: rooms crime ldist

chi2(3) = 80.11Prob > chi2 = 0.0000

. whitetst command whitetst is unrecognized r(199);

end of do-file
r(199);

6.2.1.2 Feasible Generalized Least Squares Estimation

- . use hprice2a.(Housing price data for Boston-area communities)
- . generate rooms2 = rooms^2
- . regress lprice rooms crime ldist [aw=1/rooms2]
 (sum of wgt is 13.31716591697057)

Source | SS df MS Number of obs = 506

| Model Residual Total | 39.6051883 41.426616 | 3 502 | 13.2017294 .082523139 | R-squared Adj R-square | = = = ed = = | 0.4888 0.4857 |
|--|-------------------------|---|--------------------------|---------------------------|--------------------------|--|
| lprice | Coef. | | | P> t [95% | | - |
| rooms crime ldist _cons | .2345368 0175759 | .0194432 .0016248 .027514 .1172977 | 12.06 -10.82 2.37 | 0.000 .196 0.000020 | 3367 7682 0349 | .272737 0143837 .1191483 8.680536 |

6.2.2 Heteroskedasticity Between Groups of Observations

6.2.2.1 Testing for Heteroskedasticity Between Groups of Observations

. use NEdata.. describe

Contains data from NEdata.dta

120 obs: vars:

24 Oct 2004 13:28

size: 2,640

| variable name | storage type | display format | value label | variable label |
|--|---|--|----------------|----------------|
| state year pop dpi dpipc ldpipc | long int float float float float | %8.0g %8.0g %9.0g %9.0g %9.0g %9.0g | state | |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|----------|-----|----------|-----------|----------|----------|
| state | 120 | 3.5 | 1.714986 | 1 | 6 |
| year | 120 | 1990.5 | 5.790459 | 1981 | 2000 |
| pop | 120 | 2196276 | 1931629 | 515594 | 6362076 |
| dpi | 120 | 4.33e+07 | 4.46e+07 | 4385134 | 1.93e+08 |
| dpipc | 120 | 18.15802 | 5.662848 | 8.153382 | 33.38758 |
| ldpipc | 120 | 2.848302 | .3265395 | 2.098433 | 3.508184 |

. regress dpipc year

| Source | SS | df | MS | Number of obs | = | 120 |
|--------|----|----|----|---------------|---|--------|
| + | | | | F(1, 118) | = | 440.17 |

| Model Residual | 3009.33617 806.737449 3816.07362 | | 6.83675804 | Prob > F R-squared Adj R-squared Root MSE | = = = | 0.7868 |
|---------------------------|--|----------------------|------------|--|-------|-----------------------|
| dpipc | | Std. Err. | | | | Interval] |
| year _cons | .8684582 -1710.508 | .0413941 82.39534 | 20.98 | 9.000 .78648 9.000 -1873.6 | 65 | .9504298 -1547.343 |

- . predict double eps, residual
- . robvar eps, by(state)

| | - | Mean | ory of Residua Std. Dev. | Freq. |
|-------|-----------|-----------|-----------------------------|----------------|
| | | | 1.3596266 | |
| | MA | 1.618796 | .86550138 | 20 |
| | ME - | 2.9841056 | .93797625 | 20 |
| | NH | .51033312 | .61139299 | 20 |
| | RI | 8927223 | .63408722 | 20 |
| | | | .71470977 | |
| | | | 2.6037101 | |
| W0 = | 4.3882072 | df(5, 1 | .14) Pr > | F = 0.00108562 |
| W50 = | 3.2989851 | df(5, 1 | .14) Pr > | F = 0.00806751 |
| W10 = | 4.2536245 | df(5, 1 | .14) Pr > | F = 0.00139064 |

6.2.2.2 Feasible Generalized Least Squares Estimation

. use NEdata.. regress dpipc year

| Source | SS | df | MS | Number of obs | = | 120 |
|----------|------------|----------|------------|---------------|-----|-----------|
| +- | | | | F(1, 118) | = | 440.17 |
| Model | 3009.33617 | 1 | 3009.33617 | Prob > F | = | 0.0000 |
| Residual | 806.737449 | 118 | 6.83675804 | R-squared | = | 0.7886 |
| +- | | | | Adj R-squared | = | 0.7868 |
| Total | 3816.07362 | 119 | 32.0678456 | Root MSE | = | 2.6147 |
| | | | | | | |
| dpipc | | | | P> t [95% (| | |
| +- | | | | | | |
| year | .8684582 | .0413941 | 20.98 | 9.000 .78648 | 865 | .9504298 |
| _cons | -1710.508 | 82.39534 | -20.76 | 9.000 -1873.6 | 73 | -1547.343 |
| | | | | | | |

[.] predict double eps, residual

[.] by state, sort: egen sd_eps = sd(eps)

- . generate double gw_wt = 1/sd_eps^2
- . tabstat sd_eps gw_wt, by(state)

Summary statistics: mean
by categories of: state

| state | $sd_{-}eps$ | gw_wt |
|-------|-------------|----------|
| +- | | |
| CT | 1.359627 | .5409545 |
| MA | .8655014 | 1.334948 |
| ME | .9379762 | 1.136623 |
| NH | .611393 | 2.675218 |
| RI | .6340872 | 2.48715 |
| VT | .7147098 | 1.957675 |
| +- | | |
| Total | .8538824 | 1.688761 |
| | | |

. regress dpipc year [aw = gw_wt] (sum of wgt is 202.6513649171444)

| | ce | | df | | | er of obs | | 120 |
|--------|-----|------------|-----------|-----------|---------|-----------|----|-----------|
| | + | | | | · F(1, | 118) | = | 698.19 |
| Mod | lel | 2845.55409 | 1 | 2845.5540 | 9 Prob | > F | = | 0.0000 |
| Residu | ıal | 480.921278 | 118 | 4.0756040 | 5 R-sq | uared | = | 0.8554 |
| | + | | | | . Adj I | R-squared | = | 0.8542 |
| Tot | al | 3326.47537 | 119 | 27.953574 | 5 Root | MSE | = | 2.0188 |
| | | | | | | | | |
| • | pc | | Std. Err. | | | - | | Interval] |
| | + | | | | | | | |
| ye | ear | .8444948 | .0319602 | 26.42 | 0.000 | .78120 | 49 | .9077847 |
| _CC | ns | -1663.26 | 63.61705 | -26.14 | 0.000 | -1789.2 | 39 | -1537.281 |
| | | | | | | | | |

6.2.3 Heteroskedasticity in Grouped Data

6.2.3.1 Feasible Generalized Least Squares Estimation

. use pubschl.. describe

Contains data from pubschl.dta
obs: 420
vars: 18 2 Dec 2004 12:36
size: 58,380

storage display value
variable name type format label variable label

observation_n~r float %9.0g
dist_cod float %9.0g
county str18 %18s

 $\begin{array}{lll} \mbox{district} & \mbox{str53} & \mbox{\$53s} \\ \mbox{gr_span} & \mbox{str8} & \mbox{\$8s} \end{array}$ $enrl_tot$ float %9.0g teachers
calw_pct
meal_pct
computer float %9.0g float %9.0g float %9.0g float %9.0g testscr float %9.0g comp_stu float %9.0g expn_stu float %9.0g float %9.0g str float %9.0g avginc el_pct read_scr float %9.0g float %9.0g math_scr float %9.0g

.....

Sorted by:

. summarize

| Variable | • | | Std. Dev. | | Max |
|-------------------|-----|----------|-----------|---------|----------|
| observatio~r | | 210.5 | 121.3878 | 1 | 420 |
| dist_cod | 420 | 67472.81 | 3466.995 | 61382 | 75440 |
| county |] 0 | | | | |
| district | 0 | | | | |
| gr_span | 0 | | | | |
| enrl_tot | 420 | 2628.793 | 3913.105 | 81 | 27176 |
| teachers | 420 | 129.0674 | 187.9127 | 4.85 | 1429 |
| calw_pct | 420 | 13.24604 | 11.45482 | 0 | 78.9942 |
| ${\sf meal_pct}$ | 420 | 44.70524 | 27.12338 | Θ | 100 |
| computer | 420 | 303.3833 | 441.3413 | 0 | 3324 |
| testscr | | 654.1565 | 19.05335 | 605.55 | 706.75 |
| comp_stu | 420 | .1359266 | .0649558 | 0 | .4208333 |
| expn_stu | 420 | 5312.408 | 633.9371 | 3926.07 | 7711.507 |
| str | 420 | 19.64043 | 1.891812 | 14 | 25.8 |
| avginc | • | 15.31659 | 7.22589 | 5.335 | 55.328 |
| el_pct | | 15.76816 | 18.28593 | | 85.53972 |
| read_scr | ' | 654.9705 | | 604.5 | |
| math_scr | • | 653.3426 | 18.7542 | 605.4 | 709.5 |

. regress read_scr expn_stu comp_stu meal_pct

| Source | SS | df | MS | Number of obs | = | 420 |
|----------|------------|-----|------------|---------------|---|--------|
| + | | | | F(3, 416) | = | 565.36 |
| Model | 136046.267 | 3 | 45348.7558 | Prob > F | = | 0.0000 |
| Residual | 33368.3632 | 416 | 80.2124115 | R-squared | = | 0.8030 |
| + | | | | Adj R-squared | = | 0.8016 |
| Total | 169414.631 | 419 | 404.330861 | Root MSE | = | 8.9561 |
| | | | | | | |
| | | | | | | |

read_scr | Coef. Std. Err. t P>|t| [95% Conf. Interval]

expn_stu | .0046699 .0007204 6.48 0.000 .0032538 .006086

| comp_stu | 19.88584 | 7.168347 | 2.77 | 0.006 | 5.795143 | 33.97654 |
|------------|----------|----------|--------|-------|----------|----------|
| meal_pct | 635131 | .0164777 | -38.54 | 0.000 | 667521 | 602741 |
| $_{-}cons$ | 655.8528 | 3.812206 | 172.04 | 0.000 | 648.3592 | 663.3464 |

. regress read_scr expn_stu comp_stu meal_pct [aw = enrl_tot] (sum of wgt is 1,104,093)

| Source | SS | df | MS | Number of obs | = | 420 |
|----------|------------|-----|------------|---------------|---|--------|
| +- | | | | F(3, 416) | = | 906.75 |
| Model | 123692.671 | 3 | 41230.8903 | Prob > F | = | 0.0000 |
| Residual | 18915.9815 | 416 | 45.4711093 | R-squared | = | 0.8674 |
| +- | | | | Adj R-squared | = | 0.8664 |
| Total | 142608.652 | 419 | 340.354779 | Root MSE | = | 6.7432 |

 $read_scr \mid \qquad Coef. \quad Std. \; Err. \qquad t \qquad P>|t| \qquad [95\% \; Conf. \; Interval]$ expn_stu | .0055534 .0008322 6.67 0.000 .0039176 .0071892 comp_stu | 27.26378 8.197228 3.33 0.001 11.15063 43.37693 meal_pct | -.6352229 .013149 -48.31 0.000 -.6610696 -.6093762 648.988 4.163875 155.86 0.000 640.8031 _cons | 657.1728

6.3 SERIAL CORRELATION IN THE ERROR DISTRIBUTION

6.3.1 Testing for Serial Correlation

. use ukrates.. describe

Contains data from ukrates.dta

526 obs: 3 vars: 6,312 size:

2 Dec 2004 10:43

storage display value

variable name type format label variable label ______

float %9.0g month float %tm float %9.0g

Sorted by: month

. summarize

| Variable | | Mean | | | |
|----------|-----|----------|----------|----------|-------|
| + | | | | | |
| rs | 526 | 7.651513 | 3.553109 | 1.561667 | 16.18 |
| month | 526 | 168.5 | 151.9874 | -94 | 431 |
| r20 | 526 | 8.863726 | 3.224372 | 3.35 | 17.18 |

| Model | 13.8769739 | 1 | 13.8769739 | F(1, 522) Prob > F | = = | 52.88 0.0000 |
|-----------------|--------------|-----|------------|----------------------------|--------|-----------------|
| Residual | 136.988471 | | | R-squared Adj R-squared | | |
| · | 150.865445 | | | Root MSE | | |
| | | | | P> t [95% Co | | _ |
| r20 | . 4882883 | | | 9.000 .35637 | | |
| | | | | 0.858039955 | | |
| . predict doubl | e ens. resid | ual | | | | |

- predict double eps, residual(2 missing values generated)
- . estat bgodfrey, lags(6)

Breusch-Godfrey LM test for autocorrelation

| | chi2 | Prob > chi2 |
|---|--------|-------------|
| · | 17.237 | |

H0: no serial correlation

. wntestq eps

Portmanteau test for white noise

Portmanteau (Q) statistic = 82.3882 Prob > chi2(40) = 0.0001

. ac eps

6.3.2 FGLS Estimation with Serial Correlation

. use ukrates.. regress D.rs LD.r20

| Source | SS | df | MS | Number of obs | = | 524 |
|----------|------------|-----------|------------|---------------|------|-----------|
| + | | | | F(1, 522) | = | 52.88 |
| Model | 13.8769739 | 1 | 13.8769739 | Prob > F | = | 0.0000 |
| Residual | 136.988471 | 522 | .262430021 | R-squared | = | 0.0920 |
| + | | | | Adj R-squared | = | 0.0902 |
| Total | 150.865445 | 523 | .288461654 | Root MSE | = | .51228 |
| | | | | | | |
| | | | | | | |
| D.rs | Coef. | Std. Err. | t F | P> t [95% (| onf. | Interval] |
| + | | | | | | |
| r20 | | | | | | |
| LD. | .4882883 | .0671484 | 7.27 | .3563 | 74 | .6202027 |

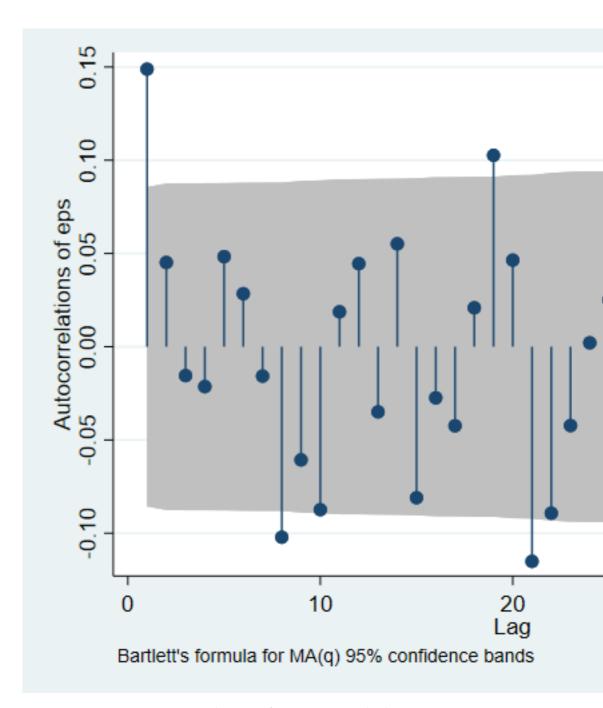


Figure 6: Autocorrelogram of regression residuals

| cons_ | .0040183 | .022384 | 0.18 | 0.858 | 039955 | 55 | .0479921 |
|----------------|---------------|----------|-------------|--------|---------|--------|----------|
| . prais D.rs L | .D.r20, nolog | | | | | | |
| Prais-Winsten | AR(1) regress | ion ite | rated estim | nates | | | |
| Source | SS | | | | | | |
| | 6.56420242 | | | | | | |
| · | 133.146932 | | | | | | |
| + | | | | | | | 0.0452 |
| Total | 139.711134 | 523 | .2671341 | L Root | MSE | = | .50505 |
| | | | | | | | |
| · · | Coef. | | | | | | |
| + | | | | | | | |
| r20 | . 3495857 | 068012 | 5.07 | 0 000 | 21/206 | :7 | 1810617 |
| LD. | . 5495657 | .000912 | 3.07 | 0.000 | .214200 | , , | .4049047 |
| _cons | .0049985 | .0272145 | 0.18 | 0.854 | 048464 | 19 | .0584619 |
| + | | | | | | | |
| rho | . 1895324 | | | | | | |
| | | | | | | | |

Durbin-Watson statistic (original) 1.702273 Durbin-Watson statistic (transformed) 2.007414

REGRESSION WITH INDICATOR VARIABLES

7.1 TESTING FOR SIGNIFICANCE OF A QUALITATIVE FACTOR

7.1.1 Regression with One Qualitative Measure

. use NEdata.. describe

Contains data from NEdata.dta

obs: 120 vars: 6 size: 2,640

: 2,640

24 Oct 2004 13:28

| variable name | storage type | display format | value label | variable label |
|--|---|--|----------------|----------------|
| state year pop dpi dpipc ldpipc | long int float float float float | %8.0g %8.0g %9.0g %9.0g %9.0g %9.0g | state | |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|----------|-----|----------|-----------|----------|----------|
| state | 120 | 3.5 | 1.714986 | 1 | 6 |
| year | 120 | 1990.5 | 5.790459 | 1981 | 2000 |
| pop | 120 | 2196276 | 1931629 | 515594 | 6362076 |
| dpi | 120 | 4.33e+07 | 4.46e+07 | 4385134 | 1.93e+08 |
| dpipc | 120 | 18.15802 | 5.662848 | 8.153382 | 33.38758 |
| + | | | | | |
| ldpipc | 120 | 2.848302 | .3265395 | 2.098433 | 3.508184 |

. mean dpipc, over(state)

Mean estimation Number of obs = 120

CT: state = CT
MA: state = MA
ME: state = ME
NH: state = NH
RI: state = RI
VT: state = VT

| | 0ver | Mean | Std. Err. | [95% Conf. | Interval] |
|-------|------|----------|-----------|------------|-----------|
| dpipc | | | | | |
| | CT | 22.32587 | 1.413766 | 19.52647 | 25.12527 |
| | MA | 19.77681 | 1.298507 | 17.20564 | 22.34798 |
| | ME | 15.17391 | .9571251 | 13.27871 | 17.06911 |
| | NH | 18.66835 | 1.193137 | 16.30582 | 21.03088 |
| | RI | 17.26529 | 1.045117 | 15.19586 | 19.33473 |
| | VT | 15.73786 | 1.020159 | 13.71784 | 17.75788 |
| | | | | | |

. tabulate state, generate(NE)

| state | Freq. | Percent | Cum. |
|------------------------------------|----------------------------------|---|---|
| CT MA ME NH RI | 20 20 20 20 20 20 | 16.67 16.67 16.67 16.67 16.67 | 16.67 33.33 50.00 66.67 83.33 |
| VT | 20 | 16.67 | 100.00 |
| Total | 120 | 100.00 | |

. regress dpipc NE2-NE6

| Source | SS | df | MS | Number of obs | = | 120 |
|----------|------------|-----|------------|---------------|---|--------|
| +- | | | | F(5, 114) | = | 5.27 |
| Model | 716.218512 | 5 | 143.243702 | Prob > F | = | 0.0002 |
| Residual | 3099.85511 | 114 | 27.1917115 | R-squared | = | 0.1877 |
| +- | | | | Adj R-squared | = | 0.1521 |
| Total | 3816.07362 | 119 | 32.0678456 | Root MSE | = | 5.2146 |

| | | | | | | |
|-----------|-----------|-----------|-------|-------|-----------|-----------|
| dpipc | Coef. | Std. Err. | t | P> t | - | Interval] |
| NE2 | -2.549057 | 1.648991 | -1.55 | 0.125 | -5.815695 | .7175814 |
| NE3 | -7.151959 | 1.648991 | -4.34 | 0.000 | -10.4186 | -3.88532 |
| NE4 | -3.65752 | 1.648991 | -2.22 | 0.029 | -6.924158 | 3908815 |
| NE5 | -5.060575 | 1.648991 | -3.07 | 0.003 | -8.327214 | -1.793937 |
| NE6 | -6.588007 | 1.648991 | -4.00 | 0.000 | -9.854646 | -3.321369 |
| _cons | 22.32587 | 1.166013 | 19.15 | 0.000 | 20.01601 | 24.63573 |
| | | | | | | |

. use NEdata.. tabulate state, generate(NE)

| state | Freq. | Percent | Cum. |
|-------|-------|---------|--------|
| | | | |
| CT | 20 | 16.67 | 16.67 |
| MA | 20 | 16.67 | 33.33 |
| ME | 20 | 16.67 | 50.00 |
| NH | 20 | 16.67 | 66.67 |
| RI | 20 | 16.67 | 83.33 |
| VT | 20 | 16.67 | 100.00 |
| | | | |
| Total | 120 | 100.00 | |

```
. forvalues i = 1/5 {
   2.   generate NE_'i' = NE'i' - NE6
   3. }
```

. regress dpipc NE_*

| Source | SS | df | MS | Number of | obs = | 120 |
|----------|------------|----------|------------|------------|--------|----------|
| + | | | | F(5, 114) | = | 5.27 |
| Model | 716.218512 | 5 | 143.243702 | Prob > F | = | 0.0002 |
| Residual | 3099.85511 | 114 | 27.1917115 | R-squared | = | 0.1877 |
| + | | | | Adj R-squa | ared = | 0.1521 |
| Total | 3816.07362 | 119 | 32.0678456 | Root MSE | = | 5.2146 |
| | | | | | | |
| | | | | | | |
| dpipc | | | | P> t [9 | | = |
| + | | | | | | |
| NE_1 | 4.167853 | 1.064419 | 3.92 | 0.000 2. | 059247 | 6.276459 |
| NE_2 | 1.618796 | 1.064419 | 1.52 | 0.131 - | .48981 | 3.727402 |
| NE_3 | -2.984106 | 1.064419 | -2.80 | 0.006 -5. | 092712 | 8754996 |
| NE_4 | .5103331 | 1.064419 | 0.48 | 0.633 -1. | 598273 | 2.618939 |
| NE_5 | 8927223 | 1.064419 | -0.84 | 0.403 -3. | 001328 | 1.215884 |
| _cons | 18.15802 | .4760227 | 38.15 | 0.000 17 | .21502 | 19.10101 |
| | | | | | | |

. $lincom - (NE_1 + NE_2 + NE_3 + NE_4 + NE_5)$

(1) - NE_1 - NE_2 - NE_3 - NE_4 - NE_5 = 0

| dpipc | Coef. | Std. Err. | t | P> t | [95% Conf. | Interval] |
|-------|-----------|-----------|-------|-------|------------|-----------|
| · | -2.420154 | 1.064419 | -2.27 | 0.025 | -4.52876 | 3115483 |

7.1.2 Regression with Two Qualitative Measures

- . use nlsw88.(NLSW, 1988 extract)
- . describe

Contains data from nlsw88.dta

 obs:
 2,246
 NLSW, 1988 extract

 vars:
 17
 21 Jun 2006 11:33

 size:
 60,642
 (_dta has notes)

| variable name | storage type | display format | value label | variable label |
|----------------------------------|-----------------------------|----------------------------------|-------------------|--|
| idcode age race married | int byte byte byte | %8.0g %8.0g %8.0g %8.0g | racelbl marlbl | NLS id age in current year race married |

| <pre>never_married grade</pre> | byte byte | %8.0g %8.0g | | never married current grade completed |
|--------------------------------|--------------|----------------|----------|---------------------------------------|
| collgrad | byte | %16.0g | gradlbl | college graduate |
| south | byte | %8.0g | | lives in south |
| smsa | byte | %9.0g | smsalbl | lives in SMSA |
| $c_{-}city$ | byte | %8.0g | | lives in central city |
| industry | byte | %23.0g | indlbl | industry |
| occupation | byte | %22.0g | occlbl | occupation |
| union | byte | %8.0g | unionlbl | union worker |
| wage | float | %9.0g | | hourly wage |
| hours | byte | %8.0g | | usual hours worked |
| ttl_exp | float | %9.0g | | total work experience |
| tenure | float | %9.0g | | job tenure (years) |

Sorted by: idcode

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|--------------|-------|----------|-----------|----------|----------|
| idcode | 2,246 | 2612.654 | 1480.864 | 1 | 5159 |
| age | 2,246 | 39.15316 | 3.060002 | 34 | 46 |
| race | 2,246 | 1.282725 | .4754413 | 1 | 3 |
| married | | .6420303 | . 4795099 | 0 | 1 |
| never_marr~d | 2,246 | | .3055687 | 0 | 1 |
| +- | | | | | |
| grade | 2,244 | 13.09893 | 2.521246 | Θ | 18 |
| collgrad | 2,246 | .2368655 | .4252538 | 0 | 1 |
| south | 2,246 | .4194123 | . 4935728 | Θ | 1 |
| smsa | 2,246 | .7039181 | . 4566292 | Θ | 1 |
| c_city | 2,246 | .2916296 | .4546139 | Θ | 1 |
| +- | | | | | |
| industry | 2,232 | 8.189516 | 3.010875 | 1 | 12 |
| occupation | 2,237 | 4.642825 | 3.408897 | 1 | 13 |
| union | 1,878 | .2454739 | . 4304825 | 0 | 1 |
| wage | 2,246 | 7.766949 | 5.755523 | 1.004952 | 40.74659 |
| hours | 2,242 | 37.21811 | 10.50914 | 1 | 80 |
| +- | | | | | |
| ttl_exp | 2,246 | 12.53498 | 4.610208 | .1153846 | 28.88461 |
| tenure | 2,231 | 5.97785 | 5.510331 | 0 | 25.91667 |

- . keep if !missing(wage + race + union)
 (368 observations deleted)
- . generate lwage = log(wage)
- . tabulate race, generate(R)

| race | Freq. | Percent | Cum. |
|-----------------------------|--------------------|------------------------|--------------------------|
| white black other | 1,353 501 24 | 72.04 26.68 1.28 | 72.04 98.72 100.00 |
| Total | 1,878 | 100.00 | |

- . test R1 R2
- (1) R1 = 0
- (2) R2 = 0

F(2, 1874) = 23.25Prob > F = 0.0000

7.1.2.1 Interaction Effects

- . use nlsw88.(NLSW, 1988 extract)
- . keep if !missing(wage + race + union)
 (368 observations deleted)
- . generate lwage = log(wage)
- . tabulate race, generate(R)

| race | Freq. | Percent | Cum. |
|-----------------------------|--------------------|------------------------|--------------------------|
| white black other | 1,353 501 24 | 72.04 26.68 1.28 | 72.04 98.72 100.00 |
| Total | 1,878 | 100.00 | |

- . generate R1u = R1 * union
- . generate R2u = R2 * union
- . regress lwage R1 R2 union R1u R2u

| Source | SS | df | MS | Number of obs | = | 1,878 |
|----------|------------|-------|------------|---------------|---|--------|
| +- | | | | F(5, 1872) | = | 26.63 |
| Model | 33.3636017 | 5 | 6.67272035 | Prob > F | = | 0.0000 |
| Residual | 469.09053 | 1,872 | .250582548 | R-squared | = | 0.0664 |
| +- | | | | Adj R-squared | = | 0.0639 |

Total | 502.454132 1,877 .267690001 Root MSE = .50058

| lw | age | | Std. Err. | t | P> t | [95% Conf. | Interval] |
|----|---------|---------|-----------|-------|-------|------------|-----------|
| | R1 | 1818955 | . 1260945 | | 0.149 | 4291962 | .0654051 |
| | R2 | 4152863 | .1279741 | -3.25 | 0.001 | 6662731 | 1642995 |
| un | ion | 2375316 | .2167585 | -1.10 | 0.273 | 6626452 | . 187582 |
| I | R1u . | 4232627 | .2192086 | 1.93 | 0.054 | 0066561 | .8531816 |
| I | R2u . | 6193578 | .2221704 | 2.79 | 0.005 | .1836302 | 1.055085 |
| _C | ons | 2.07205 | . 1251456 | 16.56 | 0.000 | 1.82661 | 2.317489 |
| | | | | | | | |

- . test R1u R2u
- (1) R1u = 0
- (2) R2u = 0

$$F(2, 1872) = 8.04$$

 $Prob > F = 0.0003$

7.2 REGRESSION WITH QUALITATIVE & QUANTITATIVE FACTORS

- . use nlsw88.(NLSW, 1988 extract)
- . keep if !missing(wage + race + union)
 (368 observations deleted)
- . generate lwage = log(wage)
- . tabulate race, generate(R)

| race | Freq. | Percent | Cum. |
|-----------------------------|--------------------|------------------------|--------------------------|
| white black other | 1,353 501 24 | 72.04 26.68 1.28 | 72.04 98.72 100.00 |
| Total | 1,878 | 100.00 | |

. regress lwage R1 R2 union tenure

| Source | SS | df | MS | Number of obs | = | 1,868 |
|----------|------------|-----------|------------|---------------|--------|-----------|
| + | | | | F(4, 1863) | = | 85.88 |
| Model | 77.1526731 | 4 | 19.2881683 | Prob > F | = | 0.0000 |
| Residual | 418.434693 | 1,863 | .224602626 | R-squared | = | 0.1557 |
| + | | | | Adj R-squared | = | 0.1539 |
| Total | 495.587366 | 1,867 | .265445831 | Root MSE | = | .47392 |
| | | | | | | |
| | | | | | | |
| lwage | Coef. | Std. Err. | t F | P> t [95% Co | onf. I | interval] |

| lwage | ı | | | | [95% Conf. | - |
|-------|---------|----------|-------|-------|------------|---------|
| • | | | | | 2619053 | |
| R2 | 2612185 | .0991154 | -2.64 | 0.008 | 4556074 | 0668297 |

union | .1871116 .0257654 7.26 0.000 .1365794 .2376438 tenure | .0289352 .0019646 14.73 0.000 .0250823 .0327882 _cons | 1.777386 .0975549 18.22 0.000 1.586058 1.968715

- . test R1 R2
- (1) R1 = 0
- (2) R2 = 0

$$F(2, 1863) = 29.98$$

 $Prob > F = 0.0000$

7.2.1 Testing for Slope Differences

- . use nlsw88.(NLSW, 1988 extract)
- . keep if !missing(wage + race + union)
 (368 observations deleted)
- . generate lwage = log(wage)
- . tabulate race, generate(R)

| race | Freq. | Percent | Cum. |
|-----------------------------|--------------------|------------------------|--------------------------|
| white black other | 1,353 501 24 | 72.04 26.68 1.28 | 72.04 98.72 100.00 |
| Total | 1,878 | 100.00 | |

- . generate uTen = union * tenure
 (10 missing values generated)
- . regress lwage R1 R2 union tenure uTen

| Source | SS | | MS | Number of obs F(5, 1862) | | 1,868 69.27 |
|-----------------------|--|----------------------------------|---|---|--------------------------|---|
| Model | 77.726069 | 5 | 15.5452138 | Prob > F | | |
| Residual | 417.861297 | 1,862 | .224415304 | R-squared | = | 0.1568 |
| +- | | | | Adj R-squared | = | 0.1546 |
| Total | 495.587366 | 1,867 | .265445831 | Root MSE | = | . 47372 |
| | | | | | | |
| | | | | | | |
| | | | | > t [95% Co | | _ |
| | | | | | | _ |
| +- | 0715443 | | -0.73 0 | | 54 | |
| R1 | 0715443 2638742 | .0976332 | -0.73 0 -2.66 0 | .464263020 | 54 93 | .1199377 |
| R1 R2 | 0715443 2638742 .2380442 | .0976332 .0990879 | -0.73 0 -2.66 0 5.81 0 | | 54 93 91 | .1199377 0695391 .3183975 |
| R1 R2 union | 0715443 2638742 .2380442 .0309616 | .0976332 .0990879 .0409706 | -0.73 0 -2.66 0 5.81 0 13.25 0 | .464263026 .008458209 .000 .15769 | 54 93 91 74 | .1199377 0695391 .3183975 .0355458 |

- . generate R1ten = R1 * tenure
 (10 missing values generated)
- . generate R2ten = R2 * tenure
 (10 missing values generated)
- . regress lwage R1 R2 union tenure R1ten R2ten

| Source | SS | df | MS | | er of obs 1861) | = | 1,868 57.26 |
|----------|------------|-----------|------------|-------|--------------------|-----|----------------|
| Model | 77.2369283 | | | ٠, , | > F | | 0.0000 |
| Residual | 418.350438 | 1,861 | .224798731 | R-sq | uared | = | 0.1558 |
| + | | | | Adj F | R-squared | = | 0.1531 |
| Total | 495.587366 | 1,867 | .265445831 | Root | MSE | = | .47413 |
| | | | | | | | |
| lwage | Coef. | Std. Err. | t I | P> t | [95% Coi | nf. | Interval] |
| R1 | | . 1395 | | 0.553 | 3563459 | | .1908398 |
| R2 | 291495 | .1422361 | -2.05 | 0.041 | 57045 | 4 | 012536 |
| union | .1876079 | .0257915 | 7.27 | 0.000 | .137024 | 6 | .2381912 |
| tenure | .0257611 | .0186309 | 1.38 | 9.167 | 010778 | 5 | .0623007 |
| R1ten | .0024973 | .0187646 | 0.13 | 0.894 | 034304 | 5 | .0392991 |
| R2ten | .0050825 | .018999 | 0.27 | 9.789 | 032179 | 9 | .0423441 |
| _cons | 1.794018 | .1382089 | 12.98 | 0.000 | 1.52295 | 7 | 2.065078 |

- . test R1ten R2ten
- (1) R1ten = 0
- (2) R2ten = 0

$$F(2, 1861) = 0.19$$

 $Prob > F = 0.8291$

. regress lwage R1 R2 union tenure uTen R1ten R2ten

| Source | SS | df | MS | Number of obs | s = | 1,868 |
|----------|------------|-----------|------------|------------------|-----|----------|
| + | | | | F(7, 1860) | = | 49.48 |
| Model | 77.8008722 | 7 | 11.1144103 | Prob > F | = | 0.0000 |
| Residual | 417.786494 | 1,860 | .224616394 | R-squared | = | 0.1570 |
| + | | | | Adj R-squared | = | 0.1538 |
| Total | 495.587366 | 1,867 | .265445831 | Root MSE | = | .47394 |
| | | | | | | |
| | | | | P> t [95% (| | |
| lwage | | | | | | _ |
| R1 l | 0697096 | | | 0.61834366 | | |
| R2 | 2795277 | . 1423788 | | .05055876 | 668 | 0002886 |
| union | .238244 | .0410597 | 5.80 0 | .000 .15771 | L61 | .3187718 |
| tenure | .0304528 | .0188572 | 1.61 0 | .10600653 | 808 | .0674364 |
| uTen | 0068628 | .0043311 | -1.58 0 | .11301535 | 572 | .0016316 |
| R1ten | 0001912 | .0188335 | -0.01 0 | .99203712 | 283 | .0367459 |
| R2ten | .0023429 | .0190698 | 0.12 0 | .90203505 | 576 | .0397433 |

_cons | 1.76904 .1390492 12.72 0.000 1.496331 2.041749

. test uTen R1ten R2ten

- (1) uTen = 0
- (2) R1ten = 0
- (3) R2ten = 0

F(3, 1860) = 0.96Prob > F = 0.4098

. regress lwage union tenure uTen

| Source | SS | df | MS Number of obs | | = | 1,868 |
|----------|------------|-------|------------------|---------------|---|---------|
| +- | | | | F(3, 1864) | = | 92.25 |
| Model | 64.0664855 | 3 | 21.3554952 | Prob > F | = | 0.0000 |
| Residual | 431.52088 | 1,864 | .231502618 | R-squared | = | 0.1293 |
| +- | | | | Adj R-squared | = | 0.1279 |
| Total | 495.587366 | 1,867 | .265445831 | Root MSE | = | . 48115 |

_cons | 1.655054 .0193938 85.34 0.000 1.617018 1.6930

. regress lwage tenure if !union

| Source | SS | df | MS | Number of obs | = | 1,408 |
|----------|------------|-------|------------|---------------|---|--------|
| + | | | | F(1, 1406) | = | 148.43 |
| Model | 36.8472972 | 1 | 36.8472972 | Prob > F | = | 0.0000 |
| Residual | 349.032053 | 1,406 | .248244703 | R-squared | = | 0.0955 |
| + | | | | Adj R-squared | = | 0.0948 |
| Total | 385.87935 | 1,407 | .274256823 | Root MSE | = | .49824 |

| lwage | Coef. | Std. Err. | t | P> t | [95% Conf. | Interval] |
|--------|----------|-----------|-------|-------|----------------------|-----------|
| tenure | .0298926 | .0024536 | 12.18 | 0.000 | .0250795 1.615659 | .0347056 |

. predict double unw if e(sample), res
(470 missing values generated)

. regress lwage tenure if union

 Source | SS
 df
 MS
 Number of obs = 460

 Number of obs = 460
 10.0775663
 F(1, 458)
 = 55.95

 Model | 10.0775663
 1 10.0775663
 Prob > F
 = 0.0000

```
Total | 92.5663941
                  459 .201669704 Root MSE =
                                         .42439
-----
           Coef. Std. Err. t P>|t|
-----
   tenure | .0242707 .0032447
                       7.48 0.000
                                 .0178944
    _cons | 1.869513 .0323515 57.79 0.000
                                 1.805937 1.933088
. predict double nunw if e(sample), res
(1,418 missing values generated)
. generate double allres = nunw
(1,418 missing values generated)
. replace allres = unw if unw <.
(1,408 real changes made)
. sdtest allres, by(union)
Variance ratio test
------
              Mean Std. Err. Std. Dev. [95% Conf. Interval]
 Group |
       0bs
nonunion | 1,408 7.51e-17 .0132735 .4980645 -.0260379
            3.33e-17
                   .0197657
       460
                          .4239271
                                 -.0388425
-----+-----
combined | 1,868 6.48e-17 .0111235
                          .4807605 -.0218157
-----
  ratio = sd(nonunion) / sd(union)
                                      f = 1.3803
Ho: ratio = 1
                            degrees of freedom = 1407, 459
  Ha: ratio < 1
                  Ha: ratio != 1
                                    Ha: ratio > 1
 Pr(F < f) = 1.0000  2*Pr(F > f) = 0.0000
                                  Pr(F > f) = 0.0000
. regress lwage union tenure uTen, robust
                            Number of obs
Linear regression
                                         1.868
                                         109.84
                            F(3, 1864)
                                     =
                            Prob > F
                                     =
                                         0.0000
                            R-squared
                                      =
                                         0.1293
                            Root MSE
                                     =
                                         . 48115
                Robust
          Coef. Std. Err.
    lwage |
                        t P>|t| [95% Conf. Interval]
union | .2144586 .0407254 5.27 0.000
                                .1345864 .2943308
   tenure | .0298926 .0023964 12.47 0.000
                                 .0251928
                                        .0345924
                      -1.46 0.146
    uTen | -.0056219 .0038631
                                 -.0131984
```

_cons | 1.655054 .0210893 78.48 0.000

.0019546

1.613693 1.696415

7.3 SEASONAL ADJUSTMENT WITH INDICATOR VARIABLES

. use turksales.. describe

Contains data from turksales.dta

320

obs: 40 vars: 2

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.....

storage display value
variable name type format label variable label

t float %tq

sales float %9.0g

Sorted by: t

size:

. summarize

| Variable | 0bs | Mean | Std. Dev | . Min | Max |
|----------|-----|----------|----------|----------|----------|
| t | | | 11.69045 | | |
| sales | 40 | 105.6178 | 4.056961 | 97.84603 | 112.9617 |

- . summarize sales, meanonly
- . generate mu = r(mean)
- . forvalues i=1/3 {
- 2. generate qseas'i'= (quarter(dofq(t)) == 'i')
- 3. }
- . regress sales qseas*

| Source | SS | df | MS | Number of obs | = | 40 |
|----------|------------|----|------------|---------------|---|--------|
| +- | | | | F(3, 36) | = | 4.03 |
| Model | 161.370376 | 3 | 53.7901254 | Prob > F | = | 0.0143 |
| Residual | 480.52796 | 36 | 13.3479989 | R-squared | = | 0.2514 |
| +- | | | | Adj R-squared | = | 0.1890 |
| Total | 641.898336 | 39 | 16.4589317 | Root MSE | = | 3.6535 |

| sales | | Std. Err. | | | - | - |
|----------|-----------|-----------|-------|-------|-----------|-----------|
| | | | | | | |
| gseas1 | -5.232047 | 1.633891 | -3.20 | 0.003 | -8.545731 | -1.918362 |
| • | | | | | | |
| qseas2 | -2.842753 | 1.633891 | -1.74 | 0.090 | -6.156437 | .4709317 |
| L Eacesn | 8969368 | 1.633891 | -0.55 | 0.586 | -4.210621 | 2.416748 |
| 436033 | 0909300 | 1.055091 | -0.55 | 0.500 | -4.210021 | 2.410/40 |
| _cons | 107.8608 | 1.155335 | 93.36 | 0.000 | 105.5177 | 110.2039 |
| | | | | | | |

. predict double salesSA, residual

. replace salesSA = salesSA + mu
(40 real changes made)

. summarize sales salesSA

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|----------|----------|----------|-----------|----------|----------|
| sales | 40 | | 4.056961 | | |
| salesSA | 40 | 105.6178 | 3.510161 | 97.49429 | 111.9563 |

- . label var salesSA "sales, seasonally adjusted"
- . tsline sales salesSA, lpattern(solid dash)

. regress sales qseas* t

| Source | SS | df | MS | Number of obs | = | 40 |
|----------|------------|----|------------|---------------|---|--------|
| +- | | | | F(4, 35) | = | 54.23 |
| Model | 552.710487 | 4 | 138.177622 | Prob > F | = | 0.0000 |
| Residual | 89.1878487 | 35 | 2.54822425 | R-squared | = | 0.8611 |
| +- | | | | Adj R-squared | = | 0.8452 |
| Total | 641.898336 | 39 | 16.4589317 | Root MSE | = | 1.5963 |

- . test qseas1 qseas2 qseas3
- (1) qseas1 = 0
- (2) qseas2 = 0
- (3) qseas3 = 0

$$F(3, 35) = 15.17$$

 $Prob > F = 0.0000$

- . predict double salesSADT, residual
- . replace salesSADT = salesSADT + mu
 (40 real changes made)
- . label var salesSADT "sales, detrende and SA"
- . tsline sales salesSADT, lpattern(solid dash) yline('mu')

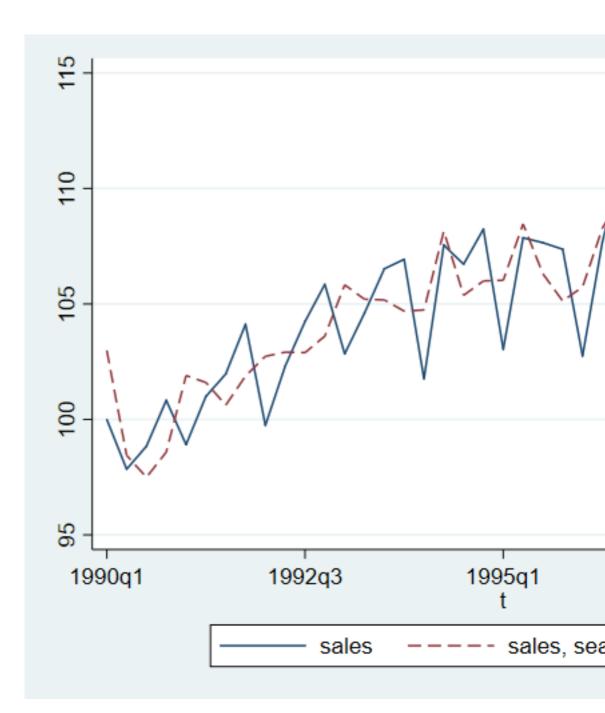


Figure 7: Seasonal adjustment of time series



Figure 8: Seasonal adjustment and deterending of time series

7.4 TESTING FOR STRUCTURAL STABILITY & STRUCTURAL CHANGE

7.4.1 Constraints of Continuity & Differentiability

. use nlsw88.(NLSW, 1988 extract)

. describe

Contains data from nlsw88.dta

 obs:
 2,246
 NLSW, 1988 extract

 vars:
 17
 21 Jun 2006 11:33

 size:
 60,642
 (_dta has notes)

| variable name | • | display format | | variable label |
|---|---|--|--|---|
| idcode age race married never_married grade collgrad south smsa c_city industry occupation union wage hours ttl_exp | byte byte byte byte byte byte byte byte | %8.0g %8.0g %8.0g %16.0g %8.0g %9.0g %8.0g %23.0g %22.0g %8.0g %9.0g | gradlbl smsalbl indlbl occlbl | married never married current grade completed college graduate lives in south lives in SMSA lives in central city industry occupation union worker hourly wage usual hours worked total work experience |
| tenure | float | %9.0g | | job tenure (years) |

Sorted by: idcode

. summarize

| Variable | 0bs | | Std. Dev. | Min | Max |
|--------------|-------|-----------|-----------|-----|------|
| idcode | 2,246 | 2612.654 | 1480.864 | 1 | 5159 |
| age | 2,246 | 39.15316 | 3.060002 | 34 | 46 |
| race | 2,246 | 1.282725 | .4754413 | 1 | 3 |
| married | 2,246 | .6420303 | .4795099 | 0 | 1 |
| never_marr~d | 2,246 | .1041852 | .3055687 | 0 | 1 |
| + | | | | | |
| grade | 2,244 | 13.09893 | 2.521246 | Θ | 18 |
| collgrad | 2,246 | . 2368655 | .4252538 | Θ | 1 |
| south | 2,246 | .4194123 | .4935728 | Θ | 1 |
| smsa | 2,246 | .7039181 | .4566292 | Θ | 1 |
| $c_{-}city$ | 2,246 | .2916296 | .4546139 | Θ | 1 |
| + | | | | | |
| industry | 2,232 | 8.189516 | 3.010875 | 1 | 12 |
| occupation | 2,237 | 4.642825 | 3.408897 | 1 | 13 |
| union | 1,878 | . 2454739 | .4304825 | Θ | 1 |

```
wage | 2,246 7.766949 5.755523 1.004952 40.74659
hours | 2,242 37.21811 10.50914 1 80
ttl_exp | 2,246 12.53498 4.610208 .1153846 28.88461
tenure | 2,231 5.97785 5.510331 0 25.91667
```

- . generate lwage = log(wage)
- . generate Ten2 = tenure <= 2
- . generate Ten7 = !Ten2 & tenure <= 7
- . generate Ten12 = !Ten2 & !Ten7 & tenure <= 12
- . generate Ten25 = !Ten2 & !Ten7 & !Ten12 & tenure < .

. generate tTen2 = tenure * Ten2
(15 missing values generated)

. generate tTen7 = tenure * Ten7
(15 missing values generated)

. generate tTen12 = tenure * Ten12
(15 missing values generated)

. generate tTen25 = tenure * Ten25
(15 missing values generated)

. regress lwage Ten* tTen*, nocons hascons

| Source | SS | df | MS | Number of obs | = | 2,231 |
|----------|------------|-------|------------|---------------|---|--------|
| +- | | | | F(7, 2223) | = | 37.12 |
| Model | 76.6387069 | 7 | 10.9483867 | Prob > F | = | 0.0000 |
| Residual | 655.578361 | 2,223 | .294907045 | R-squared | = | 0.1047 |
| +- | | | | Adj R-squared | = | 0.1018 |
| Total | 732.217068 | 2,230 | .328348461 | Root MSE | = | .54305 |

| | | | | | | |
|-----------|----------|-----------|-------|-------|------------|-----------|
| lwage | Coef. | Std. Err. | t | P> t | [95% Conf. | Interval] |
| Ten2 | 1.55662 | .0383259 | 40.62 | 0.000 | 1.481462 | 1.631778 |
| Ten7 | 1.708728 | .060084 | 28.44 | 0.000 | 1.590901 | 1.826554 |
| Ten12 | 1.870808 | .1877798 | 9.96 | 0.000 | 1.502566 | 2.23905 |
| Ten25 | 1.751961 | .1691799 | 10.36 | 0.000 | 1.420194 | 2.083728 |
| tTen2 | .0897426 | .0331563 | 2.71 | 0.007 | .0247221 | .1547631 |
| tTen7 | .0434089 | .0140739 | 3.08 | 0.002 | .0158095 | .0710083 |
| tTen12 | .0154208 | .019786 | 0.78 | 0.436 | 0233801 | .0542218 |
| tTen25 | .0238014 | .0102917 | 2.31 | 0.021 | .0036191 | .0439837 |
| | | | | | | |

. predict double lwagehat
(option xb assumed; fitted values)
(15 missing values generated)

•

- . label var lwagehat "Predicted log(wage)"
- . sort tenure

. twoway (line lwagehat tenure if tenure <= 2) (line lwagehat tenure if tenure
> > 2 & tenure <= 7) (line lwagehat tenure if tenure > 7 & tenure <= 12) (lin
> e lwagehat tenure if tenure > 12 & tenure < .), legend(off)</pre>

. mkspline sTen2 2 sTen7 7 sTen12 12 sTen25 = tenure

. regress lwage sTen*

| F(4, 2226) = 64 Model 76.1035947 |
|--|
| Model I 76 1035947 4 19 0258987 Prob > $F = 0.0$ |
| 10000 70.1055547 4 15.0250507 1100 7 1 - 0.0 |
| Residual 656.113473 2,226 .294749988 R-squared = 0.3 |
| Adj R-squared = 0.1 |
| Total 732.217068 2,230 .328348461 Root MSE = .54 |

| lwage | • | | | P> t | - | . Interval] |
|--------|----------|----------|-------|-------|----------|-------------|
| | + | | | | | |
| sTen2 | .1173168 | .0248619 | 4.72 | 0.000 | .0685619 | .1660716 |
| sTen7 | .0471177 | .009448 | 4.99 | 0.000 | .02859 | .0656455 |
| sTen12 | .0055041 | .0111226 | 0.49 | 0.621 | 0163076 | .0273158 |
| sTen25 | .0237767 | .0083618 | 2.84 | 0.005 | .007379 | .0401744 |
| _cons | 1.539985 | .0359605 | 42.82 | 0.000 | 1.469465 | 1.610505 |
| | | | | | | |

. predict double lwageSpline(option xb assumed; fitted values)

(15 missing values generated)

- . label var lwageSpline "Predicted log(wage), splined"
- . twoway line lwageSpline tenure

7.4.2 Structural Change in a Time Series Model

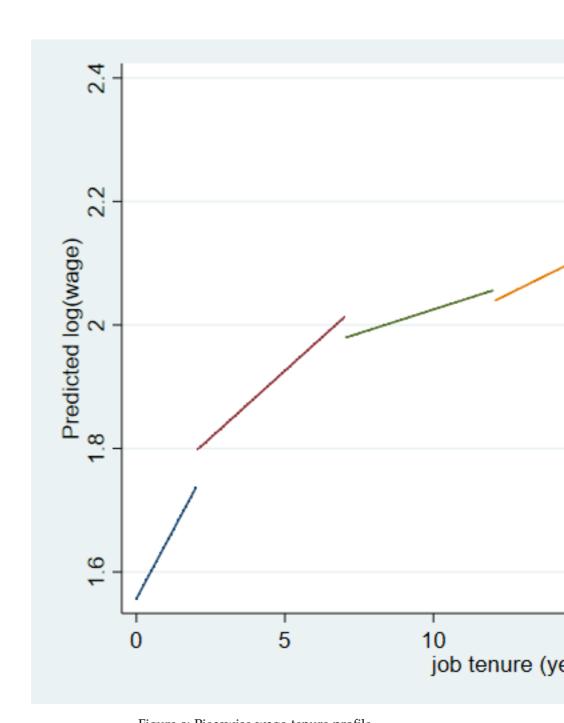


Figure 9: Piecewise wage-tenure profile

Predicted log(wage), splined 10 15 job tenure (years) 5

Figure 10: Piecewise linear wage-tenure profile

INSTRUMENTAL VARIABLES ESTIMATORS

- 8.1 INTRODUCTION
- 8.2 ENDOGENEITY IN ECONOMIC RELATIONSHIPS
- 8.3 2SLS
- 8.4 THE IVREG COMMAND
- 8.5 IDENTIFICATION & TESTS OF OVERIDENTIFYING RESTRICTIONS
- 8.6 COMPUTING IV ESTIMATES
- . use griliches.(Wages of Very Young Men, Zvi Griliches, J.Pol.Ec. 1976)
- . describe

Contains data from griliches.dta

obs: 758 Wages of Very Young Men, Zvi Griliches, J.Pol.Ec. 1976

GITTETIES, J.POT.EC. 15

vars: 26 31 Oct 2004 14:12

size: 65,188

| | storage | display | value | |
|---------------|---------|---------|-------|--------------------------------|
| variable name | type | format | label | variable label |
| | | | | |
| rns | float | %9.0g | | residency in South |
| rns80 | float | %9.0g | | |
| mrt | float | %9.0g | | marital status = 1 if married |
| mrt80 | float | %9.0g | | |
| smsa | float | %9.0g | | reside metro area = 1 if urban |
| smsa80 | float | %9.0g | | |
| med | float | %9.0g | | mother's education, years |
| iq | float | %9.0g | | iq score |
| kww | float | %9.0g | | score on knowledge in world of |
| | | | | work test |
| year | float | %9.0g | | |
| age | float | %9.0g | | |
| age80 | float | %9.0g | | |
| S | float | %9.0g | | completed years of schooling |
| s80 | float | %9.0g | | |
| expr | float | %9.0g | | experience, years |
| expr80 | float | %9.0g | | |
| tenure | float | %9.0g | | tenure, years |
| tenure80 | float | %9.0g | | - |
| | | - | | |

| float | %9.0g | log wage |
|-------|---|---|
| float | %9.0g | |
| byte | %8.0g | year==67 |
| byte | %8.0g | year==68 |
| byte | %8.0g | year==69 |
| byte | %8.0g | year==70 |
| byte | %8.0g | year==71 |
| byte | %8.0g | year==73 |
| | float byte byte byte byte byte | float %9.0g byte %8.0g byte %8.0g byte %8.0g byte %8.0g byte %8.0g byte %8.0g |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|-----------|-----|----------|-----------|-------|--------|
| rns | 758 | .2691293 | .4438001 | | 1 |
| rns80 | 758 | .292876 | . 4553825 | 0 | 1 |
| mrt | 758 | .5145119 | .5001194 | 0 | 1 |
| mrt80 | 758 | .8984169 | .3022988 | 0 | 1 |
| smsa | 758 | .7044855 | . 456575 | 0 | 1 |
| smsa80 | 758 | .7124011 | . 452942 | 0 | 1 |
| med | 758 | 10.91029 | 2.74112 | 0 | 18 |
| iq | 758 | 103.8562 | 13.61867 | 54 | 145 |
| kww | 758 | 36.57388 | 7.302247 | 12 | 56 |
| year | 758 | 69.03166 | 2.631794 | 66 | 73 |
| age | 758 | 21.83509 | 2.981756 | 16 | 30 |
| age80 | 758 | 33.01187 | 3.085504 | 28 | 38 |
| s | 758 | 13.40501 | 2.231828 | 9 | 18 |
| s80 | 758 | 13.70712 | 2.214693 | 9 | 18 |
| expr | 758 | 1.735429 | 2.105542 | 0 | 11.444 |
| expr80 | 758 | 11.39426 | 4.210745 | .692 | 22.045 |
| tenure | 758 | 1.831135 | 1.67363 | Θ | 10 |
| tenure80 | 758 | 7.362797 | 5.05024 | 0 | 22 |
| lw | 758 | 5.686739 | . 4289494 | 4.605 | 7.051 |
| lw80 | 758 | 6.826555 | .4099268 | 4.749 | 8.032 |
| _Iyear_67 | 758 | .0831135 | .2762359 | 0 | 1 |
| _Iyear_68 | 758 | .1042216 | .3057496 | Θ | 1 |
| _Iyear_69 | 758 | .1121372 | .3157435 | Θ | 1 |
| _Iyear_70 | 758 | .0844327 | .2782193 | Θ | 1 |
| _Iyear_71 | 758 | .121372 | .3267747 | 0 | 1 |
| _Iyear_73 | 758 | .2084433 | .4064636 | 0 | 1 |

[.] use griliches.(Wages of Very Young Men, Zvi Griliches, J.Pol.Ec. 1976)

. ivreg lw s expr tenure rns smsa $_{\rm I}$ * (iq=med kww age mrt), first

First-stage regressions

| Source | SS | df | MS | Number of obs | = | 758 |
|--------|----|----|----|---------------|---|-------|
| + | | | | F(15, 742) | = | 25.03 |

| Model | 47176.4676 | 15 | 3145.09784 | Prob > | · F = | 0.0000 |
|------------------------------|------------|-----------|------------|--------|----------------|-----------|
| Residual | 93222.8583 | 742 | 125.637275 | R-squa | red = | 0.3360 |
| + | | | | Adj R- | squared = | 0.3226 |
| Total | 140399.326 | 757 | 185.468066 | Root M | ISE = | 11.209 |
| | | | | | | |
| iq | Coof | Std Err | + | D- + | [95% Conf. | Intervall |
| ±4 | | Jtu. Lii. | | | | |
| s | 2.497742 | . 2858159 | 8.74 | 0.000 | 1.936638 | 3.058846 |
| expr | 033548 | . 2534458 | -0.13 | 0.895 | 5311042 | .4640082 |
| tenure | .6158215 | .2731146 | 2.25 | 0.024 | .0796522 | 1.151991 |
| rns | -2.610221 | .9499731 | -2.75 | 0.006 | -4.475177 | 7452663 |
| smsa | .0260481 | .9222585 | 0.03 | 0.977 | -1.784499 | 1.836595 |
| _Iyear_67 | . 9254935 | 1.655969 | 0.56 | 0.576 | -2.325449 | 4.176436 |
| _Iyear_68 | .4706951 | 1.574561 | 0.30 | 0.765 | -2.620429 | 3.56182 |
| $_{ m L}$ Iyear $_{ m L}$ 69 | 2.164635 | 1.521387 | 1.42 | 0.155 | 8221007 | 5.15137 |
| $_{ m _Iyear_70}$ | 5.734786 | 1.696033 | 3.38 | 0.001 | 2.405191 | 9.064381 |
| _Iyear_71 | 5.180639 | 1.562156 | 3.32 | 0.001 | 2.113866 | 8.247411 |
| _Iyear_73 | 4.526686 | 1.48294 | 3.05 | 0.002 | 1.615429 | 7.437943 |
| med | . 2877745 | .1622338 | 1.77 | 0.077 | 0307176 | .6062665 |
| kww | . 4581116 | .0699323 | 6.55 | 0.000 | .3208229 | .5954003 |
| age | 8809144 | . 2232535 | -3.95 | 0.000 | -1.319198 | 4426307 |
| mrt | 584791 | .946056 | -0.62 | 0.537 | -2.442056 | 1.272474 |
| _cons | 67.20449 | 4.107281 | 16.36 | 0.000 | 59.14121 | 75.26776 |

Instrumental variables (2SLS) regression

| Source | SS | df | MS | Number of obs | = | 758 |
|-----------|------------|----------|-------------|---------------|------|-----------|
| + | | | | F(12, 745) | = | 45.91 |
| Model | 59.2679161 | 12 | 4.93899301 | Prob > F | = | 0.0000 |
| Residual | 80.0182337 | 745 | .107407025 | R-squared | = | 0.4255 |
| + | | | | Adj R-squared | = | 0.4163 |
| Total | 139.28615 | 757 | . 183997556 | Root MSE | = | .32773 |
| | | | | | | |
| lw | | | | P> t [95% C | onf. | Interval] |
| | .0001747 | | | | 51 | .0079044 |
| s | .0691759 | .013049 | 5.30 | 0.000 .04355 | 87 | .0947931 |
| expr | .029866 | .006697 | 4.46 | 0.000 .01671 | 89 | .0430132 |
| tenure | .0432738 | .0076934 | | 0.000 .02817 | 95 | . 058377 |
| rns | 1035897 | .0297371 | -3.48 | 0.00116196 | 82 | 0452111 |
| smsa | .1351148 | .0268889 | 5.02 | 0.000 .08232 | 77 | .1879019 |
| _Iyear_67 | 052598 | .0481067 | -1.09 | 0.27514703 | 88 | .0418428 |
| _Iyear_68 | .0794686 | .0451078 | 1.76 | 0.0790090 | 85 | .1680222 |
| _Iyear_69 | .2108962 | .0443153 | 4.76 | 0.000 .12389 | 84 | . 2978939 |
| _Iyear_70 | .2386338 | .0514161 | 4.64 | 0.000 .13769 | 62 | .3395714 |
| _Iyear_71 | .2284609 | .0441236 | 5.18 | 0.000 .14183 | 96 | .3150823 |
| _Iyear_73 | .3258944 | .0410718 | 7.93 | 0.000 .24526 | 42 | .4065247 |
| _cons | 4.39955 | .2708771 | 16.24 | 0.000 3.8677 | 77 | 4.931323 |
| | | | | | | |

Instrumented: iq

Instruments: s expr tenure rns smsa _Iyear_67 _Iyear_68 _Iyear_69

_Iyear_70 _Iyear_71 _Iyear_73 med kww age mrt

- 8.7 IVREG2 & GMM ESTIMATION
- 8.7.1 *The GMM Estimator*
- 8.7.2 GMM in a Homoskedastic Context
- 8.7.3 GMM & Heteroskedasticity-Consistent Standard Errors
- . use griliches.(Wages of Very Young Men, Zvi Griliches, J.Pol.Ec. 1976)
- . ivreg2 lw s expr tenure rns smsa $_{\rm I}*$ (iq=med kww age mrt), gmm command ivreg2 is unrecognized r(199);

end of do-file
r(199);

- 8.7.4 GMM & Clustering
- 8.7.5 GMM & HAC Standard Errors
- 8.8 TESTING & OVERIDENTIFYING RESTRICTIONS IN GMM
- 8.8.1 Testing a subset of the Overidentifying Restrictions in GMM
- 8.9 TESTING FOR HETEROSKEDASTICITY IN THE IV CONTEXT
- 8.10 TESTING THE RELEVANCE OF INSTRUMENTS
- 8.11 DURBIN-WU-HAUSMAN TESTS FOR ENDOGENEITY IN IV ESTIMATION

PANEL DATA MODELS

9.1 FIXED EFFECTS & RANDOM EFFECTS MODELS

9.1.1 One-Way Fixed Effects Models

. use traffic.. describe

Contains data from traffic.dta

obs: 336 vars: 54 size: 61,152

30 Nov 2004 10:23

| | storage | display | value | |
|---------------|---------|---------|-------|--|
| variable name | type | format | | |
| state | float | %9.0g | | State ID (FIPS) Code |
| year | int | %9.0g | | Year |
| spircons | float | %9.0g | | Spirits Consumption |
| unrate | float | %9.0g | | Unemployment Rate |
| perinc | float | %9.0g | | Per Capita Personal Income |
| emppop | float | %9.0g | | Employment/Population Ratio |
| beertax | float | %9.0g | | Tax on Case of Beer |
| sobapt | float | %9.0g | | % Southern Baptist |
| mormon | float | %9.0g | | % Mormon |
| mlda | float | %9.0g | | Minimum Legal Drinking Age |
| dry | float | %9.0g | | % Residing in Dry Counties |
| yngdrv | float | %9.0g | | % of Drivers Aged 15-24 |
| vmiles | float | %9.0g | | Ave. Mile per Driver |
| breath | byte | %9.0g | | Prelim. Breath Test Law |
| jaild | byte | %9.0g | | Mandatory Jail Sentence |
| comserd | byte | %9.0g | | Mandatory Community Service |
| allmort | int | %9.0g | | <pre># of Vehicle Fatalities (#VF)</pre> |
| mrall | float | %9.0g | | Vehicle Fatality Rate (VFR) |
| allnite | int | %9.0g | | <pre># of Night-time VF (#NVF)</pre> |
| mralln | float | %9.0g | | Night-time VFR (NFVR) |
| allsvn | int | %9.0g | | <pre># of Single VF (#SVN)</pre> |
| a1517 | int | %9.0g | | #VF, 15-17 year olds |
| mra1517 | float | %9.0g | | VFR, 15-17 year olds |
| a1517n | byte | %9.0g | | #NVF, 15-17 year olds |
| mra1517n | float | %9.0g | | NVFR, 15-17 year olds |
| a1820 | int | %9.0g | | #VF, 18-20 year olds |
| a1820n | int | %9.0g | | #NVF, 18-20 year olds |
| mra1820 | float | %9.0g | | VFR, 18-20 year olds |
| mra1820n | float | %9.0g | | NVFR, 18-20 year olds |
| a2124 | int | %9.0g | | #VF, 21-24 year olds |
| mra2124 | float | %9.0g | | VFR, 21-24 year olds |
| a2124n | int | %9.0g | | #NVF, 21-24 year olds |
| mra2124n | float | %9.0g | | NVFR, 21-24 year olds |

| aidall | float | %9.0q | | # of alcohol-involved VF |
|------------|-------|-------|-----|-------------------------------|
| mraidall | float | %9.0q | | Alcohol-Involved VFR |
| pop | float | %9.0q | | Population |
| pop1517 | float | %9.0q | | Population, 15-17 year olds |
| pop1820 | float | %9.0q | | Population, 18-20 year olds |
| pop2124 | float | %9.0g | | Population, 21-24 year olds |
| miles | float | %9.0g | | total vehicle miles (millions |
| unus | float | %9.0g | | U.S. unemployment rate |
| epopus | float | %9.0q | | U.S. Emp/Pop Ratio |
| gspch | float | %9.0q | | GSP Rate of Change |
| stabrv | str2 | %9s | | |
| region | str2 | %9s | | |
| reg | long | %8.0q | reg | |
| fatal | float | %9.0q | J | |
| fatal1517 | float | %9.0g | | |
| fatal1820 | float | %9.0g | | |
| fatal2124 | float | %9.0g | | |
| fatal1517n | float | %9.0g | | |
| fatal1820n | float | %9.0g | | |
| fatal2124n | float | %9.0g | | |
| perincK | float | %9.0g | | |
| | | - | | |

.....

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max | |
|----------|-------|----------|-----------|----------|----------|--|
| state | | 30.1875 | 15.30985 | 1 | 56 | |
| year | • | 1985 | 2.002983 | 1982 | 1988 | |
| spircons | l 336 | 1.75369 | .6835745 | .79 | 4.9 | |
| unrate | • | 7.346726 | 2.533405 | 2.4 | 18 | |
| perinc | • | 13880.18 | 2253.046 | 9513.762 | | |
| per inc | | | | | | |
| emppop | 336 | 60.80568 | 4.721656 | 42.9932 | 71.26865 | |
| beertax | 336 | .513256 | . 4778442 | .0433109 | 2.720764 | |
| sobapt | 336 | 7.156925 | 9.762621 | 0 | 30.3557 | |
| mormon | 336 | 2.801933 | 9.665279 | .1 | 65.9165 | |
| mlda | 336 | 20.45563 | .8990255 | 18 | 21 | |
| + | | | | | | |
| dry | 336 | 4.267074 | 9.500901 | 0 | 45.7921 | |
| yngdrv | 336 | .1859299 | .0248736 | .073137 | .281625 | |
| vmiles | 336 | 7890.754 | 1475.659 | 4576.346 | 26148.27 | |
| breath | 336 | .4613095 | . 4992443 | 0 | 1 | |
| jaild | 335 | .280597 | . 449963 | 0 | 1 | |
| + | | | | | | |
| comserd | 335 | .1850746 | . 388939 | 0 | 1 | |
| allmort | 336 | 928.6637 | 934.0515 | 79 | 5504 | |
| mrall | 336 | .000204 | .000057 | .0000821 | .0004218 | |
| allnite | 336 | 182.5833 | 188.4311 | 13 | 1049 | |
| mralln | 336 | .0000388 | .000011 | .0000172 | .0000944 | |
| + | | | | | | |
| allsvn | 336 | 109.9494 | 108.5397 | 8 | 603 | |
| a1517 | 336 | 62.61012 | 55.72909 | 3 | 318 | |
| mra1517 | 336 | .0003034 | .0000937 | .0001163 | .0006735 | |
| a1517n | 336 | 12.2619 | 12.25341 | Θ | 76 | |
| mra1517n | 336 | .0000598 | .000033 | Θ | .0002571 | |

336 106.6607 104.2236 7 a1820 | 601 336 33.52679 33.23834 a1820n | 0 196 336 .0004728 .0001522 .0001855 .0010952 mra1820 | mra1820n | 336 .0001436 .0000613 0 .0005238 336 126.872 131.7886 a2124 | 12 ------

 336
 .0004091
 .0001225
 .0002
 .0008922

 336
 41.37798
 42.93031
 1
 249

 mra2124 | a2124n | 336 .0001284 .0000422 .0000222 .0003143 336 293.3332 303.5807 24.6 2094.9 mra2124n | aidall | mraidall | 336 .0000659 .000026 .0000234 .0001772 4930272 5073704 478999.7 2.83e+07 336 pop | 336 230815.5 229896.3 21000.02 1172000 pop1517 | 336 249090.4 249345.6 20999.96 1321004 pop1820 | 336 336389.9 345304.4 30000.16 1892998 pop2124 | miles | 336 37101.49 37454.37 241575 3993 336 7.528571 1.479376 unus | 5.5 9.7 epopus | 336 59.97143 1.585048 57.8 336 .0253135 .0431732 -.1236415 .1423609 gspch | stabrv | 0 0 region | ------1 336 2.395833 1.16971 reg | 336 2.040444 .5701938 fatal | .82121 4.21784 336 3.033664 .9374229 1.16279 6.73469 fatal1517 | 336 4.727677 1.521962 1.854922 10.9524 fatal1820 | 336 4.091315 1.224812 2 8.92157 fatal2124 | -----+-----.597681 .3297287 fatal1517n | 336 0 2.57143 0 336 1.436218 .6128804 fatal1820n | 5.2381 336 1.284335 .4224502 .22222 3.14286 fatal2124n |

. xtsum fatal beertax spircons unrate perincK state year

perincK |

| Variable | | | | Std. Dev. | | | | | |
|----------|---------|------|----------|-----------|----------|----------|-------|---|-----|
| | • | | | .5701938 | | · | | | |
| | between | | | .5461407 | 1.110077 | 3.653197 | n | = | 48 |
| | within | | | . 1794253 | 1.45556 | 2.962664 | Т | = | 7 |
| beertax | overall | | .513256 | . 4778442 | .0433109 | 2.720764 | N | = | 336 |
| | between | | | .4789513 | .0481679 | 2.440507 | n | = | 48 |
| | within | l | | .0552203 | .1415352 | .7935126 | Т | = | 7 |
| | | | | | | | | | |
| spircons | overall | | 1.75369 | .6835745 | .79 | 4.9 | N | = | 336 |
| | between | | | .6734649 | .8614286 | 4.388572 | n | = | 48 |
| | within | | | .147792 | 1.255119 | 2.265119 | Т | = | 7 |
| | | | | | | | | | |
| unrate | overall | 7 | 7.346726 | 2.533405 | 2.4 | 18 | N | = | 336 |
| | between | | | 1.953377 | 4.1 | 13.2 | n | = | 48 |
| | within | | | 1.634257 | 4.046726 | 12.14673 | T | = | 7 |
| | | | | | | | | | |
| perincK | overall | : | 13.88018 | 2.253046 | 9.513762 | 22.19345 | N | = | 336 |

336 13.88018 2.253046 9.513762 22.19345

| | between | | 2.122712 | 9.9 | 5087 1 | 19.51582 | r | 1 = | 48 |
|--|--------------------|-------------|---------------|---------|----------|--------------|---------|------------|---------|
| | within | | .8068546 | 11.43 | 3261 1 | 16.55782 | | Γ = | 7 |
| | | | | | | | | | |
| state | overall | 30.1875 | 15.30985 | | 1 | 56 | 1 | 1 = | 336 |
| | between | | 15.44883 | | 1 | 56 | 1 | 1 = | 48 |
| | within | | 0 | 30.3 | 1875 | 30.1875 | 1 7 | Γ = | 7 |
| | | | | | | | ! . | | |
| year | overall | • | | | 1982 | 1988 | • | I = | 336 |
| | between | | 0 2.002983 | | 1985 | 1985 | • |) = - | 48 7 |
| | within | | 2.002983 | | 1982 | 1988 | I | Γ = | , |
| | | | | | | | | | |
| | | | | | | | | | |
| . use tr | affic x | treg fatal | beertax spi | rcons u | nrate pe | erincK, f | е | | |
| | | | | | | | | | |
| · · · · · · · · · · · · · · · · · · · | | | | | | | | | 336 |
| Group variable: state Number of groups = 48 | | | | | | | | | |
| D. ca. | | | | | | | | | |
| R-sq: Obs per group: within = 0.3526 min = 7 | | | | | | | | | |
| | ween = 0.3 | | | | | | avg = | | 7.0 |
| | erall = 0.0 | | | | | | max = | | 7.0 |
| | | | | | | • | | | • |
| | | | | | F(4,28 | 34) | = | | 38.68 |
| corr(u_i | , Xb) = | -0.8804 | | | Prob > | > F | = | | 0.0000 |
| | | | | | | | | | |
| | | | | | | | | | |
| | • | | Std. Err. | t | P> t | [95% | Conf. | Int | terval] |
| | + ertax - | 4840728 | .1625106 | -2.98 | 0.003 | 803 | 9508 | | 1641948 |
| | • | .8169652 | .0792118 | 10.31 | 0.000 | | | | 9728819 |
| • | | | .0090274 | -3.22 | 0.001 | | 8191 | | 9112808 |
| | incK | | .0205986 | 5.08 | 0.000 | | 4165 | | 1452555 |
| • | • | 383783 | .4201781 | -0.91 | 0.362 | -1.21 | | | 4432754 |
| | | | | | | | | | |
| sig | ıma_u 1 | . 1181913 | | | | | | | |
| sig | gma_e . ∶ | 15678965 | | | | | | | |
| | rho .9 | 98071823 | (fraction o | f varia | nce due | to $u_{-}i)$ | | | |
| | | | | | | | | | |
| F test t | :hat all u | _i=0: F(47, | 284) = 59. | 77 | | Р | rob > I | = | 0.0000 |

9.1.2 Time Effects & Two-Way Fixed Effects Models

. use traffic.. tabulate year, generate(yr)

| Year | Freq. | Percent | Cum. |
|------|-------|---------|--------|
| 1982 | 48 | 14.29 | 14.29 |
| 1983 | 48 | 14.29 | 28.57 |
| 1984 | 48 | 14.29 | 42.86 |
| 1985 | 48 | 14.29 | 57.14 |
| 1986 | 48 | 14.29 | 71.43 |
| 1987 | 48 | 14.29 | 85.71 |
| 1988 | 48 | 14.29 | 100.00 |
| + | | | |

```
Total | 336 100.00
```

(6) yr87 = 0

```
. local j 0
. forvalues i=82/87 {
 2. local ++j
 rename yr'j' yr'i'
 4. quietly replace yr'i' = yr'i' - yr7
 5. }
. drop yr7
. xtreg fatal beertax spircons unrate perincK yr*, fe
                                                           336
                                       Number of obs =
Fixed-effects (within) regression
Group variable: state
                                       Number of groups =
                                                            48
R-sq:
                                       Obs per group:
                                                             7
    within = 0.4528
                                                  min =
                                                  avg =
    between = 0.1090
                                                           7.0
    overall = 0.0770
                                                  max =
                                       F(10,278)
                                                    =
                                                          23.00
                                                    = 0.0000
corr(u_i, Xb) = -0.8728
                                       Prob > F
     fatal | Coef. Std. Err. t P>|t| [95% Conf. Interval]
beertax | -.4347195 .1539564 -2.82 0.005 -.7377878 -.1316511
   spircons | .805857 .1126425 7.15 0.000 .5841163 1.027598
unrate | -.0549084 .0103418 -5.31 0.000 -.0752666 -.0345502
    perincK | .0882636 .0199988 4.41 0.000 .0488953 .1276319
      yr82 | .1004321 .0355629
                                2.82 0.005
                                              .0304253 .170439
      yr83 | .0470609 .0321574 1.46 0.144 -.0162421 .1103638
      yr84 | -.0645507 .0224667 -2.87 0.004 -.1087771 -.0203243
      yr85 | -.0993055 .0198667 -5.00 0.000 -.1384139 -.0601971
      yr86 | .0496288 .0232525
                                2.13 0.034
                                              .0038554 .0954021
      yr87 | .0003593 .0289315 0.01 0.990 -.0565933 .0573119
     _cons | .0286246 .4183346 0.07 0.945 -.7948812 .8521305
------
    sigma_u | 1.0987683
    sigma_e | .14570531
     rho | .98271904 (fraction of variance due to u_{-}i)
F test that all u_i=0: F(47, 278) = 64.52
                                                Prob > F = 0.0000
. test yr82 yr83 yr84 yr85 yr86 yr87
 (1) yr82 = 0
 (2) yr83 = 0
 (3) yr84 = 0
 (4) yr85 = 0
 (5) yr86 = 0
```

```
F(6, 278) = 8.48

Prob > F = 0.0000
```

9.1.3 The Between Estimator

| use traffic xtreg fatal beertax spircons unrate perincK, | | use | traffic | xtrea | fatal | beertax | spircons | unrate | perincK. | be |
|--|--|-----|---------|-------|-------|---------|----------|--------|----------|----|
|--|--|-----|---------|-------|-------|---------|----------|--------|----------|----|

| Between regress Group variable: | | | of obs = of groups = | 336 48 | | |
|--|--|--|-------------------------------|----------------------------------|--------------------------------|---|
| between = | R-sq: within = 0.0479 between = 0.4565 overall = 0.2583 | | | | | 7 7.0 7 |
| $sd(u_i + avg(e_i)$ | F(4,43) Prob > F | | | | | |
| fatal | | | | | [95% Conf. | = |
| beertax spircons unrate perincK | .0740362 .2997517 .0322333 1841747 | | 0.51 2.66 0.85 -4.36 | 0.614 0.011 0.401 0.000 | 2196614 .0722417 0444111 | .3677338 .5272618 .1088776 0990218 |

9.1.4 One-Way Random Effects Models

. use traffic.. xtreg fatal beertax spircons unrate perincK, re

| Random-effects Group variable: | | f obs = f groups = | | | | | | |
|-----------------------------------|----------|-----------------------|-------|----------------|---------------|-----------|--|--|
| R-sq: | | | | Obs per group: | | | | |
| within = | | min = | • | | | | | |
| between = | | avg = | 7.0 | | | | | |
| overall = | 0.0042 | | | | max = | 7 | | |
| | | | | | | | | |
| | | | | Wald chi | 2(4) = | 49.90 | | |
| $corr(u_i, X)$ | | hi2 = | | | | | | |
| fatal | Coef. | Std. Err. | z | P> z | [95% Conf | | | |
| +- | | | | | | | | |
| beertax | .0442768 | .1204613 | 0.37 | 0.713 | 191823 | .2803765 | | |
| spircons | .3024711 | .0642954 | 4.70 | 0.000 | .1764546 | . 4284877 | | |
| unrate | 0491381 | .0098197 | -5.00 | 0.000 | 0683843 | 0298919 | | |
| perincK l | 0110727 | .0194746 | -0.57 | 0.570 | 0492423 | .0270968 | | |

_cons | 2.001973 .3811247 5.25 0.000 1.254983 2.748964

sigma_u | .41675665
sigma_e | .15678965
rho | .87601197 (fraction of variance due to u_i)

9.1.5 Testing the appropriateness of Random Effects Model

- . use traffic.. quietly xtreg fatal beertax spircons unrate perincK, fe
- . estimates store fix
- . quietly xtreg fatal beertax spircons unrate perincK, re
- . estimates store ran
- . hausman fix ran

| | Coeffi | cients | | |
|----------|----------|----------|------------|---------------------|
| | (b) | (B) | (b-B) | sqrt(diag(V_b-V_B)) |
| | fix | ran | Difference | S.E. |
| beertax | 4840728 | .0442768 | 5283495 | .1090815 |
| spircons | .8169652 | .3024711 | .514494 | .0462668 |
| unrate | 0290499 | 0491381 | .0200882 | |
| perincK | .1047103 | 0110727 | .115783 | .0067112 |

b = consistent under Ho and Ha; obtained from xtreg B = inconsistent under Ha, efficient under Ho; obtained from xtreg

Test: Ho: difference in coefficients not systematic

$$chi2(4) = (b-B)'[(V_b-V_B)^{-1}](b-B)$$

= 130.93
Prob>chi2 = 0.0000
 (V_b-V_B) is not positive definite)

- 9.1.6 Prediction from One-Way Fixed Effects Model & Random Effects Model
- 9.2 IV MODELS FOR PANEL DATA
- 9.3 DYNAMIC PANEL DATA MODELS

. use traffic.. tsset

panel variable: state (strongly balanced)

time variable: year, 1982 to 1988

delta: 1 unit

. xtabond2 fatal L.fatal spircons year, gmmstyle(beertax spircons unrate perinc

> K) ivstyle(year) twostep robust nolevelq

command xtabond2 is unrecognized

r(199);

end of do-file

r(199);

9.4 SEEMINGLY UNRELATED REGRESSION MODELS

9.4.1 SUR with Identical Regressors

9.5 MOVING-WINDOW REGRESSION ESTIMATES

MODELS OF DISCRETE & LIMITED DEPENDENT VARIABLES

10.1 BINOMIAL LOGIT & PROBIT MODELS

- 10.1.1 The Latent Variable Approach
- 10.1.2 Marginal Effects & Predictions
- 10.1.2.1 Binomial Probit
- . use womenwk.. describe

Contains data from womenwk.dta

obs: 2,000 vars: 15 size: 134,000

9 Nov 2004 20:23

| | | | | |
|--|--------------------------------------|---|--------------------|--|
| variable name | • | display format | variable label | |
| c1 c2 u v county age education married children select wagefull wage | double double double double | %10.0g %10.0g %10.0g %10.0g %9.0g %8.0g %8.0g | variable tabet | |
| lw | float | %9.0g %9.0g | | |
| work lwf | float float | %9.0g %9.0g | | |
| | | | | |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | | Max |
|----------|-------|---------|-----------|-----------|----------|
| c1 | | 0023069 | | -3.500514 | 3.614182 |
| c2 | 2,000 | 0077596 | 1.006025 | -3.410111 | 3.423961 |
| u | 2,000 | 0011535 | .4940286 | -1.750257 | 1.807091 |
| V | 2,000 | 0071367 | 1 | -3.954782 | 3.229851 |
| county | 2,000 | 4.5 | 2.873 | 0 | 9 |

| | + | | | | |
|-----------|-------|----------|-----------|-----------|----------|
| age | 2,000 | 36.208 | 8.28656 | 20 | 59 |
| education | 2,000 | 13.084 | 3.045912 | 10 | 20 |
| married | 2,000 | . 6705 | .4701492 | 0 | 1 |
| children | 2,000 | 1.6445 | 1.398963 | 0 | 5 |
| select | 2,000 | 35.78556 | 14.98163 | -14.45688 | 89.63869 |
| | + | | | | |
| wagefull | 2,000 | 21.31176 | 7.012038 | -1.680425 | 45.80979 |
| wage | 1,343 | 23.69217 | 6.305374 | 5.88497 | 45.80979 |
| lw | 1,343 | 3.126703 | .2865111 | 1.772402 | 3.824498 |
| work | 2,000 | .6715 | . 4697852 | 0 | 1 |
| lwf | 2,000 | 2.099581 | 1.487519 | 0 | 3.824498 |
| | | | | | |

. probit work age married children education, nolog

| Probit regression | Number of obs | = | 2,000 |
|-----------------------------|---------------|---|--------|
| | LR chi2(4) | = | 478.32 |
| | Prob > chi2 | = | 0.0000 |
| Log likelihood = -1027.0616 | Pseudo R2 | = | 0.1889 |

| work | | | Z | | [95% Conf. | - |
|-----------|-----------|-----------|--------|-------|------------|-----------|
| age | .0347211 | .0042293 | 8.21 | 0.000 | .0264318 | .0430105 |
| married | .4308575 | .074208 | 5.81 | 0.000 | .2854125 | .5763025 |
| children | .4473249 | .0287417 | 15.56 | 0.000 | .3909922 | .5036576 |
| education | .0583645 | .0109742 | 5.32 | 0.000 | .0368555 | .0798735 |
| _cons | -2.467365 | . 1925635 | -12.81 | 0.000 | -2.844782 | -2.089948 |
| | | | | | | |

. mfx compute

Marginal effects after probit
y = Pr(work) (predict)
= .71835948

| variable | dy/dx | | Z | P> z | [95% | C.I.] | Х |
|----------|--------------------------------|----------------------------|-------------------------------|-------------------------|--------------------|------------------------------|-------------------------------------|
| age | .011721 .150478 .1510059 | .00142 .02641 .00922 | 8.25 5.70 16.38 5.32 | 0.000 0.000 0.000 | .008935 .098716 | .014507 .20224 .169073 | 36.208 .6705 1.6445 13.084 |

(*) dy/dx is for discrete change of dummy variable from 0 to 1

. margeff, dummies(married) count command margeff is unrecognized r(199);

end of do-file
r(199);

10.1.3 Evaluating Specification & Goodness of Fit

- . use womenwk..
- . logit work age married children education, nolog

| Logistic regress | ion | | | Number of | obs = | 2,000 |
|------------------|-----------|-----------|------|-----------|------------|----------------------|
| | | | | LR chi2(4 |) = | 476.62 |
| | | | | Prob > ch | i2 = | 0.0000 |
| Log likelihood = | -1027.914 | 4 | | Pseudo R2 | = | 0.1882 |
| | | | | | | |
| | | | | | | |
| work | Coef. | Std. Err. | Z | P> z | [95% Conf. | <pre>Interval]</pre> |
| + | | | | | | |
| age | .0579303 | .007221 | 8.02 | 0.000 | .0437773 | .0720833 |
| married | .7417775 | .1264705 | 5.87 | 0.000 | . 4938998 | .9896552 |

_cons | -4.159247 .3320401 -12.53 0.000 -4.810034 -3.508461

 children | .7644882
 .0515289
 14.84
 0.000
 .6634935
 .865483

 education | .0982513
 .0186522
 5.27
 0.000
 .0616936
 .134809

. mfx compute

Marginal effects after logit

y = Pr(work) (predict)

= .72678588

| variable | • | | | | - | = | |
|--------------------------------|---------------------------------|-------------------------------------|-----------------------|-------------------------|-------------------------------|-------------------------------|-------------------------------------|
| age married* children | .0115031 .1545671 .151803 | .00142 .02703 .00938 .0037 | 8.08 5.72 16.19 | 0.000 0.000 0.000 | .008713 .101592 .133425 | .014293 .207542 .170181 | 36.208 .6705 1.6445 13.084 |

- (*) dy/dx is for discrete change of dummy variable from 0 to 1
- . mfx compute, at(children = 0)

warning: no value assigned in at() for variables age married education; means used for age married education

Marginal effects after logit

y = Pr(work) (predict)

= .43074191

| variable | dy/dx | Std. Err. | Z | P> z | [95% | C.I.] | Х |
|----------|---------------------------------|-----------|---|-------|------------------------------|--------|--------------------------------|
| · | .0142047 .1762562 .187455 | | | 0.000 | .01071 .120897 .165609 | .0177 | 36.208 .6705 0 13.084 |

(*) dy/dx is for discrete change of dummy variable from 0 to 1

10.2 ORDERED LOGIT & PROBIT MODELS

. use panel84extract.. describe

Contains data from panel84extract.dta

obs: 98 vars: 38

1 Dec 2004 23:14

size: 14,896

| | storage | display format | value | |
|---------------|---------|-------------------|----------|-------------------|
| variable name | | Tormat | | variable label |
| s77 | float | %9.0g | | |
| s78 | float | %9.0g | | |
| s79 | float | %9.0g | | |
| s80 | float | %9.0g | | |
| s81 | float | %9.0g | | |
| s82 | float | %9.0g | | |
| s83 | float | %9.0g | | |
| b77 | float | %9.0g | | |
| b78 | float | %9.0g | | |
| b79 | float | %9.0g | | |
| b80 | float | %9.0g | | |
| b81 | float | %9.0g | | |
| b82 | float | %9.0g | | |
| b83 | float | %9.0g | | |
| is77 | float | %9.0g | | |
| is78 | float | %9.0g | | |
| is79 | float | %9.0g | | |
| is80 | float | %9.0g | | |
| is81 | float | %9.0g | | |
| is82 | float | %9.0g | | |
| is83 | float | %9.0g | | |
| ia77 | float | %9.0g | | |
| ia78 | float | %9.0g | | |
| ia79 | float | %9.0g | | |
| ia80 | float | %9.0g | | |
| ia81 | float | %9.0g | | |
| ia82 | float | %9.0g | | |
| ia83 | float | %9.0g | | |
| r77 | float | %9.0g | | |
| r78 | float | %9.0g | | |
| r79 | float | %9.0g | | |
| r80 | float | %9.0g | | |
| r81 | float | %9.0g | | |
| r82 | float | %9.0g | | |
| r83 | float | %9.0g | | |
| rating83 | float | %9.0g | | |
| rating83c | float | %9.0g | bondrati | ng |
| 3 | | 3 | | Bond rating, 1983 |
| dia | float | %9.0g | | 3, |

Sorted by:

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|-----------|--------|----------|-----------|-----------|----------|
| s77 | 98 | 2399226 | 6835850 | 28188 | 5.50e+07 |
| s78 | 98 | 2757617 | 7812873 | 48432 | 6.32e+07 |
| s79 | 98 | 3084304 | 8191207 | 103004 | 6.63e+07 |
| s80 | 98 | 3081560 | 7295521 | 114562 | 5.77e+07 |
| s81 | 98 | 3328268 | 7815398 | 167239 | 6.27e+07 |
| + | | | | | |
| s82 | | 3126689 | 7479808 | 115376 | 6.00e+07 |
| s83 | • | 3409835 | 9008834 | 108397 | 7.46e+07 |
| b77 | • | 1.019388 | .2675541 | . 29 | 2.27 |
| b78 | • | 1.035 | . 2937677 | . 29 | 2.27 |
| b79 | 98 | 1.008265 | .2601666 | .29 | 2.27 |
| b80 | 98 | 1.032245 | .2380095 | .29 | 2 |
| b81 | • | 1.015612 | .2452773 | .29 | 2.08 |
| b82 | 98 | .9689796 | .2139394 | . 29 | 1.65 |
| | 98 | .9932653 | .2011348 | . 46 | 1.65 |
| is77 | 98 | 11.24923 | 5.615914 | -16.54562 | 26.56383 |
| + | · · | | | | |
| is78 | 98 | 11.65585 | 4.528279 | .5114942 | 26.84296 |
| is79 | 98 | 11.36975 | 4.527863 | . 045454 | 23.66122 |
| is80 | 98 | 11.04448 | 9.335074 | -4.169284 | 88.73264 |
| is81 | 98 | 10.34562 | 4.951754 | -2.06479 | 25.77528 |
| is82 | 98 | 7.204422 | 6.152187 | -12.69882 | 24.89959 |
| ·+ | | 7.046005 | | | |
| is83 | • | 7.846085 | 6.495875 | -16.40544 | 26.55174 |
| ia77 | 98 | 14.95777 | 6.958497 | -18.13673 | 37.67006 |
| ia78 | 98 | 15.53974 | 5.104227 | 1.188427 | 29.87578 |
| ia79 | 98 | 15.48429 | 5.302939 | .0836498 | 28.42293 |
| ia80 | 98 | 13.8006 | 6.085422 | -6.350523 | 29.75574 |
| ia81 | J 98 | 13.82875 | 5.952011 | -2.480706 | 30.5371 |
| ia82 | | 9.407202 | 7.306611 | -12.54386 | 32.30307 |
| ia83 | | 10.11473 | 7.441946 | -13.08016 | 30.74564 |
| r77 | • | 2.34113 | 1.874804 | 0 | 6 |
| r78 | 98 | 2.419951 | 1.911573 | 0 | 6 |
| + | | | | | |
| r79 | 98 | 2.373979 | 1.912279 | 0 | 6 |
| r80 | 98 | 2.315973 | 1.855837 | 0 | 6 |
| r81 | 98 | 2.399914 | 1.841927 | 0 | 6 |
| r82 | 98 | 2.342588 | 1.851536 | 0 | 6 |
| r83 | 98 | 2.391481 | 1.850125 | 0 | 6 |
| + | | 2 226725 | 1 707070 | | |
| rating83 | | 2.336735 | 1.787378 | 0 | 6 |
| rating83c | | 3.479592 | 1.17736 | 10 70014 | 5 |
| dia | 98 | .7075242 | 4.711211 | -10.79014 | 20.05367 |

. tabulate rating83c

| Bond rating, | | | | |
|-----------------|-------|-------|---------|-------|
| 1983 | | Freq. | Percent | Cum. |
| BA_B_C | + | 26 | 26.53 | 26.53 |
| BAA | i | 28 | 28.57 | 55.10 |

| AA_A | 15 | 15.31 | 70.41 |
|-------|----|--------|--------|
| AAA | 29 | 29.59 | 100.00 |
| Total | 98 | 100.00 | |

. ologit rating83c ia83 dia, nolog

| Ordered logistic regression | Number of obs | = | 98 |
|-------------------------------|---------------|---|--------|
| | LR chi2(2) | = | 11.54 |
| | Prob > chi2 | = | 0.0031 |
| Log likelihood = -127.27146 | Pseudo R2 | = | 0.0434 |

| rating83c | Coef. | | Z | P> z | [95% Conf. | - |
|-----------|----------|----------|-------|-------|------------|-----------|
| ia83 | .0939166 | .0296196 | 3.17 | 0.002 | .0358633 | . 1519699 |
| dia | 0866925 | .0449789 | -1.93 | 0.054 | 1748496 | .0014646 |
| +- | | | | | | |
| /cut1 | 1853053 | .3571432 | | | 8852932 | .5146826 |
| /cut2 | 1.185726 | .3882099 | | | .4248488 | 1.946604 |
| /cut3 | 1.908412 | .4164896 | | | 1.092108 | 2.724717 |
| | | | | | | |

. predict spBA_B_C spBAA spAA_A spAAA
(option pr assumed; predicted probabilities)

- . summarize spAAA, mean
- . list sp* rating83c if spAAA == r(max)

| | + | | | | + |
|-----|----------|----------|----------|----------|----------|
| | spBA_B_C | spBAA | spAA_A | spAAA | rati~83c |
| | | | | | |
| 31. | .0388714 | .0985567 | .1096733 | .7528986 | AAA |
| | + | | | | + |

- . summarize spBA_B_C, mean
- . list sp* rating83c if spBA_B_C == r(max)

| i | spBA_B_C | spBAA | spAA_A | spAAA | rati~83c |
|-----|----------|----------|----------|----------|----------|
| 67. | .7158453 | .1926148 | .0449056 | .0466343 | AAA |

10.3 TRUNCATED REGRESSION & TOBIT MODELS

10.3.1 Truncation

. use laborsub.. describe

Contains data from laborsub.dta

obs: 250 vars: 6

25 Sep 2004 18:36

size: 1,750

| variable name | storage type | display format | value label | variable label |
|---------------|-----------------|-------------------|----------------|---|
| lfp | byte | %9.0g | | 1 if woman worked in 1975 |
| whrs | int | %9.0g | | Wife's hours of work |
| kl6 | byte | %9.0g | | <pre># of children younger than 6</pre> |
| k618 | byte | %9.0g | | # of children between 6 and 18 |
| wa | byte | %9.0g | | Wife's age |
| we | byte | %9.0g | | Wife's educational attainment |
| | | | | |

Sorted by:

. summarize

| Variable | Obs | Mean | Std. Dev. | Min | Max |
|----------|-----|--------|-----------|-----|------|
| lfp | 250 | .6 | .4908807 | 0 | 1 |
| whrs | 250 | 799.84 | 915.6035 | 0 | 4950 |
| kl6 | 250 | . 236 | .5112234 | 0 | 3 |
| k618 | 250 | 1.364 | 1.370774 | 0 | 8 |
| wa | 250 | 42.92 | 8.426483 | 30 | 60 |
| | + | | | | |
| we | 250 | 12.352 | 2.164912 | 5 | 17 |

. regress whrs $\,$ kl6 k618 wa we if whrs > 0 $\,$

| SS | df | MS | Number of obs | = | 150 |
|------------|--------------------------|--------------------------------|--|----------------------|------------|
| | | | F(4, 145) | = | 2.80 |
| 7326995.15 | 4 | 1831748.79 | Prob > F | = | 0.0281 |
| 94793104.2 | 145 | 653745.546 | R-squared | = | 0.0717 |
| | | | Adj R-squared | = | 0.0461 |
| 102120099 | 149 | 685369.794 | Root MSE | = | 808.55 |
| | 7326995.15 94793104.2 | 7326995.15 4 94793104.2 145 | 7326995.15 4 1831748.79 94793104.2 145 653745.546 | F(4, 145) 7326995.15 | 7326995.15 |

| whrs | | | | | [95% Conf | _ |
|-------|-----------|----------|-------|-------|-----------|-----------|
| kl6 | -421.4822 | 167.9734 | -2.51 | 0.013 | -753.4748 | -89.48953 |
| k618 | -104.4571 | 54.18616 | -1.93 | 0.056 | -211.5538 | 2.639668 |
| wa | -4.784917 | 9.690502 | -0.49 | 0.622 | -23.9378 | 14.36797 |
| we | 9.353195 | 31.23793 | 0.30 | 0.765 | -52.38731 | 71.0937 |
| _cons | 1629.817 | 615.1301 | 2.65 | 0.009 | 414.0371 | 2845.597 |

. truncreg whrs kl6 k618 wa we, ll(0) nolog

(note: 100 obs. truncated)

Truncated regression

| Limit: | lower = | 0 |
|--------|---------|------|
| | upper = | +inf |

Number of obs = 150Wald chi2(4) = 10.05

9 Nov 2004 20:23

| whrs | Coef. | | | P> z | - | . Interval] |
|--------|-----------|----------|-------|-------|-----------|-------------|
| kl6 | -803.0042 | 321.3614 | -2.50 | 0.012 | -1432.861 | -173.1474 |
| k618 | -172.875 | 88.72898 | -1.95 | 0.051 | -346.7806 | 1.030579 |
| wa | -8.821123 | 14.36848 | -0.61 | 0.539 | -36.98283 | 19.34059 |
| we | 16.52873 | 46.50375 | 0.36 | 0.722 | -74.61695 | 107.6744 |
| _cons | 1586.26 | 912.355 | 1.74 | 0.082 | -201.9233 | 3374.442 |
| +- | | | | | | |
| /sigma | 983.7262 | 94.44303 | 10.42 | 0.000 | 798.6213 | 1168.831 |
| | | | | | | |

10.3.2 Censoring

. use womenwk.. describe

Contains data from womenwk.dta

obs: 2,000

vars: 15

size: 134,000

| V | ariable name | _ | display format | variable | label |
|---|--------------|--------|-------------------|----------|-------|
| c | | double | %10.0g | | |
| С | 2 | double | %10.0g | | |
| u | | | %10.0g | | |
| V | | double | %10.0g | | |
| С | ounty | float | %9.0g | | |
| а | ge | int | %8.0g | | |
| е | ducation | int | %8.0g | | |
| m | arried | byte | %8.0g | | |
| С | hildren | int | %8.0g | | |
| S | elect | float | %9.0g | | |
| W | agefull | float | %9.0g | | |
| W | age | float | %9.0g | | |
| ŀ | W | float | %9.0g | | |
| W | ork | float | %9.0g | | |
| ŀ | wf | float | %9.0g | | |

Sorted by:

. summarize

| Variable | 0bs | Mean | Std. Dev. | | Max |
|-------------------|-------------------------|-------------------------------|----------------------|-------------------------------------|----------------------------------|
| c1 c2 u | 2,000 2,000 2,000 | 0023069 0077596 0011535 | .9880571 1.006025 | -3.500514 -3.410111 -1.750257 | 3.614182 3.423961 1.807091 |
| v | 2,000 | 0071367 | 1 | -3.954782 | 3.229851 |
| county | 2,000 | 4.5 | 2.873 | 0 | 9 |

| age | 2,000 | 36.208 | 8.28656 | 20 | 59 |
|-----------|-------|----------|----------|-----------|----------|
| education | 2,000 | 13.084 | 3.045912 | 10 | 20 |
| married | 2,000 | .6705 | .4701492 | 0 | 1 |
| children | 2,000 | 1.6445 | 1.398963 | 0 | 5 |
| select | 2,000 | 35.78556 | 14.98163 | -14.45688 | 89.63869 |
| | + | | | | |
| wagefull | 2,000 | 21.31176 | 7.012038 | -1.680425 | 45.80979 |
| wage | 1,343 | 23.69217 | 6.305374 | 5.88497 | 45.80979 |
| lw | 1,343 | 3.126703 | .2865111 | 1.772402 | 3.824498 |
| work | 2,000 | .6715 | .4697852 | 0 | 1 |
| lwf | 2,000 | 2.099581 | 1.487519 | 0 | 3.824498 |
| | | | | | |

.

. regress lwf age married children education

| Source | SS | df | MS | Number of obs | = | 2,000 |
|----------|------------|-------|------------|---------------|---|--------|
| +- | | | | F(4, 1995) | = | 134.21 |
| Model | 937.873188 | 4 | 234.468297 | Prob > F | = | 0.0000 |
| Residual | 3485.34135 | 1,995 | 1.74703827 | R-squared | = | 0.2120 |
| +- | | | | Adj R-squared | = | 0.2105 |
| Total | 4423.21454 | 1,999 | 2.21271363 | Root MSE | = | 1.3218 |

| lwf | Coef. | Std. Err. | t | P> t | - | Interval] |
|--|---|---|--|---|---|---|
| age married children education _cons | .0363624 .3188214 .3305009 .0843345 -1.077738 | .003862 .0690834 .0213143 .0102295 .1703218 | 9.42 4.62 15.51 8.24 -6.33 | 0.000 0.000 0.000 0.000 0.000 | .0287885 .1833381 .2887004 .0642729 -1.411765 | .0439362 .4543046 .3723015 .1043961 7437105 |
| | | | | | | |

. tobit lwf age married children education, ll(0)

Refining starting values:

Grid node 0: log likelihood = -3563.7251

Fitting full model:

Iteration 0: log likelihood = -3563.7251
Iteration 1: log likelihood = -3368.4259
Iteration 2: log likelihood = -3350.1512
Iteration 3: log likelihood = -3349.9689
Iteration 4: log likelihood = -3349.9685
Iteration 5: log likelihood = -3349.9685

| Tobit regression | Number of obs | = | 2,000 |
|-----------------------------|----------------|-----|--------|
| | Uncensored | = | 1,343 |
| Limits: lower = 0 | Left-censored | = | 657 |
| upper = +inf | Right-censored | = t | 0 |
| | LR chi2(4) | = | 461.85 |
| | Prob > chi2 | = | 0.0000 |
| log likelihood = -3349.9685 | Pseudo R2 | = | 0.0645 |

```
Coef. Std. Err. t P>|t| [95% Conf. Interval]
age | .052157 .0057457 9.08 0.000
married | .4841801 .1035191 4.68 0.000
                                     .0408888
                                    .2811633
                                    .4238228 .5481814
  children | .4860021 .0317055 15.33 0.000
  education | .1149492 .0150913 7.62 0.000
                                    .0853528
    _cons | -2.807696 .2632573 -10.67 0.000 -3.323984 -2.291408
var(e.lwf)| 3.507421 .1498785
                                    3.225466 3.814024
______
. mfx compute, predict(pr(0,.))
Marginal effects after tobit
   y = Pr(lwf>0) (predict, pr(0,.))
     = .81920975
        dv/dx Std. Err. z P>|z| [ 95% C.I. ]
.00083 8.84 0.000 .005703 .008952 36.208
   age | .0073278
                .01576 4.48 0.000 .039803 .101596
married*| .0706994
                                             .6705
                .00479 14.26 0.000 .058899 .077663 1.6445
children | .0682813
educat~n | .0161499 .00216 7.48 0.000 .011918 .020382 13.084
-----
(*) dy/dx is for discrete change of dummy variable from 0 to 1
. mfx compute, predict(e(0,.))
Marginal effects after tobit
   y = E(lwf|lwf>0) (predict, e(0,.))
     = 2.3102021
```

| variable | dy/dx | Std. Err. | Z | P> z | [95% | C.I.] | Х |
|--------------------------------|----------------------|--------------------------------------|-----------------------|--------|-------------------------------|------------------------------|-------------------------------------|
| age married* children | .0314922 .2861047 | .00347 .05982 .01908 .00912 | 9.08 4.78 15.38 | | .024695 .168855 .256041 | .03829 .403355 .330852 | 36.208 .6705 1.6445 13.084 |

(*) dy/dx is for discrete change of dummy variable from 0 to 1

10.4 INCIDENTAL TRUNCATION & SAMPLE SELECTION MODELS

- . use womenwk..
- . heckman lw education age children, select(age married children education) no
- > log

| Heckman selection model | Number of obs | = | 2,000 |
|--|---------------|------|--------|
| (regression model with sample selection) | Selected | = | 1,343 |
| | Nonselecte | ed = | 657 |
| | Wald chi2(3) | = | 454.78 |
| Log likelihood = -1052.857 | Prob > chi2 | = | 0.0000 |

| lw | Coef. | | | | [95% Conf. | . Interval] |
|--|---|--|--|--|---|--|
| + | | | | | | |
| lw | 0207100 | 0024525 | 16 20 | 0.000 | 0240121 | 0445256 |
| | .0397189 .0075872 | | | | .0349121 | |
| | .0075872 | | | | .0056767 0306981 | |
| _cons | | | 35.30 | | 2.177509 | |
| | • | .0055024 | | 0.000 | 2.177309 | 2.43349 |
| select | | | | | | |
| | .0350233 | .0042344 | 8.27 | 0.000 | .0267241 | .0433225 |
| married | • | .0735876 | 6.18 | 0.000 | .3105434 | |
| children | • | .0288398 | 15.74 | 0.000 | .3973122 | |
| education | .0565136 | .0110025 | 5.14 | 0.000 | .0349492 | |
| _cons | • | .1927823 | | 0.000 | -2.855901 | |
| + | · · | | | | | |
| /athrho | .3377674 | .1152251 | 2.93 | 0.003 | .1119304 | .5636045 |
| /lnsigma | -1.375543 | .0246873 | -55.72 | 0.000 | -1.423929 | -1.327156 |
| + | | | | | | |
| rho | .3254828 | .1030183 | | | .1114653 | .5106469 |
| sigma | . 2527024 | .0062385 | | | . 2407662 | .2652304 |
| lambda | .0822503 | .0273475 | | | .0286501 | .1358505 |
| | | | | | | |
| LR test of ind | dep. eqns. (rl | ho = 0): c | hi2(1) = | 5.53 | Prob > chi | i2 = 0.0187 |
| | | | | | | |
| • | | | | | | |
| . heckman lw e | education age | children, | select(a | ge marrie | d children ed | ducation) tw |
| | | | | | | |
| > ostep | | | | | | |
| · | | | | | | 2 000 |
| Heckman select | | | | | of obs = | • |
| · | | | | S | elected = | 1,343 |
| Heckman select | | | | S | | 1,343 |
| Heckman select | | | | S. N | elected = onselected = | 1,343 657 |
| Heckman select | | | | S N Wald ch | elected = onselected = i2(3) = | 1,343 657 405.68 |
| Heckman select | | | | S N Wald ch | elected = onselected = | 1,343 657 405.68 |
| Heckman select | | | | S N Wald ch | elected = onselected = i2(3) = | 1,343 657 405.68 |
| Heckman select | odel with samp | ole selectio | n) | S N Wald ch Prob > | elected = onselected = i2(3) = chi2 = | 1,343 657 405.68 0.0000 |
| Heckman select | odel with samp | ole selectio | n) z | Solution No. 1 No. | elected = onselected = i2(3) = | 1,343 657 405.68 0.0000 |
| Heckman select (regression mo | odel with samp | ole selectio | n) z | Solution No. 1 No. | elected = onselected = i2(3) = chi2 = | 1,343 657 405.68 0.0000 |
| Heckman select (regression mo | odel with samp | ole selectio | n) z | Solution No. 1 No. 1 No. 2 No. | elected = onselected = i2(3) = chi2 = [95% Conf. | 1,343 657 405.68 0.0000 |
| Heckman selection model (regression model) | odel with samp Coef. | Std. Err. | n) z 13.75 | S N Wald ch Prob > P z | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] |
| Heckman select (regression mo | codel with samp Coef. Coef. Coef. Coef. Coef. Coef. | Std. Err003106 | z 13.75 6.50 | Solution No. | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 |
| Heckman select (regression mo | Coef. .0427067 .009322 | Std. Err003106 .0014343 .0115202 | 13.75 6.50 -0.17 | Solution N N N N N N N N N N N N N N N N N N N | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 |
| Heckman select (regression mo | codel with samp Coef. Coef. Coef. Coef. Coef. Coef. | Std. Err003106 .0014343 .0115202 .1249789 | z 13.75 6.50 -0.17 17.00 | Solution No. 100 No. 1 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 |
| Heckman select (regression mo | Coef. .0427067 .009322 0019549 2.124787 | Std. Err003106 .0014343 .0115202 .1249789 | z 13.75 6.50 -0.17 17.00 | Solution No. 100 No. 1 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 |
| Heckman select (regression mo | Coef. .0427067 .009322 0019549 2.124787 | Std. Err003106 .0014343 .0115202 .1249789 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.000 0.865 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] |
| Heckman select (regression mo | Coef. .0427067 .009322 0019549 2.124787 | Std. Err003106 .0014343 .0115202 .1249789 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.865 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 |
| lw education age children _cons | Coef. Coef. .0427067 .009322 0019549 2.124787 .0347211 .4308575 | Std. Err003106 .0014343 .0115202 .1249789 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.865 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 .0430105 .5763025 |
| lw education age children _cons | Coef. Coef. .0427067 .009322 0019549 2.124787 .0347211 .4308575 | Std. Err003106 .0014343 .0115202 .1249789 | 13.75 6.50 -0.17 17.00 | Solution No. 100 No. 1 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 |
| lw education age children _cons select age married children education | Coef. Coef. .0427067 .009322 0019549 2.124787 .4308575 .4473249 .0583645 | Std. Err003106 .0014343 .0115202 .1249789 | 13.75 6.50 -0.17 17.00 8.21 5.81 15.56 5.32 | Solution No. 100 No. 1 | elected = onselected = i2(3) = chi2 = [95% Conf | 1,343 657 405.68 0.0000 |
| Heckman select (regression mode) lw education age children cons select age married children education cons | Coef. Coef. .0427067 .009322 0019549 2.124787 .4308575 .4473249 .0583645 | Std. Err. .003106 .0014343 .0115202 .1249789 .0042293 .074208 .0287417 .0109742 .1925635 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.000 0.865 0.000 0.000 0.000 0.000 0.000 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf0366191 .00651080245341 1.879833 | 1,343 657 405.68 0.0000 |
| Heckman select (regression mode) lw education age children cons select age married children education cons | Coef. Coef. .0427067 .009322 0019549 2.124787 .0347211 .4308575 .4473249 .0583645 -2.467365 | Std. Err. .003106 .0014343 .0115202 .1249789 .0042293 .074208 .0287417 .0109742 .1925635 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.000 0.865 0.000 0.000 0.000 0.000 0.000 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf0366191 .00651080245341 1.879833 | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 .0430105 .5763025 .5036576 .0798735 -2.089948 |
| lw education age children cons select age married children education _cons+ /mills | Coef. Coef. .0427067 .009322 0019549 2.124787 .0347211 .4308575 .4473249 .0583645 -2.467365 | Std. Err003106 .0014343 .0115202 .1249789 .0042293 .074208 .0287417 .0109742 .1925635 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.000 0.865 0.000 0.000 0.000 0.000 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf0366191 .00651080245341 1.879833 | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 .0430105 .5763025 .5036576 .0798735 -2.089948 |
| lw education age children cons select age married children education _cons+ /mills | Coef. Coef. .0427067 .009322 0019549 2.124787 .4308575 .4473249 .0583645 -2.467365 | Std. Err003106 .0014343 .0115202 .1249789 .0042293 .074208 .0287417 .0109742 .1925635 | 13.75 6.50 -0.17 17.00 | S. N Wald ch Prob > P> z 0.000 0.000 0.865 0.000 0.000 0.000 0.000 0.000 0.000 | elected = onselected = i2(3) = chi2 = [95% Conf0366191 .00651080245341 1.879833 | 1,343 657 405.68 0.0000 Interval] .0487944 .0121333 .0206242 2.369741 |

rho | 0.66698

10.5 BIVARIATE PROBIT & PROBIT WITH SELECTION

- . use hmda..
- . replace fanfred = . if deny (285 real changes made, 285 to missing)
- . rename s6 loanamt
- . rename vr vacancy
- . rename mi med_income
- . rename s50 appr_value
- . rename s17 appl_income
- . replace appl_income = appl_income/1000 (2,379 real changes made)
- . rename s46 debt_inc_r

. summarize

| Variable | 0bs | Mean | Std. Dev. | Min | Max |
|-------------|-------|----------|-----------|------|----------|
| seq | 2,380 | 2328.934 | 1293.337 | 2 | 4509 |
| s3 | 2,380 | 1 | Θ | 1 | 1 |
| s4 | 2,380 | 1 | 0 | 1 | 1 |
| s5 | 2,380 | 1.036555 | .2027758 | 1 | 3 |
| loanamt | 2,380 | 139.1353 | 83.42097 | 2 | 980 |
| s7 | 2,380 | 1.268908 | .6605115 | 1 | 3 |
| s9 | 2,380 | 1120 | Θ | 1120 | 1120 |
| s11 | 2,380 | .1743697 | .3795069 | Θ | 1 |
| s13 | 2,380 | 4.715126 | .6991424 | 3 | 5 |
| s14 | 2,379 | 5.702816 | 1.580592 | 1 | 8 |
| s15 | 2,380 | 1.223109 | .4342082 | 1 | 3 |
| s16 | 2,379 | 2.513241 | .9831064 | 1 | 8 |
| appl_income | 2,380 | 13.9406 | 116.9485 | 0 | 999.9994 |
| s18 | 2,265 | 1.954967 | 3.044967 | Θ | 9 |
| s19a | 0 | | | | |
| s19b | 0 | | | | |
| s19c | 0 | | | | |
| s19d | 0 | | | | |
| s20 | 2,380 | 1261.521 | 35488.59 | 1 | 999999.4 |
| s23a | 0 | | | | |
| s24a | 2,380 | .7605042 | 1.104747 | 0 | 8 |

| s25a | 2,380 | 19758.32 | 139160.4 | 0 | 999999.4 |
|---------------|-------|----------------------|----------------------|------------|-------------------|
| s26a | 2,380 | 22275.29 | 147586.7 | 0 | 999999.4 |
| s27a | 2,380 | .1163866 | .3207553 | 0 | 1 |
| s30a | 2,380 | 4332.712 | 4663.801 | 0 | 81000 |
| + | | 1260 250 | 2122 56 | | 41.667 |
| s30c | | 1360.359 | 2130.56 | 0 | 41667 |
| s31a | • | 4914.016 | 5162.458 | 0 | 81000 |
| s31c | | 1471.81 | 2358.868 | 0 | 41667 |
| s32 | • | 1457.293 | 854.9721 | 0 | 10798 |
| s33 | 2,380 | 1870.239 | 40962.55 | | 999999.4 |
| s34 | 2,380 | 4.507271 | 87.8796 | 0 | 3908 |
| s35 | 2,380 | 5134.612 | 70836.87 | 0 | 999999.4 |
| s39 | 2,380 | 3782.993 | 61390.39 | 0 | 999999.4 |
| s40 | 2,380 | 1.75 | 23.60479 | 0 | 666 |
| s41 | 2,380 | 433.6915 | 20497.71 | 0 | 999999.4 |
| + s42 | 2,380 | 1.721008 | .5372816 | 1 | 4 |
| s43 | • | 2.116387 | 1.666721 | 1 | 6 |
| s44 | ' | .0735294 | .2610584 | 0 | 1 |
| s45 | • | 25.53461 | 9.665561 | 0 | 300 |
| debt_inc_r | | 33.08136 | 10.72573 | 0 | 300 |
| + | | | | | |
| s47 | 2,380 | 1.670588 | .4736616 | 1 | 3 |
| s48 | 2,380 | 2022.587 | 40956.14 | 6 | 999999.4 |
| s49 | 0 | | | | |
| appr_value | 2,380 | 198.5426 | 152.9863 | 25 | 4316 |
| s51 | 2,380 | 1.711765 | .4530364 | 1 | 2 |
| + | | | | | |
| s52 | 2,380 | .2256303 | .4180845 | 0 | 1 |
| s53 | 2,380 | .0201681 | .1406045 | 0 | 1 |
| s54 | 0 | | | | |
| s55 | 2,380 | .0294118 | .1689932 | Θ | 1 |
| s56 | 2,380 | .047479 | .2127058 | 0 | 1 |
| + | | 102102 E | 202042 7 | | 000000 4 |
| s57 | | 102102.5 | 302843.7 1072.576 | 0 -7919 | 999999.4 28023 |
| netw | | 253.0412 3.774496 | 2.027062 | 1.8 | |
| uria rtdum | | | | | 10.6 |
| rtdum | 2,380 | .0726891 .444958 | .25968 | 0 | 1 |
| bd + | 2,380 | .444936 | .4970656 | | 1 |
| med_income | 2,380 | .8294118 | .3762278 | Θ | 1 |
| old | | . 4630252 | .5054334 | Θ | 2 |
| vacancy | • | . 4365546 | .4960626 | Θ | 1 |
| school | | 9679.206 | 97847.79 | 5 | 999999.4 |
| chval | 2,380 | 6993.166 | 81710.9 | -8.333333 | 999999.4 |
| + | | | | | |
| dnotown | • | .0336134 | .1802699 | 0 | 1 |
| dprop | • | 0 | 0 | 0 | 0 |
| deny | | .1197479 | .3247347 | 0 | 1 |
| fanfred | • | .3331742 | .4714608 | 0 | 1 |
| approve | 2,380 | .8802521 | .3247347 | 0 | 1 |
| black | 2,380 | .142437 | .3495712 | 0 | 1 |

[.] heckprob fanfred loanamt vacancy med_income appr_value, select(approve = blac

| Probit model with sample selection | | | | Se | f obs = elected = enselected = | 2,095 |
|------------------------------------|----------|-----------|-------|----------|--------------------------------|-----------|
| | | | | | | 80.69 |
| Log likelihood = -2063.066 | | | | Prob > c | :hi2 = | 0.0000 |
| | Coef. | Std. Err. | z | P> z | [95% Conf. | Interval] |
| fanfred | | | | | | |
| loanamt | 0026434 | .0008029 | -3.29 | 0.001 | 0042169 | 0010698 |
| vacancy | 2163307 | .0609798 | -3.55 | 0.000 | 3358489 | 0968125 |
| med_income | .2671341 | .0893349 | 2.99 | 0.003 | .0920409 | .4422273 |
| appr_value | 0014358 | .0005099 | -2.82 | 0.005 | 0024351 | 0004364 |
| _cons | .1684824 | | | | | |
| approve | | | | | | |
| black | 7343534 | .081858 | -8.97 | 0.000 | 8947921 | 5739148 |
| $appl_income$ | | | | | | |
| debt_inc_r | 0262367 | .0036441 | -7.20 | 0.000 | 0333791 | 0190944 |
| _cons | 2.236424 | | | | | |
| | 6006599 | . 2712535 | -2.21 | 0.027 | -1.132307 | 0690128 |
| rho | 537519 | . 1928812 | | | 8118074 | 0689034 |
| LR test of ind | | | | | | |

10.5.1 Binomial Probit with Selection

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