

Statistical Inference

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References

- TEXTBOOK: Statistical Inference, 2nd edition; Casella and Berger, 2002.
- Other References:
 - Introduction to Mathematical Statistics, 6th edition; Hogg, McKean and Craig.
 - 高等数理统计(第2版), 茆诗松, 王静龙



Chapter 1 ~ 4

- Chapter 1. Probability Theory
- Chapter 2. Transformations and Expectations
- Chapter 3. Common Families of Distributions
- Chapter 4. Multiple Random Variables

Chapter 1. Probability Theory

Terminology

- **Sample space**
- **random variable**
- **Distribution functions**
 - **Cumulative distribution function (cdf)**
 - Probability mass function (pmf)
 - Probability density function(pdf)

Chapter 2 Transformations and Expectations

- $Y=g(X)$ and X with cdf $F_X(x)$
- X : discrete random variable, $Y=g(X)$

$$f_Y(y)=P(Y=y)=\sum_{x \in \{x: g(x)=y\}} P(X=x) = \sum_{x \in \{x: g(x)=y\}} f_X(x)$$

- X : continuous random variable

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(g(X) \leq y) \\ &= P(\{x: g(x) \leq y\}) = \int_{\{x: g(x) \leq y\}} f_X(x) dx \end{aligned}$$

- Theorem 2.1.8 $f_X(x)$, $Y=g(X)$ with g is continuous, suppose there exist monotone functions $g_1(x), \dots, g_k(x)$, defined on A_1, \dots, A_k
- 分段单调

$$f_Y(y) = \begin{cases} \sum_{i=1}^k f_X(g_i^{-1}(y)) \left| \frac{d}{dy} g_i^{-1}(y) \right| \\ 0 \end{cases}$$

Expectation of Random Variable

Let X be a r.v. If X is a discrete or continuous r.v. with pdf $f(x)$ (or pmf $p(x)$ if X is a discrete type r. v.), and

$$\int_{-\infty}^{\infty} |g(x)| f(x) dx \text{ (or } \sum_{x \in S} |g(x)| p(x))$$

exists. Then the expectation of $Y=g(X)$ is

$$E[Y] = E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx \quad \text{(or } \sum_{x \in S} g(x) p(x))$$

Moments and Moment Generating Functions

- The n th moment, $E(X^n)$. (e.g. $\mu = EX$)
- The n th central moment, $E(X - \mu)^n$,
 - Variance: $\text{Var}(X) = E(X - \mu)^2$
 - Standard deviation: $\sqrt{\text{var}(X)}$
- Moment generating function(mgf),

$$M_X(t) = E(e^{tX})$$

$$\frac{d^n}{dt^n} M_X(t) \big|_{t=0} = EX^n \quad \text{e.g.} \quad \frac{d}{dt} M_X(t) \big|_{t=0} = EX$$

Chapter 3 Common Families of Distributions

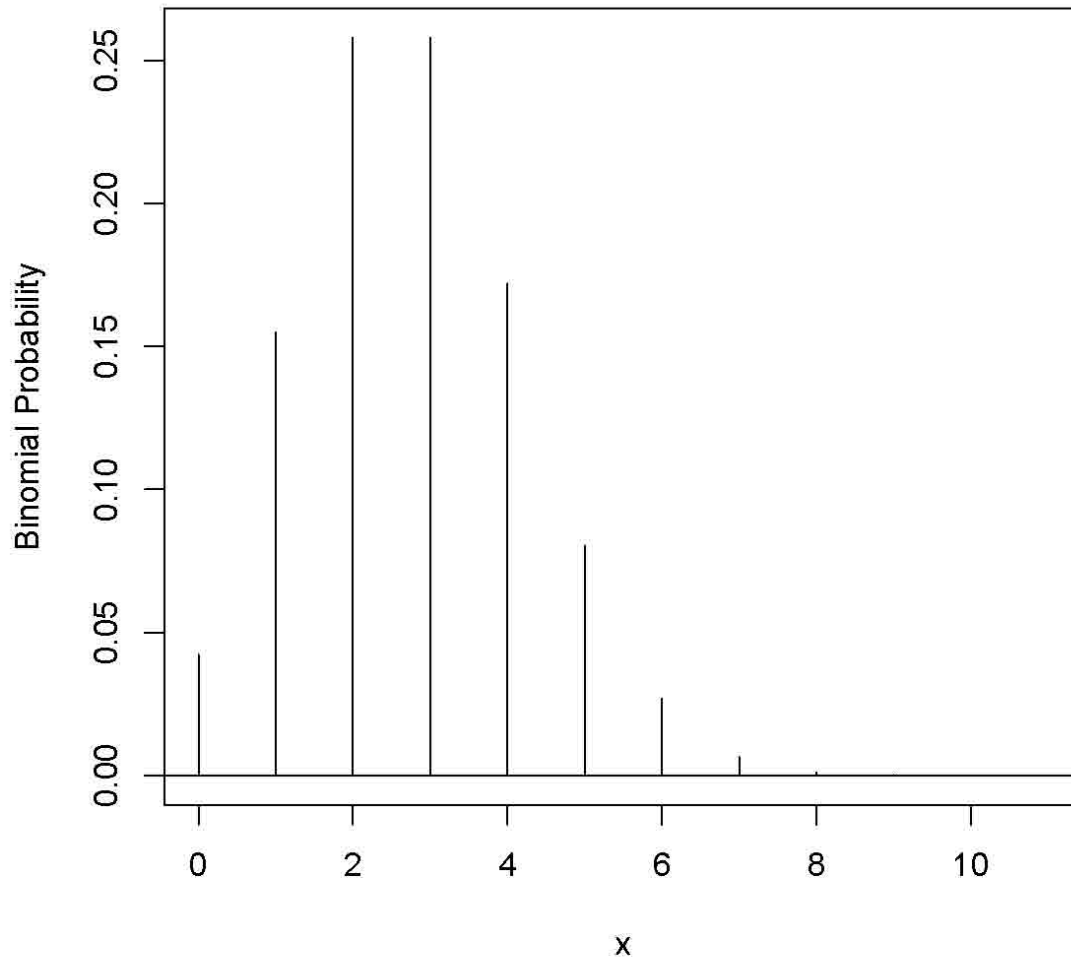
- Discrete Distributions
 - Discrete Uniform
 - Binomial,
 - Poisson,
 - Geometric,
 - Negative binomial (Parskal),
 - Hyper-geometric,
 -
- Continuous Distributions
 - Continuous Uniform
 - Exponential
 - Normal
 - Gamma
 - Beta
 - Cauchy
 - Lognormal
 - Double Exponential
 -

Discrete Uniform Distribution

$$P(X = x \mid N) = \frac{1}{N}, x = 1, 2, \dots, N,$$

Binomial Distribution

$$P(X = x | n, p) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, 2, \dots, n,$$



the number of observed
successes in a sequence
of n times Bernoulli trials

Poisson Distribution

$$P(X = x \mid \lambda) = \frac{e^{-\lambda} \lambda^x}{x!}, x = 0, 1, \dots$$

Negative Binomial Distribution

The number of trials observed until the **r-th** success occurs in a sequence of Bernoulli trials

$$P(X = x | r, p) = \binom{x-1}{r-1} p^r (1-p)^{x-r}, x = r, r+1, \dots$$

The number of fails observed until the **r-th** success occurs in a sequence of Bernoulli trials

$$P(X = x | r, p) = \binom{x+r-1}{r-1} p^r (1-p)^x = \binom{-r}{x} (-1)^x p^r (1-p)^x, x = 0, 1, 2, \dots$$

Geometric Distribution is a special case of Negative Binomial Distribution with **r=1**. $P(X = x | p) = p(1-p)^{x-1}, x = 1, 2, \dots$

$$P(X=x/p)=p (1-p)^x, \quad x=0, 1, 2, \dots$$

Hypergeometric Distribution

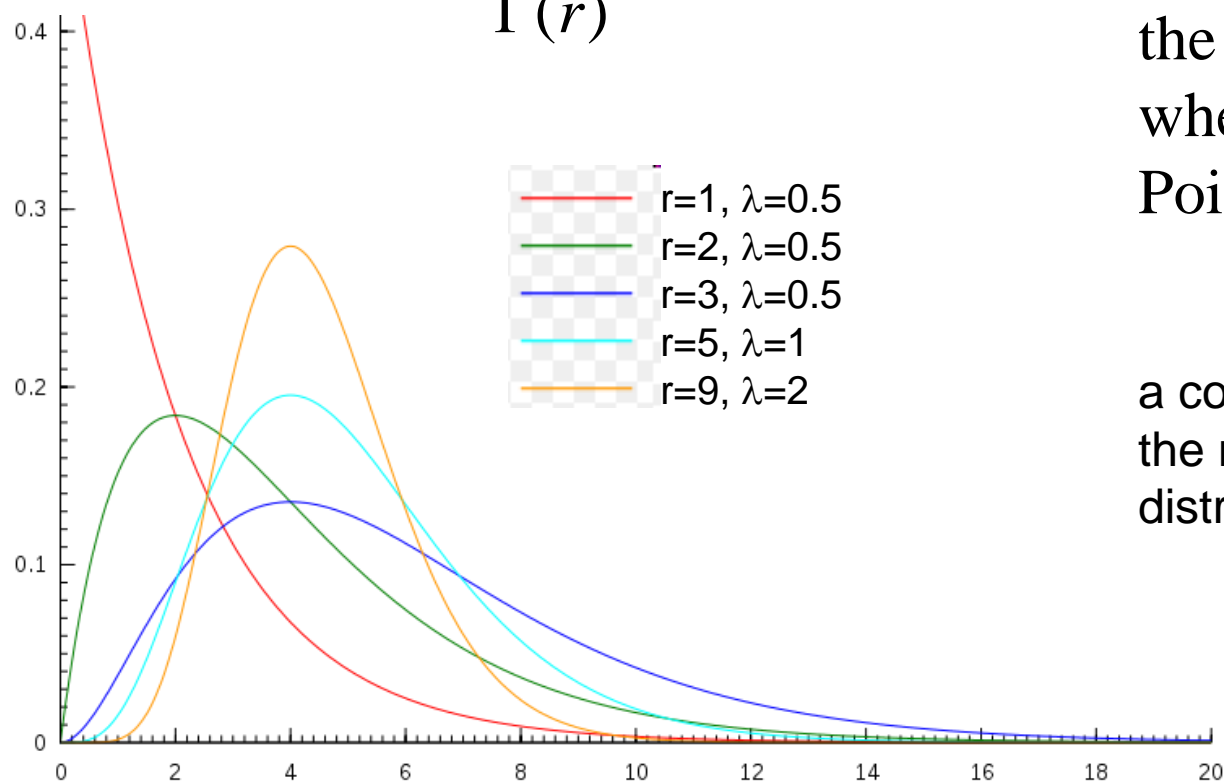
$$P(X = x | N, M, K) = \frac{\binom{M}{x} \binom{N-M}{K-x}}{\binom{N}{k}}, x = 0, 1, \dots, K.$$

Continuous Uniform Distribution

$$f(x | a, b) = \begin{cases} \frac{1}{b-a}, & \text{if } x \in (a, b) \\ 0, & \text{otherwise} \end{cases}$$

Gamma Distribution

$$f(x | r, \lambda) = \frac{\lambda^r}{\Gamma(r)} x^{r-1} e^{-\lambda x}, \quad x \geq 0$$



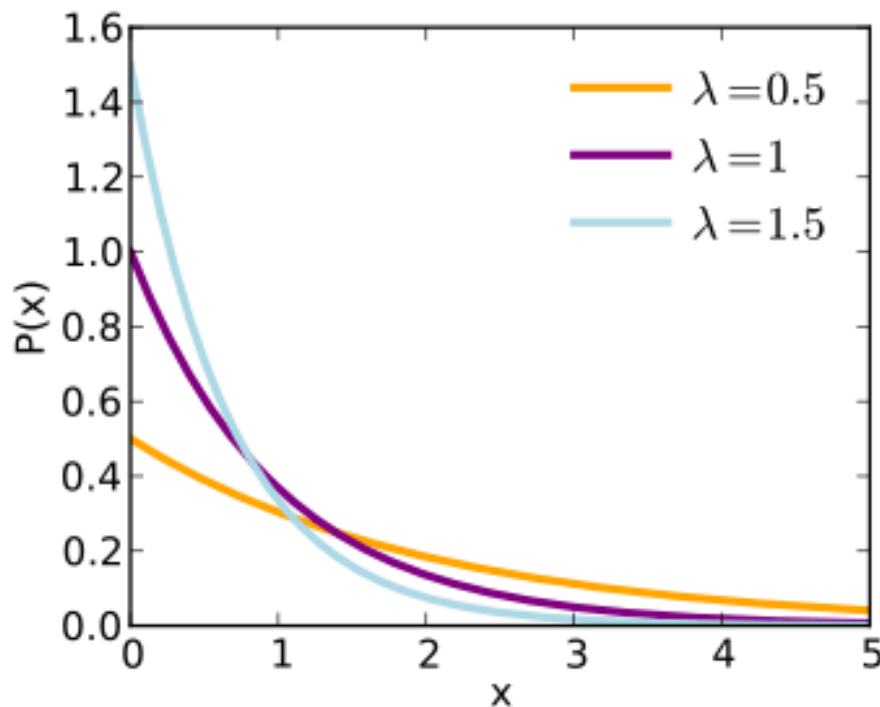
The waiting time until the ***r*-th** change occurs when observing a Poisson process $P(\lambda)$.

a continuous counterpart of the negative-binomial distribution

$$f(x | \alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}, \quad 0 < x < \infty, \alpha > 0, \beta > 0.$$

Exponential distribution

$$f(x|\lambda) = \lambda e^{-\lambda x}, \quad x \geq 0$$



the waiting time until the **first** change occurs when observing a Poisson process $P(\lambda)$.

a continuous counterpart of the geometric distribution

Exponential Distribution is a special case of Gamma Distribution with $r=1$.

Chi-square Distribution

$$f(x | n) = \frac{1}{2^{\frac{n}{2}} \Gamma(\frac{n}{2})} x^{\frac{n}{2}-1} e^{-\frac{x}{2}}, \quad x \geq 0$$

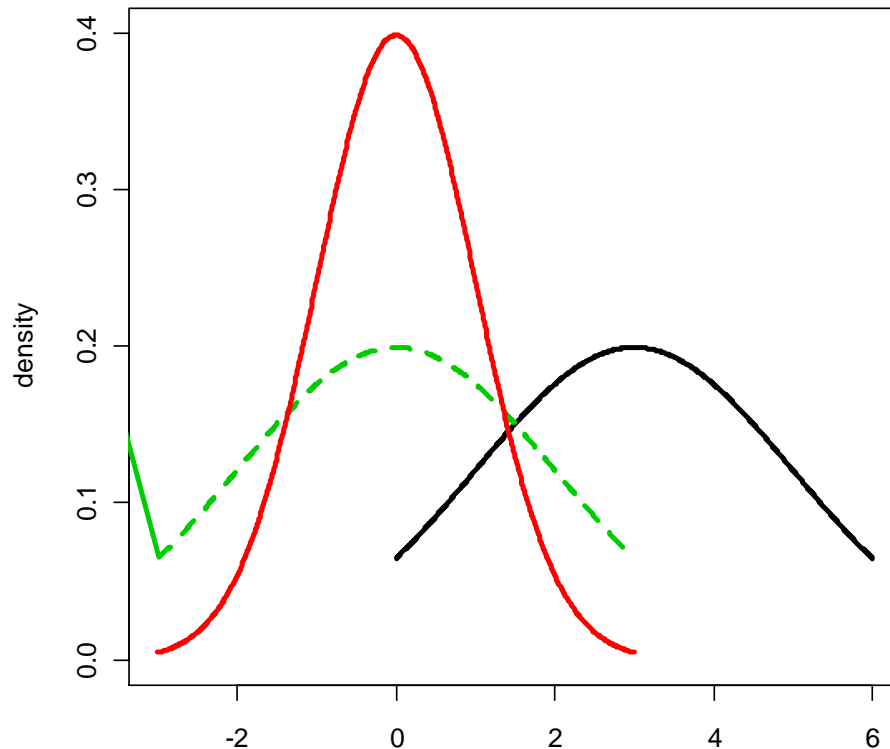
Exponential Distribution is a special case of Gamma Distribution with $r=n/2$ and $\lambda=1/2$ (or $\alpha=n/2$ and $\beta=2$).

Remark: the **sum of the squares** of k independent **standard normal** random variables

$$Z_1^2 + Z_2^2 + \cdots + Z_k^2 \sim \chi_{(k)}^2$$

Normal Distribution

$$f(x | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-(x-\mu)^2 / (2\sigma^2)}, -\infty < x < \infty$$



$$X \sim N(\mu, \sigma^2)$$

↓

$$Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

Beta Distribution

$$f(x | \alpha, \beta) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1}, 0 < x < 1, \alpha > 0, \beta > 0.$$

$$B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}$$

Remark: Suppose the independent r. vs $X_1 \sim \Gamma(\alpha, \lambda)$,
and $X_2 \sim \Gamma(\beta, \lambda)$, then $X_1 + X_2 \sim \Gamma(\alpha + \beta, \lambda)$ is independent $X_1 / (X_1 + X_2) \sim \text{Beta}(\alpha, \beta)$.

Cauchy Distribution

$$f(x | \theta) = \frac{1}{\pi} \frac{1}{1 + (x - \theta)^2}, -\infty < x < \infty, -\infty < \theta < \infty$$

$$E|X| = \infty$$

Lognormal Distribution

$$f(x | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \frac{1}{x} e^{-(\log x - \mu)^2 / (2\sigma^2)}, 0 < x < \infty, -\infty < \mu < \infty, \sigma > 0$$

$$EX = e^{\mu + (\sigma^2 / 2)}$$

Remark: If $X \sim \text{LogN}(\mu, \sigma^2)$, then $\log X \sim N(\mu, \sigma^2)$.

Double Exponential Distribution (Laplace Distribution)

$$f(x | \mu, \sigma) = \frac{1}{2\sigma} e^{-|x-\mu|/\sigma}, -\infty < x < \infty, -\infty < \mu < \infty, \sigma > 0.$$

$$EX = \mu$$

$$VarX = 2\sigma^2$$

Exponential Families

$$f(x | \theta) = h(x)c(\theta) \exp\left(\sum_{i=1}^k w_i(\theta)t_i(x)\right).$$

- Continuous families
 - Normal, Beta, Gamma,
- Discrete families
 - Binomial, Poisson, Negative binomial,

Location and Scale Families

$$g(x|\mu)=f(x-\mu)$$

$$g(x|\sigma)=\frac{1}{\sigma} f\left(\frac{x}{\sigma}\right)$$

$$g(x|\mu,\sigma)=\frac{1}{\sigma} f\left(\frac{x-\mu}{\sigma}\right)$$

Chapter 4 Multiple Random Variables

- Joint distribution
 - Cdf $F(x,y)=P(X\leq x, Y\leq y)$
 - joint pmf (probability mass function)
 - joint pdf (probability density function)

- Marginal distributions
 - Marginal pmfs of discrete bivariate (X,Y)
 - Marginal pdfs of continuous bivariate

Conditional Distributions

- For discrete bivariate,

$$f(y|x) = P(Y=y | X=x) = f(x,y)/f_x(x)$$

providing that $f_x(x) > 0$

$$f(x|y) = P(X=x | Y=y) = f(x,y)/f_y(y)$$

providing that $f_y(y) > 0$

- For continuous bivariate random vector,
replace the pmf function with pdf function

Conditional Expectations

- Expectations

for discrete bivariate random vector

$$E(g(Y) | x) = \sum_y g(y) f(y | x)$$

for continuous bivariate random vector

$$E(g(Y) | x) = \int_{-\infty}^{\infty} g(y) f(y | x) dy$$

Conditional Distributions and Independence

- X and Y are called independent random variables if for every x and y
$$f(x,y)=f_x(x)f_Y(y)$$
- if X and Y are independent, then
$$f(y|x)=f(x,y)/f_x(x)=f_x(x)f_Y(y)/f_x(x)=f_Y(y)$$
- The knowledge $X=x$ doesn't give us any information about Y

Bivariate Transformations

- If (X,Y) is a continuous random vector with joint pdf $f_{X,Y}(x,y)$, and $U=g_1(X,Y)$, $V=g_2(X,Y)$,

Suppose $(g_1, g_2): A \rightarrow B$ is a one to one transformation

- $x=h_1(u,v)$, $y=h_2(u,v)$, then

$$f_{U,V}(u,v)=f_{X,Y}(h_1(u,v),h_2(u,v))|J|$$

$$|J| = \begin{vmatrix} \frac{\partial x}{\partial u} & \frac{\partial x}{\partial v} \\ \frac{\partial y}{\partial u} & \frac{\partial y}{\partial v} \end{vmatrix}$$

- Extending to **not one-to-one transformation**

$$f_{U,V}(u,v) = \sum f_{X,Y}(h_{1i}(u,v),h_{2i}(u,v))|J_i|$$

Covariance and Correlation

- The covariance of X and Y is the number defined by

$$\text{Cov}(X, Y) = E((X - u_X)(Y - u_Y))$$

The correlation of X and Y is the number defined by

$$\rho_{XY} = \text{Cov}(X, Y) / \sigma_X \sigma_Y$$

Covariance and Correlation

- Theorem 4.5.3 For any random variables X and Y , $\text{Cov}(X, Y) = EXY - u_X u_Y$
- Theorem 4.5.5 If X and Y are independent random variables, then $\text{Cov}(X, Y) = 0$ and $\rho_{XY} = 0$
- However $\text{Cov}(X, Y) = 0$ and $\rho_{XY} = 0$ doesn't mean the two are independent (P171)

Covariance and Correlation

• Theorem 4.5.7 For any random variables X and Y ,

a. $-1 \leq \rho_{XY} \leq 1$

b. $|\rho_{XY}| = 1$ if and only if there exist numbers $a \neq 0$ and b such that $P(Y = aX + b) = 1$.

If $\rho_{XY} = 1$, then $a > 0$, and if $\rho_{XY} = -1$, then $a < 0$