

## About exchangeable random vectors

Let  $(z_1, \dots, z_n)$  be a random vector taking its values in  $\mathbb{Z}^n$ .

Definition:  $(z_1, \dots, z_n)$  is exchangeable iff, for any permutation  $\sigma \in S_n$ , it holds that:

$$(z_1, \dots, z_n) \stackrel{(d)}{=} (z_{\sigma(1)}, \dots, z_{\sigma(n)}).$$

Proposition (conditioning on the order statistics): assume that  $\mathbb{Z} = \mathbb{R}$ , and define the order statistics

$z_{(1)} \leq \dots \leq z_{(n)}$  (i.e. the sorted version of the vector  $(z_1, \dots, z_n)$ ).

If  $(z_1, \dots, z_n)$  is exchangeable, it holds that:

$$(z_1, \dots, z_n) \mid z_{(1)}, \dots, z_{(n)} \sim \frac{1}{n!} \sum_{\sigma \in S_n} S_{(z_{(\sigma(1)}), \dots, z_{(\sigma(n)})})$$

$$\text{i.e. } (z_1, \dots, z_n) \mid z_1 = v_1, \dots, z_n = v_n \sim \frac{1}{n!} \sum_{\sigma \in S_n} S_{(v_{\sigma(1)}, \dots, v_{\sigma(n)})}.$$

Remark: Let  $(v_1, \dots, v_n) \in \mathbb{Z}^n$ . When drawing  $(U_1, \dots, U_n) \sim \frac{1}{n!} \sum_{\sigma \in S_n} S_{(v_{\sigma(1)}, \dots, v_{\sigma(n)})}$

note that the  $(U_i)_{i=1}^n$  are not necessarily unique! Thus, for any  $(w_1, \dots, w_n) \in \mathbb{Z}^n$ ,

it leads to:

$$P(U_1 = w_1, \dots, U_n = w_n) = \frac{1}{n!} \sum_{\sigma \in S_n} \prod_{i=1}^n \mathbb{I}_{\{U_{\sigma(i)} = w_i, \dots, U_{\sigma(n)} = w_n\}} \quad \begin{matrix} \rightarrow \text{multiple indicator can} \\ \text{take the value 1!} \end{matrix}$$

What if  $\mathbb{Z} \neq \mathbb{R}$ ? In fact, knowing the order statistics only means knowing the values of the variables without knowing in which order they appear in the vector. This is captured in the definition of a **bag / multiset**.

Definition (bag / multiset): let  $(v_1, \dots, v_n) \in \mathbb{Z}^n$ . The bag (or multiset)  $\{v_1, \dots, v_n\}$

contains all the elements of  $(v_1, \dots, v_n)$  unordered.

It is an object of size  $n$ : it contains duplicates when there are in  $(v_1, \dots, v_n)$ .

for a bag  $\{v_1, \dots, v_n\}$ , we define  $\Psi_{\{v_1, \dots, v_n\}}: \mathbb{Z}^n \rightarrow \mathbb{N}$

$$(w_1, \dots, w_n) \mapsto \#\{\sigma \in S_n : v_{\sigma(1)} = w_1, \dots, v_{\sigma(n)} = w_n\}$$

$$= \sum_{\sigma \in S_n} \mathbb{I}_{\{v_{\sigma(1)} = w_1, \dots, v_{\sigma(n)} = w_n\}}$$

Ex.: \* for  $(v_1, \dots, v_n) = (1, 3, 2, 3)$ , we have  $\mathcal{G}_{v_1, \dots, v_n} = \mathcal{G}_{1, 2, 3, 3}$

$$*\mathcal{G}_{1, 2, 3, 3} = \mathcal{G}_{3, 2, 1, 3} \neq \mathcal{G}_{1, 3, 2}$$

$$*\Psi_{\mathcal{G}_{1, 2, 3, 3}}((1, 3, 2, 3)) = 2 \quad \Psi_{\mathcal{G}_{1, 2, 3, 3}}((1, 2, 1, 3)) = 1 \quad \Psi_{\mathcal{G}_{1, 2, 3, 3}}((3, 3)) = 0$$

Proposition (conditioning on the bag): If  $(z_1, \dots, z_n)$  is exchangeable, it holds that:

$$(z_1, \dots, z_n) | \mathcal{G}_{z_1, \dots, z_n} \sim \frac{1}{n!} \sum_{\sigma \in S_n} \delta_{(z_{\sigma(1)}, \dots, z_{\sigma(n)})}.$$

$$\underbrace{\mathcal{U}(\mathcal{G}_{(z_{\sigma(1)}, \dots, z_{\sigma(n)})}, \sigma \in S_n)}$$

Proof: first, remark that  $\text{supp}((z_1, \dots, z_n) | \mathcal{G}_{z_1, \dots, z_n}) = \{(z_{\sigma(1)}, \dots, z_{\sigma(n)}), \sigma \in S_n\}$ .

thus, let  $(w_1, \dots, w_n) \in \{(z_{\sigma(1)}, \dots, z_{\sigma(n)}), \sigma \in S_n\}$ .

To prove the result, we are going to use Bayes theorem.  $P(A|B) = \frac{P(A, B)}{P(B)}$

thus, let  $R \subseteq \mathcal{G}_{z_1, \dots, z_n}, (z_1, \dots, z_n) \in \mathbb{Z}^n$ .

$$P(z_1 = w_1, \dots, z_n = w_n | \mathcal{G}_{z_1, \dots, z_n} \in R)$$

$$= \frac{P(z_1 = w_1, \dots, z_n = w_n, \mathcal{G}_{z_1, \dots, z_n} \in R)}{P(\mathcal{G}_{z_1, \dots, z_n} \in R)}$$

by definition of exchangeability  $\Rightarrow = \frac{P(z_{\sigma(1)} = w_1, \dots, z_{\sigma(n)} = w_n, \mathcal{G}_{z_1, \dots, z_n} \in R)}{P(\mathcal{G}_{z_1, \dots, z_n} \in R)}$  for any  $\sigma \in S_n$

bags are permutation-invariant  
 $\mathcal{G}_{z_1, \dots, z_n} = \mathcal{G}_{z_{\sigma(1)}, \dots, z_{\sigma(n)}} \Rightarrow = \frac{P(z_{\sigma(1)} = w_1, \dots, z_{\sigma(n)} = w_n, \mathcal{G}_{z_1, \dots, z_n} \in R)}{P(\mathcal{G}_{z_1, \dots, z_n} \in R)}$  for any  $\sigma \in S_n$

$$= \frac{1}{n!} \sum_{\sigma \in S_n} P(z_{\sigma(1)} = w_1, \dots, z_{\sigma(n)} = w_n | \mathcal{G}_{z_1, \dots, z_n} \in R)$$

$$= \frac{1}{n!} E \left[ \sum_{\sigma \in S_n} \mathbb{1}_{\{\mathcal{G}_{z_{\sigma(1)}, \dots, z_{\sigma(n)}} = \mathcal{G}_{w_1, \dots, w_n}\}} | \mathcal{G}_{z_1, \dots, z_n} \in R \right]$$

$$= \frac{1}{n!} \Psi_{\mathcal{G}_{z_1, \dots, z_n}}(w_1, \dots, w_n)$$

□

We now understand that the only requirement for the result to hold, other than exchangeability, is to condition on a random variable that is permutation-invariant. Hence, we have the following generalized result:

Proposition (conditioning on any permutation-invariant function): Let  $g$  be a function with argument in  $\mathbb{Z}^n$ ,

that is permutation-invariant (i.e.  $g(z_1, \dots, z_n) = g(z_{\sigma(1)}, \dots, z_{\sigma(n)})$  for any  $\sigma \in S_n$   $\left| (z_1, \dots, z_n) \in \mathbb{Z}^n \right.$ ).

Then, if  $(z_1, \dots, z_n)$  is exchangeable, it holds that:

$$(z_1, \dots, z_n) | g(z_1, \dots, z_n) \sim U(g(z_{\sigma(1)}, \dots, z_{\sigma(n)}), \sigma \text{ and } \mathcal{S}).$$



Corollary (conditioning on the empirical measure): Denote  $\hat{D}_n = \frac{1}{n} \sum_{i=1}^n \delta_{z_i}$  the empirical distribution.

If  $(z_1, \dots, z_n)$  is exchangeable, it holds that:

$$(z_1, \dots, z_n) | \hat{D}_n \sim U(g(z_{\sigma(1)}, \dots, z_{\sigma(n)}), \sigma \text{ and } \mathcal{S}).$$



Given the characterization of the conditional distribution of  $(z_1, \dots, z_n)$  in the exchangeable setting, we can deduce the conditional distribution of each  $z_i : i \in \{1, n\}$ .

Proposition: Let  $g$  be a permutation-invariant function with argument in  $\mathbb{Z}^n$ .

or

of order statistics, bag, empirical measure?

If  $(z_1, \dots, z_n)$  is exchangeable, for any  $i \in \{1, n\}$  we have:

$$z_i | g(z_1, \dots, z_n) \sim U(g(z_1, \dots, z_n), \mathcal{S}) = \frac{1}{n} \sum_{i=1}^n \delta_{z_i}.$$



Corollary (quantile lemma): If  $(z_1, \dots, z_n)$  is exchangeable, for any  $i \in \{1, n\}$

and  $p \in (0, 1)$ , we have:

$$P(z_i \leq q_p(z_1, \dots, z_n)) \in [p; p + \frac{1}{n}],$$

where the upper bound holds if the  $(z_i)_{i=1}^n$  are a.s. distinct.



Proof: Let  $\beta \in \mathbb{C}_D$  and  $i \in \mathbb{U}, n \in \mathbb{N}$

$$\begin{aligned} P(z \in q_{\beta}(z_1, \dots, z_n)) &= \mathbb{E} \left[ \mathbb{E} \left[ \mathbb{E} \left[ \mathbb{1}_{\{z_i \in q_{\beta}(z_1, \dots, z_n)\}} \mid q_{\beta}(z_1, \dots, z_n) \right] \right] \right] \\ &\stackrel{\text{by the previous proposition}}{=} \mathbb{E} \left[ \mathbb{E} \left[ \mathbb{E} \left[ \mathbb{1}_{\{z_j \in q_{\beta}(z_1, \dots, z_n)\}} \mid q_{\beta}(z_1, \dots, z_n) \right] \right] \right] \text{ for any } j \\ &= P(z_j \in q_{\beta}(z_1, \dots, z_n)) \quad \text{for any } j \in \mathbb{U}, n \\ &= \frac{1}{n} \sum_{j=1}^n P(z_j \in q_{\beta}(z_1, \dots, z_n)) \\ &= \frac{1}{n} \mathbb{E} \left[ \sum_{j=1}^n \mathbb{1}_{\{z_j \in q_{\beta}(z_1, \dots, z_n)\}} \right] \\ &\stackrel{(*)}{\geq} \frac{1}{n} \mathbb{E}[F_{B^n}] \geq \beta \end{aligned}$$

If the  $(z_i)_{i \in \mathbb{U}}$  are a.s. distinct,  
then it becomes an equality:  $= \frac{1}{n} \mathbb{E}[F_{B^n}] \leq \beta + 1/n$

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