

The Value of Location: What Matters Most for Older Individuals Considering Relocation in Sweden?

Nick Christie^a, Björn Slaus^a, Jonas Björk^b, Magnus Zingmark^a, Susanne Iwarsson^a

^a*Department of Health Sciences, Lund University, P.O. Box 7080, 22100, Lund, Sweden*

^b*Division of Occupational and Environmental Medicine, Lund University, Scheelevägen 2, 22363, Lund, Sweden*

Abstract

We examine heterogeneity in housing preferences among older adults in Sweden using discrete choice experiment data from the Prospective RELOC-AGE study (n = 957; mean age = 71.9; 55.3 percent women). Respondents assessed trade-offs between key residential attributes, including proximity to services, access to public transportation, green space, and dedicated parking, and planned monthly expenses. We estimate mixed logit models to recover marginal willingness to pay estimates for each attribute and include interactions with age, gender, and income terciles to capture systematic variation in preferences. Our results show that individuals in the oldest age groups express significantly higher willingness to pay for several attributes, up to three times that of younger respondents. We also identify meaningful differences by gender and tenure status, reflecting underlying patterns of social inequality in later life. These findings contribute policy-relevant evidence to support the development of age-inclusive housing strategies that address both diverse preferences and structural disparities in residential choice.

Email addresses: nick.christie@med.lu.se (Nick Christie), bjorn.slaus@med.lu.se (Björn Slaus), jonas.bjork@med.lu.se (Jonas Björk), magnus.zingmark@med.lu.se (Magnus Zingmark), susanne.iwarsson@med.lu.se (Susanne Iwarsson)

1. Introduction

As populations age and the expansion of housing associated with boundless population growth maintains pace, understanding the housing preferences of the ageing demographic becomes increasingly important for future societies. It is well known that societies across the globe are facing an increasing proportion of older individuals. Research suggests that most older adults prefer to age in place, changing households to a lesser extent compared to younger demographics ([Abramsson and Andersson 2015](#)). In Sweden, the majority of senior citizens live in their own homes with 94% of the population aged 65+ remaining in ordinary housing ([Jennbert 2009](#)). As this population segment grows, appropriate housing options are essential to provide viable options for an ageing society in need of housing. When relocations occur, what matters most for older individuals in their pursuit of accommodation becomes essential for understanding our future needs as a society.

This study employs a discrete choice experiment (DCE) to delve into the critical factors influencing the housing choices of older individuals in Sweden who are considering relocation. Using key housing attributes identified in the nation-wide Prospective RELOC-AGE project, we are able to take a closer look at the many factors which matter most for this segment. Our tests use a diverse sample of older individuals across Sweden considering relocation, presenting them with hypothetical scenarios that vary in locational attributes, including proximity to green spaces, access to public transportation, shops, and parking availability.

In our analyses, we are able to calculate marginal willingness to pay estimates, allowing a monetary interpretation and providing valuable insights for rural planners, policy-makers, and healthcare providers, offering guidance on creating age-friendly environments that may cater to the unique needs and desires of older home adults in their communities.

This study is related to a number of research areas. First, we contribute to the growing literature related to housing and ageing. Second, on a methodological level, we contribute to the literature involving stated choice experiments. To the best of our knowledge, our study is the first to examine housing preferences in the Swedish context. Lastly, we contribute to ageing in place research by identifying differences in housing preferences among older adults. With a larger proportions of older adults in most parts of the world, knowledge into what may influence independence at home is valuable to governments, policy setters, urban planners, and private companies. We contribute to this literature with evidence towards preferred housing attributes of older adults.

In a similar study, [Ossokina et al. \(2020\)](#) run a discrete choice experiment utilizing predominately building characteristics in their model.

In the existing literature, [Ossokina and Arentze \(2022\)](#) is closest to our study in the choice of locational attributes. The authors find that proximity to public transport has the highest effect on utility in their sample, followed by proximity to shops. Our paper differs from theirs in a few ways. First, our sample is larger and encompasses a larger age group. They have 441 home-owners in a nine year age group (65-75). Second, our study comprises both home owners and rental occupants, representing a larger diversity of individuals searching for housing. Finally, instead of a purely hypothetical experiment, the respondents in our study have signed up for housing services, where we can assume they have the intention, or at least a strong consideration, to relocate.

2. Method

2.1. Participants

This paper utilizes survey and experiment data derived from the Prospective RELOC-AGE project, a longitudinal two-tiered mixed-method cohort investigation conducted in Sweden. The study was registered under the identifier NCT04765696 on ClinicalTrials.gov (U.S. National Library of Medicine, 2021).¹ Data collection from this study was in conjunction with the second follow-up survey administered in May 2024. A geographical diverse sample of individuals aged 55 and above was recruited for this study across Sweden (see Figure 1).

The primary objective of the Prospective RELOC-AGE study was to explore the long-term dynamics associated with housing choices, relocation, and active and healthy ageing, focusing on individuals across the ageing process. Eligible participants were individuals aged 55 or older, residing in Sweden, and actively registered for relocation with one of three housing companies: two local public housing providers and a national provider of tenant-owned dwellings.

2.2. Ethics

The Prospective RELOC-AGE study was approved by the Swedish Ethical Review Authority (No. 2020-03457), in alignment with the Declaration of Helsinki and current national ethical regulations for research involving human participants. Potential participants received written information highlighting that participation was voluntary and that declining would not affect their access to housing options or public services. Informed consent was considered given upon the completion and return of the survey.

2.3. Experiment design and attributes

Stated choice experiments include a range of techniques in which respondents indicate their preferences by explicitly stating their choices. In contrast to revealed choice experiments, where preferences are inferred from past behaviour, stated choice methods allow for the evaluation of decisions in a controlled environment. This controlled setting enables the researcher to systematically manipulate attributes and isolate the impact of specific factors on decision-making. A Discrete Choice Experiment (DCE) is a type of stated choice model that presents individuals with hypothetical scenarios, allowing researchers to quantify how much value respondents place on different attributes of a product, service, or housing option.

DCE's have been used in a variety of housing and ageing studies. [Ossokina et al. \(2020\)](#) estimate a stated choice experiment to study the residential preferences of elderly homeowners in the Netherlands, reporting that residential attributes connecting to safety and social cohesion play an important role for elderly. ?

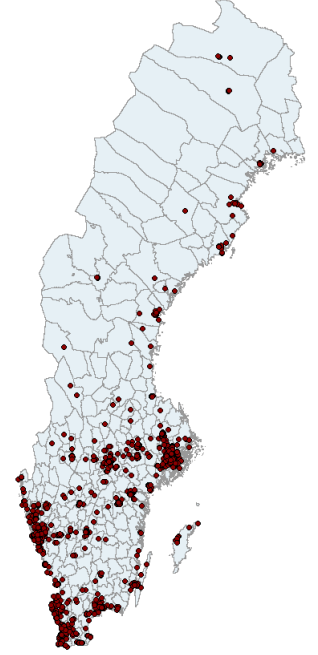


Figure 1: Distribution of respondents

¹For comprehensive insights into the study's procedures, please refer to the study protocol [Zingmark et al. \(2021\)](#)

Respondents were asked to choose the most desirable housing option from a set of alternatives which contained varying levels of attributes. Each respondent was then presented with multiple scenarios, or choice sets, where each choice set contained random attribute levels. To reduce the dimensionality of a full factorial design, a D-optimal subset was generated that reduces dimensionality while maintaining statistical power.

The choice of attributes and associated levels in our study was a combination of factors identified from the Prospective RELOC-AGE follow-up study and attributes found in the housing literature. Proximity to green areas has been examined in numerous contexts including improved cardiometabolic and general health (Paquet et al. 2013, Maas 2006), lower stress (Nielsen and Hansen 2007), and improved mental health (Cohen-Cline et al. 2015, Sturm and Cohen 2014). Proximity to shops and services represent not only distance to frequent amenities which may become more burdensome to transverse, but also also may constitute an integral social experience to participate in the social live of communities (Lucas et al. 2016). Distance to transportation has been shown to affect accessibility levels of populations, with significant differences identified in senior cohorts (Ricciardi et al. 2015, Hildebrand 2003, Alsnih and Hensher 2003) Available parking facilities may also affect acceptability, particularly for our sample where over 90% of respondents indicated access to an automobile.

The attribute *greenspace* is defined as the distance in kilometres to green areas including parks, forests, hiking areas, and open spaces. Similarly, *shops* represents the distance to shopping amenities such as grocery stores, malls, boutiques, and shopping centres. The attribute *transport* is the distance to transportation, such as a bus stop, metro station, or train station. *price* represents the percentage change in price the respondents for each alternative, where planned housing costs are k

We select levels that allow for interpretability and viability of our willingness to pay estimates while holding certain attributes constant to make meaningful comparisons (Hensher et al. 2015).
known a priori from the initial questions on the survey.

Table 1 shows the attributes and their corresponding levels used in the experiment.

Table 1: Attributes and levels

Attribute	Attribute levels
Distance to green area (<i>greenspace</i>)	<ul style="list-style-type: none"> – within 500 m – within 10 km – within 15 km
Distance to shops (<i>shops</i>)	<ul style="list-style-type: none"> – within 500 m – within 10 km – within 15 km
Distance to transportation (<i>transport</i>)	<ul style="list-style-type: none"> – within 300 m – within 600 m – within 900 m
Parking (<i>parking</i>)	<ul style="list-style-type: none"> – no reserved parking – reserved parking place – reserved garage place
Price(<i>price</i>)	<ul style="list-style-type: none"> – 20% less than planned costs – 10% less than planned costs – same as planned cost – 10% more than planned costs – 20% more than planned costs

Note. Table shows attributes and their corresponding attribute levels utilized in the DCE. Variable names given in parentheses.

Before commencing the experiment, respondents were given a definition of each attribute, as well as an example to clarify any ambiguity in interpretation of the attributes. Respondents were also instructed to base each choice on the assumption that the alternative housing options were identical in every way aside from the attribute levels. Figure 2 depicts a typical choice set presented to the respondents.

If these were the only housing options available to you and differ only by the attributes shown, which would you choose?



Figure 2: Example choice set

The number of choice sets was limited to 9 in order to minimize the cognitive burden of the DCE while maximizing the statistical power of our tests (Mangham et al. 2009, DeShazo and Fermo

2002)².

2.4. Development and data

Prior to administering the DCE, we conducted an internal review and one pilot study ($n = 20$; $n = 56$). The internal review involved researchers and staff within the University’s network. The second pilot was administered to participants in Lund University’s Intressentpoolen, a network of individuals placed in collaboration with academia and society in the areas of ageing and health³. In both studies, respondents completed the experiment and a structured feedback form assessing the relevance of attributes, clarity of wording and levels, and overall task comprehension and burden. Feedback informed several revisions: clarifying attribute descriptions and levels; removing redundant instructional text; fine-tuning attribute ranges; and reducing the number of choice sets from 11 to 9 to manage respondent burden. Pilot data were also used to confirm that the planned sample size would provide adequate statistical precision (e.g., stable parameter estimates and predicted choice shares).

The choice experiment was administered in conjunction with the follow-up Prospective RELOC-AGE survey in May 2024, where 1,247 respondents agreed to participate in the survey via a web-based platform. Following the initial survey administration, two reminder mails were sent out to encourage participation in the DCE. The first in September 2024, to the entire respondent group, and another reminder in October 2024, targeting only those who had not taken the experiment portion of the survey. This resulted in a 80% increase from initial turn out. The final sample size of 957 individuals was finalized, representing a 73% response rate. The experiment portion of the survey was administered within the flexible survey framework Formr, which was hosted on university servers.

Among the sample, 55% were female, 46% were male, and the majority (64%) reported being married or cohabiting with a partner. Additionally, 73% of respondents fell within the age range of 54-74, and the majority rated their health as either good (34%) or very good (33%).

2.5. Statistical analyses

Utilizing the discrete choice data, respondent’s choices may be modelled within a random utility theory framework, which assumes individuals choose options which maximize their utility based upon available options (Lancsar and Louviere 2008). An underlying utility model can then be estimated where the utility that individual i derives from alternative j in a choice set t is given by:

$$U_{ijt} = V_{ijt} + \varepsilon_{ijt} \quad (1)$$

where V_{ijt} is the systematic component of utility, modelled as a function of the attributes of the alternative, and ε_{ijt} is an unobserved random error term. The systematic utility is specified as:

$$V_{ijt} = \beta_1 X_{ijt1} + \beta_2 X_{ijt2} + \dots + \beta_k X_{ijtk} \quad (2)$$

where X_{ijtk} represents the level of attribute k for alternative j in task t , and β_k are the corresponding utility coefficients to be estimated. Here, X_{ijtk} represent the respective attributes found in Table 1.

²Himmler et al. (2021) highlights increased age would tend to exacerbate the cognitive burden of a discrete choice experiment, suggesting complex designs would lead to unreliable results.

³For more information see: <https://www.case.lu.se/intressentpoolen>

Multinomial logit (MNL) models have seen a popularity in utility modelling, where the model assumes homogeneous preferences across respondents and independently and identically distributed (i.i.d.) error terms. We estimate and report MNL specifications in our baseline models for comparison purposes. In MNL models, the choice probability is computed using the standard logit formula::

$$P_{ijt} = \frac{\exp(V_{ijt})}{\sum_{l=1}^J \exp(V_{ilt})} \quad (3)$$

The homogeneous preferences assumption may be troublesome in modelling, most notably with the aim of uncovering preference heterogeneity as in our study. Recent studies estimate models that allow utility coefficients to vary groups or individuals (Aitken et al. 2024, Zhao et al. 2023, Caplan et al. 2021).

We follow this line of research and estimate mixed logit models (ML), in which the utility coefficients β_i are allowed to vary randomly across individuals to account for unobserved heterogeneity in preferences (McFadden and Train 2000) :

$$\beta_i = \beta + \Sigma \eta_i \quad (4)$$

where $\eta_i \sim \mathcal{N}(0, I)$ and Σ is the covariance matrix of the random parameters. The mixed logit model captures heterogeneity by integrating over the distribution of random coefficients using simulated maximum likelihood.

We use our price attribute to compute monetary trade-offs for non-cost attributes. First, we estimate marginal rate of substitution (MRS) for each attribute, where the MRS for attribute k is calculated as:

$$\text{MRS}_k = -\frac{\hat{\beta}_k}{\hat{\beta}_{\text{cost}}} \quad (5)$$

where $\hat{\beta}_{\text{cost}}$ is the estimated coefficient on the cost attribute.

Because the cost attribute is specified as a percentage change from the respondent's expected housing cost, we convert marginal rates of substitution (MRS) into marginal willingness to pay (MWTP) in monetary terms by scaling with 10 percent of the mean reported monthly housing cost:

$$\text{MWTP}_k = \text{MRS}_k \times 0.10 \times \bar{\text{Planned}_{\text{cost}}} \quad (6)$$

where \bar{C} denotes the sample mean of respondents' stated planned monthly housing cost.

These MWTP estimates represent the amount, in SEK per month, that respondents are willing to pay for improvements in each housing attribute, relative to their baseline housing cost expectations.

All models are estimated in R using the `logitr` package (Helveston 2023), which provides flexible estimation routines for both MNL and ML models using simulated maximum likelihood. For the ML models, we use 100 Halton draws and multiple random starting values to ensure convergence to the global maximum, following best practice in DCE estimation (Train 2003).

3. Empirical Results

Our discussion begins with results from the baseline model, which excludes interaction terms and serves as a point of comparison for subsequent models that account for heterogeneity across household types. Table 2 presents the estimation results for respondents who own their housing unit. In the proceeding specifications, a positive (negative) coefficient indicates an increase (decrease) in average utility associated with the attribute level, relative to its reference level. Reference levels are specified as the least desirable alternative to aid in interpretation of the results. We report coefficient estimates from the multinomial logit model in column two and mixed logit estimates in column three to facilitate comparison across the model specifications.

Table 2: Baseline results - Owner

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.57*** (0.05)	1.16*** (0.12)	0.12 (0.23)	0.22*** (0.02)	232.29
Green space: 500 m (vs 15 km)	1.12*** (0.05)	2.21*** (0.16)	1.24*** (0.25)	0.42*** (0.03)	440.66
Shops: 5 km (vs 15 km)	0.58*** (0.05)	1.01*** (0.13)	0.10 (0.20)	0.19*** (0.03)	200.82
Shops: 500 m (vs 15 km)	1.68*** (0.05)	3.12*** (0.18)	1.05*** (0.21)	0.59*** (0.04)	622.34
Transit stop: 600 m (vs 900 m)	0.20*** (0.05)	0.35** (0.11)	-0.76*** (0.20)	0.07** (0.02)	69.56
Transit stop: 300 m (vs 900 m)	0.51*** (0.05)	1.13*** (0.12)	0.79*** (0.18)	0.21*** (0.02)	225.03
Parking: reserved garage (vs none)	1.53*** (0.06)	3.08*** (0.20)	0.56* (0.25)	0.59*** (0.04)	615.13
Parking: reserved space (vs none)	1.30*** (0.06)	2.56*** (0.17)	-2.29*** (0.17)	0.49*** (0.03)	511.17
Price	-2.83*** (0.16)	-5.26*** (0.35)			
Num. obs.	7110	7110			
Log Likelihood	-3570.76	-3136.51			
AIC	7159.53	6363.02			
BIC	7221.35	6672.14			
McFadden R ²	0.28	0.36			
LR χ^2 (df=9)	2715.02	3583.53			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Although both model specifications yield similar patterns in coefficient signs and levels of statistical significance, the mixed logit estimates (column three) are systematically larger in magnitude than those from the multinomial logit model. In terms of model performance, the mixed logit specification demonstrates superior fit across all criteria, including log-likelihood, McFadden's pseudo-R², AIC, and BIC. Additionally, several standard deviation estimates for the random parameters (column four) are statistically significant, confirming the presence of preference heterogeneity for certain attributes. We therefore adopt the mixed logit specification as our preferred model for all subsequent tests and discussion.

Marginal rate of substitution (MRS) estimates are presented in column five and are calculated as described in Equation (5). To obtain marginal willingness to pay (MWTP) estimates, MRS

values are multiplied by 10 percent of the average planned housing cost associated with the group under analysis, per Equation (6). These monetary estimates are presented in column six.

As shown in Table 2, owners are, on average, willing to pay up to 440 SEK per month to avoid residing 15 kilometres from the nearest green area. Proximity to shops is valued to a greater extent, with owners willing to pay over 620 SEK per month to avoid greater distances from retail services. Table 3 presents results for respondents who do not own their current housing. Compared to owners, renters place a greater value on proximity to green space, with a MWTP that is 177 SEK greater. Both groups place increased importance of proximity to shops, with Renters showing an average MWTP of 636 SEK for locations within 500 meters, compared to 622 SEK for owners. Reserved parking emerges as a essential attribute across both tenure groups. However, owners are willing to pay substantially more for access to a reserved garage (615 SEK vs 337 SEK), suggesting stronger preferences for secure or private vehicle storage among this group.

Table 3: Baseline results - Renter

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.75*** (0.11)	4.80*** (1.29)	2.30*** (0.65)	0.36*** (0.06)	377.54
Green space: 500 m (vs 15 km)	1.25*** (0.12)	7.85*** (2.07)	6.64*** (1.85)	0.59*** (0.08)	618.10
Shops: 5 km (vs 15 km)	0.77*** (0.12)	3.67** (1.15)	0.36 (0.46)	0.27*** (0.06)	288.71
Shops: 500 m (vs 15 km)	1.68*** (0.12)	8.09*** (2.12)	6.47*** (1.81)	0.61*** (0.09)	636.95
Transit stop: 600 m (vs 900 m)	0.30* (0.12)	1.57** (0.55)	0.99 (0.63)	0.12** (0.04)	123.73
Transit stop: 300 m (vs 900 m)	0.68*** (0.12)	2.01*** (0.52)	1.64** (0.57)	0.15*** (0.03)	158.21
Parking: reserved garage (vs none)	0.92*** (0.12)	4.28*** (0.97)	-0.02 (0.32)	0.32*** (0.04)	337.04
Parking: reserved space (vs none)	0.94*** (0.13)	3.90*** (0.82)	-4.64*** (1.35)	0.29*** (0.03)	306.98
Price	-4.10*** (0.37)	-13.34*** (2.61)			
Num. obs.	1458	1458			
Log Likelihood	-721.69	-612.41			
AIC	1461.39	1314.82			
BIC	1508.95	1552.64			
McFadden R ²	0.29	0.39			
LR χ^2 (df=9)	577.83	796.40			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 4: Baseline results - Male

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.65*** (0.07)	1.33*** (0.16)	-0.25 (0.24)	0.23*** (0.03)	227.78
Green space: 500 m (vs 15 km)	1.01*** (0.07)	2.05*** (0.20)	1.16*** (0.32)	0.35*** (0.04)	351.54
Shops: 5 km (vs 15 km)	0.69*** (0.08)	1.26*** (0.18)	0.32 (0.26)	0.22*** (0.04)	215.57
Shops: 500 m (vs 15 km)	1.76*** (0.07)	3.45*** (0.25)	0.17 (0.30)	0.59*** (0.05)	590.67
Transit stop: 600 m (vs 900 m)	0.20** (0.08)	0.32* (0.16)	-1.14*** (0.23)	0.05 (0.03)	54.05
Transit stop: 300 m (vs 900 m)	0.41*** (0.07)	0.85*** (0.16)	0.21 (0.20)	0.15*** (0.03)	145.00
Parking: reserved garage (vs none)	1.71*** (0.08)	3.65*** (0.29)	-1.54*** (0.24)	0.62*** (0.05)	624.07
Parking: reserved space (vs none)	1.43*** (0.08)	3.12*** (0.26)	2.66*** (0.25)	0.53*** (0.04)	533.73
Price	-3.01*** (0.22)	-5.84*** (0.52)	—	—	—
Num. obs.	3852	3852		—	—
Log Likelihood	-1889.12	-1648.41		—	—
AIC	3796.25	3386.81		—	—
BIC	3852.55	3668.35		—	—
McFadden R ²	0.29	0.38		—	—
LR χ^2 (df=9)	1561.76	2043.19		—	—
p-value (LR)	0.00	0.00		—	—

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 5: Baseline results - Female

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.56*** (0.06)	1.25*** (0.17)	0.25 (0.54)	0.23*** (0.03)	242.97
Green space: 500 m (vs 15 km)	1.24*** (0.06)	2.76*** (0.24)	1.24** (0.43)	0.51*** (0.05)	537.22
Shops: 5 km (vs 15 km)	0.55*** (0.06)	1.07*** (0.21)	0.86** (0.29)	0.20*** (0.04)	207.50
Shops: 500 m (vs 15 km)	1.61*** (0.06)	3.10*** (0.24)	1.33*** (0.39)	0.57*** (0.05)	602.77
Transit stop: 600 m (vs 900 m)	0.24*** (0.07)	0.50*** (0.14)	−0.09 (0.32)	0.09*** (0.03)	98.21
Transit stop: 300 m (vs 900 m)	0.65*** (0.06)	1.38*** (0.16)	1.03*** (0.22)	0.26*** (0.03)	268.90
Parking: reserved garage (vs none)	1.21*** (0.07)	2.43*** (0.20)	0.41 (0.26)	0.45*** (0.04)	473.25
Parking: reserved space (vs none)	1.10*** (0.07)	1.96*** (0.18)	−2.17*** (0.25)	0.36*** (0.03)	381.55
Price	−3.07*** (0.19)	−5.40*** (0.44)			
Num. obs.	4761	4761			
Log Likelihood	−2431.00	−2122.98			
AIC	4880.00	4335.96			
BIC	4938.21	4627.03			
McFadden R ²	0.26	0.36			
LR χ^2 (df=9)	1738.15	2354.18			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 6: Mixed Logit Estimates for 55-64 : Base Specification

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.63*** (0.10)	1.95*** (0.40)	-0.63 (0.48)	0.19*** (0.03)	200.97
Green space: 500 m (vs 15 km)	1.45*** (0.11)	4.81*** (0.73)	4.35*** (0.82)	0.47*** (0.05)	495.74
Shops: 5 km (vs 15 km)	0.65*** (0.12)	1.67*** (0.41)	0.40 (0.39)	0.16*** (0.04)	172.18
Shops: 500 m (vs 15 km)	1.76*** (0.11)	4.94*** (0.65)	1.07* (0.44)	0.48*** (0.05)	509.16
Transit stop: 600 m (vs 900 m)	0.10 (0.11)	0.59 (0.33)	1.90*** (0.44)	0.06 (0.03)	60.54
Transit stop: 300 m (vs 900 m)	0.60*** (0.11)	1.89*** (0.38)	-0.08 (0.47)	0.19*** (0.03)	194.66
Parking: reserved garage (vs none)	1.31*** (0.11)	3.84*** (0.68)	-1.66*** (0.45)	0.38*** (0.05)	395.29
Parking: reserved space (vs none)	1.04*** (0.12)	3.19*** (0.54)	3.84*** (0.70)	0.31*** (0.04)	328.33
Price	-3.99*** (0.36)	-10.19*** (1.42)			
Num. obs.	1728	1728			
Log Likelihood	-817.95	-689.10			
AIC	1653.91	1468.20			
BIC	1703.00	1713.66			
McFadden R ²	0.32	0.42			
LR χ^2 (df=9)	759.61	1017.32			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 7: Mixed Logit Estimates for 65-74 : Base Specification

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.62*** (0.07)	1.38*** (0.22)	0.38 (0.24)	0.20*** (0.03)	214.30
Green space: 500 m (vs 15 km)	1.18*** (0.07)	2.66*** (0.33)	0.53 (0.40)	0.39*** (0.04)	411.97
Shops: 5 km (vs 15 km)	0.54*** (0.07)	1.05*** (0.24)	0.52 (0.38)	0.16*** (0.04)	163.48
Shops: 500 m (vs 15 km)	1.70*** (0.08)	3.72*** (0.39)	2.22*** (0.29)	0.55*** (0.05)	576.94
Transit stop: 600 m (vs 900 m)	0.20** (0.08)	0.24 (0.16)	-0.67** (0.24)	0.04 (0.02)	36.99
Transit stop: 300 m (vs 900 m)	0.54*** (0.07)	1.10*** (0.19)	1.19*** (0.25)	0.16*** (0.03)	170.35
Parking: reserved garage (vs none)	1.51*** (0.08)	3.61*** (0.37)	1.08*** (0.25)	0.53*** (0.05)	559.50
Parking: reserved space (vs none)	1.26*** (0.08)	2.69*** (0.28)	2.55*** (0.38)	0.40*** (0.03)	416.65
Price	-3.09*** (0.22)	-6.77*** (0.65)			
Num. obs.	3744	3744			
Log Likelihood	-1868.23	-1646.37			
AIC	3754.45	3382.73			
BIC	3810.50	3662.99			
McFadden R ²	0.28	0.37			
LR χ^2 (df=9)	1453.83	1897.56			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 8: Mixed Logit Estimates for 75+ : Base Specification

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.55*** (0.07)	1.10*** (0.21)	-0.17 (0.28)	0.23*** (0.05)	238.16
Green space: 500 m (vs 15 km)	0.95*** (0.08)	2.01*** (0.25)	0.40 (0.36)	0.41*** (0.06)	434.70
Shops: 5 km (vs 15 km)	0.66*** (0.08)	1.21*** (0.22)	0.47 (0.28)	0.25*** (0.06)	261.18
Shops: 500 m (vs 15 km)	1.59*** (0.08)	3.34*** (0.33)	0.13 (0.25)	0.69*** (0.08)	722.79
Transit stop: 600 m (vs 900 m)	0.32*** (0.08)	0.55*** (0.16)	-0.70*** (0.21)	0.11** (0.04)	119.57
Transit stop: 300 m (vs 900 m)	0.52*** (0.08)	1.20*** (0.20)	0.57* (0.25)	0.25*** (0.04)	258.91
Parking: reserved garage (vs none)	1.42*** (0.08)	3.04*** (0.29)	0.61* (0.24)	0.63*** (0.08)	658.01
Parking: reserved space (vs none)	1.33*** (0.09)	2.76*** (0.26)	-2.61*** (0.31)	0.57*** (0.06)	597.03
Price	-2.51*** (0.23)	-4.86*** (0.55)			
Num. obs.	3141	3141			
Log Likelihood	-1632.69	-1425.97			
AIC	3283.38	2941.93			
BIC	3337.85	3214.29			
McFadden R ²	0.25	0.35			
LR χ^2 (df=9)	1088.97	1502.42			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 9: Mixed Logit Estimates for retired : Base Specification

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.60*** (0.05)	1.00*** (0.12)	0.01 (0.23)	0.20*** (0.03)	208.42
Green space: 500 m (vs 15 km)	1.10*** (0.05)	1.86*** (0.15)	0.27 (0.43)	0.37*** (0.03)	388.08
Shops: 5 km (vs 15 km)	0.58*** (0.06)	1.04*** (0.15)	1.07*** (0.20)	0.21*** (0.03)	216.03
Shops: 500 m (vs 15 km)	1.65*** (0.06)	3.10*** (0.20)	0.43 (0.29)	0.62*** (0.05)	646.36
Transit stop: 600 m (vs 900 m)	0.26*** (0.06)	0.49*** (0.11)	1.00*** (0.19)	0.10*** (0.02)	103.05
Transit stop: 300 m (vs 900 m)	0.55*** (0.06)	1.05*** (0.12)	0.18 (0.17)	0.21*** (0.03)	218.75
Parking: reserved garage (vs none)	1.50*** (0.06)	2.92*** (0.18)	1.11*** (0.20)	0.58*** (0.05)	609.32
Parking: reserved space (vs none)	1.35*** (0.06)	2.52*** (0.19)	2.14*** (0.18)	0.50*** (0.04)	525.65
Price	-2.88*** (0.16)	-5.03*** (0.38)			
Num. obs.	6642	6642			
Log Likelihood	-3364.20	-2978.20			
AIC	6746.39	6046.40			
BIC	6807.60	6352.46			
McFadden R ²	0.27	0.35			
LR ² (df=9)	2479.37	3251.36			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 10: Mixed Logit Estimates for notretired : Base Specification

	MNL	MXL		MRS	MWTP (SEK/mo)
		Mean	SD		
Green space: 5 km (vs 15 km)	0.58*** (0.09)	1.40*** (0.27)	0.52 (0.28)	0.16*** (0.03)	171.51
Green space: 500 m (vs 15 km)	1.29*** (0.10)	3.27*** (0.42)	−0.86** (0.29)	0.38*** (0.04)	399.24
Shops: 5 km (vs 15 km)	0.69*** (0.11)	1.63*** (0.36)	0.65* (0.33)	0.19*** (0.04)	199.17
Shops: 500 m (vs 15 km)	1.74*** (0.10)	4.48*** (0.57)	1.51*** (0.36)	0.52*** (0.06)	546.79
Transit stop: 600 m (vs 900 m)	0.11 (0.11)	0.37 (0.25)	0.26 (0.43)	0.04 (0.03)	44.95
Transit stop: 300 m (vs 900 m)	0.54*** (0.10)	1.61*** (0.27)	1.39*** (0.25)	0.19*** (0.03)	196.08
Parking: reserved garage (vs none)	1.20*** (0.10)	3.16*** (0.40)	−1.60*** (0.33)	0.37*** (0.05)	385.74
Parking: reserved space (vs none)	0.88*** (0.11)	2.32*** (0.38)	3.38*** (0.51)	0.27*** (0.04)	283.63
Price	−3.60*** (0.33)	−8.60*** (1.10)			
Num. obs.	1935	1935			
Log Likelihood	−946.20	−800.34			
AIC	1910.40	1690.69			
BIC	1960.51	1941.24			
McFadden R ²	0.29	0.40			
LR ² (df=9)	790.08	1081.79			
p-value (LR)	0.00	0.00			

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

3.1. Heterogeneity Models

We next estimate mixed logit models introducing cross-product terms to identify sources of preference heterogeneity in our sample.

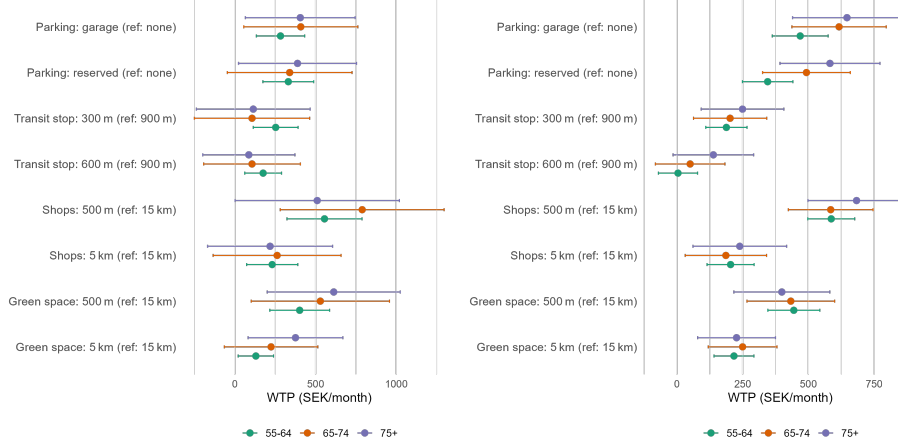


Figure 3: WTP estimates from model 2

4. Discussion

This study explored preference heterogeneity among older individuals considering relocation in Sweden. To our knowledge, this is the first study to utilize a discrete choice experiment to examine locational preferences.

5. Limitations

6. Conclusions

As the global population continues to age, understanding the housing preferences of older demographics becomes increasingly crucial for the planning and development of future societies. In Sweden, where a significant portion of older individuals prefer to live in their own homes, the need for appropriate housing options for this segment is becoming more pronounced as this demographic group expands.

This study utilized a discrete choice experiment to delve into factors influencing the housing choices of older home owners in Sweden considering relocation. By presenting them with hypothetical scenarios that varied in locational attributes such as healthcare facilities, public transportation, green spaces, social amenities, and natural surroundings, we were able to identify key housing preferences.

We find that respondents in older age groups demonstrate substantial differences in preferred housing attributes. Individual aged 75+ are willing to pay more to be closer to public transportation compared to individuals aged 55-64.

Across all tests, access to public transportation and proximity to green spaces emerged as paramount factors influencing housing choices. These findings offer valuable insights for rural planners, policy makers, and healthcare providers, providing guidance on creating age-friendly

environments that cater to the unique needs and desires of older home owners while promoting resilient and vibrant communities. Understanding the interplay between these key factors is essential for optimizing the rural living experience for the older population, ensuring their well-being, and fostering their continued contribution to the vitality of future societies. As populations continue to age, these insights will be instrumental in shaping the future of housing and community development for older adults.

References

- Abramsson M, Andersson EK (2015) Changing locations: Central or peripheral moves of seniors? *Journal of Housing and the Built Environment* 30(4):535–551, ISSN 1566-4910, 1573-7772, URL <http://dx.doi.org/10.1007/s10901-014-9427-0>.
- Aitken D, Willis K, Gilroy R (2024) Do older homebuyers prefer dwellings with accessibility and adaptability features? Findings from an exploratory study. *Housing Studies* 39(3):608–630, ISSN 0267-3037, 1466-1810, URL <http://dx.doi.org/10.1080/02673037.2022.2060944>.
- Alsnih R, Hensher DA (2003) The mobility and accessibility expectations of seniors in an aging population. *Transportation Research Part A: Policy and Practice* 37(10):903–916, ISSN 0965-8564, URL [http://dx.doi.org/10.1016/S0965-8564\(03\)00073-9](http://dx.doi.org/10.1016/S0965-8564(03)00073-9).
- Caplan AJ, Akhundjanov SB, Toll K (2021) Measuring heterogeneous preferences for residential amenities. *Regional Science and Urban Economics* 87:103646, ISSN 0166-0462, URL <http://dx.doi.org/10.1016/j.regsciurbeco.2021.103646>.
- Cohen-Cline H, Turkheimer E, Duncan GE (2015) Access to green space, physical activity and mental health: A twin study. *Journal of Epidemiology and Community Health* 69(6):523–529, ISSN 0143-005X, 1470-2738, URL <http://dx.doi.org/10.1136/jech-2014-204667>.
- DeShazo J, Fermo G (2002) Designing Choice Sets for Stated Preference Methods: The Effects of Complexity on Choice Consistency. *Journal of Environmental Economics and Management* 44(1):123–143, ISSN 0095-0696, URL <http://dx.doi.org/10.1006/jeem.2001.1199>.
- Helveston JP (2023) **Logitr** : Fast Estimation of Multinomial and Mixed Logit Models with Preference Space and Willingness-to-Pay Space Utility Parameterizations. *Journal of Statistical Software* 105(10), ISSN 1548-7660, URL <http://dx.doi.org/10.18637/jss.v105.i10>.
- Hensher DA, Rose JM, Greene WH (2015) *Applied Choice Analysis* (Cambridge University Press), 2 edition, ISBN 978-1-107-09264-8 978-1-107-46592-3 978-1-316-13623-2, URL <http://dx.doi.org/10.1017/CBO9781316136232>.
- Hildebrand ED (2003) Dimensions in elderly travel behaviour: A simplified activity-based model using lifestyle clusters .
- Himmler S, Soekhai V, van Exel J, Brouwer W (2021) What works better for preference elicitation among older people? Cognitive burden of discrete choice experiment and case 2 best-worst scaling in an online setting. *Journal of Choice Modelling* 38:100265, ISSN 1755-5345, URL <http://dx.doi.org/10.1016/j.jocm.2020.100265>.
- Jennbert K (2009) Developments in Elderly Policy in Sweden .
- Lancsar E, Louviere J (2008) Conducting Discrete Choice Experiments to Inform Healthcare Decision Making: A User's Guide. *PharmacoEconomics* 26(8):661–677, ISSN 1170-7690, URL <http://dx.doi.org/10.2165/00019053-200826080-00004>.
- Lucas K, van Wee B, Maat K (2016) A method to evaluate equitable accessibility: Combining ethical theories and accessibility-based approaches. *Transportation* 43(3):473–490, ISSN 1572-9435, URL <http://dx.doi.org/10.1007/s11116-015-9585-2>.
- Maas J (2006) Green space, urbanity, and health: How strong is the relation? *Journal of Epidemiology & Community Health* 60(7):587–592, ISSN 0143-005X, URL <http://dx.doi.org/10.1136/jech.2005.043125>.
- Mangham LJ, Hanson K, McPake B (2009) How to do (or not to do) ... Designing a discrete choice experiment for application in a low-income country. *Health Policy and Planning* 24(2):151–158, ISSN 1460-2237, 0268-1080, URL <http://dx.doi.org/10.1093/heapol/czn047>.
- McFadden D, Train K (2000) Mixed MNL models for discrete response. *Journal of Applied Econometrics* 15(5):447–470, ISSN 1099-1255, URL [http://dx.doi.org/10.1002/1099-1255\(200009/10\)15:5<447::AID-JAE570>3.0.CO;2-1](http://dx.doi.org/10.1002/1099-1255(200009/10)15:5<447::AID-JAE570>3.0.CO;2-1).

- Nielsen TS, Hansen KB (2007) Do green areas affect health? Results from a Danish survey on the use of green areas and health indicators. *Health & Place* 13(4):839–850, ISSN 13538292, URL <http://dx.doi.org/10.1016/j.healthplace.2007.02.001>.
- Ossokina IV, Arentze TA (2022) Reference-dependent housing choice behaviour: Why are older people reluctant to move? *Housing Studies* 1–19, ISSN 0267-3037, 1466-1810, URL <http://dx.doi.org/10.1080/02673037.2022.2151984>.
- Ossokina IV, Arentze TA, Van Gasteren D, Van Den Heuvel D (2020) Best living concepts for elderly homeowners: Combining a stated choice experiment with architectural design. *Journal of Housing and the Built Environment* 35(3):847–865, ISSN 1566-4910, 1573-7772, URL <http://dx.doi.org/10.1007/s10901-019-09716-5>.
- Paquet C, Orschulok TP, Coffee NT, Howard NJ, Hugo G, Taylor AW, Adams RJ, Daniel M (2013) Are accessibility and characteristics of public open spaces associated with a better cardiometabolic health? *Landscape and Urban Planning* 118:70–78, ISSN 01692046, URL <http://dx.doi.org/10.1016/j.landurbplan.2012.11.011>.
- Ricciardi AM, Xia J, Currie G (2015) Exploring public transport equity between separate disadvantaged cohorts: A case study in Perth, Australia. *Journal of Transport Geography* 43:111–122, ISSN 09666923, URL <http://dx.doi.org/10.1016/j.jtrangeo.2015.01.011>.
- Sturm R, Cohen D (2014) Proximity to Urban Parks and Mental Health .
- Train K, ed. (2003) *Discrete Choice Methods with Simulation* (New York: Cambridge University Press), ISBN 978-0-521-81696-0 978-0-511-12085-5 978-0-511-07834-7 978-0-511-07677-0 978-0-511-75393-0.
- Zhao L, Szajnarfarber Z, Broniatowski D, Helveston JP (2023) Using conjoint analysis to incorporate heterogeneous preferences into multimodal transit trip simulations. *Systems Engineering* 26(4):438–448, ISSN 1520-6858, URL <http://dx.doi.org/10.1002/sys.21670>.
- Zingmark M, Björk J, Granbom M, Gefenaite G, Nordeström F, Schmidt SM, Rantanen T, Slaug B, Iwarsson S (2021) Exploring Associations of Housing, Relocation, and Active and Healthy Aging in Sweden: Protocol for a Prospective Longitudinal Mixed Methods Study. *JMIR Research Protocols* 10(9):e31137, ISSN 1929-0748, URL <http://dx.doi.org/10.2196/31137>.