The Value of Location: What Matters Most for Older Individuals Considering Relocation in Sweden?

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Abstract

We examined heterogeneity in housing preferences among older adults in Sweden using discrete choice experiment data (n = 957; mean age = 72; 55.3% women). Respondents assessed trade-offs between key residential attributes, including proximity to shops and services, green spaces, access to public transportation, dedicated parking, and planned monthly expenses. We estimate mixed logit models to recover marginal willingness to pay estimates for each attribute, including interactions with age, gender, and income to capture systematic variation in preferences. The results show that individuals in the oldest age groups expressed significantly higher willingness to pay for several attributes, with values around 1.5 times greater than those of younger respondents. We identified meaningful differences by gender and tenure status, reflecting underlying patterns of social inequality in later life. These findings contribute policy-relevant evidence to support the development of age-inclusive housing strategies that address both diverse preferences and structural disparities in residential choice.

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1. Introduction

As populations age and the expansion of housing associated with boundless population growth maintains pace, understanding the housing preferences of the ageing demographic becomes increasingly important for future societies. It is well known that societies across the globe are facing an increasing proportion of older individuals. Research suggests that most older adults prefer to age in place, changing households to a lesser extent compared to younger demographics (Abramsson and Andersson 2015). In Sweden, the majority of senior citizens live in their own homes with 94% of the population aged 65+ remaining in ordinary housing (Jennbert 2009). As this population segment grows, appropriate housing options are essential to provide viable options for an ageing society in need of housing. When relocations occur, what matters most for older individuals in their pursuit of accommodation becomes essential for understanding our future needs as a society.

This study employs a stated choice experiment to delve into the critical factors influencing the housing choices of older individuals in Sweden who are considering relocation. Our tests use a diverse sample of older individuals across Sweden considering relocation, presenting them with hypothetical scenarios that vary in locational attributes, including proximity to green spaces, access to public transportation, shops, and parking availability.

We calculate marginal willingness to pay estimates, allowing a monetary interpretation and providing valuable insights for rural planners, policy-makers, and healthcare providers, offering guidance on creating age-friendly environments that may cater to the unique needs and desires of older home adults in their communities.

This study is related to a number of research areas. First, we contribute to the growing literature related to housing and ageing. Second, on a methodological level, we contribute to the literature involving stated choice experiments. To the best of our knowledge, our study is the first to examine housing preferences in the Swedish context. Lastly, we contribute to ageing in place research by identifying differences in housing preferences among older adults. With a larger proportions of older adults in most parts of the world, knowledge into what may influence independence at home is valuable to governments, policy setters, urban planners, and private companies. We contribute to this literature with evidence towards preferred housing attributes of older adults.

In a similar study, Ossokina et al. (2020) run a discrete choice experiment utilizing predominately building characteristics in their model. In the existing literature, Ossokina and Arentze (2022) is closest to our study in the choice of locational attributes. The authors find that proximity to public transport has the highest effect on utility in their sample, followed by proximity to shops. Our paper differs from theirs in a few ways. First, our sample is larger an encompasses a larger age group. They have 441 home-owners in a nine year age group (65-75). Second, our study comprises both home owners and rental occupants, representing a larger diversity of individuals searching for housing. Finally, instead of a purely hypothetical experiment, the respondents in our study have signed up for housing services, where we can assume they have the intention, or at least a strong consideration, to relocate.

2. Method

2.1. Participants

This paper utilizes survey and experiment data derived from the Prospective RELOC-AGE project, a longitudinal two-tiered mixed-method cohort investigation conducted in Sweden. The study was registered under the identifier NCT04765696 on ClinicalTrials.gov (U.S. National Library

of Medicine, 2021). [1] Data collection from this study was in conjunction with the second follow-up survey administered in May 2024. A geographical diverse sample of individuals aged 55 and above was recruited for this study across Sweden (see Figure 1).

The primary objective of the Prospective RELOC-AGE study was to explore the long-term dynamics associated with housing choices, relocation, and active and healthy ageing, focusing on individuals across the ageing process. Eligible participants were individuals aged 55 or older, residing in Sweden, and actively registered for relocation with one of three housing companies: two local public housing providers and a national provider of tenant-owned dwellings.

2.2. Ethics

The Prospective RELOC-AGE study was approved by the Swedish Ethical Review Authority (No. 2020–03457), in alignment with the Declaration of Helsinki and current national ethical regulations for research involving human participants. Potential participants received written information highlighting that participation was voluntary and that declining would not affect their access to housing options or public services. Informed consent was considered given upon the completion and return of the survey.

2.3. Experiment design and attributes

Stated choice experiments include a range of techniques in which respondents indicate their preferences by explicitly stating their choices. In contrast to revealed choice experiments, where preferences are inferred from past behaviour, stated choice methods allow for the evaluation of decisions in a controlled environment. This controlled setting enables the researcher to systematically manipulate attributes and isolate the impact of specific factors on decision-making. A Discrete Choice Experiment (DCE) is a type of stated choice model that presents individuals with hypothetical scenarios, allowing researchers to quantify how much value respondents place on different attributes of a product, service, or housing option.

Figure 1: Distribution of respondents

DCE's have been used in a variety of housing and ageing studies. Ossokina et al. (2020) estimate a stated choice experiment to study the residential preferences of elderly homewoners in the Netherlands, reporting that residential attributes connecting to safety and social cohesion play an important role for elderly. Ossokina and Arentze (2022)

Respondents were asked to choose the most desirable housing option from a set of alternatives which contained varying levels of attributes. Each respondent was then presented with multiple scenarios, or choice sets, where each choice set contained random attribute levels. To reduce the dimensionality of a full factorial design, a D-optimal subset was generated that reduces dimensionality while maintaining statistical power.
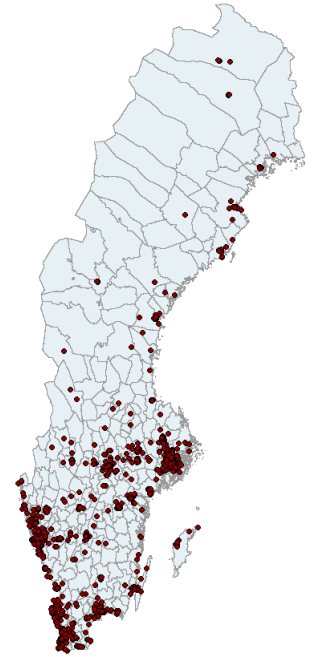
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[1] For comprehensive insights into the study's procedures, please refer to the study protocol Zingmark et al. (2021)

The choice of attributes and associated levels in our study was a combination of factors identified from the Prospective RELOC-AGE follow-up study and attributes found in the housing literature. Proximity to green areas has been examined in numerous contexts including improved cardiometabolic and general health (Paquet et al. 2013, Maas 2006), lower stress (Nielsen and Hansen 2007), and improved mental health (Cohen-Cline et al. 2015, Sturm and Cohen 2014). Proximity to shops and services represent not only distance to frequent ammenities which may become more burdensome to transverse, but also also may constitute an integral social experience to participate in the social live of communities (Lucas et al. 2016). Distance to transportation has been shown to affect accessibility levels of populations, with significant differences identified in senior cohorts (Ricciardi et al. 2015, Hildebrand 2003, Alsnih and Hensher 2003) Available parking facilities may also affect acceptability, particularly for our sample where over 90% of respondents indicated access to an automobile.

We select levels that allow for interpretability and viability of our willingness to pay estimates while holding certain attributes constant to make meaningful comparisons (Hensher et al. 2015). The attribute *greenspace* is defined as the distance in kilometres to green areas including parks, forests, hiking areas, and open spaces. Similarly, *shops* represents the distance to shopping amenities such as grocery stores, malls, boutiques, and shopping centres. The attribute *transport* is the distance to transportation, such as a bus stop, metro station, or train station. *price* represents the percentage change from respondent's planned housing cost in 10% intervals. Table 1 shows the attributes and their corresponding levels used in the experiment.

Table 1: Attributes and levels used in the DCE

Attribute (variable)	Attribute levels
Distance to green area (*greenspace*)	Within 500 m Within 10 km Within 15 km
Distance to shops (*shops*)	Within 500 m Within 10 km Within 15 km
Distance to transportation (*transport*)	Within 300 m Within 600 m Within 900 m
Parking (*parking*)	No reserved parking Reserved parking place Reserved garage place
Price (*price*)	20% less than planned cost 10% less than planned cost Same as planned cost 10% more than planned cost 20% more than planned cost

Note. Attributes and levels used in the discrete choice experiment (DCE). Variable names used in the model are shown in parentheses.

Before commencing the experiment, respondents where given a definition of each attribute, as well as an example to clarify any ambiguity in interpretation of the attributes. Respondents were also instructed to base each choice on the assumption that the alternative housing options where identical in every way aside from the attribute levels. Figure 2 depicts a typical choice set presented to the respondents.

Option 1

Distance to green area: 10km
Distance to shops: 5km
Distance to transportation : 300m
Parking: No reserved parking
Price: 10% more

Option 2

Distance to green area: 5km
Distance to shops: 15km
Distance to transportation : 900m
Parking: Garage parking
Price: 10% less

If these were the only housing options available to you and differ only by the attributes shown, which would you choose?

Figure 2: Example choice set

The number of choice sets was limited to 9 in order to minimize the cognitive burden of the DCE while maximizing the statistical power of our tests (Mangham et al. 2009, DeShazo and Fermo 2002) [2].

2.4. Development and data

Prior to administering the DCE, we conducted and internal review and one pilot study (n = 20; n = 56). The internal review involved researchers and staff within the University's network. The second pilot was administered to participants in Lund University's Intressentpoolen, a network of individuals placed in collaboration with academia and society in the areas of ageing and health[3]. In both studies, respondents completed the experiment and a structured feedback form assessing the relevance of attributes, clarity of wording and levels, and overall task comprehension and burden. Feedback informed several revisions: clarifying attribute descriptions and levels; removing redundant instructional text; fine-tuning attribute ranges; and reducing the number of choice sets from 11 to 9 to manage respondent burden. Pilot data were also used to confirm that the planned sample size would provide adequate statistical power with stable parameter estimates.

The choice experiment was administered in conjunction with the follow-up Prospective RELOC-AGE survey in May 2024, where 1,247 respondents agreed to participate in the survey via a web-based platform. Following the initial survey administration, two reminder mails were sent out to encouraged participation in the DCE. The first in September 2024, to the entire respondent group, and another reminder in October 2024, targeting only those who had not taken the experiment portion of the survey. This resulted in a 80% increase from initial turn out. The final sample size of 957 individuals was finalized, representing a 73% response rate.

Table 2 provides descriptive statistics for the study sample, disaggregated by housing tenure. Of the 957 respondents, 790 were homeowners and 167 were renters. The sample was slightly majority female (55.3%), with most participants aged between 65–74 years (43.5%) followed by those aged 75+ (36.5%). The majority of respondents were partnered (60.6%), retired (77.1%), and lived in

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[2]Himmler et al. (2021) highlights increased age would tend to exacerbated the cognitive burden of a discrete choice experiment, suggesting complex designs would lead to unreliable results.

[3]For more information see: https://www.case.lu.se/intressentpoolen

apartment or condominium housing (63.0%). Most had attained at least three years of university education (47.6%), and lived in city or town locations (62.2%). Households predominantly consisted of one or two members (95.2%), with only 3.8% reporting three or more. The mean monthly household income was 46,700 SEK, and the average planned monthly housing cost was 10,500 SEK.

Table 2: Descriptive statistics

	Owner	Renter	Overall
	(N = 790)	(N = 167)	(N = 957)
Sex			
Female	419 (53.0%)	110 (65.9%)	529 (55.3%)
Male	371 (47.0%)	57 (34.1%)	428 (44.7%)
Age group			
55–64	154 (19.5%)	38 (22.8%)	192 (20.1%)
65–74	344 (43.5%)	72 (43.1%)	416 (43.5%)
75+	292 (37.0%)	57 (34.1%)	349 (36.5%)
Civil status			
Not partnered	287 (36.3%)	90 (53.9%)	377 (39.4%)
Partnered	503 (63.7%)	77 (46.1%)	580 (60.6%)
Education			
Elementary school	42 (5.3%)	17 (10.2%)	59 (6.2%)
2 years upper secondary	68 (8.6%)	17 (10.2%)	85 (8.9%)
3 or 4 years upper secondary	127 (16.1%)	30 (18.0%)	157 (16.4%)
University < 3 years	161 (20.4%)	37 (22.2%)	198 (20.7%)
University ≥ 3 years	390 (49.4%)	66 (39.5%)	456 (47.6%)
Retired			
Retired	612 (77.5%)	126 (75.4%)	738 (77.1%)
Not retired	174 (22.0%)	41 (24.6%)	215 (22.5%)
Current housing type			
Apartment/Condo	449 (56.8%)	154 (92.2%)	603 (63.0%)
House	341 (43.2%)	13 (7.8%)	354 (37.0%)
Current housing location			
City/town	478 (60.5%)	117 (70.1%)	595 (62.2%)
Urban area	221 (28.0%)	41 (24.6%)	262 (27.4%)
Countryside	81 (10.3%)	6 (3.6%)	87 (9.1%)
Number in household			
1 inhabitant	240 (30.4%)	81 (48.5%)	321 (33.5%)
2 inhabitants	509 (64.4%)	81 (48.5%)	590 (61.7%)
3 or more inhabitants	34 (4.3%)	2 (1.2%)	36 (3.8%)
Monthly household income	48,700 (34,700)	37,700 (27,900)	46,700 (33,900)
Planned housing costs	10,400 (4,060)	10,500 (5,570)	10,500 (4,370)

Note: Table reports descriptive statistics of variable in the study. Values are frequencies with percentages in parentheses. Monthly household income and planned housing costs are reported as mean (standard deviation) in SEK (Swedish Krona).

2.5. Statistical analyses

Utilizing the discrete choice data, respondent's choices may be modelled within a random utility theory framework, which assumes individuals choose options which maximize their utility based upon available options (Lancsar and Louviere 2008). An underlying utility model can then be estimated where the utility that individual i derives from alternative j in a choice set t is given by:

$$U_{ijt} = V_{ijt} + \varepsilon_{ijt} \tag{1}$$

where V_{ijt} is the systematic component of utility, modelled as a function of the attributes of the alternative, and ε_{ijt} is an unobserved random error term. The systematic utility is specified as:

$$V_{ijt} = \beta_1 X_{ijt1} + \beta_2 X_{ijt2} + \ldots + \beta_k X_{ijtk} \tag{2}$$

where X_{ijtk} represents the level of attribute k for alternative j in task t, and β_k are the corresponding utility coefficients to be estimated. Here, X_{ijtk} represent the respective attributes found in Table 1.

Multinomial logit (MNL) models have seen a popularity in utility modelling, where the model assumes homogeneous preferences across respondents and independently and identically distributed (i.i.d.) error terms. We estimate and report MNL specifications in our baseline models for comparison purposes. In MNL models, the choice probability is computed using the standard logit formula::

$$P_{ijt} = \frac{\exp(V_{ijt})}{\sum_{l=1}^{J} \exp(V_{ilt})} \tag{3}$$

The assumption of homogeneous preferences is problematic when the objective is to uncover heterogeneity in preferences, as in our study. Recent research addresses this limitation by allowing utility coefficients to vary across individuals or groups (Aitken et al. 2024, Zhao et al. 2023, Caplan et al. 2021). We follow this line of research and estimate mixed logit (ML) models, in which the utility coefficients $\boldsymbol{\beta}_i$ are allowed to vary randomly across individuals to account for unobserved heterogeneity in preferences (McFadden and Train 2000):

$$\boldsymbol{\beta}_i = \boldsymbol{\beta} + \Sigma \eta_i \tag{4}$$

where $\eta_i \sim \mathcal{N}(0, I)$ and Σ is the covariance matrix of the random parameters. The mixed logit model captures heterogeneity by integrating over the distribution of random coefficients using simulated maximum likelihood. Because this integral has no closed-form solution, we approximate it using simulation:

$$\mathcal{L}^S(\boldsymbol{\beta}, \Sigma) = \sum_{i=1}^{N} \ln \left[\frac{1}{R} \sum_{r=1}^{R} P_i(y_i \mid \boldsymbol{\beta}_i^r) \right], \tag{5}$$

where each simulated draw $\boldsymbol{\beta}_i^r = \boldsymbol{\beta} + \Sigma \eta_i^r$, and R denotes the number of Sobol draws used to approximate the integral over the distribution of random parameters (Train 2003). In our specification, all non-monetary attributes are treated as random and normally distributed, while the price coefficient is fixed to preserve consistency in the computation of willingness-to-pay (WTP) measures. We also allow for full covariance among the random parameters to capture potential correlations in preferences.

We then use our price coefficient to compute monetary trade-offs for non-cost attributes. First, we estimate marginal rate of substitution (MRS) for each attribute, where the MRS for attribute k is calculated as:

$$\text{MRS}_k = -\frac{\hat{\beta}_k}{\hat{\beta}_{\text{cost}}} \tag{6}$$

where $\hat{\beta}_{\text{cost}}$ is the estimated coefficient on the cost attribute. Because the cost attribute is specified as a percentage change from the respondent's expected housing cost in 10 percent intervals, we convert marginal rates of substitution (MRS) into marginal willingness to pay (MWTP) in monetary terms by scaling with 10 percent of the mean reported monthly housing cost:

$$\text{MWTP}_k = \text{MRS}_k \times 0.10 \times \bar{PC} \tag{7}$$

where $\bar{PC}$ denotes the mean of respondents' stated planned monthly housing cost. These MWTP estimates represent the amount, in SEK per month, that respondents are willing to pay for improvements in each housing attribute, relative to their baseline housing cost expectations.

All models are estimated in R using the `logitr` package (Helveston 2023), which provides flexible estimation routines for both MNL and ML models using simulated maximum likelihood. For the ML models, we use 100 Sobol draws and multiple random starting values to ensure convergence to the global maximum (Train 2003).

3. Empirical Results

Our discussion begins with results from the baseline model, which excludes interaction terms and serves as a point of comparison for subsequent models that account for heterogeneity across household types. Table 3 presents the estimation results for respondents who own their housing unit. In the proceeding specifications, a positive (negative) coefficient indicates an increase (decrease) in average utility associated with the attribute level, relative to its reference level. Reference levels are specified as the least desirable alternative to aid in interpretation of the results. We report coefficient estimates from the multinomial logit model in column two and mixed logit estimates in column three to facilitate comparison across the model specifications.

Table 3: Baseline results - Owner

	MNL	ML		MRS	MWTP
		Mean	SD		
Green space: 5 km (vs 15 km)	0.57***	1.16***	0.12	0.22***	232.29
	(0.05)	(0.12)	(0.23)	(0.02)	
Green space: 500 m (vs 15 km)	1.12***	2.21***	1.24***	0.42***	440.66
	(0.05)	(0.16)	(0.25)	(0.03)	
Shops: 5 km (vs 15 km)	0.58***	1.01***	0.10	0.19***	200.82
	(0.05)	(0.13)	(0.20)	(0.03)	
Shops: 500 m (vs 15 km)	1.68***	3.12***	1.05***	0.59***	622.34
	(0.05)	(0.18)	(0.21)	(0.04)	
Transit stop: 600 m (vs 900 m)	0.20***	0.35**	−0.76***	0.07**	69.56
	(0.05)	(0.11)	(0.20)	(0.02)	
Transit stop: 300 m (vs 900 m)	0.51***	1.13***	0.79***	0.21***	225.03
	(0.05)	(0.12)	(0.18)	(0.02)	
Parking: reserved garage (vs none)	1.53***	3.08***	0.56*	0.59***	615.13
	(0.06)	(0.20)	(0.25)	(0.04)	
Parking: reserved space (vs none)	1.30***	2.56***	−2.29***	0.49***	511.17
	(0.06)	(0.17)	(0.17)	(0.03)	
Price	−2.83***	−5.26***			
	(0.16)	(0.35)			
Num. obs.	7110	7110			
Log Likelihood	−3570.76	−3136.51			
AIC	7159.53	6363.02			
BIC	7221.35	6672.14			
McFadden R^2	0.28	0.36			
LR $\chi 2$ (df=9)	2715.02	3583.53			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Although both model specifications yield similar patterns in coefficient signs and levels of statistical significance, the mixed logit estimates (column three) are systematically larger in magnitude than those from the multinomial logit model. In terms of model performance, the mixed logit specification demonstrates superior fit across all criteria, including log-likelihood, McFadden's pseudo-R2, AIC, and BIC. Additionally, several standard deviation estimates for the random parameters (column four) are statistically significant, confirming the presence of preference heterogeneity for certain attributes. We therefore adopt the mixed logit specification as our preferred model for all subsequent tests and discussion.

Estimates of MRS are presented in column five and are calculated as described in Equation (6). To obtain MWTP estimates, MRS values are multiplied by 10 percent of the average planned housing cost associated with the group under analysis, per Equation (7). These monetary estimates are presented in column six.

As shown in Table 3, owners are, on average, willing to pay up to 440 SEK per month to avoid residing 15 kilometres from the nearest green area. Proximity to shops is valued to a greater extent, with owners willing to pay over 620 SEK per month to avoid greater distances from retail services. Table 4 presents results for respondents who do not own their current housing. Compared to owners, renters place a greater value on proximity to green space, with a MWTP that is 177 SEK greater. Both groups place increased importance of proximity to shops, with renters showing an average MWTP of 636 SEK for locations within 500 meters, compared to 622 SEK for owners. Reserved parking emerges as a essential attribute for both tenure groups. However, owners are

willing to pay substantially more for access to a reserved garage (615 SEK vs 337 SEK), suggesting stronger preferences for secure or private vehicle storage among this group.

Table 4: Baseline results - Renter

	MNL	ML Mean	ML SD	MRS	MWTP)
Green space: 5 km (vs 15 km)	0.75***	4.80***	2.30***	0.36***	377.54
	(0.11)	(1.29)	(0.65)	(0.06)	
Green space: 500 m (vs 15 km)	1.25***	7.85***	6.64***	0.59***	618.10
	(0.12)	(2.07)	(1.85)	(0.08)	
Shops: 5 km (vs 15 km)	0.77***	3.67**	0.36	0.27***	288.71
	(0.12)	(1.15)	(0.46)	(0.06)	
Shops: 500 m (vs 15 km)	1.68***	8.09***	6.47***	0.61***	636.95
	(0.12)	(2.12)	(1.81)	(0.09)	
Transit stop: 600 m (vs 900 m)	0.30*	1.57**	0.99	0.12**	123.73
	(0.12)	(0.55)	(0.63)	(0.04)	
Transit stop: 300 m (vs 900 m)	0.68***	2.01***	1.64**	0.15***	158.21
	(0.12)	(0.52)	(0.57)	(0.03)	
Parking: reserved garage (vs none)	0.92***	4.28***	−0.02	0.32***	337.04
	(0.12)	(0.97)	(0.32)	(0.04)	
Parking: reserved space (vs none)	0.94***	3.90***	−4.64***	0.29***	306.98
	(0.13)	(0.82)	(1.35)	(0.03)	
Price	−4.10***	−13.34***			
	(0.37)	(2.61)			
Num. obs.	1458	1458			
Log Likelihood	−721.69	−612.41			
AIC	1461.39	1314.82			
BIC	1508.95	1552.64			
McFadden R²	0.29	0.39			
LR χ2 (df=9)	577.83	796.40			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

3.1. Heterogenity Models

Thus far our result point towards the presence of heterogeneous effects for housing preferences in regard to housing tenure. We next expand our testing by introducing cross-product terms in our mixed logit models framework to identify sources of preference heterogeneity in our sample.

3.1.1. Age effects

To investigate whether preferences vary systematically across the ageing process, we interacted all attribute levels with three age categories (55–64, 65–74, and 75+ years). The youngest group (55–64) serves as the reference category, and therefore the main effects represent their preferences. The interaction terms capture differences in preferences for the older groups relative to this baseline.

Table 5 shows that respondents aged 55–64 exhibit strong positive utilities and high MWTP values across all attributes. Among owners, the 55–64 group was willing to pay around 490 SEK/month to live within 500 m of shops, 429 SEK/month for green space within 500 m, and 453 SEK/month for a reserved garage. Among renters, corresponding values were higher, approximately 545 SEK, 498 SEK, and 323 SEK, respectively.

For the 65–74 group, MWTP values remain largely comparable to those of the 55–64 reference group. Differences are small, suggesting stability in preferences across early old age. By contrast, respondents aged 75 and above display notable shifts in emphasis. Among homeowners, MWTP for proximity to green areas and shops declines modestly to roughly 410 SEK and 240 SEK, while the value of parking amenities increases substantially, up to 700 SEK/month for a reserved garage and 640 SEK/month for a reserved space.

Among renters, MWTP in the oldest age group remains high for proximity and comfort attributes: approximately 580 SEK/month for nearby green areas, 400 SEK for shops, and 390–400 SEK for reserved parking. Figure 3 illustrates MWTP values for across the age groups.

Results suggest that while all age groups value proximity and amenities, the relative importance of these attributes evolves across the ageing spectrum. Younger and mid-age respondents emphasize spatial accessibility, being near shops, services, and green areas, whereas the oldest group increasingly private parking and local environmental quality.

Table 5: Hetergeneity models: Age groups

	Owners		Renters	
	Coef.	MWTP	Coef.	MWTP
Green space: 5 km (vs 15 km)	1.34***	207.89***	4.56**	267.46*
	(0.24)	(41.42)	(1.50)	(115.84)
Green space: 500 m (vs 15 km)	2.77***	428.67***	8.48***	497.75**
	(0.31)	(59.48)	(2.11)	(166.14)
Shops: 5 km (vs 15 km)	1.24***	192.54***	4.44**	260.62*
	(0.28)	(51.64)	(1.62)	(114.51)
Shops: 500 m (vs 15 km)	3.16***	490.07***	9.29***	545.39***
	(0.30)	(62.12)	(1.94)	(119.10)
Transit stop: 600 m (vs 900 m)	−0.06	−9.81	2.61*	153.07*
	(0.24)	(36.17)	(1.04)	(67.29)
Transit stop: 300 m (vs 900 m)	1.13***	175.41***	3.83***	224.71**
	(0.26)	(42.69)	(1.08)	(73.92)
Parking: reserved garage (vs none)	2.93***	453.25***	4.51***	264.94***
	(0.32)	(62.52)	(1.26)	(60.67)
Parking: reserved space (vs none)	2.17***	336.81***	5.49***	322.26***
	(0.29)	(45.36)	(1.29)	(64.76)
Price	−6.78***		−17.89***	
	(0.71)		(4.53)	
Green space: 5 km (vs 15 km) × Age 65–74	−0.08	237.96***	−1.06	256.82***
	(0.26)	(35.32)	(1.28)	(58.54)
Green space: 500 m (vs 15 km) × Age 65–74	−0.47	432.51***	−2.66	427.55***
	(0.34)	(52.38)	(1.58)	(70.44)
Shops: 5 km (vs 15 km) × Age 65–74	−0.34	169.27***	−1.01	252.39***
	(0.34)	(44.61)	(1.43)	(67.91)
Shops: 500 m (vs 15 km) × Age 65–74	−0.33	533.49***	−1.71	556.94***
	(0.31)	(54.13)	(1.35)	(74.88)
Transit stop: 600 m (vs 900 m) × Age 65–74	0.27	37.93	−1.98	46.54
	(0.27)	(30.21)	(1.16)	(42.94)
Transit stop: 300 m (vs 900 m) × Age 65–74	−0.02	208.46***	−2.90*	68.64
	(0.31)	(37.30)	(1.14)	(35.73)
Parking: reserved garage (vs none) × Age 65–74	0.35	616.50***	−1.44	226.28***
	(0.36)	(60.90)	(1.28)	(45.21)
Parking: reserved space (vs none) × Age 65–74	0.44	490.85***	−2.20	242.12***
	(0.33)	(47.56)	(1.34)	(45.88)
Price × Age 65–74	1.19		3.60	
	(0.80)	(182.90)	(4.18)	(388.70)
Green space: 5 km (vs 15 km) × Age 75+	−0.38	226.54***	0.33	477.80***
	(0.27)	(45.28)	(1.22)	(101.36)
Green space: 500 m (vs 15 km) × Age 75+	−1.02**	411.97***	−2.54	580.64***
	(0.34)	(66.71)	(1.72)	(127.64)
Shops: 5 km (vs 15 km) × Age 75+	−0.21	242.51***	−1.39	298.16**
	(0.32)	(60.27)	(1.45)	(96.09)
Shops: 500 m (vs 15 km) × Age 75+	−0.18	703.83***	−5.18**	402.18***
	(0.32)	(88.20)	(1.62)	(99.04)
Transit stop: 600 m (vs 900 m) × Age 75+	0.66*	139.76**	−2.46	14.04
	(0.29)	(46.13)	(1.30)	(68.86)
Transit stop: 300 m (vs 900 m) × Age 75+	0.02	270.91***	−1.80	198.01**
	(0.30)	(54.04)	(1.15)	(70.29)
Parking: reserved garage (vs none) × Age 75+	0.08	708.28***	−0.39	402.94***
	(0.37)	(86.08)	(1.23)	(86.56)
Parking: reserved space (vs none) × Age 75+	0.55	642.81***	−1.47	392.44***
	(0.34)	(71.37)	(1.25)	(75.99)
Price × Age 75+	2.32**		7.14	
	(0.82)	(229.50)	(4.46)	(504.58)
Num. obs.	7110		1458	
Log Likelihood	−3119.74		−604.05	
AIC	6365.47		1334.09	
BIC	6798.23		1667.04	
McFadden R²	0.37		0.40	
LR χ2 (df=63)	3617.08		813.13	
p-value (LR)	0.00		0.00	

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$
Note: MWTP denotes marginal willingness to pay. Interaction terms indicate differential effects by age group.
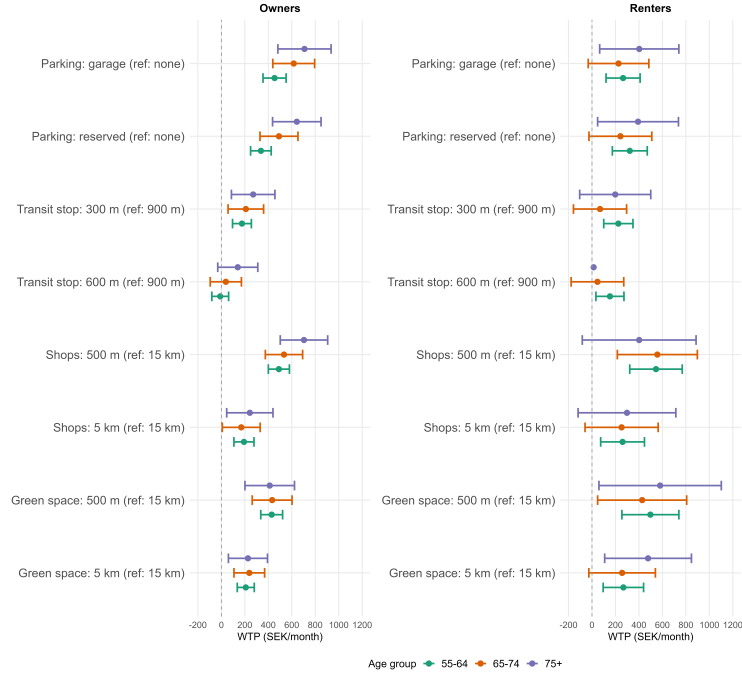
12

Figure 3: WTP estimates from model 2

3.1.2. Income effects

3.1.3. Education effects

4. Discussion

This study explored preference heterogeneity among older individuals considering relocation in Sweden. To our knowledge, this is the first study to utilize a discrete choice experiment to examine locational preferences in Sweden.

5. Limitations

6. Conclusions

As the global population continues to age, understanding the housing preferences of older demographics becomes increasingly crucial for the planning and development of future societies. In Sweden, where a significant portion of older individuals prefer to live in their own homes, the need for appropriate housing options for this segment is becoming more pronounced as this demographic group expands.

This study utilized a discrete choice experiment to delve into factors influencing the housing choices of older individuals in Sweden considering relocation. By presenting hypothetical scenarios that varied in locational attributes such as healthcare facilities, public transportation, green spaces, social amenities, and natural surroundings, we were able to estimate willingness-to-pay values and compare differences across various groups.

7. Appendix

Table 6: Baseline results - Men

| | MNL | ML | | MRS | MWTP |
		Mean	SD		
Green space: 5 km (vs 15 km)	0.65***	1.10***	−0.03	0.20***	206.44
	(0.07)	(0.17)	(0.32)	(0.03)	
Green space: 500 m (vs 15 km)	1.01***	1.78***	0.20	0.32***	334.09
	(0.07)	(0.21)	(0.53)	(0.04)	
Shops: 5 km (vs 15 km)	0.69***	1.14***	0.13	0.20***	213.22
	(0.08)	(0.18)	(0.33)	(0.04)	
Shops: 500 m (vs 15 km)	1.76***	3.46***	0.76	0.62***	649.92
	(0.07)	(0.29)	(0.41)	(0.06)	
Transit stop: 600 m (vs 900 m)	0.20**	0.33*	0.01	0.06*	62.86
	(0.08)	(0.15)	(0.31)	(0.03)	
Transit stop: 300 m (vs 900 m)	0.41***	1.01***	0.51*	0.18***	188.91
	(0.07)	(0.18)	(0.24)	(0.03)	
Parking: reserved garage (vs none)	1.71***	3.42***	1.04***	0.61***	642.86
	(0.08)	(0.28)	(0.25)	(0.06)	
Parking: reserved space (vs none)	1.43***	2.85***	2.43***	0.51***	534.90
	(0.08)	(0.25)	(0.25)	(0.05)	
Price	−3.01***	−5.59***			
	(0.22)	(0.49)			
Num. obs.	3852	3852			
Log Likelihood	−1889.12	−1654.13			
AIC	3796.25	3398.26			
BIC	3852.55	3679.79			
McFadden R^2	0.29	0.38			
LR $\chi 2$ (df=9)	1561.76	2031.75			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 7: Baseline results - Women

| | MNL | ML | | MRS | MWTP |
		Mean	SD		
Green space: 5 km (vs 15 km)	0.56***	1.25***	0.25	0.23***	242.97
	(0.06)	(0.17)	(0.54)	(0.03)	
Green space: 500 m (vs 15 km)	1.24***	2.76***	1.24**	0.51***	537.22
	(0.06)	(0.24)	(0.43)	(0.05)	
Shops: 5 km (vs 15 km)	0.55***	1.07***	0.86**	0.20***	207.50
	(0.06)	(0.21)	(0.29)	(0.04)	
Shops: 500 m (vs 15 km)	1.61***	3.10***	1.33***	0.57***	602.77
	(0.06)	(0.24)	(0.39)	(0.05)	
Transit stop: 600 m (vs 900 m)	0.24***	0.50***	−0.09	0.09***	98.21
	(0.07)	(0.14)	(0.32)	(0.03)	
Transit stop: 300 m (vs 900 m)	0.65***	1.38***	1.03***	0.26***	268.90
	(0.06)	(0.16)	(0.22)	(0.03)	
Parking: reserved garage (vs none)	1.21***	2.43***	0.41	0.45***	473.25
	(0.07)	(0.20)	(0.26)	(0.04)	
Parking: reserved space (vs none)	1.10***	1.96***	−2.17***	0.36***	381.55
	(0.07)	(0.18)	(0.25)	(0.03)	
Price	−3.07***	−5.40***			
	(0.19)	(0.44)			
Num. obs.	4761	4761			
Log Likelihood	−2431.00	−2122.98			
AIC	4880.00	4335.96			
BIC	4938.21	4627.03			
McFadden R²	0.26	0.36			
LR χ2 (df=9)	1738.15	2354.18			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 8: Baseline results - 55-64

| | MNL | ML | | MRS | MWTP |
		Mean	SD		
Green space: 5 km (vs 15 km)	0.63***	1.95***	−0.63	0.19***	200.97
	(0.10)	(0.40)	(0.48)	(0.03)	
Green space: 500 m (vs 15 km)	1.45***	4.81***	4.35***	0.47***	495.74
	(0.11)	(0.73)	(0.82)	(0.05)	
Shops: 5 km (vs 15 km)	0.65***	1.67***	0.40	0.16***	172.18
	(0.12)	(0.41)	(0.39)	(0.04)	
Shops: 500 m (vs 15 km)	1.76***	4.94***	1.07*	0.48***	509.16
	(0.11)	(0.65)	(0.44)	(0.05)	
Transit stop: 600 m (vs 900 m)	0.10	0.59	1.90***	0.06	60.54
	(0.11)	(0.33)	(0.44)	(0.03)	
Transit stop: 300 m (vs 900 m)	0.60***	1.89***	−0.08	0.19***	194.66
	(0.11)	(0.38)	(0.47)	(0.03)	
Parking: reserved garage (vs none)	1.31***	3.84***	−1.66***	0.38***	395.29
	(0.11)	(0.68)	(0.45)	(0.05)	
Parking: reserved space (vs none)	1.04***	3.19***	3.84***	0.31***	328.33
	(0.12)	(0.54)	(0.70)	(0.04)	
Price	−3.99***	−10.19***			
	(0.36)	(1.42)			
Num. obs.	1728	1728			
Log Likelihood	−817.95	−689.10			
AIC	1653.91	1468.20			
BIC	1703.00	1713.66			
McFadden R²	0.32	0.42			
LR χ2 (df=9)	759.61	1017.32			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 9: Baseline results - 65-74

	MNL	ML Mean	ML SD	MRS	MWTP
Green space: 5 km (vs 15 km)	0.62***	1.38***	0.38	0.20***	214.30
	(0.07)	(0.22)	(0.24)	(0.03)	
Green space: 500 m (vs 15 km)	1.18***	2.66***	0.53	0.39***	411.97
	(0.07)	(0.33)	(0.40)	(0.04)	
Shops: 5 km (vs 15 km)	0.54***	1.05***	0.52	0.16***	163.48
	(0.07)	(0.24)	(0.38)	(0.04)	
Shops: 500 m (vs 15 km)	1.70***	3.72***	2.22***	0.55***	576.94
	(0.08)	(0.39)	(0.29)	(0.05)	
Transit stop: 600 m (vs 900 m)	0.20**	0.24	−0.67**	0.04	36.99
	(0.08)	(0.16)	(0.24)	(0.02)	
Transit stop: 300 m (vs 900 m)	0.54***	1.10***	1.19***	0.16***	170.35
	(0.07)	(0.19)	(0.25)	(0.03)	
Parking: reserved garage (vs none)	1.51***	3.61***	1.08***	0.53***	559.50
	(0.08)	(0.37)	(0.25)	(0.05)	
Parking: reserved space (vs none)	1.26***	2.69***	2.55***	0.40***	416.65
	(0.08)	(0.28)	(0.38)	(0.03)	
Price	−3.09***	−6.77***			
	(0.22)	(0.65)			
Num. obs.	3744	3744			
Log Likelihood	−1868.23	−1646.37			
AIC	3754.45	3382.73			
BIC	3810.50	3662.99			
McFadden R²	0.28	0.37			
LR χ2 (df=9)	1453.83	1897.56			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 10: Baseline results - 75+

	MNL	ML		MRS	MWTP
		Mean	SD		
Green space: 5 km (vs 15 km)	0.55***	1.10***	−0.17	0.23***	238.16
	(0.07)	(0.21)	(0.28)	(0.05)	
Green space: 500 m (vs 15 km)	0.95***	2.01***	0.40	0.41***	434.70
	(0.08)	(0.25)	(0.36)	(0.06)	
Shops: 5 km (vs 15 km)	0.66***	1.21***	0.47	0.25***	261.18
	(0.08)	(0.22)	(0.28)	(0.06)	
Shops: 500 m (vs 15 km)	1.59***	3.34***	0.13	0.69***	722.79
	(0.08)	(0.33)	(0.25)	(0.08)	
Transit stop: 600 m (vs 900 m)	0.32***	0.55***	−0.70***	0.11**	119.57
	(0.08)	(0.16)	(0.21)	(0.04)	
Transit stop: 300 m (vs 900 m)	0.52***	1.20***	0.57*	0.25***	258.91
	(0.08)	(0.20)	(0.25)	(0.04)	
Parking: reserved garage (vs none)	1.42***	3.04***	0.61*	0.63***	658.01
	(0.08)	(0.29)	(0.24)	(0.08)	
Parking: reserved space (vs none)	1.33***	2.76***	−2.61***	0.57***	597.03
	(0.09)	(0.26)	(0.31)	(0.06)	
Price	−2.51***	−4.86***			
	(0.23)	(0.55)			
Num. obs.	3141	3141			
Log Likelihood	−1632.69	−1425.97			
AIC	3283.38	2941.93			
BIC	3337.85	3214.29			
McFadden R^2	0.25	0.35			
LR χ2 (df=9)	1088.97	1502.42			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 11: Baseline results - Retired

| | MNL | ML | | MRS | MWTP |
		Mean	SD		
Green space: 5 km (vs 15 km)	0.60***	1.00***	0.01	0.20***	208.42
	(0.05)	(0.12)	(0.23)	(0.03)	
Green space: 500 m (vs 15 km)	1.10***	1.86***	0.27	0.37***	388.08
	(0.05)	(0.15)	(0.43)	(0.03)	
Shops: 5 km (vs 15 km)	0.58***	1.04***	1.07***	0.21***	216.03
	(0.06)	(0.15)	(0.20)	(0.03)	
Shops: 500 m (vs 15 km)	1.65***	3.10***	0.43	0.62***	646.36
	(0.06)	(0.20)	(0.29)	(0.05)	
Transit stop: 600 m (vs 900 m)	0.26***	0.49***	1.00***	0.10***	103.05
	(0.06)	(0.11)	(0.19)	(0.02)	
Transit stop: 300 m (vs 900 m)	0.55***	1.05***	0.18	0.21***	218.75
	(0.06)	(0.12)	(0.17)	(0.03)	
Parking: reserved garage (vs none)	1.50***	2.92***	1.11***	0.58***	609.32
	(0.06)	(0.18)	(0.20)	(0.05)	
Parking: reserved space (vs none)	1.35***	2.52***	2.14***	0.50***	525.65
	(0.06)	(0.19)	(0.18)	(0.04)	
Price	−2.88***	−5.03***			
	(0.16)	(0.38)			
Num. obs.	6642	6642			
Log Likelihood	−3364.20	−2978.20			
AIC	6746.39	6046.40			
BIC	6807.60	6352.46			
McFadden R^2	0.27	0.35			
LR χ2 (df=9)	2479.37	3251.36			
p-value (LR)	0.00	0.00			

***$p < 0.001$; **$p < 0.01$; *$p < 0.05$

Table 12: Baseline results - not Retired

| | MNL | ML | | MRS | MWTP |
		Mean	SD		
Green space: 5 km (vs 15 km)	0.58***	1.40***	0.52	0.16***	171.51
	(0.09)	(0.27)	(0.28)	(0.03)	
Green space: 500 m (vs 15 km)	1.29***	3.27***	−0.86**	0.38***	399.24
	(0.10)	(0.42)	(0.29)	(0.04)	
Shops: 5 km (vs 15 km)	0.69***	1.63***	0.65*	0.19***	199.17
	(0.11)	(0.36)	(0.33)	(0.04)	
Shops: 500 m (vs 15 km)	1.74***	4.48***	1.51***	0.52***	546.79
	(0.10)	(0.57)	(0.36)	(0.06)	
Transit stop: 600 m (vs 900 m)	0.11	0.37	0.26	0.04	44.95
	(0.11)	(0.25)	(0.43)	(0.03)	
Transit stop: 300 m (vs 900 m)	0.54***	1.61***	1.39***	0.19***	196.08
	(0.10)	(0.27)	(0.25)	(0.03)	
Parking: reserved garage (vs none)	1.20***	3.16***	−1.60***	0.37***	385.74
	(0.10)	(0.40)	(0.33)	(0.05)	
Parking: reserved space (vs none)	0.88***	2.32***	3.38***	0.27***	283.63
	(0.11)	(0.38)	(0.51)	(0.04)	
Price	−3.60***	−8.60***			
	(0.33)	(1.10)			
Num. obs.	1935	1935			
Log Likelihood	−946.20	−800.34			
AIC	1910.40	1690.69			
BIC	1960.51	1941.24			
McFadden R²	0.29	0.40			
LR $\chi 2$ (df=9)	790.08	1081.79			
p-value (LR)	0.00	0.00			

$^{***}p < 0.001$; $^{**}p < 0.01$; $^{*}p < 0.05$

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