Statistical Computation and Simulation

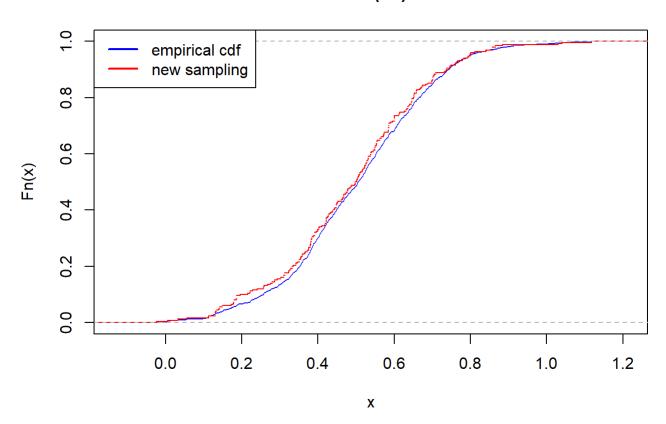
Generating empirical cdf from some data (using distr package) using it to generate new random samples

```
# install.packages('distr')
n <- 1e3
xs \leftarrow rnorm(n, .5, .2) # a random sample
library(distr)
## Warning: package 'distr' was built under R version 3.5.1
## Loading required package: startupmsg
## Warning: package 'startupmsg' was built under R version 3.5.1
## Utilities for Start-Up Messages (version 0.9.5)
## For more information see ?"startupmsg", NEWS("startupmsg")
## Loading required package: sfsmisc
## Warning: package 'sfsmisc' was built under R version 3.5.1
## Object Oriented Implementation of Distributions (version 2.7.0)
## Attention: Arithmetics on distribution objects are understood as operations on corresponding
random variables (r.v.s); see distrARITH().
## Some functions from package 'stats' are intentionally masked ---see distrMASK().
## Note that global options are controlled by distroptions() ---c.f. ?"distroptions".
## For more information see ?"distr", NEWS("distr"), as well as
    http://distr.r-forge.r-project.org/
## Package "distrDoc" provides a vignette to this package as well as to several extension packag
es; try vignette("distr").
##
## Attaching package: 'distr'
## The following objects are masked from 'package:stats':
##
```

df, qqplot, sd

```
emp.cdf <- DiscreteDistribution(xs)</pre>
                                             # fit an empirical cdf
new.xs <- emp.cdf@r(250)
                                             # generate new samples
# plot original samples vs new samples
plot(ecdf(xs), col="blue")
plot(ecdf(new.xs), col="red", pch=".", add=T)
legend("topleft",c("empirical cdf","new sampling"), col=c("blue","red"), lwd=2)
```

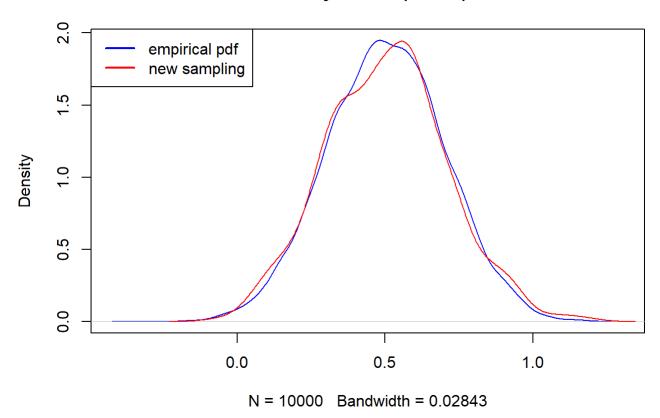
ecdf(xs)



Another alternative is using the density function which uses (by default gaussian) kernels to estimate the density of the sample

```
n <- 1e4
xs \leftarrow rnorm(n, .5, .2)
d <- density(xs)</pre>
new_xs <- sample(d$x, replace = TRUE, prob=d$y)</pre>
#plot original sample vs new sample
plot(density(xs), col= "blue")
lines(density(new xs), col = "red")
legend("topleft",c("empirical pdf", "new sampling"), col = c("blue","red"), lwd =2)
```

density.default(x = xs)



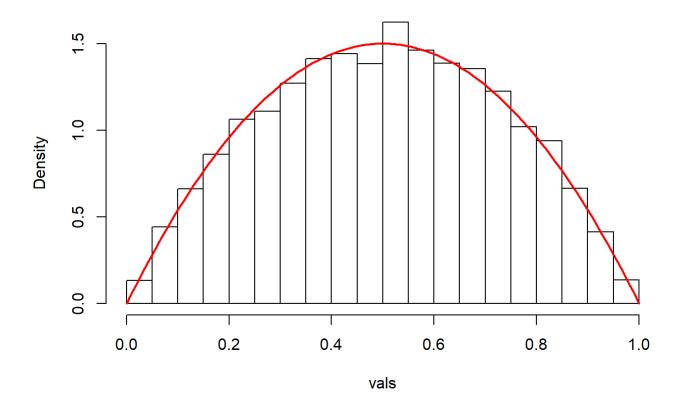
Acceptance - Rejection method

Generate n samples from f using rejection sampling with g (rg samples from g)

```
accept.reject <- function(f, c, g, rg, n) {</pre>
  n.accepts
                  <- 0
  result.sample <- rep(NA, n)</pre>
  while (n.accepts < n) {</pre>
    y \leftarrow rg(1)
                                 # step 1
    u <- runif(1,0,1)
                                # step 2
    if (u < f(y)/(c*g(y))) \{ # step 3 (accept) \}
       n.accepts <- n.accepts+1
       result.sample[n.accepts] = y
    }
  result.sample
}
```

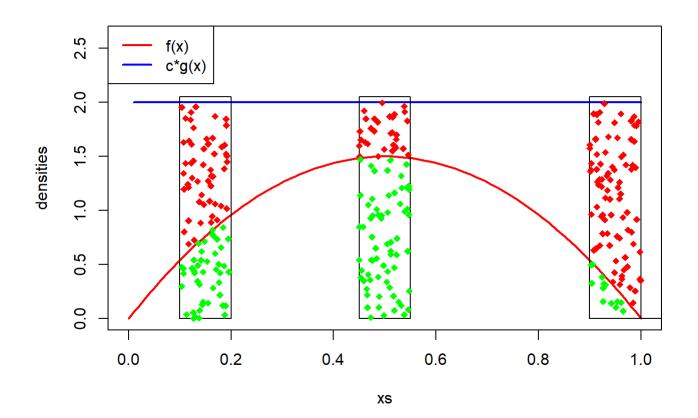
```
f <- function(x) 6*x*(1-x) # pdf of Beta(2,2), maximum density is 1.5
g \leftarrow function(x) x/x
                                \# q(x) = 1 but in vectorized version
rg <- function(n) runif(n,0,1) # uniform, in this case
                                # c=2 since f(x) <= 2 g(x)
vals <- accept.reject(f, c, g, rg, 10000)</pre>
                                          # generating sample values
# Checking if it went well
hist(vals, breaks=30, freq=FALSE, main="Sample vs true Density")
xs <- seq(0, 1, len=100)
lines(xs, dbeta(xs,2,2), col="red", lwd=2) # fitting beta distribution
```

Sample vs true Density



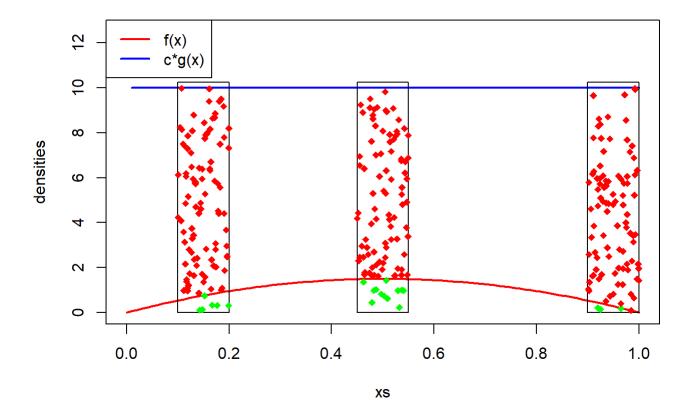
Visualizing the method of accepting (green dots) and rejecting (red dots) at some specified segments:

```
xs <- seq(0,1, len =100)
#plot(xs)
#plot(dbeta(xs,2,2))
                                # Plotting y vs x; xs sets the x range from 0 -1
\#plot(xs,dbeta(xs,2,2))
plot(xs, dbeta(xs,2,2), ylim=c(0,c*1.3), type="1", col="red", lwd=2, ylab="densities")
lines(xs)
lines(xs, c*g(xs), type="l", col="blue", lwd=2)
legend("topleft",c("f(x)","c*g(x)"), col=c("red","blue"), lwd=2)
draw.segment <- function(begin.segment, end.segment) {</pre>
  segments(c(begin.segment,end.segment,end.segment,begin.segment), c(0,0,c*1.025,c*1.025),
           c(end.segment,end.segment,begin.segment), c(0,c*1.025,c*1.025,0))
  n.pts <- 100
  us <- runif(n.pts, 0, 1)
 ys <- begin.segment + rg(n.pts)*(end.segment-begin.segment)</pre>
  accepted <- us < f(ys)/(c*g(ys))
  points(ys, c*us, col=ifelse(accepted, "green", "red"), pch=18)
}
draw.segment(0.10, 0.20)
draw.segment(0.45, 0.55)
draw.segment(0.90, 1.00)
```



The higher the density of f the more points are accepted however if cg(x) >> f(x) we will have lots of rejections, which will decrease the quality of simulation for same amount of points, Checking for c = 10:

```
c <- 10
xs \leftarrow seq(0,1, len = 100)
plot(xs, dbeta(xs,2,2), ylim=c(0,c*1.25), type="l", col="red", lwd=2, ylab="densities")
lines(xs, c*g(xs), type="l", col="blue", lwd=2)
legend("topleft",c("f(x)","c*g(x)"), col=c("red","blue"), lwd=2)
draw.segment(0.10, 0.20)
draw.segment(0.45, 0.55)
draw.segment(0.90, 1.00)
```



Adaptive Rejection Sampling

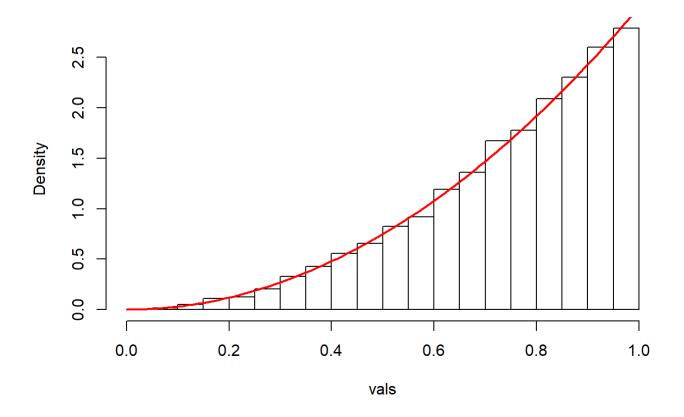
This number of pints is not enough to get an estimate with the quality of previous example.

```
#install.packages("ars")
library(ars)
```

```
## Warning: package 'ars' was built under R version 3.5.1
```

```
\log f \leftarrow function(x) \log(3*x^2) \# \log(f(x))
log.f.dx \leftarrow function(x) 2/x
                               \# d/dx \log(f(x))
vals <- ars(1e4,
                                # how many points are required
            log.f, log.f.dx,
                                # the needed functions
            lb=TRUE, ub=TRUE, # there are a lower and upper bounds for the pdf
            xlb=0, xub=1,
                                # and which are those bounds
            x=c(.1,.5,.9))
                                # some initial points inside the pdf
# Checking if it went well
hist(vals, breaks=30, freq=FALSE, main="Sample vs true Density")
xs <- seq(0, 1, len=100)
curve(3*x^2, 0, 1, col="red", lwd=2, add=T)
```

Sample vs true Density



Monte Carlo Integration

pre-condition: a < b

```
MC.simple.est <- function(g, a, b, n=1e4) {</pre>
                            # step 1
  xi <- runif(n,a,b)</pre>
  g.mean \leftarrow mean(g(xi)) # step 2
                             # step 3
  (b-a)*g.mean
}
g \leftarrow function(x) exp(-x)
MC.simple.est(g,2,4)
```

```
## [1] 0.1169242
```

Importance Sampling

Two problems with the previous method: does not apply to unbounded intervals performs poorly if the pdf is not very uniform, namely at distribution tails

Eg: $X \sim N(0,1)$, estimate P(X>4.5)

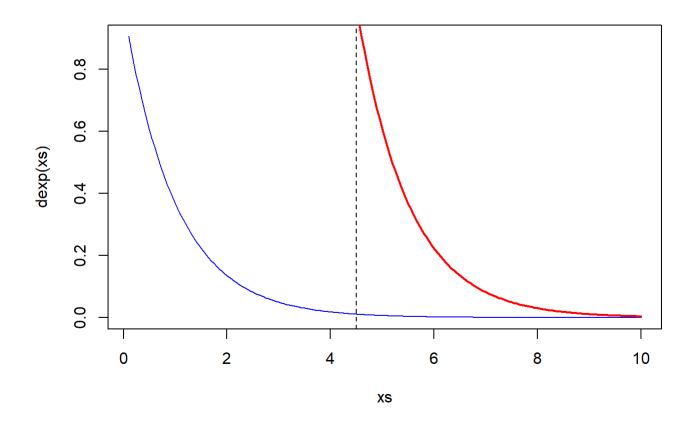
```
# True - value
pnorm(4.5, lower.tail=FALSE)
                                      # theta (could also be computed by 1-pnorm(4.5))
```

```
## [1] 3.397673e-06
```

```
# MC estimation
n <- 1e4
indicators <- rnorm(n) > 4.5
sum(indicators) / n
                   # hat.theta
```

```
## [1] 0
```

```
# rh generates sample from candidate pdf
i.sampling <- function(f, g, h, rh, n=1e4) {</pre>
  ys \leftarrow rh(n)
  mean(g(ys)*f(ys)/h(ys))
}
xs \leftarrow seq(0.1,10,by=0.05)
plot(xs,dexp(xs),col="blue", type="l") # the exponential pdf
lines(xs,exp(-(xs-4.5)),col="red",lwd=2) # the truncated pdf
abline(v=4.5,lty=2)
```



```
\# plot of target pdf g (in blue) and the candidate pdf h (in red)
g <- dnorm
h \leftarrow function(x) exp(-(x-4.5))
xs <- seq(4.5,20,by=0.05)
plot(xs,g(xs),col="blue", type="l", ylim=c(0,0.5e-4))
lines(xs,h(xs),col="red",lwd=2)
```

