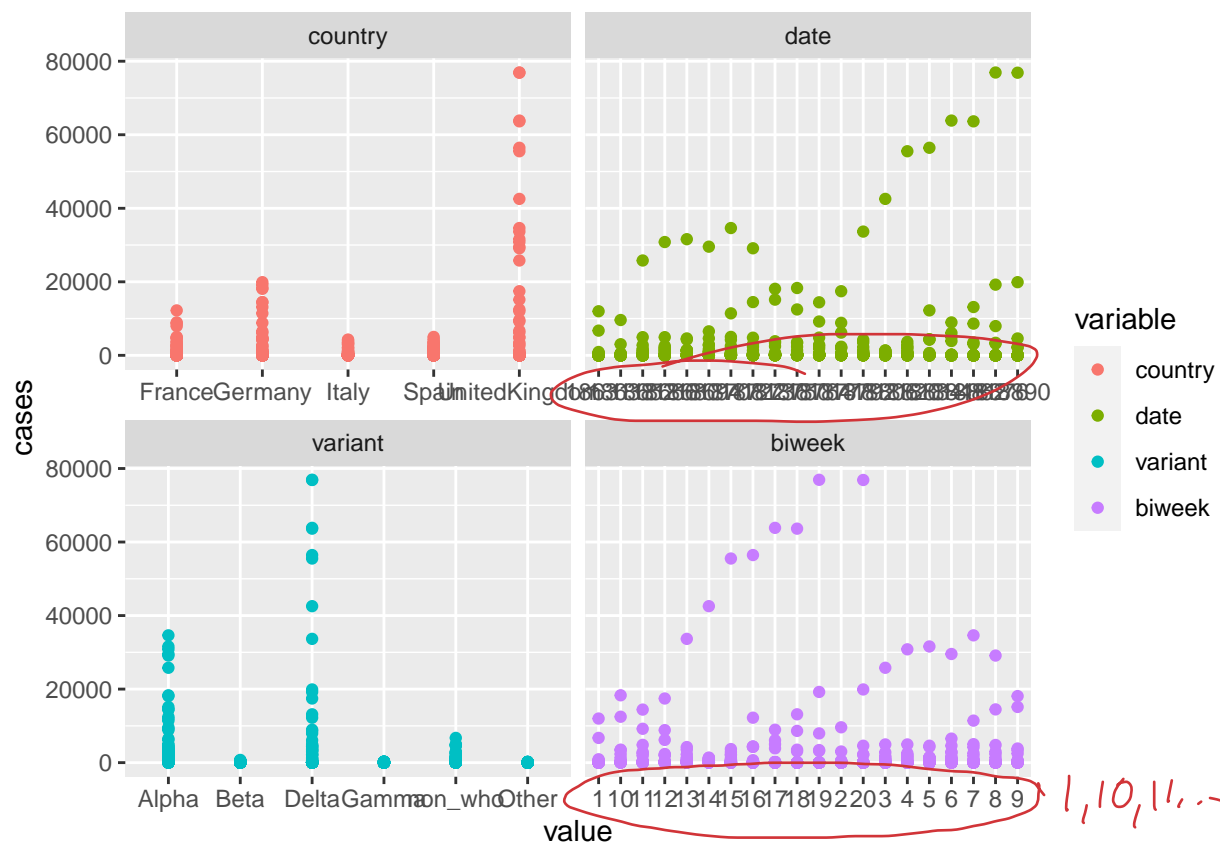


# MTH3041 Bayesian Statistics, Philosophy and Practice CourseWork

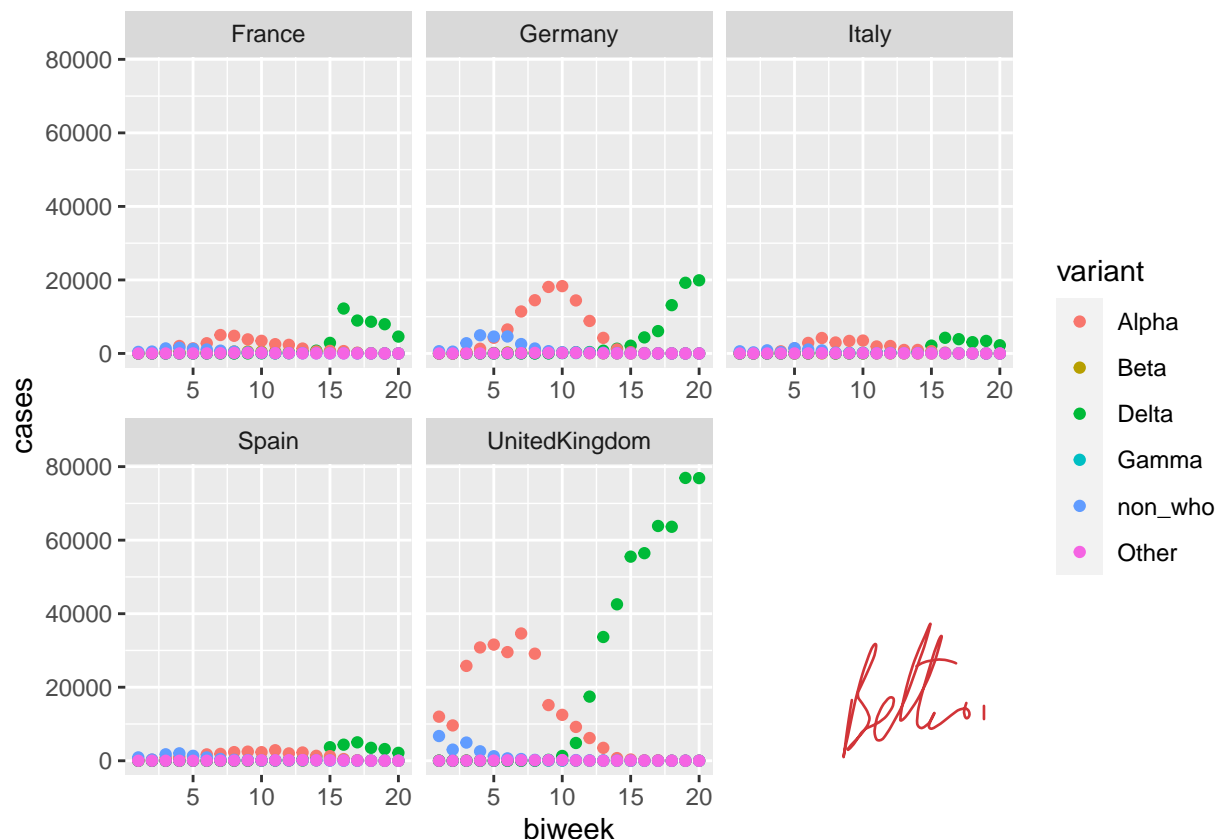
198728

1)

```
variants_5Largest =
  rbind(variants_Germany, variants_UnitedKingdom, variants_France, variants_Italy, variants_Spain)
# view(variants_5Largest)
variants_5Largest %>% melt(id.vars="cases") %>%
  ggplot()+
  geom_point(aes(x=value, y=cases, colour=variable))+
  facet_wrap(~variable, scales="free_x")
```



```
ggplot(variants_5Largest)+
  geom_point(aes(x = biweek, y = cases, colour = variant))+
  facet_wrap(~country, scales = "free_x")
```



```
ggplot(variants_5Largest)+
  geom_point(aes(x = date, y = cases, colour = variant))+
  facet_wrap(~country, scales = "free_x")
```

```
ggplot(variants_5Largest)+
  geom_point(aes(x = biweek, y = cases, colour = country))+
  facet_wrap(~variant, scales = "free_x")
```

These plots show that the data of 'date' and 'biweek' are the exact same data, since they are a one to one translation of each other, so there is no need to consider both of them in the model. So only 'biweek' will be considered as a variable in the model. ✓

Cases can only be whole numbers so we are modelling counts. So let  $y = \text{cases}$ , so for a count distribution, can use poisson or negative binomial. But if poisson then  $Mean[y] = Var[y]$ . Which makes the likelihoods flat everywhere if there's no poisson where the mean and variance are equal that actually explains the data, so a poisson regression should not be used. So instead use a negative binomial regression, since it allows  $Mean(y) \neq Var(y)$ . So  $y | \beta, x \sim NegBin(\lambda(\beta, x), \phi)$ . So  $P(y = K | \lambda, \phi) \propto (\frac{\mu}{\phi + \mu})^K (\frac{\phi}{\phi + \mu})^\phi$ . So  $E[y] = \lambda \geq 0$  and  $Var[y] \neq \lambda$  but  $Var[y] = \lambda(1 + \frac{\lambda}{\phi}) > \lambda$  for  $\phi \geq 0$ . So  $log(\lambda(\beta, x)) = \sum_{i=0}^p \beta_i x_i$ . So for posterior distribution:  $\pi(\beta, \Sigma) \neq \pi(B, \sigma, \phi)$ . So let  $y_i = \text{cases}_i$  ✓

```
nrow(variants_5Largest)
```

So  $i = 1, \dots, 600$ . So  $y_i | \beta \sim NegBin(\lambda(\beta, x), \phi)$ , where  $\lambda$  is the inverse link. So  $log(\lambda(\beta, x)) = \beta_0 + \beta_1 x_1 + \dots + \beta_p x_p$ . I believe cases will change drastically based on country because of the different population sizes and each one having a different government. So I wish to group by country, and so make country a grouping variable in my model. So because 'date' is just a replicate of 'biweek':  $x_1$  is variant,  $x_2$  is biweek.

```
variants_5Largest$variant
```

So for grouping by country:  $j = \{1, 2, 3, 4, 5\}$ . So 1 is France, 2 is Germany, 3 is Italy, 4 is Spain, 5 is United Kingdom.

```
sum(variants_5Largest$country == "France")
sum(variants_5Largest$country == "Germany")
sum(variants_5Largest$country == "Italy")
sum(variants_5Largest$country == "Spain")
sum(variants_5Largest$country == "UnitedKingdom")
```

So  $n_1 = 120, n_2 = 120, n_3 = 120, n_4 = 120, n_5 = 120$ . So  $y_{ij} \mid \beta_j, x_{ij} \sim \text{NegBin}(\lambda(\beta, x), \phi)$ . So  $\log(\lambda(\beta, x)) = \beta_{0j} + \beta_{1j}x_{\text{variant},j} + \beta_{2j}x_{\text{biweek},j}$ . So model for  $\beta$ 's are:  $\beta_j \sim N(B, \Sigma)$ . So prior  $\pi(B, \Sigma)$ . So  $\log(\lambda(\beta, x_{ij})) = b_0 + b_1x_{\text{variant},ij} + b_2x_{\text{biweek},ij} + \beta_{0j} + \beta_{1j}x_{\text{variant},j} + \beta_{2j}x_{\text{biweek},j}$ . So  $\beta_j \sim N(0, \Sigma)$ . And need prior for  $\pi(b, \Sigma)$ . So  $E[y] = \lambda = \exp(b_0 + b_1x_{\text{variant}} + b_2x_{\text{biweek}} + \beta_0 + \beta_1x_{\text{variant}} + \beta_2x_{\text{biweek}}) = e^{b_0}e^{b_1x_{\text{variant}}}e^{b_2x_{\text{biweek}}}e^{\beta_0}e^{\beta_1x_{\text{variant}}}e^{\beta_2x_{\text{biweek}}}$ . And  $\Sigma = \text{vec}(\underline{\sigma}) \times \Omega_k \times \text{vec}(\underline{\sigma})$ , where  $\underline{\sigma} = \sigma_0, \sigma_{\text{variant}}, \sigma_{\text{biweek}}$ . The cases in the data range from a lower bound of a count of 1, all the way up to 77000, with the vast majority of the counts being very small. So  $e^{b_0} \in [1, 77000]$ .

```
log(c(1, 77000))
```

```
## [1] 0.00000 11.25156
```

```
log(77000)/2
```

```
## [1] 5.62578
```

So set weakly informative prior on the intercept as  $N(0, 4)$ , and on the grouping as  $N(0, 10)$ .

```
Intercept_Prior <- set_prior("normal(0,4)", class="Intercept")
sd_priors_Intercept <- set_prior("normal(0, 10)", class="sd", group="country", coef="Intercept")
```

In the data, biweek ranges from 1 to 20.

```
exp(0.15*c(1, 20))
```

```
## [1] 1.161834 20.085537
```

So make b prior for biweek  $N(0, 0.01)$  as well as for the grouping. And make the b priors as well as the groupings, for variant very weakly informative as  $N(0, 100)$ .

```
biweek_priors <- set_prior("normal(0, 0.01)", class="b", coef = "biweek")
sd_priors_biweek <- set_prior("normal(0, 0.01)", class = "sd", group = "country", coef = "biweek")
variant_priors <- c(set_prior("normal(0, 100)", class="b", coef="variantBeta"),
  set_prior("normal(0, 100)", class="b", coef="variantDelta"),
  set_prior("normal(0, 100)", class="b", coef="variantGamma"),
  set_prior("normal(0, 100)", class="b", coef="variantnon_who"),
  set_prior("normal(0, 100)", class="b", coef="variantOther"))
sd_priors_variant <- c(set_prior("normal(0, 100)", class="sd", group="country", coef="variantBeta"),
  set_prior("normal(0, 100)", class="sd", group="country", coef="variantDelta"),
  set_prior("normal(0, 100)", class="sd", group="country", coef="variantGamma"),
  set_prior("normal(0, 100)", class="sd", group="country", coef="variantnon_who"),
  set_prior("normal(0, 100)", class="sd", group="country", coef="variantOther"))
```

So based on the variance formula:

```
lambdas <- c(1, 77000)
vargs <- function(phi){
  lambdas *(1+lambdas/phi) }
sqrt(vargs(800))
```

```
## [1] 1.000625 2736.466700
```

```
2700/77000
```

```
## [1] 0.03506494
```

So, to make reasonable prediction error bars, for about 35% error bars, make the shape prior  $N(400, 200)$ .

```
shape_prior <- set_prior("normal(400, 200)", class="shape")
```

```
model_1<-brm(cases ~ biweek+ variant+ (biweek+ variant | country),
data = variants_5Largest, family = negbinomial(), prior=c(Intercept_Prior, sd_priors_Intercept,
biweek_priors, sd_priors_biweek, variant_priors, sd_priors_variant, shape_prior),
iter=6000, warmup = 5000, inits="0") # is my first model
```

```
summary_model_1 <- summary(model_1)
rbind(summary_model_1$random$country[-2:-4], summary_model_1$fixed[-2:-4],
summary_model_1$spec_pars[-2:-4])
```

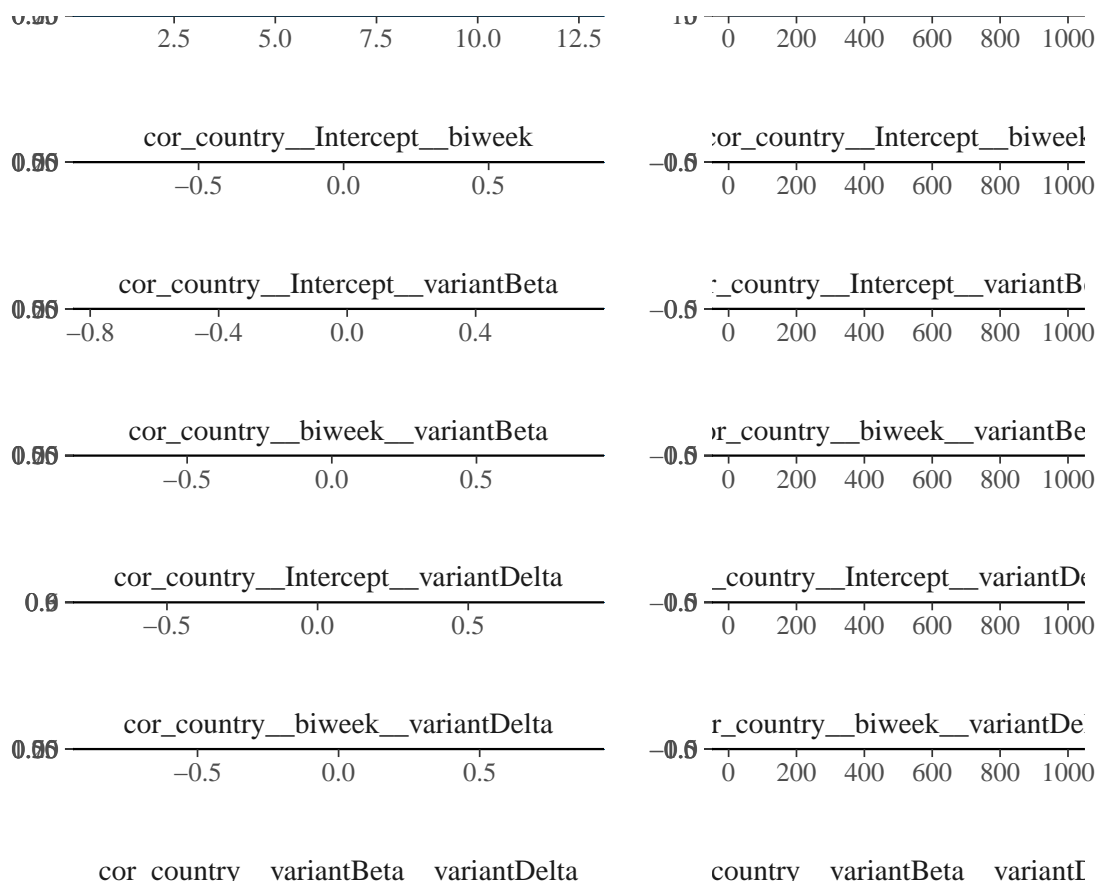
##	Estimate	Rhat	Bulk_ESS	Tail_ESS
## sd(Intercept)	1.178018820	1.0184545	381.7273	146.1426
## sd(biweek)	0.009359852	1.0016202	1638.5602	2034.2631
## sd(variantBeta)	2.670555351	1.0005355	1118.5845	1460.6956
## sd(variantDelta)	1.030040643	1.0021051	1506.8777	1483.3983
## sd(variantGamma)	2.244572870	1.0051071	897.9396	441.3606
## sd(variantnon_who)	0.530122838	1.0049380	865.6571	1788.5258
## sd(variantOther)	0.982551827	1.0131778	338.0679	134.5648
## cor(Intercept,biweek)	-0.102476413	1.0027181	3737.3875	2415.4632
## cor(Intercept,variantBeta)	-0.035146039	1.0056547	1032.7487	225.8081
## cor(biweek,variantBeta)	0.064092421	1.0036660	1532.4334	2428.5260
## cor(Intercept,variantDelta)	0.144906464	1.0003813	2739.2121	2452.2025
## cor(biweek,variantDelta)	-0.090747299	1.0053205	1689.6774	2534.9643
## cor(variantBeta,variantDelta)	-0.039903989	1.0047781	1399.8894	1384.1790
## cor(Intercept,variantGamma)	-0.333899045	1.0014279	1478.5543	566.2504
## cor(biweek,variantGamma)	0.061799911	1.0004939	1879.2013	2392.8335
## cor(variantBeta,variantGamma)	0.011162502	1.0000485	2995.5186	2662.9233
## cor(variantDelta,variantGamma)	-0.186605758	1.0011477	2505.2352	3164.8757
## cor(Intercept,variantnon_who)	-0.092212605	1.0037338	1653.0279	841.0911
## cor(biweek,variantnon_who)	0.057502097	0.9998067	3221.5491	2948.9471
## cor(variantBeta,variantnon_who)	0.054577559	1.0062807	962.5171	378.6696
## cor(variantDelta,variantnon_who)	-0.058708971	1.0053763	760.5779	329.9412
## cor(variantGamma,variantnon_who)	0.090742221	1.0005533	2666.5365	3054.7605
## cor(Intercept,variantOther)	-0.214808410	1.0049661	2146.9193	2496.8637
## cor(biweek,variantOther)	0.080492145	1.0007705	2755.3530	2951.8334

Part of the  
Jellie sand  
rearing  
here.

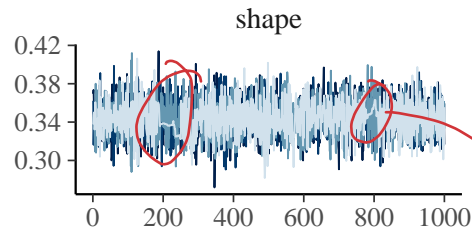
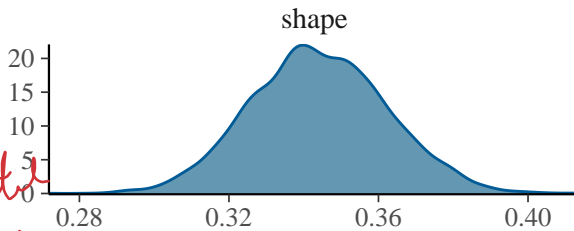
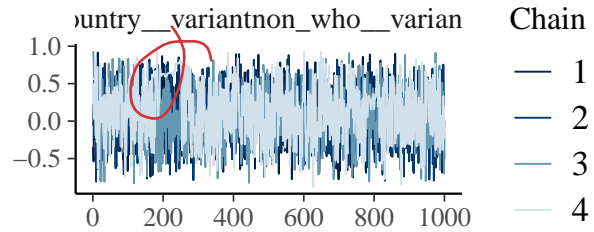
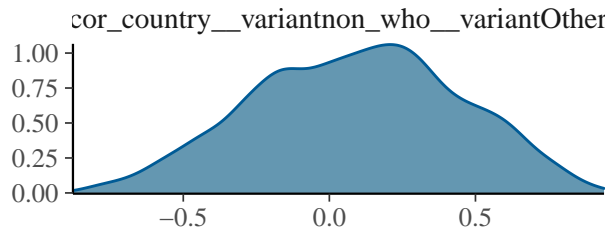
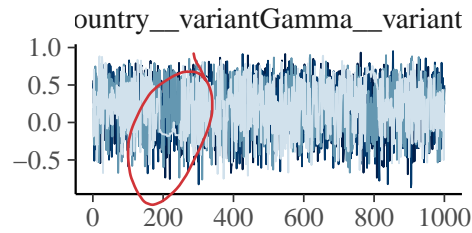
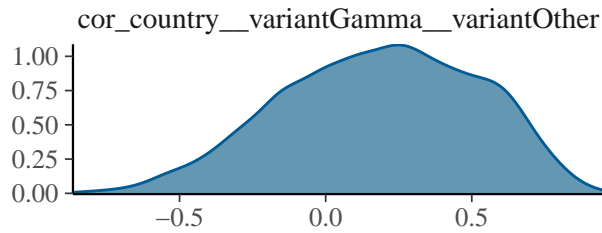
```
## cor(variantBeta,variantOther)    0.116344047  1.0021083 3130.2481 2717.8934
## cor(variantDelta,variantOther) -0.080059835  1.0074768  536.7563  186.4523
## cor(variantGamma,variantOther)  0.183455828  1.0002016 1972.8309 2944.3245
## cor(variantnon_who,variantOther) 0.090909612  1.0049336 1132.0486 1376.4759
## Intercept                        7.977300958  1.0020156 1009.0879 1398.0636
## biweek                          -0.003250502  1.0038026 4048.0442 2745.6956
## variantBeta                     -3.960871695  1.0026749 1275.3823 1652.1023
## variantDelta                     0.039212875  1.0041501 1225.3079 1520.4159
## variantGamma                     -4.113235710  1.0068501  688.6339  418.1550
## variantnon_who                   -1.593045561  1.0042711 1409.3378  864.7554
## variantOther                     -4.230312564  1.0114443  403.2153  133.7641
## shape                           0.344008562  1.0018202 4528.0343 2806.9581
```

All the Rhat's are 1, and all the effective sample size's for the Bulk and the Tail are very big. Which is an extremely good sign.

```
plot(model_1, N=33, ask=FALSE)
```



*Perhaps it states the obvious that these traceplots are difficult to interpret as displayed!*

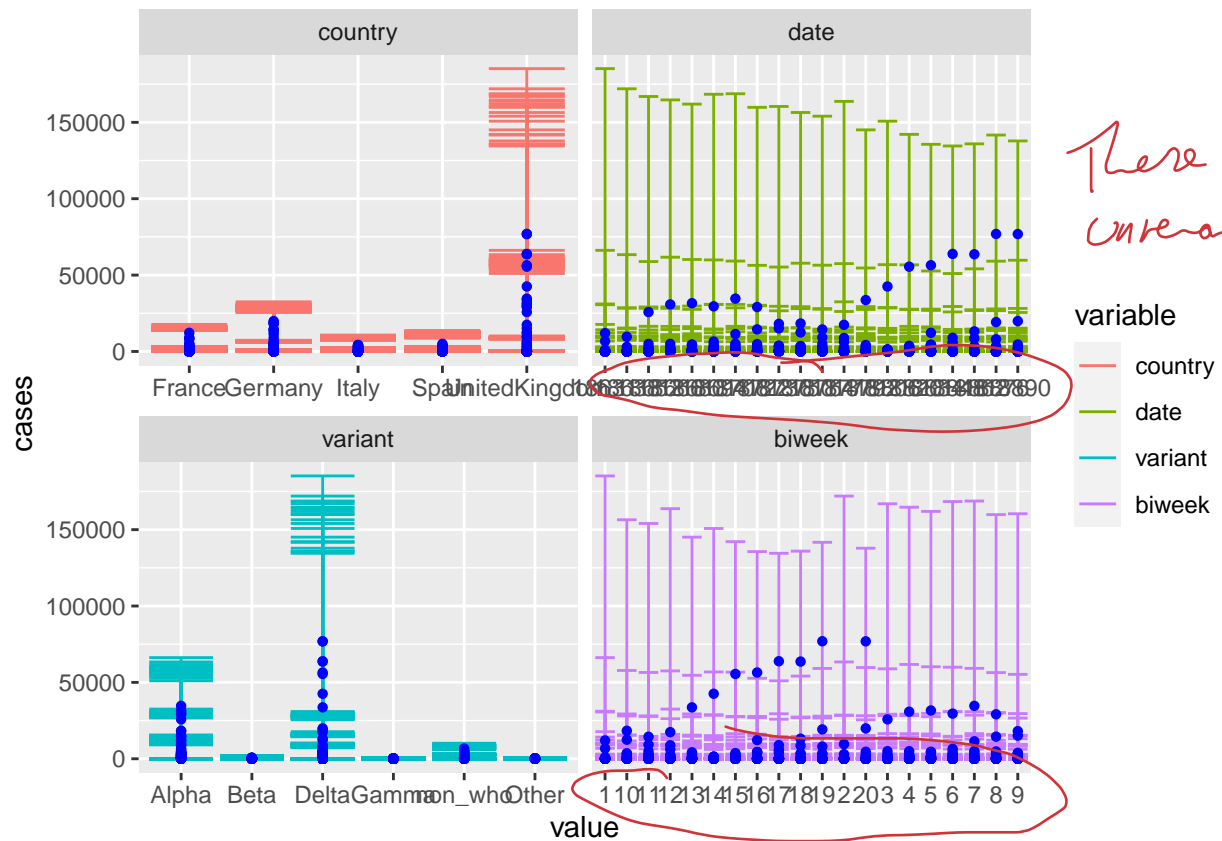


None of the trace plots got stuck anywhere and they are all very spiky, so they have all converged. This, together with the analysis of the summary, means the model has converged very well. The intercept has moved from the prior by an amount that is fairly believable. The shape mean and mode is about 0.34, which may lead to big error bars. Both biweek and variant have moved away from my prior, by a fairly believable amount, which is a good sign. The model is reasonably small which is good.

```
case_preds <- predict(model_1, newdata = variants_5Largest)
preds <- cbind(variants_5Largest, as.tibble(case_preds)) %>%
  dplyr::select(-Est.Error) %>%
  melt(id.vars = c("Estimate", "Q2.5", "Q97.5", "cases")) %>%
  ggplot()+ geom_errorbar(aes(x=value, ymin=Q2.5, ymax=Q97.5, colour = variable))+
  geom_point(aes(x=value, y=cases), colour="blue", pch=16)+
  facet_wrap(~variable, scale="free_x")
preds
```

what does "fairly believable" mean

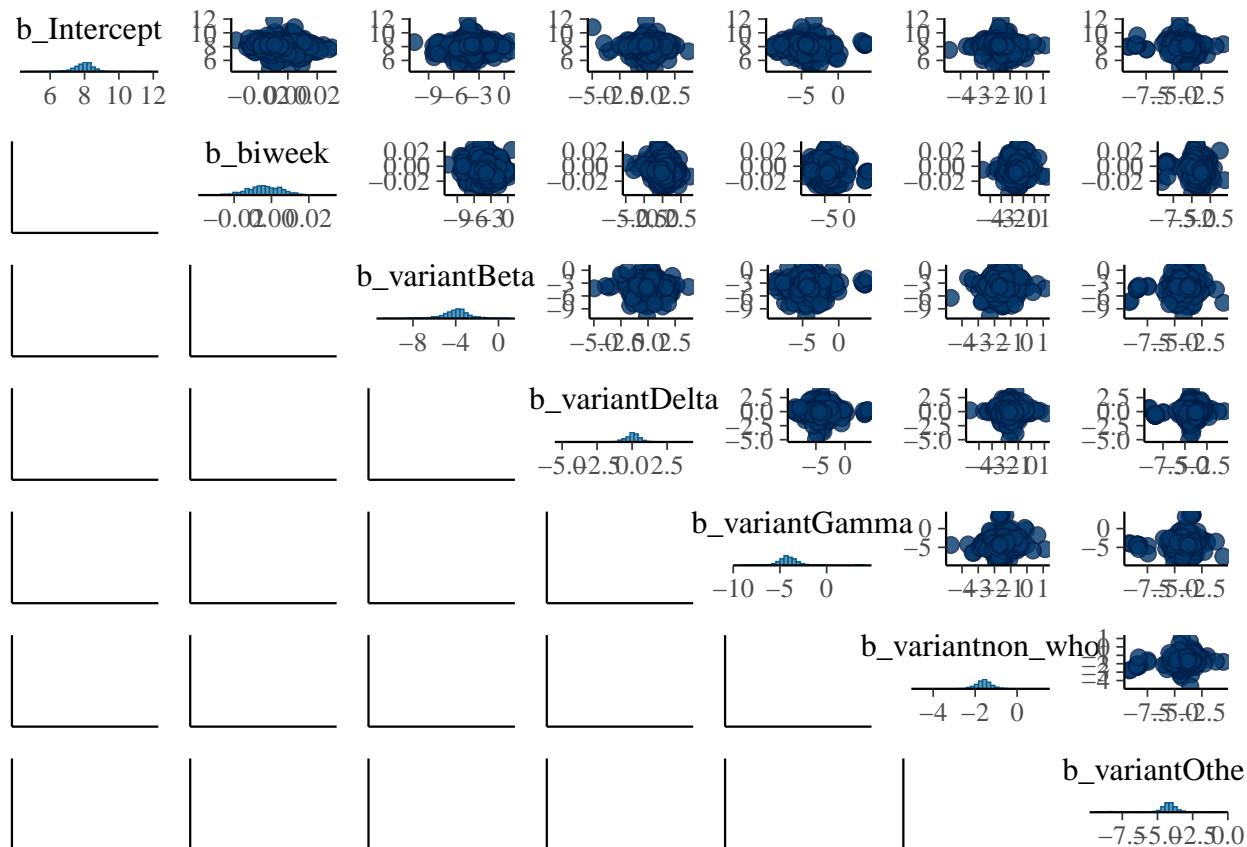
Not true



The error bars seem reasonably skillful and none of the data points are outside the error bars. Which is good considering how wide ranging the data is.

*You can't assess this from these plots.*

```
samples <- posterior_samples(model_1, subset=seq(from=1,to=4000,by=4))
mcmc_pairs(samples[,1:7])
```



looking at the mcmc plots for biweek, it is easy to see, there are no correlations between biweek and variant.

*Not easy given size of plots.*

```
sum(samples$b_biweek<0)/length(samples$b_biweek)
```

```
## [1] 0.65
```

Is the  $P(\text{biweek} < 0)$  which shows signs, that the mean of biweek is not zero, which means biweek has moved away from my prior. ✓

```
ranef(model_1)
```

```
## $country
## , , Intercept
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      -0.2540797 0.6264088 -1.43964485 1.1074297
## Germany       0.4995146 0.6277490 -0.62585988 1.8473218
## Italy        -0.6177521 0.6340276 -1.80778609 0.6697288
## Spain       -0.4510390 0.6393902 -1.76744573 0.8316282
## UnitedKingdom 1.2165054 0.6758361 0.03026119 2.6522811
##
## , , biweek
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      0.001773964 0.009568604 -0.01638736 0.02555707
```



```

## Germany      -0.001094240 0.009046730 -0.02248043 0.01633117
## Italy        -0.000387160 0.011053218 -0.02409864 0.02283731
## Spain        0.004574718 0.010458604 -0.01177012 0.03084174
## UnitedKingdom -0.007861679 0.015035949 -0.04769505 0.01194004
##
## , , variantBeta
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      1.9603371  1.311773 -0.5109685 4.821087
## Germany     0.3147633  1.320850 -2.2914071 3.166180
## Italy      -1.4778838  1.310647 -4.1576524 1.271872
## Spain       0.8417316  1.323846 -1.8105061 3.646232
## UnitedKingdom -1.1954779 1.328160 -3.9132751 1.464348
##
## , , variantDelta
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      0.02748791 0.6513216 -1.2796636 1.3547947
## Germany    -0.12138325 0.6564354 -1.4924087 1.1614889
## Italy      -0.28113631 0.6526608 -1.6086322 0.9761512
## Spain     -0.33553560 0.6741141 -1.8151605 0.8659412
## UnitedKingdom 0.91160712 0.7637863 -0.2308664 2.6114289
##
## , , variantGamma
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      0.4519899  1.339945 -1.9646651 2.845881590
## Germany    -0.4758315  1.334619 -3.0039714 1.904619459
## Italy       1.5416428  1.343741 -0.8702836 4.134217127
## Spain       0.6833346  1.362420 -1.7469203 3.123152556
## UnitedKingdom -2.3780860 1.367580 -4.9228634 -0.006374396
##
## , , variantnon_who
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      0.02309038 0.4421880 -0.9131979 0.9667116
## Germany     0.10012445 0.4272180 -0.6228816 1.0779445
## Italy       0.08032578 0.4547156 -0.7534626 1.1684189
## Spain       0.14640297 0.4607050 -0.6293169 1.2434922
## UnitedKingdom -0.31029755 0.5026857 -1.4749951 0.4413056
##
## , , variantOther
##
##           Estimate Est.Error      Q2.5      Q97.5
## France      0.2922804 0.8211968 -0.7992016 2.0648593
## Germany    -0.1006695 0.7522104 -1.3606147 1.5425474
## Italy       0.2246743 0.8088114 -0.9235219 2.0173403
## Spain       0.5587096 0.8764608 -0.4174780 2.3477649
## UnitedKingdom -0.5891482 0.8029012 -2.0342347 0.9112287

```

The random effects also back up this hypothesis that biweek has no correlation to variant. And the effect of biweek, and the effect of variant, has moved away from my prior's, of them having no effect, to them having an effect on my model. So on this first model, I will do a sensitivity step and analysis. The sensitivity step will be to remove the variable: 'variant' from the  $\beta$  variable's of the model. So the second model will be the

*This model doesn't remove variant as you find above?*

sensitivity model. So  $\log(\lambda(\beta, x_{ij})) = b_0 + b_1 x_{variant,ij} + b_2 x_{biweek,ij} + \beta_{0j} + \beta_{1j} x_{variant,j}$  is my sensitivity model.

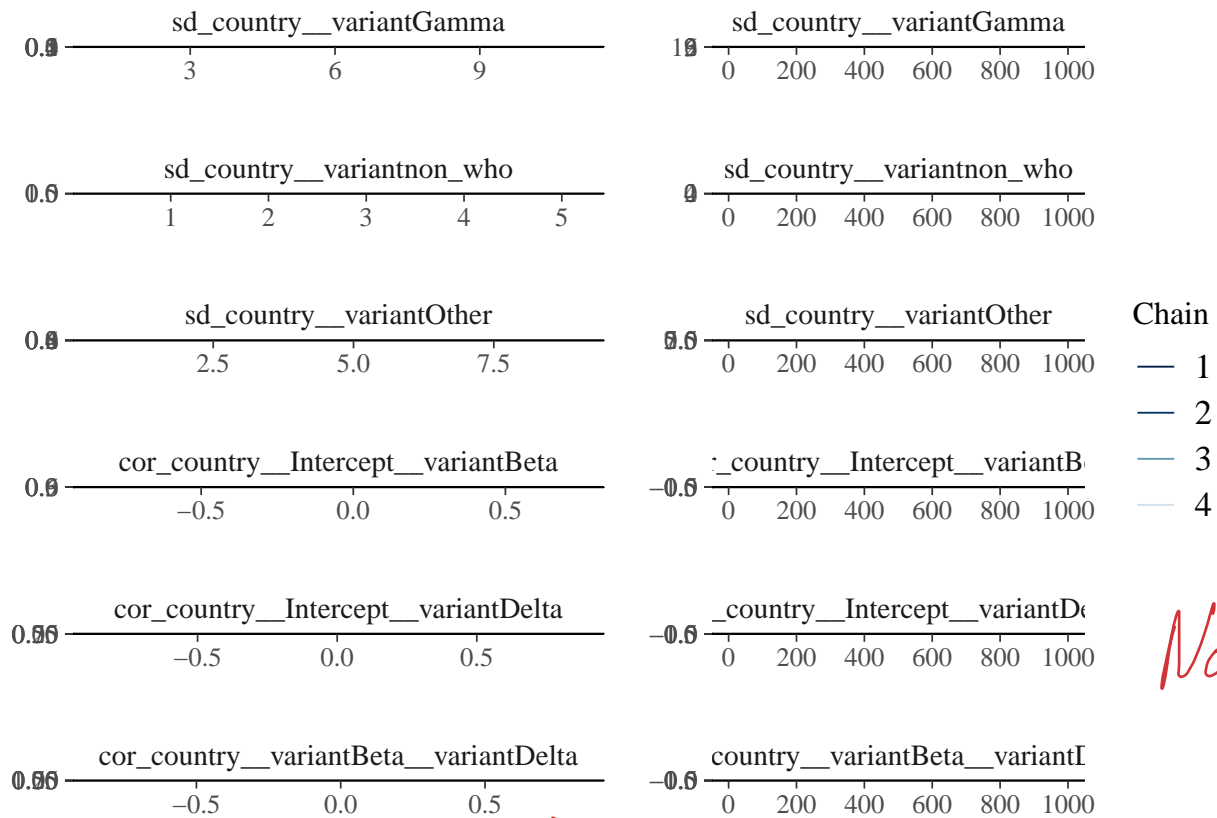
```
model_2<-brm(cases ~ biweek+ variant+ (variant | country), data=variants_5Largest,
family = negbinomial(), prior=c(Intercept_Prior, sd_priors_Intercept,
biweek_priors, variant_priors, sd_priors_variant, shape_prior), iter=6000,
warmup = 5000, inits="0") # is my sensitivity model
```

```
summary_model_2 <- summary(model_2)
rbind(summary_model_2$random$country[-2:-4], summary_model_2$fixed[-2:-4],
summary_model_2$spec_pars[-2:-4])
```

##	Estimate	Rhat	Bulk_ESS	Tail_ESS
## sd(Intercept)	1.129644686	1.0027252	859.5246	1004.9130
## sd(variantBeta)	2.912246574	1.0032548	1391.8405	1527.5365
## sd(variantDelta)	0.950384187	1.0009566	1135.6366	732.3082
## sd(variantGamma)	2.314306503	1.0049157	733.8799	464.9261
## sd(variantnon_who)	0.513019589	1.0031862	1403.3087	1811.0667
## sd(variantOther)	0.890686967	1.0005169	1134.2783	1536.3338
## cor(Intercept,variantBeta)	-0.046126442	1.0010381	1809.8329	2409.0861
## cor(Intercept,variantDelta)	0.148110032	1.0026593	2350.2085	1362.4162
## cor(variantBeta,variantDelta)	-0.036331784	1.0004665	3318.8057	2739.4096
## cor(Intercept,variantGamma)	-0.370825317	1.0010041	1742.9217	1430.2355
## cor(variantBeta,variantGamma)	0.026792097	1.0032211	3032.9879	2879.0052
## cor(variantDelta,variantGamma)	-0.214025560	1.0020114	1856.2066	2165.3566
## cor(Intercept,variantnon_who)	-0.110063870	1.0018321	3542.5879	2565.2510
## cor(variantBeta,variantnon_who)	0.052547120	1.0008143	3420.3867	2729.7107
## cor(variantDelta,variantnon_who)	-0.062650310	1.0029544	3000.2901	2798.2735
## cor(variantGamma,variantnon_who)	0.102767249	1.0008724	2221.0581	2175.1269
## cor(Intercept,variantOther)	-0.255302086	1.0006715	2790.0191	2860.3425
## cor(variantBeta,variantOther)	0.142137903	1.0009563	3490.6545	1976.8580
## cor(variantDelta,variantOther)	-0.093027629	1.0017800	2940.8790	3354.8698
## cor(variantGamma,variantOther)	0.210310627	0.9997915	2005.5474	1540.7964
## cor(variantnon_who,variantOther)	0.089731054	1.0004221	1960.2584	1703.2765
## Intercept	7.949438149	1.0063149	1011.9016	950.5655
## biweek	-0.003429046	1.0019325	4483.9699	2876.9214
## variantBeta	-3.901050328	1.0090140	863.9308	1177.2538
## variantDelta	0.070179353	1.0041852	1203.1023	1685.5330
## variantGamma	-4.142103160	1.0085216	824.1618	523.4327
## variantnon_who	-1.576407966	1.0006785	1825.1836	1543.3800
## variantOther	-4.150766218	1.0012704	1528.1004	1204.9333
## shape	0.343486989	1.0010569	3512.1798	903.2146

All the Rhat's are 1, and all the effective sample size's for the Bulk and the Tail are very big. Which is an extremely good sign.

```
plot(model_2, N=26, ask=FALSE)
```



No plots!

Makes (having) end of Q1).

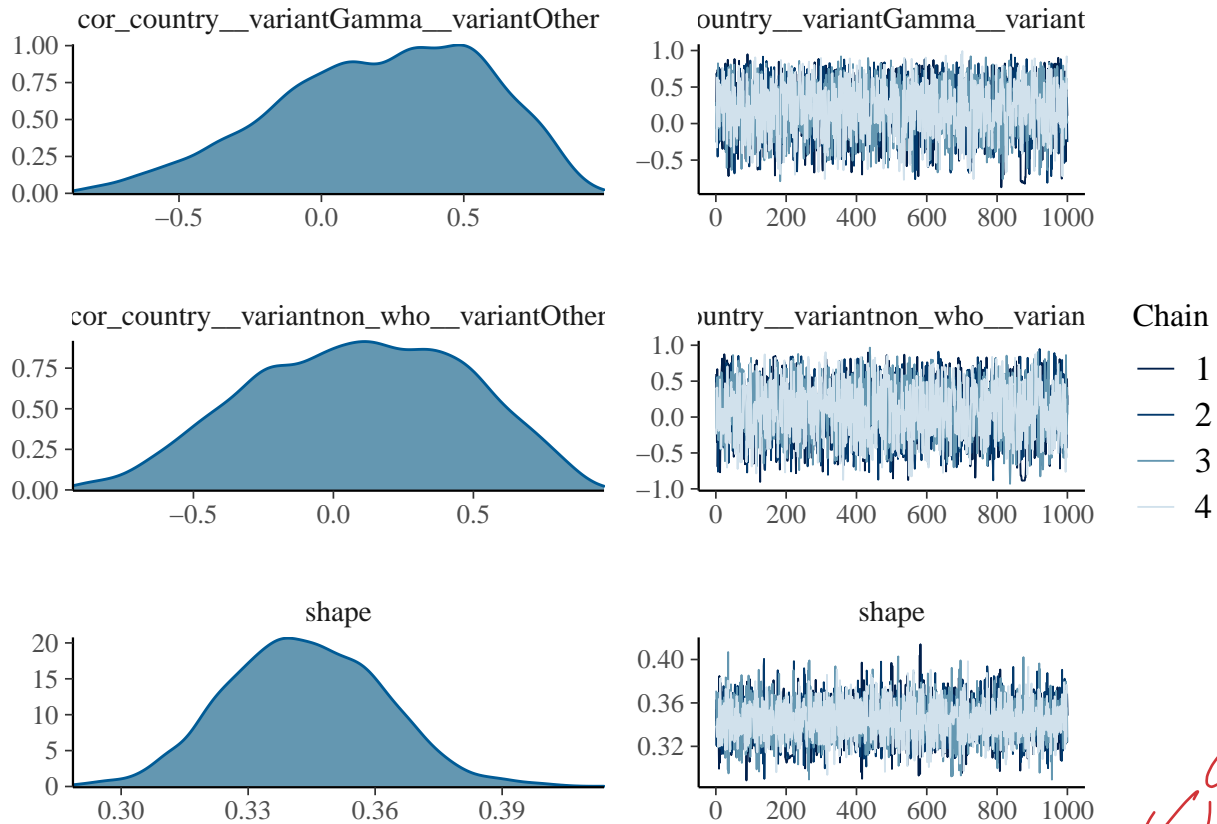
Modelling: I don't follow the logic of having a static variant effect but the same temporal pattern for each variant. Mathematically mostly correct up to vague m's. NB well used. Convert call to biom. Sensitivity features into model choice. 11/14.

Prior: Subjective (even a couple of nearly informative) priors given. Didn't get logic of prior on  $\sigma^2$  or of  $N(0,10)$  for group intercepts given  $N(0,4)$  on main intercept. Investigated into distributions not included. 11/14

Convergence: Mixing, trace, Rhats, Neff all covered. 2 problems (erroneous claim that no chains got stuck appeared below a chain that got stuck, and > 50% of traceplots were included AND invisible! 9/14.

Validation: In single validation plots uninterpretable (chain data within error bars not immediately verifiable from plots. Sensitivity to grouping on binch looked at but

Critique: *Can't assess based on plots. Load at prior moving was good. Condition in chain load ok. Quite loose with long range ("fairly believable" "pretty insensitive".* 5/14

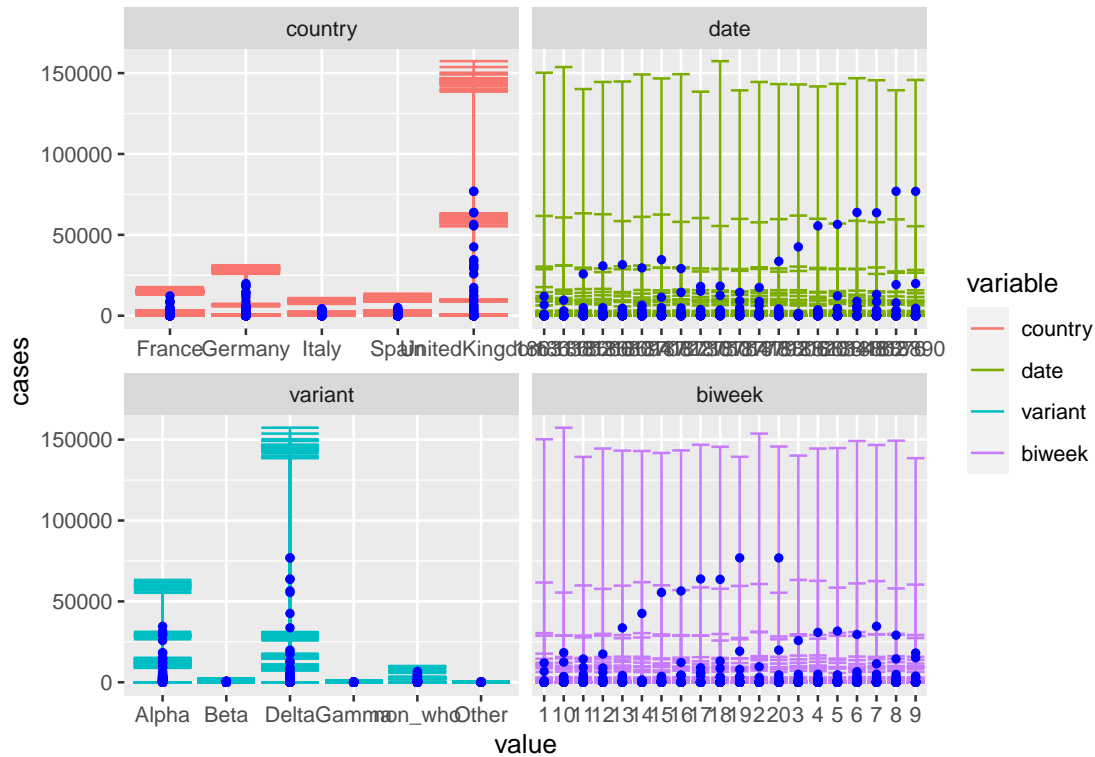


None of the trace plots got stuck anywhere and they are all very spiky, so they have all converged. This, together with the analysis of the summary, means the model has converged very well. The intercept and variant variables have not changed in this sensitivity step. ✓

```
case_preds <- predict(model_2, newdata = variants_5Largest)
preds <- cbind(variants_5Largest, as.tibble(case_preds)) %>%
  dplyr::select(-Est.Error) %>%
  melt(id.vars = c("Estimate", "Q2.5", "Q97.5", "cases")) %>%
  ggplot()+ geom_errorbar(aes(x=value, ymin=Q2.5, ymax=Q97.5, colour = variable))+
  geom_point(aes(x=value, y=cases), colour="blue", pch=16)+
  facet_wrap(~variable, scale="free_x")
preds
```

Presentation: *Blank plots and illegible verification plots are main problems. Code is fine, mathematics should be better spaced (like individual lines for key equations.)* 4/7.

Q1: 44/70



Again how  
can this  
be interpreted?

The predictions and error bars have not changed by much at all, in this sensitivity step. So my conclusions seem pretty insensitive to this main choice I have made. So my final model is my first model.

2a)

```
pred_UnitedKingdom_Delta_90 <- predict(model_1, newdata = filter(variants_5Largest,
  variant == 'Delta', country == "UnitedKingdom"),
  allow_new_levels=TRUE, prob = c(0.05, 0.95))
max(pred_UnitedKingdom_Delta_90[, "Estimate"])
```

What about biweekly 21:26.

```
## [1] 28768.01
```

Is the estimate for the peak number of sequenced Delta cases in the UK in 2021.

```
max_value <- which.max(pred_UnitedKingdom_Delta_90[, "Estimate"])
c(pred_UnitedKingdom_Delta_90[max_value, "Q5"], pred_UnitedKingdom_Delta_90[max_value, "Q95"])
```

```
##      Q5      Q95
## 10.0 129293.2
```

Missing next 4 2021

Is the 90% prediction interval for the peak number of sequenced Delta cases in the UK in 2021.

2b)

```

UnitedKingdom_Delta_Pred_90 <- variants_5Largest %>%
  filter(variant== 'Delta', country == "UnitedKingdom") %>%
  mutate(Cases = pred_UnitedKingdom_Delta_90[, "Estimate"], Q5 = pred_UnitedKingdom_Delta_90[, "Q5"],
         Q95=pred_UnitedKingdom_Delta_90[, "Q95"])
UnitedKingdom_Delta_Pred_90[max_value, "date"]

```

```

## # A tibble: 1 x 1
##   date
##   <date>
## 1 2021-01-04

```

3 Delta didn't exist in January.

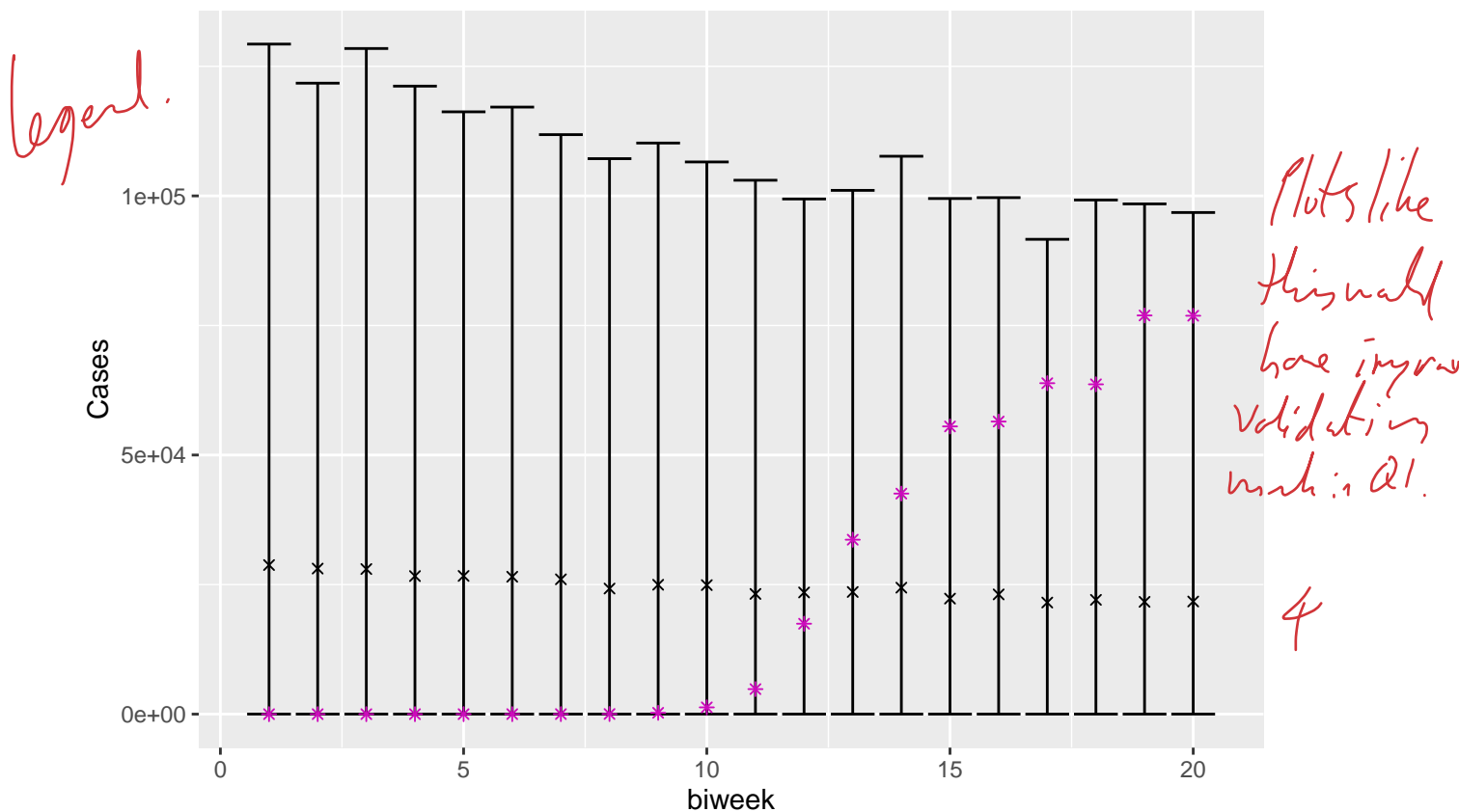
Is the expected date of the peak.

2c)

```

ggplot(UnitedKingdom_Delta_Pred_90) +
  geom_errorbar(aes(x=biweek, ymin=Q5, ymax=Q95)) +
  geom_point(aes(x=biweek, y=Cases), col=1, pch=4)+
  geom_point(aes(x=biweek, y=cases), col=6, pch=8)

```



The graph shows the model has an extremely low efficacy, shown by how by the error bars are at the beginning of the year. The model is predicting the highest number of cases at the very start of the year, for the delta variant in the UK. And predicts only a decreases in the number of its cases, as the date increases.

12/15

Where as the data clearly shows a trend, that at the date increases, the number of cases for the delta variant in the UK increases. With the highest number of cases recorded, being for the most recent date. So the model predicts the complete opposite of what the true data shows, showing an extremely low efficacy. So there was no need to try to go beyond biweek 20, since the predicted estimate's from the model, only decreases as biweek increases.

*You either need to show that mathematically or demonstrate it by going beyond 20.*

3a)

```
Ess_Model_1 <- rbind(summary_model_1$random$country[-1:-5], summary_model_1$fixed[-1:-5],
                     summary_model_1$spec_pars[-1:-5])
Ess_Bulk_Model_1 <- Ess_Model_1[-2]
Bulk_Ess <- min(Ess_Bulk_Model_1)
pred_samples_France_Beta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Beta', country == "France"), allow_new_levels=TRUE, summary = FALSE)
pred_samples_France_non_who <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'non_who', country == "France"), allow_new_levels=TRUE, summary = FALSE)
n_France_Beta <- ncol(pred_samples_France_Beta)
prob_France_Beta_more_non_who <- numeric(n_France_Beta)
sd_France_Beta_more_non_who <- numeric(n_France_Beta)
for (i in 1:n_France_Beta) {
  prob_France_Beta_more_non_who[i]<-mean(pred_samples_France_Beta[,i]>pred_samples_France_non_who[,i])
  sd_France_Beta_more_non_who[i] <- sd(pred_samples_France_Beta[,i]>pred_samples_France_non_who[,i])
}
MC_error_France_Beta_more_non_who <- sd_France_Beta_more_non_who/Bulk_Ess
MC_error_France_Beta_more_non_who
```

```
## [1] 0.001470364 0.001467010 0.001472628 0.001472417 0.001467200 0.001466625
## [7] 0.001468844 0.001467389 0.001468219 0.001467106 0.001469871 0.001469788
## [13] 0.001470445 0.001473242 0.001469364 0.001470119 0.001472130 0.001467946
## [19] 0.001465838 0.001469871
```

```
MC_error_France_Beta_more_non_who<0.01
```

```
## [1] TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE
## [16] TRUE TRUE TRUE TRUE TRUE
```

Estimates of dominance probabilities all have Monte Carlo error less than 0.01.

```
prob_France_Beta_more_non_who
```

```
## [1] 0.44550 0.43600 0.45300 0.45225 0.43650 0.43500 0.44100 0.43700 0.43925
## [10] 0.43625 0.44400 0.44375 0.44575 0.45525 0.44250 0.44475 0.45125 0.43850
## [19] 0.43300 0.44400
```

```
prob_France_Beta_more_non_who>0.5
```

```
## [1] FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE
## [13] FALSE FALSE FALSE FALSE FALSE FALSE FALSE FALSE
```

```
prob_France_Beta_more_non_who_True_or_False <- as.numeric(prob_France_Beta_more_non_who>0.5)
as.numeric(Position(function(x) x>0, prob_France_Beta_more_non_who_True_or_False))
```

```
## [1] NA
```

Is the first biweek, my model predicts; the Beta variant first dominates the original Covid-19 strain, in France, if the output is 'NA', then the model predicts that Beta variant never dominates the original Covid-19 strain. So my model never predicts that Beta will dominate the original strain in France.

3b)

```
pred_samples_Italy_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "Italy"), allow_new_levels=TRUE, summary = FALSE)
pred_samples_Italy_Gamma <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Gamma', country == "Italy"), allow_new_levels=TRUE, summary = FALSE)
n_Italy_Delta <- ncol(pred_samples_Italy_Delta)
prob_Italy_Delta_more_Gamma <- numeric(n_Italy_Delta)
sd_Italy_Delta_more_Gamma <- numeric(n_Italy_Delta)
for (i in 1:n_Italy_Delta) {
  prob_Italy_Delta_more_Gamma[i]<-mean(pred_samples_Italy_Delta[,i]>pred_samples_Italy_Gamma[,i])
  sd_Italy_Delta_more_Gamma[i] <- sd(pred_samples_Italy_Delta[,i]>pred_samples_Italy_Gamma[,i])
}
MC_error_Italy_Delta_more_Gamma <- sd_Italy_Delta_more_Gamma/Bulk_Ess
MC_error_Italy_Delta_more_Gamma
```

```
## [1] 0.001255263 0.001286089 0.001302054 0.001311853 0.001301655 0.001296826
## [7] 0.001302453 0.001300856 0.001293972 0.001302453 0.001295606 0.001293562
## [13] 0.001321320 0.001318326 0.001309532 0.001291917 0.001295606 0.001297636
## [19] 0.001304041 0.001300054
```

```
MC_error_Italy_Delta_more_Gamma<0.01
```

```
## [1] TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE
## [16] TRUE TRUE TRUE TRUE TRUE
```

Estimates of dominance probabilities all have Monte Carlo error less than 0.01.

```
prob_Italy_Delta_more_Gamma
```

```
## [1] 0.76450 0.74700 0.73725 0.73100 0.73750 0.74050 0.73700 0.73800 0.74225
## [10] 0.73700 0.74125 0.74250 0.72475 0.72675 0.73250 0.74350 0.74125 0.74000
## [19] 0.73600 0.73850
```

```
prob_Italy_Delta_more_Gamma>0.5
```

```
## [1] TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE TRUE
## [16] TRUE TRUE TRUE TRUE TRUE
```



```
prob_IItaly_Delta_more_Gamma_True_or_False <- as.numeric(prob_IItaly_Delta_more_Gamma>0.5)
as.numeric(Position(function(x) x>0, prob_IItaly_Delta_more_Gamma_True_or_False))
```

```
## [1] 1
```

Is the first biweek, my model predicts; the Delta variant dominates the Gamma variant, in Italy, if the output is 'NA', then the model predicts that Delta variant never dominates the Gamma variant. So my model predicts that Gamma variant will never dominate the Delta variant in Italy.

*4*  
*not NA so date?*

3c)

```
pred_samples_France_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "France"), allow_new_levels=TRUE, summary = FALSE)
variant_names <- c("Alpha", "Beta", "Gamma", "non_who", "Other")
variant_NotDelta_length <- length(variant_names)
France_Delta_dominates <- numeric(variant_NotDelta_length)
for (j in 1:variant_NotDelta_length) {
  pred_samples_France_NotDelta <- predict(model_1, newdata = filter(variants_5Largest,
    variant== variant_names[j], country == "France"), allow_new_levels=TRUE, summary = FALSE)
  n_France_Delta <- ncol(pred_samples_France_Delta)
  prob_France_Delta_MoreThan <- numeric(n_France_Delta)
  for (i in 1:n_France_Delta) {
    prob_France_Delta_MoreThan[i] <-
      mean(pred_samples_France_Delta[,i]>pred_samples_France_NotDelta[,i])
  }
  prob_France_Delta_MoreThan_True_or_False <- as.numeric(prob_France_Delta_MoreThan>0.5)
  France_Delta_dominates[j]<-
    as.numeric(Position(function(x) x>0, prob_France_Delta_MoreThan_True_or_False))
}
max(France_Delta_dominates)
```

```
## [1] 2
```

Is the first biweek, my model predicts; the Delta variant first dominates all the other Covid-19 variants, in France, if the output is 'NA', then the model predicts that Delta variant never dominates all the other Covid-19 variants, in France.

```
pred_samples_Germany_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "Germany"), allow_new_levels=TRUE, summary = FALSE)
variant_names <- c("Alpha", "Beta", "Gamma", "non_who", "Other")
variant_NotDelta_length <- length(variant_names)
Germany_Delta_dominates <- numeric(variant_NotDelta_length)
for (j in 1:variant_NotDelta_length) {
  pred_samples_Germany_NotDelta <- predict(model_1, newdata = filter(variants_5Largest,
    variant== variant_names[j], country == "Germany"), allow_new_levels=TRUE, summary = FALSE)
  n_Germany_Delta <- ncol(pred_samples_Germany_Delta)
  prob_Germany_Delta_MoreThan <- numeric(n_Germany_Delta)
  for (i in 1:n_Germany_Delta) {
    prob_Germany_Delta_MoreThan[i] <-
      mean(pred_samples_Germany_Delta[,i]>pred_samples_Germany_NotDelta[,i])
  }
}
```

```

}
prob_Germany_Delta_MoreThan_True_or_False <- as.numeric(prob_Germany_Delta_MoreThan>0.5)
Germany_Delta_dominates[j] <-
  as.numeric(Position(function(x) x>0, prob_Germany_Delta_MoreThan_True_or_False))
}
max(Germany_Delta_dominates)

```

```
## [1] 11
```

Is the first biweek, my model predicts; the Delta variant first dominates all the other Covid-19 variants, in Germany, if the output is 'NA', then the model predicts that Delta variant never dominates all the other Covid-19 variants, in Germany.

```

pred_samples_Italy_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "Italy"), allow_new_levels=TRUE, summary = FALSE)
variant_names <- c("Alpha", "Beta", "Gamma", "non_who", "Other")
variant_NotDelta_length <- length(variant_names)
Italy_Delta_dominates <- numeric(variant_NotDelta_length)
for (j in 1:variant_NotDelta_length) {
  pred_samples_Italy_NotDelta <- predict(model_1, newdata = filter(variants_5Largest,
    variant== variant_names[j], country == "Italy"), allow_new_levels=TRUE, summary = FALSE)
  n_Italy_Delta <- ncol(pred_samples_Italy_Delta)
  prob_Italy_Delta_MoreThan <- numeric(n_Italy_Delta)
  for (i in 1:n_Italy_Delta) {
    prob_Italy_Delta_MoreThan[i] <-
      mean(pred_samples_Italy_Delta[,i]>pred_samples_Italy_NotDelta[,i])
  }
  prob_Italy_Delta_MoreThan_True_or_False <- as.numeric(prob_Italy_Delta_MoreThan>0.5)
  Italy_Delta_dominates[j] <-
    as.numeric(Position(function(x) x>0, prob_Italy_Delta_MoreThan_True_or_False))
}
max(Italy_Delta_dominates)

```

```
## [1] NA
```

Is the first biweek, my model predicts; the Delta variant first dominates all the other Covid-19 variants, in Italy, if the output is 'NA', then the model predicts that Delta variant never dominates all the other Covid-19 variants, in Italy.

```

pred_samples_Spain_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "Spain"), allow_new_levels=TRUE, summary = FALSE)
variant_names <- c("Alpha", "Beta", "Gamma", "non_who", "Other")
variant_NotDelta_length <- length(variant_names)
Spain_Delta_dominates <- numeric(variant_NotDelta_length)
for (j in 1:variant_NotDelta_length) {
  pred_samples_Spain_NotDelta <- predict(model_1, newdata = filter(variants_5Largest,
    variant== variant_names[j], country == "Spain"), allow_new_levels=TRUE, summary = FALSE)
  n_Spain_Delta <- ncol(pred_samples_Spain_Delta)
  prob_Spain_Delta_MoreThan <- numeric(n_Spain_Delta)
  for (i in 1:n_Spain_Delta) {
    prob_Spain_Delta_MoreThan[i] <-
      mean(pred_samples_Spain_Delta[,i]>pred_samples_Spain_NotDelta[,i])
  }
}

```

```

}
prob_Spain_Delta_MoreThan_True_or_False <- as.numeric(prob_Spain_Delta_MoreThan>0.5)
Spain_Delta_dominates[j] <-
  as.numeric(Position(function(x) x>0, prob_Spain_Delta_MoreThan_True_or_False))
}
max(Spain_Delta_dominates)

```

```
## [1] NA
```

Is the first biweek, my model predicts; the Delta variant first dominates all the other Covid-19 variants, in Spain, if the output is 'NA', then the model predicts that Delta variant never dominates all the other Covid-19 variants, in Spain.

```

pred_samples_UnitedKingdom_Delta <- predict(model_1, newdata = filter(variants_5Largest,
  variant== 'Delta', country == "UnitedKingdom"), allow_new_levels=TRUE, summary = FALSE)
variant_names <- c("Alpha", "Beta", "Gamma", "non_who", "Other")
variant_NotDelta_length <- length(variant_names)
UnitedKingdom_Delta_dominates <- numeric(variant_NotDelta_length)
for (j in 1:variant_NotDelta_length) {
  pred_samples_UnitedKingdom_NotDelta <- predict(model_1, newdata = filter(variants_5Largest,
    variant==variant_names[j], country=="UnitedKingdom"), allow_new_levels=TRUE, summary = FALSE)
  n_UnitedKingdom_Delta <- ncol(pred_samples_UnitedKingdom_Delta)
  prob_UnitedKingdom_Delta_MoreThan <- numeric(n_UnitedKingdom_Delta)
  for (i in 1:n_UnitedKingdom_Delta) {
    prob_UnitedKingdom_Delta_MoreThan[i] <-
      mean(pred_samples_UnitedKingdom_Delta[,i]>pred_samples_UnitedKingdom_NotDelta[,i])
  }
  prob_UnitedKingdom_Delta_MoreThan_True_or_False <-
    as.numeric(prob_UnitedKingdom_Delta_MoreThan>0.5)
  UnitedKingdom_Delta_dominates[j] <-
    as.numeric(Position(function(x) x>0, prob_UnitedKingdom_Delta_MoreThan_True_or_False))
}
max(UnitedKingdom_Delta_dominates)

```

```
## [1] 1
```

Is the first biweek, my model predicts; the Delta variant first dominates all the other Covid-19 variants, in the United Kingdom, if the output is 'NA', then the model predicts that Delta variant never dominates all the other Covid-19 variants, in the United Kingdom.

*Code is fine but no dates. 4.*  
*13/15*