## Bayesian inference for covariance matrix

Ignacio Alvarez

IESTA, Universidad de la Repblica, Uruguay

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### Introduction

#### Covariance matrix estimation

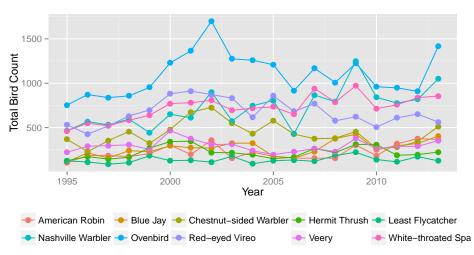
- Multivariate normal sampling models
- random-intercept, random-slope models

$$\begin{array}{lcl} y_{ij} & = & \beta_{0j} + \beta_{1j} x_{ij} + \beta_{2j} z_{ij} + \epsilon_{ij} \\ \begin{pmatrix} \beta_{0j} \\ \beta_{1j} \\ \beta_{2j} \end{pmatrix} & \sim & N\left(\begin{pmatrix} \mu_0 \\ \mu_1 \\ \mu_2 \end{pmatrix}, \Sigma\right), \ \epsilon_{ij} \sim N(0, \sigma^2) \end{array}$$

 Ovaskainen et al. (2017) includes covariance matrix priors for comunity data hierarchical models

# Bird counts on Superior forests

The Natural Resources Research Institute (University of Minnesota Duluth) carry out monitoring program for study regional population trends of forest birds.

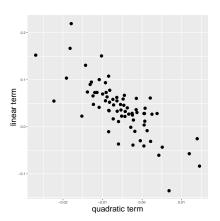


## Quadratic trend model

- $y_{st}$ : bird count for spacies s in year t.
- OLS regression model for each species

$$\mathit{y_{st}} \sim \textit{N}(\beta_{0s} + \beta_{1s}t + \beta_{2t}t^2, \sigma^2)$$

•  $Corr(\hat{\beta}_{1s}, \hat{\beta}_{2s}) = -.77$ 



## Quadratic trend model

Hierarchical linear model with IW prior.

$$y_{st} \sim N(eta_{0s} + eta_{1s}t + eta_{2t}t^2, \sigma^2)$$
 $egin{pmatrix} eta_{0j} \ eta_{1j} \ eta_{2j} \end{pmatrix} \sim N egin{pmatrix} 0 \ 0 \ 0 \ \end{pmatrix}, \Sigma \ & \sum \sim IW(d+1,I) \end{pmatrix}$ 

Define 
$$\rho = \Sigma_{23}/\sqrt{\Sigma_{22}\Sigma_{33}}$$
.

## Quadratic trend model

Hierarchical linear model with IW prior.

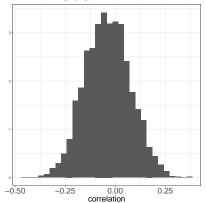
$$y_{st} \sim N(\beta_{0s} + \beta_{1s}t + \beta_{2t}t^2, \sigma^2)$$

$$\begin{pmatrix} \beta_{0j} \\ \beta_{1j} \\ \beta_{2j} \end{pmatrix} \sim N\begin{pmatrix} \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \Sigma \end{pmatrix}$$

$$\Sigma \sim N(d+1, I)$$

Define  $\rho = \Sigma_{23}/\sqrt{\Sigma_{22}\Sigma_{33}}$ .

### How does $p(\rho|y)$ look like?



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3 Separation strategy in linear model

Discussion

Consider *n* observations from  $Y_i \sim N_d(0, \Sigma)$  distribution.

Likelihood:

$$\rho(y|\mu,\Sigma) \propto |\Sigma|^{-n/2} e^{-\frac{1}{2}\sum_{i=1}^n y_i' \Sigma^{-1} y_i} = |\Sigma|^{-n/2} e^{-\frac{1}{2}tr(\Sigma^{-1}S_0)}$$

where  $y_i \in R^d$  a realization of  $Y_i$ ,  $S_0 = \sum_{i=1}^n y_i y_i^{'}$ , individual entries of  $\Sigma$  are  $\Sigma_{ii} = \sigma_i^2$  and  $\Sigma_{ij} = \sigma_i \sigma_i \rho_{ij}$ .

Alvarez-Castro et al. (2014) compare alternative priors for  $\Sigma$  in this context.

- using simulations
- with the bird count data set (not shown)

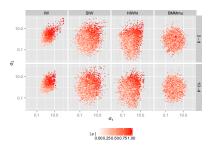
### Alternative $\Sigma$ priors

• Inverse Wishart:  $IW(v, \Lambda)$ .

$$p(\Sigma) \propto |\Sigma|^{-(\nu+d+1)/2} e^{-\frac{1}{2}tr(\Lambda\Sigma^{-1})}$$

- Scaled Inverse Wishart
- Hierarchical inverse Wishart
- Separation strategy

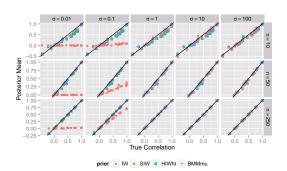
# Samples from prior



Positive relationship among  $\sigma_1$  and  $\sigma_2$ , also large  $|\rho_{12}|$  values appear when the two variances are high.

## Impact on posterior inference

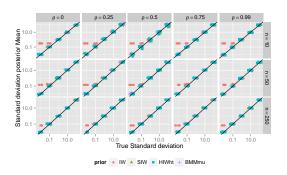
### Inference for $\rho_{12}$



When standard deviation is small,  $\sigma=0.01$  or  $\sigma=0.1$  the IW prior heavily shrinks the posterior correlation towards 0 even if the true correlation is close to 1.

## Impact on posterior inference

#### Inference for $\sigma_1$



*IW* prior overestimate the standard deviation when its true value is very low

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# Separation strategy in linear model

$$y_{st} \sim N(eta_{0s} + eta_{1s}t + eta_{2t}t^2, \sigma^2)$$

$$\begin{pmatrix} eta_{0s} \\ eta_{1s} \\ eta_{2s} \end{pmatrix} \sim N\begin{pmatrix} \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}, \Sigma \end{pmatrix}$$

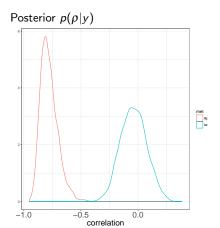
$$\Sigma = \operatorname{diag}(\lambda) \Omega \operatorname{diag}(\lambda)$$

$$\lambda_k \sim Ca^+(0,1)$$

$$\Omega \sim LKJ(3)$$

# Separation strategy in linear model

$$y_{st} \sim \mathcal{N}(eta_{0s} + eta_{1s}t + eta_{2t}t^2, \sigma^2)$$
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#### Discussion

Prior choice (based on multivariate normal data results)

- Separation strategy gives modeling flexibility and good inferences properties. (Barnard et al., 2000)
- ② In Gibbs base samplers a prior which maintain conjugacy might be preferable (scaled inverse Wishart or hierarchical inverse Wishart)
- $\odot$  If we are constraint to use IW, we may recommend to scale the data first.

#### Future steps

Different Model Hierarchical linear model context.

Different Priors Use LKJ prior for the correlation matrix.

Use other distributions for IW parameters v and  $\Lambda$ .

#### References

- Alvarez-Castro, I., Niemi, J., Simpson, M., Alvarez, I., Niemi, J., and Simpson, M. (2014), "BAYESIAN INFERENCE FOR A COVARIANCE MATRIX Bayesian Inference for a Covariance Matrix," in *Annual Conference on Applied Statistics in Agriculture*, New Prairie Press, pp. 71 – 82
- Barnard, J., McCulloch, R., and Meng, X.-L. (2000), "Modeling covariance matrices in terms of standard deviations and correlations, with application to shrinkage," *Statistica Sinica*, 10, 1281–1312
- Ovaskainen, O., Tikhonov, G., Norberg, A., Guillaume Blanchet, F., Duan, L., Dunson, D., Roslin, T., and Abrego, N. (2017), "How to make more out of community data? A conceptual framework and its implementation as models and software." *Ecology letters*, 20, 561–576.