

Reproducibility and Applied Econometrics - The Effect of Studying on Grades

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Abstract

In this paper we establish a framework for reproducible empirical research. We use a non-standard 2SLS model to estimate the marginal effects of studying on grades. The paper is split into two distinct sections. The first part is the econometric analysis on the causal impact of studying. The second part details the steps taken to ensure reproducibility and suggests how to easily integrate these methods into a researcher's future projects.

Git Repository: https://github.com/nadavtadelis/Reproducible_Metrics

1 Introduction

In recent years there has been a strong push to increase the reproducibility and replicability of scientific research. Unfortunately this movement seems to have been centered on the hard sciences and has not yet become standard practice in the social sciences. It is possible that this is partly due to a lack of reproducibly researched papers in these fields. This paper explores what responsible and reproducible research practices look like in applied econometrics. We present an instrumental variables approach to estimating the causal effect of studying on grades and develop custom python scripts to implement an unusual 2SLS set up that allows for nonlinearity in our endogenous predictor. The latter part of this work discusses the current state of reproducibility in econometric research, and explains in detail the techniques implemented in the analysis.

This paper was written as my honors thesis for undergraduate studies in statistics at UC Berkeley. I would like to thank my thesis advisor, Professor Fernando Pérez, for his generosity with his time and deep knowledge of the reproducibility movement. Due to deadlines for submission I was unable to spend an appropriate amount of time on the econometric analysis, and will point out some of the weak points in my models (specifically the instruments). Any comments or suggestions would be greatly appreciated.

The rough idea for the econometric analysis in the paper is adapted from an independent project I completed during my Junior year. In a subsequent class the project was modified and rebuilt in a reproducible fashion. Sarah Johnson created the original intermediate functions in `p3functions.py` [p3functions.py](#), and the associated tests and Travis integration. Chitwan Kaudan created the original Makefile for running the individual Jupyter Notebooks. All aspects of those original analyses have been altered significantly, and the history of the alterations is fully documented in the commit history of the git repo.

[Maybe more stuff needs to go here before we dive in]

2 Data

The [data](#) being used are from the public archive of UCI's machine learning repository and were collected by Paulo Cortez of the University of Minho, Portugal in the 2005 - 2006 academic year (Cortez and Silva [2008]). The data were collected in two secondary schools in the Alentejo region of Portugal, using school reports and questionnaires. The data were cleaned to only include students for which all the variables are known - and a further 111 students were discarded because of mismatched information between the surveys and the school reports. The data come with a file containing attribute information which can be found [here](#); these include school, course, and many individual level characteristics. The data include 649 students from a Portuguese Language course of study, and 395 students from a Mathematics course.

[Need more background here?]

2.1 Data Exploration

There is some pre-analysis data cleaning and exploration that we run, which can be found in the data exploration notebook - [nbviewer](#), [git](#); some figures from this notebook are included in the Appendix [\[MAKE SURE TO INCLUDE THESE\]](#). The notebook contains the full data exploration process, as well as explanations and commentary detailing the figure generation process.

We draw attention to two attributes of special import: there are stark differences in the distributions of the data between the two schools, and across the two courses of study. The school level differences seem to be fairly consistent; perhaps the two schools have slightly different grading policies. However, the within school course level differences are highly variable; this is perhaps capturing the heterogeneity between students who choose Portuguese Language and Mathematics. This strongly suggests that even our simplest models should include controls for school to account for different policies, and should estimate the two courses as separate entities to account for different types of students.

[Very Rough]

2.2 Data Issues

As is often the case in the education space, the data have some large issues that must be considered. As any undergraduate taking their first econometrics course would point out, much of the data are stated preferences (as opposed to observed). Happily, the test

scores and number of absences are provided by the schools, so we can be certain in their accuracy. The issue with self-reported data is the potential for inaccuracy. Even if a student is not maliciously providing misinformation, it is likely that personal biases are affecting the response. With variables like weekly hours of free time, this self reported data might actually be a benefit, as we are getting information about the student's perception of reality rather than the truth. If a student reports that they have very little free time then that tells us something about how they view their current time allocations. As such, these variables might be useful as controls for student level heterogeneity, but should be considered with a healthy dose of skepticism, and estimated coefficients should be interpreted with care. [Think about this more.] The issue of self reported data becomes more problematic when it comes to our variable of interest - hours of studying per week. While our central research goal is identifying a relationship between hours of studying and academic success, our data on hours of studying cannot be trusted as accurate [expand more here].

In addition to the issue of self reporting, our data on studying time (and other variables) suffers from another issue; categorical mappings of quantitative measures. Many of the quantitative variables in the data are reported as categorical bins. For example, weekly studying time is coded as four distinct levels (0-2 hours, 2-5 hours, 5-10 hours, and 10+ hours). In the case of studying time, we explore two different re-mapping schemes¹. But for the other variables with this format we treat them as categorical variables and include indicators for each level (thus allowing for some nonlinearity). [expand more here]

Lastly, the data are cross-sectional rather than longitudinal, and the specific sampling time is unspecified. [Expand more here, talk about why this is bad when studytime changes through the academic year. This might be a good place to include the simultaneous causality issue, or at least allude to it so that it is more familiar when we go over it later (?)]

[Very Rough]

¹Mapping both as discrete and continuous, detailed in the data exploration notebook - [nbviewer](#), [git](#)

3 Models

We need to first set up our structural equation [Maybe this isn't actually a structural equation? Need to double check this terminology] defining the relationship between grades and studying (Card and Krueger [1992]). Let an individual's grade be g_i and weekly hours of studying be s_i and their "ability" be a_i . Then our model is:

$$g_i = \beta_0 + \beta_1 s_i + \beta_2 s_i^2 + \beta_3 a_i + \beta_{4:k} X_i + \varepsilon_i$$

Where X_i is a matrix of school and course level characteristics and ε_i captures some unobserved heterogeneity and disturbances. Note that grades are nonlinear in weekly study time. While this nonlinearity complicates our model, it seems necessary because assuming that marginal returns to studying must vary depending on the initial level of studying.

Clearly, there are issues with this model. How are we defining grades? The ideal set up would have a course specific set of simultaneous equations, where the number of equations is equal to the number of classes. The next best set up would involve estimating one equation for each type of class (quantitative, literary, historical, etc.) within each course. Another alternative would be to define grades as cumulative GPA. In this analysis, due to the limitations of the data, we define g_i as the student's score on the final test for their course of study (G3 in the data).

Another issue with this model is: how are we defining ability? By its very nature, ability is unobserved. We can proxy for ability using other individual level characteristics (intelligence, age, parents' education, etc.) but we cannot fully capture ability because it does not have a clear measurable meaning. Hence, we must keep in mind that the model is never going to be fully specified.

We can think of a student's utility maximization problem as being some function of grades, free time (let's consider everything that is not studying, sleeping, or class as "free time"), and studying time. We would expect that the coefficients on grades and free time would be positive, and the coefficient on study time would be negative, with magnitudes of these coefficients being determined by an individual's preferences. For example, a student who cares very little about grades, enjoys constantly partying, and hates studying, would have a small positive coefficient on grades, a large positive coefficient on free time, and a large

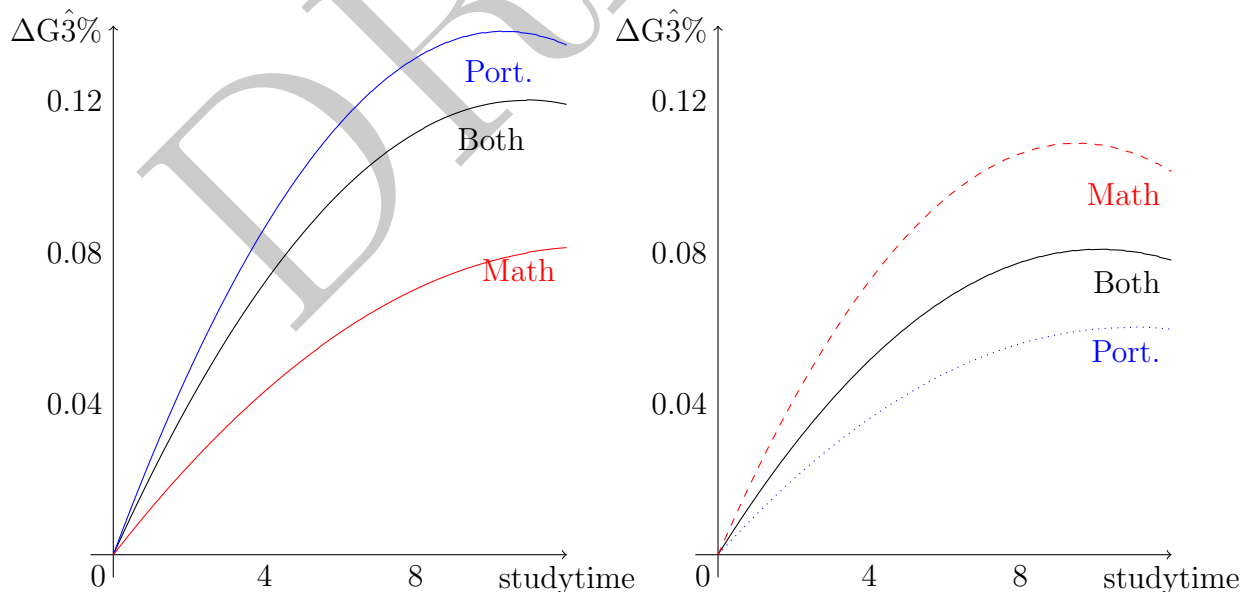
negative coefficient on studying time². This student would maximize their utility and choose how to allocate their time, and would probably end up spending very little time studying. Notice that before maximizing their utility, an individual would replace grades with the previously defined model for grades (dependent on studying and ability); so someone who heavily values grades could end up having a positive coefficient on studying after including the model for grades into their utility function, even if they do not intrinsically value studying.

Establishing this utility function gives motivation for including variables that might introduce multicollinearity. For example, weekly amount of free time might not improve our estimate of the marginal effect of studying, and would be collinear with the amount of weekly studying. However, when we think of our observations as realizations of a decision making process that involves utility maximization, there is an argument for including free time in the final model estimation procedure.

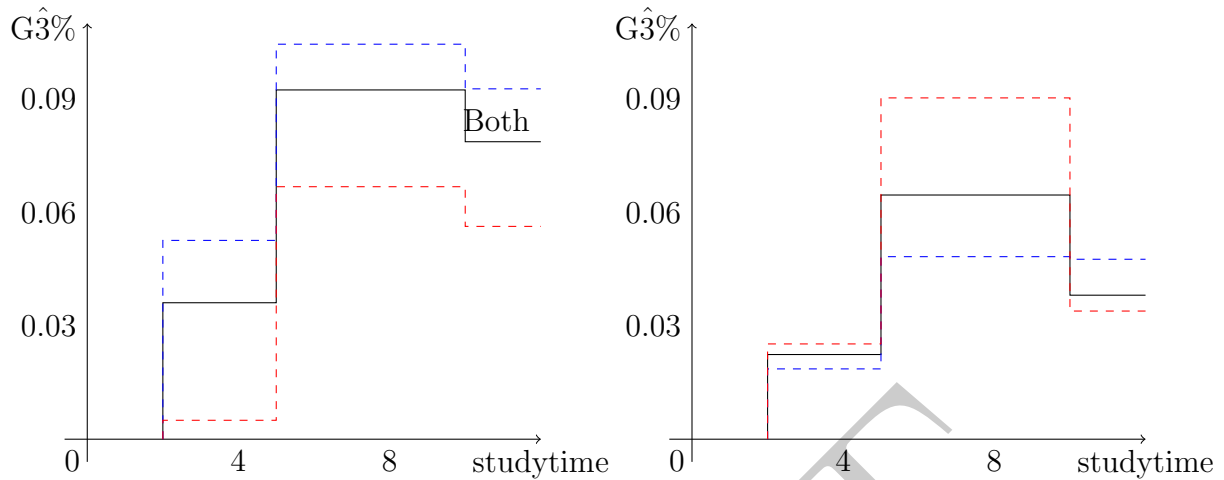
3.1 Naive OLS

[Report the four fits from the naive OLS - first one with only 'studytime' and school level characteristics, then one with individual level characteristics (for each course of study)]

[Maybe include plot of estimated relationship between 'studytime' and 'G3' (nonlinear, one line for math, one for Portuguese, line type can denote individual level controls - or will that make the graph to crowded?)]



²Of course, for some people and some subject matters of study, the coefficient on studying time may be positive with a decreasing marginal utility. We simplify the specification here dramatically.



[Plots need finishing (better labels, axis names, titles for with/without controls, choose line type and color, etc.) when explaining the plots, maybe call the y axis studytime's contribution to the predicted G3 percent?]

[Maybe only include the first two plots in this section, and include the discontinuous mapping only in appendix?]

3.1.1 Additional Analysis

[Lasso-flavor penalization and VIF discussion here (maybe no numbers, just general take-aways and link to the relevant notebook), point out that Lasso keeps the variable of interest and also explore the other variables that it weights highly]

note: in lasso section cite: (Cameron [2017])

note: in VIF section cite: (O'Brien [2007]) and (Mansfield and Helms [1982])

Make sure to include discussion of why it seems okay to include parental education even though it is a bit collinear - because they are considered good proxies for ability

3.1.2 Simultaneous Causality

There is an additional issue with this model specification that we did not explore in the original discussion. The data include three test scores: G1, G2, and G3. G1 and G2 are midterms, and G3 is the final. It is plausible that part of how students inform their study allocation decision comes from how well they did on their previous exam and how many hours they studied in the time leading up to the exam. The data is taken at a single undefined point in time, so it is unclear at what point the students are reporting their weekly amount of studying. This complicates things because if the survey was administered after the midterms, then there is clear simultaneous causality between study time, G1 and G2.

This is the motivation for excluding G1 and G2 from the model estimation. There may be a novel way of

[Need to edit this section, right now its really rough]

[Maybe move this to right before the Naive OLS section?]

Actually: [Maybe move this to the Data Issues as its own subsection?]

3.1.3 Endogeneity

In the second run of naive regressions we add many covariates as proxies representing unobserved ability (parental education, individual characteristics, and others detailed in the model fitting 1 notebook - [nbviewer](#), [git](#)). We pointed out that since ability is unobservable and undefinable we cannot expect to fully capture it through our proxies. Let q_i be the unobserved portion of ability that we are not capturing through our imperfect proxies. Then we have:

$$g_i = \beta_0 + \beta_1 s_i + \beta_2 s_i^2 + \beta_3 school_i + \beta_4 course_i + \beta_5 x_{i,5} + \dots + \beta_k x_{i,k} + \gamma q_i + \nu_i$$

Where $x_{i,5}, \dots, x_{i,k}$ are proxies and ν_i is the structural error term. Let $x_{i,j}$ be the j 'th covariate in the design matrix ($x_{1i} = 1$, $x_{2i} = s_i$, etc.) and let \mathbf{x}_i be the vector of covariates for i . We are mainly interested in β_1 and β_2 the coefficients on studying. It is reasonable to assume that $E[\nu_i | \mathbf{x}_i, q_i] = 0$ unfortunately we have no choice but to stick q_i into the error term which gives us the new structural equation:

$$g_i = \beta_0 + \beta_1 s_i + \beta_2 s_i^2 + \beta_3 school_i + \beta_4 course_i + \beta_5 x_{i,5} + \dots + \beta_k x_{i,k} + \varepsilon_i$$

Where $\varepsilon_i = \gamma q_i + \nu_i$. We know that ν_i is well behaved (uncorrelated with $x_{i,j} \forall j \in \{1, k\}$) but ε_i is uncorrelated with the covariates iff q_i is uncorrelated with the covariates. Lets look at the consistency of this set up (starting with the linear projection of q_i onto \mathbf{x}_i):

$$g_i = \delta_0 + \delta_1 s_i + \delta_2 s_i^2 + \delta_3 school_i + \delta_4 course_i + \delta_5 x_{i,5} + \dots + \delta_k x_{i,k} + \eta_i$$

$$E[\eta_i] = 0 \tag{1}$$

$$Cov[x_{i,j}, \eta_i] = 0 \tag{2}$$

Where (1) and (2) follow from the definition of linear projections. Now we can represent our original model in terms of this linear projection:

$$\begin{aligned}
g_i &= \beta_0 + \beta_1 s_i + \beta_2 s_i^2 + \beta_3 school_i + \beta_4 course_i + \beta_5 x_{i,5} + \dots + \beta_k x_{i,k} \\
&\quad + \gamma \delta_0 + \gamma \delta_1 s_i + \dots + \gamma \delta_k x_{i,k} + \gamma \eta_i + \nu_i \\
g_i &= (\beta_0 + \gamma \delta_0) + (\beta_1 + \gamma \delta_1) s_i + (\beta_2 + \gamma \delta_2) s_i^2 + \dots + (\beta_k + \gamma \delta_k) x_{i,k} + \gamma \eta_i + \nu_i \\
&\Rightarrow E[\gamma \eta_i + \nu_i] = 0 \\
&\text{Cov}[x_{i,j}, \gamma \eta_i + \nu_i] = 0
\end{aligned}$$

This shows that we can run OLS and get consistent estimates of $(\beta_j + \gamma \delta_j)$, but we are actually interested in β . OLS does not allow us to recover β in this case because:

$$\hat{\beta}_j \xrightarrow{p} \beta_j + \underbrace{\gamma \delta_j}_{\text{omitted variable bias}}$$

If $\gamma \neq 0$ and $\delta_j \neq 0$ then we cannot get a consistent estimate of β_j using OLS. Rather, we get an endogenous estimator with a bias of $\gamma \delta_j$ (notice that the better our proxies identify ability, the smaller the bias term will be). Often times in practice researchers will assume that all δ_j 's are zero except for the ones on the variables of interest. The implicit assumption behind setting $\delta_j = 0$ is that the unobserved q_i and $x_{i,j}$ are independent (i.e. $\text{Cov}[x_{i,j}, q_i] = 0$). In our setting this is actually a very reasonable assumption; we have already defined $x_{i,5}, \dots, x_{i,k}$ as proxies for ability, meaning that the unobserved portion of ability should be orthogonal to each of these variables. It is also reasonable to assume that there is independence between school and course level characteristics and the unobserved q_i . This simplifies³ our endogeneity problem so that the only endogenous variables are s_i and s_i^2 ($\text{Cov}[s_i, \varepsilon_i], \text{Cov}[s_i^2, \varepsilon_i] \neq 0$). In the next sections we first give a brief overview of the instrumental variables approach to addressing endogeneity, and explain why standard two stage least squares is intractable in our setting. We then explore an instrumentation approach with nested generated regressors that may solve the problem.

³If we had only one endogenous variable we could simplify further because $\hat{\beta}_j \xrightarrow{p} \beta_j + \gamma \frac{\text{Cov}[x_j, q]}{\text{Var}[x_j]}$. In this case we would be able to make some guesses about the direction of the bias.

3.2 Q2SLS

[Introduce idea of instrumental variables really briefly, maybe just go over the conditions for validity of instruments (see page 45 of 212 notes)]

Lets first give a brief refresher of the instrumental variables approach to addressing endogeneity. In the last section we showed that for an endogenous variable $x_{i,j}$ OLS is no longer a consistent estimator for β_j . The instrumental variables approach allows us to solve the endogeneity problem. We require an instrumental variable z_i which is observable and satisfies the following conditions:

- (1) The instrument is uncorrelated with disturbances: $\text{Cov}[z_i, \varepsilon_i] = 0$
- (2) In the reduced form of the endogenous variable $x_{i,j} = \phi_0 x_{i,1} + \phi_1 \text{school}_i + \dots + \phi_{k-2} x_{i,k} + \theta_1 z_i + \vartheta_i$ the coefficient on our instrument is nonzero: $\theta_1 \neq 0$
- (3) And z_i is not one of our exogenous variables in the original estimation

If (1) – (3) hold, then we call z_i a valid instrument for $x_{i,j}$. In settings with multiple valid instruments z_1, z_2, \dots, z_M the standard procedure for identifying β is 2SLS. In the first stage of 2SLS we regress the endogenous variable $x_{i,j}$ on the exogenous variables and the instruments to get the fitted values $\hat{x}_{i,j} = \sum_{k \neq j} \hat{\phi}_k x_{i,k} + \hat{\theta}_1 z_{i,1} + \dots + \hat{\theta}_M z_{i,M}$. We use $\hat{x}_{i,j}$ as an estimate of the exogenous part of $x_{i,j}$ and in the second stage we regress our dependent variable on our exogenous variables and $\hat{x}_{i,j}$ to identify $\hat{\beta}$. This approach can be easily extended to cases with multiple endogenous variables (where you would have as many first stage regressions as there are endogenous variables). However, with multiple endogenous variables, you must have at least as many instruments as endogenous variables, otherwise the fitted values would be collinear and the second stage estimators would lose consistency. In the next section we discuss issues with the 2SLS method in our setting.

3.2.1 Motivation

[Point out issue with trying to use 2SLS when you want to account for diminishing marginal returns in the endogenous variable, especially when the instruments are binary (so squaring them does nothing). Lead to discussion of Q2SLS procedure from Wooldridge]

In section 3.1.3 we established that the causal relationship between our grades and studying time is quadratic. [NOTE: for now I'm going to assume we only end up using the 'good' binary instrument, but even if we end up using more binary instruments, the procedure still holds but it is less clear that its the right approach.] This would usually not be a barrier to

using 2SLS; the standard approach⁴ is to use the original instrument and its square as instruments in the two first stage equations (Angrist and Pischke [2009]). However, this approach is unusable when the only instrument available is binary. In fact, even if we have multiple binary instruments it is not convincing to think that the exogenous portion of the squared endogenous variable will be correctly identified because interacting the binary instruments gives very little additional variation [THINK ABOUT THIS MORE].

3.2.2 Estimation Procedure

[describe how to for q2sls] or maybe this should be in above section?

3.2.3 Properties of Q2SLS

[Explain the testing that I did for Q2SLS, mention that it looks like the coeff.s on endog_sq_hat and the exog vars are consistent and unbiased, but that the coeff on endog_hat is consistent but biased. Point out the oddness of this behavior and note that I plan on researching this further.] Note: pretty sure my instruments are weak (need to check), and the simulation seems to suggest that in the case of weak instruments the estimates are still okay (maybe) but that the SE's are going to be too large for the endogenous vars.

[Question:] Since both fitted values from the first stage are linear combinations of the instruments (albeit the second part being built with an added instrument), could there be collinearity between them? The only reason I'm skeptical is that it doesn't seem like we have a collinearity issue in the simulation (if they were collinear then wouldn't we expect to see the range of the coefficient to be very wide, and include the true coefficient within it? not the consistent bias that we see?). I think perhaps the inclusion of the new instrument in the second part actually means that there isn't collinearity. Mostly Harmless (p.192) suggests that in the case of squared endogenous vars, you can use the two part first stage set up, and then in the second part just use the instrument and the instrument squared, which should mean that as long as there is a new instrument in the second part *that is a nonlinear transformation* of the first part, collinearity shouldn't be an issue (this also fits with the intuition I have).

[Discuss difficulty with finding the asymptotically consistent estimator for variance in Q2SLS because of the nested generated regressors]

⁴Assuming only one instrument is available.

[Give brief overview of bootstrapping and explain how its being used here to estimate variance of coeff. estimates in 2nd stage] Also add note about future integration of a wild bootstrap approach to variance estimation (stricter assumptions, but some really nice properties for endogenous models - see slide 15 from the link below)

FOR BOOTSTRAPPING COV MATRIX: See 8th slide of <https://www.math.kth.se/matstat/gru/sf293>
It also has notes on hypothesis testing and p values for bootstrapping with regression coefficients. SLIDE 14 HAS DISCUSSION OF BOOTSTRAPPING WITH ENDOGENOUS VARS!!

NOTE: this procedure only requires one instrument because the fitted vals from part a act as a second instrument so theres for sure no collinearity, this means we could potentially only use one binary instrument!!! maybe this is the direction we should take rather than using a less than stellar second instrument. If this is the direction we go then I should repeat the function testing using binary instruments.

3.2.4 Application

[Discuss chosen instruments and give motivation for why they might be okay to use, give strong disclaimer that even if they are valid, they are probably weak instruments]

Note: Include discussion of validity of exclusion restriction - <http://econ.lse.ac.uk/staff/spischke/ec533/Weak%20IV.pdf>

[Report results]

[Discuss results]

Maybe: [Reference appendix, and in the appendix include the results from Q2SLS with only 'essential' controls in the extra results subsection. Similar to the approach for the naive OLS at the beginning]

4 Reproducibility

The importance of reproducible research has been understood for many years. In 1980 Thomas Mayer said “Neither originality, logical rigor or any other criterion is ranked as ‘essential’ by so many natural scientists as is replicability”. Mayer was focusing on replicable research; unfortunately in the social science replicability is often far harder than reproducibility. When it takes five years and tens of thousands of dollars to collect the data for a single study, it is a difficult proposition to attempt to replicate the study from scratch. Additionally, it is unlikely that Mayer could have foreseen the degree to which data wrangling and programming would penetrate into academic research. This reliance on complex data pipelines and analysis with hundreds of dependencies makes replication a beast of a whole new nature. In the current state of empirical research, reproducing the results is half the battle. This opaqueness has led to many published papers that included results that were ‘p-hacked’ or just included mistakes. Ensuring reproducibility would help alleviate many of these issues, and is a big step on the road to replicability.

Reproducibility has become especially important in the current political environment. The National Association for Scholars’ recently published a report in which they advocate for use of the Secret Science Reformation Act (renamed the HONEST Act), which would forbid the Environmental Protection Agency from using any research that is not “substantially reproducible” Schulson [n.d.]. The NAS report goes even further and suggests that the bill should encompass all federal agencies and courts. Putting aside the various issues with the act⁵, if HONEST were to be expanded to other federal agencies, then policy-oriented research in the social sciences would be forced to adapt. Policy makers would need to see “substantial reproducibility” before allowing a study to influence their policy decisions. Currently, very few empirical papers could be labeled as even somewhat reproducible, and part of the reason why might be that researchers believe implementing reproducible methods into their projects would be too costly in time and resources Feigenbaum and Levy [1993]. Hopefully this paper provides a strong argument showing that conducting empirical research using reproducible methods is not only easy, but beneficial. And that were HONEST to be implemented to a wide degree, research need not be impeded⁶.

[Need to clean up this section a lot]

⁵The act does not define “substantially reproducible” with any rigor. It is possible that the flexibility of the act could lead to climate change deniers (and others) to call almost any study not substantially reproducible and bar federal agencies from acting on good research.

⁶Ignoring the issue of data confidentiality.

4.1 Current Resources

[Discuss current approaches to reproducibility in econometrics]

[Especially point out Gentzkow & Shapiro's guide for research methods, and its helpfulness in ensuring that research workflow makes sense, but it's lack of open source reproducibility]

4.2 Basic Workflow

[Discuss git version control, environments, makefiles, using notebooks as a way to document any analysis that was not included in the final paper, but had an impact on the direction of the results (this increases confidence that there is not any unintentional p-hacking style things going on)] [Definitely link to the stat 159 page because we don't want to go into the nitty gritty details of everything, just an overarching view of the tools and how they're used together. also explain the benefits of all of these tools (ability to see edit history, ease of remote collaboration, ensuring compatibility on multiple machines, automatic function testing, etc.)]

4.3 Custom Functions

[Discuss the process of creating your own function in a way that other people can also use it without hassle: docstrings, comments, readability, etc.]

[Cover the long process of testing a new function that isn't necessarily included in the final tests.py script (function_testing.ipynb)]

[Explain Travis and CI]

4.4 Optional/Additional Elements

[Cover optional but useful things that the open source community has begun using: Binder, Sphinx, etc.] and discuss things I didn't have time to implement

[FOR SURE: mention that I did not have time to set up a way of porting the coeff.s and SE's directly from the notebooks to the tex file, and that ideally this would be fully automatic, so if the results changed in the jupyter notebook, then they would change in the tex file and the pdf (although since we are careful with the environments and version control, the coeffs should never change unless the model specifications change)]

5 Conclusion (? - maybe not needed)

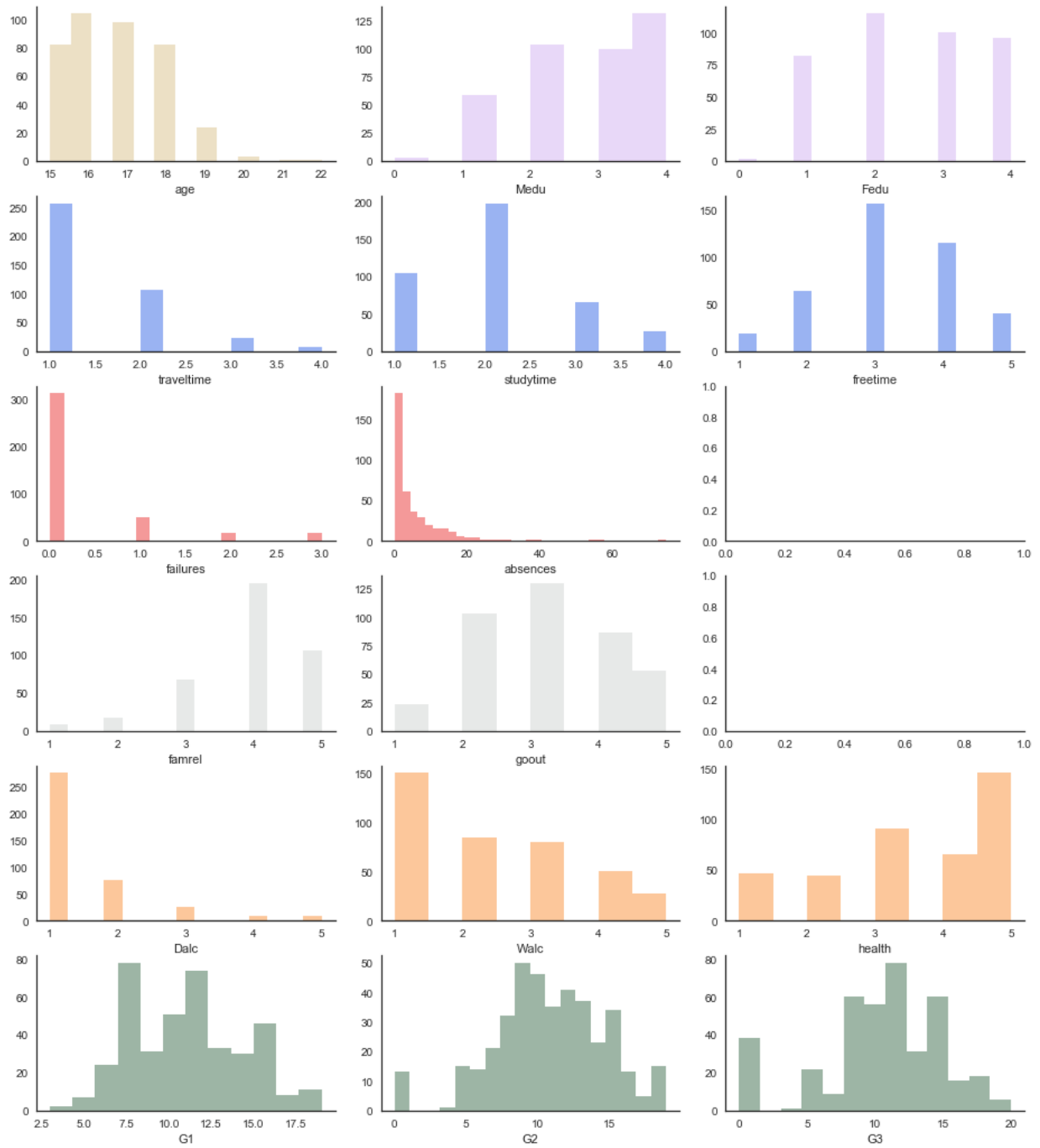
Maybe include a conclusion in the Q2SLS section, but not here.

DRAFT

6 Appendix

6.1 Data Exploration Figures

Histograms Grouped By Related Variables (Mathematics Course)



6.2 Naive OLS Extra Models

[PUTTING ALL TABLES HERE FOR NOW BECAUSE HAVING TROUBLE PUTTING THEM IN SPECIFIC SECTION]

QUESTION: [Should we report these tables in the main paper and then report tables with all the covariates in the appendix? These would be extra long, and are viewable in the model fitting 1 notebook, so maybe it'd be better to just link to them?]

Table 1: Naive OLS, Discrete Mapping

	<i>Dependent variable: G3 Grade (Percent)</i>					
	<i>Both</i>	<i>Portuguese</i>	<i>Mathematics</i>	<i>Both</i>	<i>Portuguese</i>	<i>Mathematics</i>
	(1)	(2)	(3)	(4)	(5)	(6)
Constant	0.5120*** (0.0137)	0.4942*** (0.0143)	0.4784*** (0.0366)	0.5611*** (0.1392)	0.3397** (0.1381)	1.0857** (0.4946)
Studytime_2	0.0361*** (0.0136)	0.0526*** (0.0143)	0.0050 (0.0285)	0.0224 (0.0142)	0.0186 (0.0149)	0.0252 (0.0327)
Studytime_3	0.0924*** (0.0177)	0.1045*** (0.0166)	0.0668* (0.0380)	0.0646*** (0.0183)	0.0483*** (0.0184)	0.0903** (0.0426)
Studytime_4	0.0787*** (0.0285)	0.0927*** (0.0278)	0.0563 (0.0575)	0.0381 (0.0293)	0.0476* (0.0269)	0.0339 (0.0672)
School_GP	0.0742*** (0.0133)	0.0856*** (0.0142)	0.0283 (0.0337)	0.0371** (0.0150)	0.0605*** (0.0160)	-0.0309 (0.0453)
Course_math	-0.0954*** (0.0128)			-0.0958*** (0.0150)		
Individual Controls	<i>no</i>	<i>no</i>	<i>no</i>	<i>yes</i>	<i>yes</i>	<i>yes</i>
Observations	1,044	649	395	1,044	649	395
R ²	0.094	0.131	0.015	0.328	0.410	0.352
Adjusted R ²	0.090	0.126	0.005	0.282	0.342	0.219
F Statistic	24.62***	22.95***	1.469	7.04***	5.72***	2.834***

Note:

*p<0.1; **p<0.05; ***p<0.01

Table 2: Naive OLS, Continuous Mapping

	<i>Dependent variable: G3 Grade (Percent)</i>					
	<i>Both</i>	<i>Portuguese</i>	<i>Mathematics</i>	<i>Both</i>	<i>Portuguese</i>	<i>Mathematics</i>
	(1)	(2)	(3)	(4)	(5)	(6)
Constant	0.4879*** (0.0161)	0.4686*** (0.0165)	0.4572*** (0.0402)	0.5497*** (0.1384)	0.3312** (0.1365)	1.0605** (0.4971)
Studytime_continuous	0.0220*** (0.0053)	0.0268*** (0.0052)	0.0128 (0.0113)	0.0160*** (0.0056)	0.0110** (0.0056)	0.0229* (0.0126)
Studytime_continuous_sq	-0.0010*** (0.0004)	-0.0013*** (0.0004)	-0.0005 (0.0007)	-0.0008** (0.0004)	-0.0005 (0.0004)	-0.0012 (0.0008)
School_GP	0.0737*** (0.0133)	0.0856*** (0.0142)	0.0262 (0.0335)	0.0372** (0.0150)	0.0605*** (0.0161)	-0.0312 (0.0453)
Course_math	-0.0956*** (0.0128)			-0.0957*** (0.0150)		
Individual Controls	<i>no</i>	<i>no</i>	<i>no</i>	<i>yes</i>	<i>yes</i>	<i>yes</i>
Observations	1,044	649	395	1,044	649	395
R ²	0.09	0.13	0.01	0.33	0.41	0.35
Adjusted R ²	0.09	0.13	0.00	0.28	0.34	0.22
F Statistic	30.11***	30.61***	1.41	7.09***	5.84***	2.97***

Note:

*p<0.1; **p<0.05; ***p<0.01

6.3 Q2SLS Function Testing Procedure

[Can use some of the markdown from the functions testing notebook here, and obvs link to the notebook as well]

6.4 Q2SLS Extra Models

6.5 Miscellanea

References

- Angrist, J. D., Imbens, G. W. and Rubin, D. B. [1996], ‘Identification of causal effects using instrumental variables’, *Journal of the American Statistical Association* **91**, 444–455.
- Angrist, J. D. and Pischke, J.-S. [2009], *Mostly Harmless Econometrics: An Empiricist’s Companion*, Princeton University Press.
- Cameron, A. C. [2017], ‘Machine learning for microeconometrics’.
URL: <http://cameron.econ.ucdavis.edu/e240f/trmachinelearningseminar.pdf>
- Card, D. and Krueger, A. B. [1992], ‘Does school quality matter? returns to education and the characteristics of public schools in the united states’, *Journal of Political Economy* **100**(1), 1–40.
- Cortez, P. and Silva, A. [2008], ‘Using data mining to predict secondary school student performance’, *A. Brito and J. Teixeira Eds., Proceedings of 5th FUTURE BUSINESS TECHNOLOGY Conference (FUBUTEC 2008)* pp. 5–12.
URL: <http://www3.dsi.uminho.pt/pcortez/student.pdf>
- Feigenbaum, S. and Levy, D. M. [1993], ‘The market for (ir)reproducible econometrics’, *Social Epistemology* **7**(3), 215–232.
- Greene, W. H. [2011], *Econometric Analysis*, 7 edn, Prentice Hall.
- MacKinnon, J. G. and White, H. [1985], ‘Some heteroskedasticity-consistent covariance matrix estimators with improved finite sample properties’, *Journal of Econometrics* **29**(3), 305–325.
- Mansfield, E. R. and Helms, B. P. [1982], ‘Detecting multicollinearity’, *The American Statistician* **36**(3a), 158–160.
- O’Brien, R. M. [2007], ‘A caution regarding rules of thumb for variance inflation factors’, *Quality & Quantity* **41**, 673–690.
- Schulson, M. [n.d.], ‘Science’s “reproducibility crisis” is being used as political ammunition’.
URL: <https://www.wired.com/story/sciences-reproducibility-crisis-is-being-used-as-political-ammunition/>
- Tibshirani, R. [1996], ‘Regression shrinkage and selection via the lasso’, *Journal of the Royal Statistical Society. Series B (Methodological)* **58**(1), 267–288.
- Wooldridge, J. M. [2001], *Econometric Analysis of Cross Section and Panel Data*, MIT Press Books, 2 edn, The MIT Press.