

5.8) Joint distribution can be factorized into $\theta = (\theta_1, \theta_2)$

a) $p(x, y | \theta) = p(y | x, \theta_2) p(x | \theta_1)$

So, table becomes

	$y=0$	$y=1$
$x=0$	$\theta_2(1-\theta_1)$	$(1-\theta_1)(1-\theta_2)$
$x=1$	$\theta_1(1-\theta_2)$	$\theta_1\theta_2$

Data

x	y
1	1
1	0
0	0
1	0
1	1
0	0
0	1

b) MLE Solution for likelihood given

$x = (1, 1, 0, 1, 1, 0, 0)$ $y = (1, 0, 0, 0, 1, 0, 1)$ in

$$\hat{\theta} = \underset{\theta_1, \theta_2}{\operatorname{argmax}} \log \prod_{i=1}^n p(x_i | \theta)$$

$$= \underset{\theta}{\operatorname{argmax}} \cdot \log \left((\theta_1 \cdot \theta_2)^4 (\theta_1(1-\theta_2))^3 (\theta_2(1-\theta_1))^4 (\theta_1(1-\theta_2)) \right)$$

$$\cdot (\theta_1 \cdot \theta_2) \cdot (\theta_2(1-\theta_1)) \cdot (1-\theta_1)(1-\theta_2)$$

$$= \underset{\theta}{\operatorname{argmax}} \cdot \log \theta_1^4 \cdot (1-\theta_1)^3 \cdot \theta_2^4 (1-\theta_2)^3$$

$$= \underset{\theta}{\operatorname{argmax}} \cdot (4 \log \theta_1 + 3 \log(1-\theta_1) + 4 \log \theta_2 + 3 \log(1-\theta_2))$$

Let call this maximization as $J(\theta)$ and take partial derivatives w.r.t θ_1, θ_2 and set 0

$$\frac{\partial J}{\partial \theta_1} = \frac{4}{\theta_1} - \frac{3}{1-\theta_1} = 0 \Rightarrow 4(1-\theta_1) - 3\theta_1 = 0$$

$$\Rightarrow \theta_1 = 4/7$$

$$\frac{\partial J}{\partial \theta_2} = \frac{4}{\theta_2} - \frac{3}{1-\theta_2} = 0 \Rightarrow \theta_2 = 4/7$$

the marginal likelihood for this data is

$$P(\mathbf{y}|\theta) = \prod_{i=1}^n P(y_i|\theta)$$

$$= \left(\frac{4}{7} \cdot \frac{4}{7}\right) \left(\frac{4}{7} \cdot \frac{3}{7}\right) \left(\frac{4}{7} \cdot \frac{3}{7}\right) \left(\frac{4}{7} \cdot \frac{3}{7}\right) \left(\frac{4}{7} \cdot \frac{4}{7}\right) \left(\frac{4}{7} \cdot \frac{3}{7}\right) \left(\frac{3}{7} \cdot \frac{3}{7}\right)$$

MLE with 4 parameters representing all combinations

$$\hat{\theta} = \underset{\theta}{\operatorname{argmax}} \log \prod_{i=1}^n P(\mathbf{y}_i|\theta) \quad \text{s.t.} \quad \sum_{i=1}^4 \theta_{i,i} = 1$$

$$= \underset{\theta}{\operatorname{argmax}} \log (\theta_{1,1} \cdot \theta_{1,0} \cdot \theta_{0,0} \cdot \theta_{1,0} \cdot \theta_{1,1} \cdot \theta_{0,0} \cdot \theta_{0,1})$$

$$\text{s.t. } \theta_{0,0} + \theta_{1,1} + \theta_{1,0} + \theta_{0,1} = 1$$

$$= \underset{\theta}{\operatorname{argmax}} (2 \log \theta_{1,1} + 2 \log \theta_{1,0} + 2 \log \theta_{0,0} + 1 \log \theta_{0,1})$$

Let call this as $J(\theta)$ and do same as above,

lets write this as using Lagrangian

$$\frac{\partial J}{\partial \theta_{0,0}} = \frac{2}{\theta_{0,0}} \quad \text{multiplier}$$

$$J(\theta) = 2 \log \theta_{1,1} + 2 \log \theta_{1,0} + 2 \log \theta_{0,0} + 1 \log \theta_{0,1} + \lambda (\theta_{0,0} + \theta_{1,1} + \theta_{1,0} + \theta_{0,1} - 1) = 0$$

$$\frac{\partial J}{\partial \theta_{0,0}} = \frac{2}{\theta_{0,0}} - \lambda = 0 \Rightarrow \theta_{0,0} = \frac{2}{\lambda} \rightarrow 0$$

$$\text{Same like this } \theta_{1,0} = \theta_{1,1} = \frac{2}{\lambda}, \theta_{0,1} = \frac{1}{\lambda} \rightarrow 0$$

$$\text{we can solve } \lambda \text{ by using } \sum_{i=1}^4 \theta_{i,i} = 1$$

$$\Rightarrow 0.7/\lambda = 1 \Rightarrow \lambda = 0.7$$

By substituting $\lambda = 7$ in above
 $\theta_{0,0} = \theta_{1,0} = \theta_{1,1} = \frac{2}{7}$ and $\theta_{0,1} = \frac{1}{7}$
 and marginal likelihood with 4-parameters;

$$P(\mathbf{y}/\hat{\theta}, \mathbf{m}_y) = \prod_{i=1}^m p(x_i, y_i/\hat{\theta})$$

$$= \frac{2}{7} \cdot \frac{2}{7} \cdot \frac{2}{7} \cdot \frac{1}{7} = \frac{8}{7^4}$$

d) For leave-one-out cross validation, we have to remove one data point at every time and see how well the new parameters explain remaining data. Let's see first for model 1:

$$L(1) = \sum_{i=1}^m \log p(x_i, y_i / \hat{\theta}_{-i})$$

Here sample size is 7 $\Rightarrow m = 7$

So, for $i=1$, $\hat{\theta}_{-1} = \{\theta_1 = \frac{3}{6}, \theta_2 = \frac{3}{6}\}$ and $p(x_1, y_1 / \theta_{-1}) = \frac{3}{6} \log \frac{3}{6}$
 $= \log \frac{1}{4}$

for $i=2$, $\hat{\theta}_{-2} = \{\theta_1 = \frac{3}{6}, \theta_2 = \frac{4}{6}\}$ and $p(x_2, y_2 / \theta_{-2}) = \log \frac{3}{6} \cdot \frac{4}{6} = \log \frac{1}{3}$

for $i=3$, $\hat{\theta}_{-3} = \{\theta_1 = \frac{4}{6}, \theta_2 = \frac{3}{6}\}$ and $p(x_3, y_3 / \theta_{-3}) = \log \frac{1}{3}$

for $i=4$, $\hat{\theta}_{-4} = \{\theta_1 = \frac{3}{6}, \theta_2 = \frac{4}{6}\}$ and $p(x_4, y_4 / \theta_{-4}) = \log \frac{1}{3}$

for $i=5$, $\hat{\theta}_{-5} = \{\theta_1 = \frac{3}{6}, \theta_2 = \frac{3}{6}\}$ and $p(x_5, y_5 / \theta_{-5}) = \log \frac{1}{4}$

for $i=6$, $\hat{\theta}_{-6} = \{\theta_1 = \frac{4}{6}, \theta_2 = \frac{3}{6}\}$ and $p(x_6, y_6 / \theta_{-6}) = \log \frac{1}{3}$

for $i=1$, $\hat{\theta}_{D-1} = \left[\theta_{1,1} = \frac{1}{6}, \theta_{1,0} = \frac{1}{6} \right]$ and $p(x_1, y_1 / \hat{\theta}_{D-1}) = \log \frac{1}{6} \cdot \frac{1}{6} = \log \frac{1}{9}$

the cross validated likelihood for model M_2

so,

$$L(1) = 2 \log \frac{1}{9} + 4 \log \frac{1}{3} + \log \frac{1}{9}$$

is

lets do same for model 2:

$$L(2) = \sum_{i=1}^m \log p(x_i, y_i / \hat{\theta}_{D-1})$$

for $i=1$, $\hat{\theta}_{D-1} = \left[\theta_{1,1} = \frac{1}{6}, \theta_{1,0} = \frac{2}{6}, \theta_{0,0} = \frac{2}{6}, \theta_{0,1} = \frac{1}{6} \right]$

and $p(x_1, y_1 / \hat{\theta}_{D-1}) = \log \frac{1}{6}$

for $i=2$, $\hat{\theta}_{D-1} = \left[\theta_{1,1} = \frac{2}{6}, \theta_{1,0} = \frac{1}{6}, \theta_{0,0} = \frac{2}{6}, \theta_{0,1} = \frac{1}{6} \right]$

and $p(x_2, y_2 / \hat{\theta}_{D-1}) = \log \frac{1}{6}$

like this for $i=3, 4, 5, 6$ we get $p(x_i, y_i / \hat{\theta}_{D-1}) = \log \frac{1}{6}$

for $i=7$, $\hat{\theta}_{D-1} = \left[\theta_{1,1} = \frac{2}{6}, \theta_{1,0} = \frac{2}{6}, \theta_{0,0} = \frac{2}{6}, \theta_{0,1} = 0 \right]$

and $p(x_7, y_7 / \hat{\theta}_{D-1}) = \log 0$
= undefined.

So, for model M_4 , as we saw training examples in predictions that are not there in training samples, solution is undefined. Then it because in model M_4 we overfit the data and set $\theta_{0,1} = 0$, as we didn't see $x=0, y=1$ in training samples. We'll choose M_2 for this reason.

b)) we have completely random dataset with N_1 examples from class 1, and N_2 from class 2, (with equal proportions i.e. $N_1 = N_2$). Based on this intuition as there is no learning here,

$$P(\hat{y} = \text{class 1}) = \frac{N_1}{N_1 + N_2} = \frac{1}{2}$$

$$P(\hat{y} = \text{class 2}) = \frac{N_2}{N_1 + N_2} = \frac{1}{2}$$

misclassification rate $E(\text{error}) = 1 \cdot P(\hat{y} \neq \text{correct class}) + 0 \cdot P(\hat{y} = \text{correct class})$

The best misclassification rate is $\frac{1}{2}$.

Suppose we use leave one out - cross validation then

for all i ,
if i th ~~class~~^{example} belongs to class 1, then prediction for

$$P(\hat{y} = \text{class 1} / i \in \text{class 1}) = \frac{N_1 - 1}{N - 1}$$

and for class 2 $\Rightarrow P(\hat{y} = \text{class 2} / i \in \text{class 2}) = \frac{N_2 - 1}{N_1 + N_2 - 1}$

and suppose if i th class belongs to class 2,

then

$$P(\hat{y} = \text{class 1} / i \in \text{class 2}) = \frac{N_1}{N - 1}$$

$$\& P(\hat{y} = \text{class 2} / i \in \text{class 2}) = \frac{N_2 - 1}{N - 1}$$

Estimated misclassification rate

ith example is $p(\hat{y}/i \in \text{class 1}) = \frac{N_1 - 1}{N - 1}$

in same way, if ith example belongs to class 2 then

$$p(\hat{y}/i \in \text{class 2}) = \frac{N_2 - 1}{N - 1}$$

then total misclassification rate can be

$$= \frac{N_2}{N} \left(1 - \frac{N_2 - 1}{N - 1} \right) + \frac{N_1}{N} \left(1 - \frac{N_1 - 1}{N - 1} \right)$$

$$= \frac{N_2 \cdot (N_1 - 1)}{N(N - 1)} + \frac{N_1 \cdot (N_2 - 1)}{N(N - 1)}$$

$$= \frac{2N_1N_2 - N}{N(N - 1)}$$

b2) $p(y_i/\theta_i) \sim N(\theta_i, \sigma^2)$ and $p(\theta) \sim N(m_0, T_0^{-1})$
 $\eta = (m_0, T_0^{-1})$

σ^2 is given 2500

ML-II estimate for m_0, T_0^{-1} is

$$\hat{\eta} = \underset{\eta}{\operatorname{argmax}} \int p(y/\theta) \cdot p(\theta) d\theta$$

$$= \underset{\eta}{\operatorname{argmax}} \int \cdot e^{-\frac{1}{2} \sum_{i=1}^n \frac{(y_i - \theta)^2}{\sigma^2}} \cdot e^{-\frac{1}{2} \cdot \left(\frac{\theta - m_0}{T_0^{-1}} \right)^2} d\theta$$

$$= \underset{\eta}{\operatorname{argmax}} \int \cdot e^{-\frac{1}{2}}$$

i.e. find out the quantity from given data

$$\sum_{i=1}^n (y_i - 0)^2 = \frac{1}{500} ((150.5 - 0)^2 + (150.5 - 0)^2 + \dots)$$

Based on equations from below

$$p(\theta|y) = N(\bar{y}/\sigma, \sigma^2/n)$$

$$\text{and } p(\theta) = N(0/m_0, \tau_0^2) \text{ then}$$

$$p(\theta|y) = N(0/m_N, \tau_N^2)$$

$$m_N = \tau_N^2 \left(\frac{\bar{y}}{\tau_0^2} + \frac{m_0}{\tau_0^2} \right)$$

$$\frac{1}{\tau_N^2} = \frac{1}{\tau_0^2} + \frac{n}{\sigma^2}$$

$$\text{and } m_N = \tau_N^2 \left(\frac{n\bar{y}}{\sigma^2} + \frac{m_0}{\tau_0^2} \right)$$

As we are trying to minimize marginal likelihood

$$p(y) = \int p(y|\theta) \cdot p(\theta) d\theta$$

$$= \int N(y/\theta, \sigma^2) \cdot N(0/m_N, \tau_N^2) d\theta$$

$$= \sqrt{2\pi\tau_N^2}$$

$$\text{So } \arg\max_{\tau_N^2} p(y|\theta) = \arg\max_{\tau_N^2} \sqrt{2\pi\tau_N^2}$$

$$= \arg\max_{\tau_N^2} \sqrt{2\pi} \cdot \left(\frac{1}{\tau_0^2} + \frac{n}{\sigma^2} \right)^{1/2}$$