

Multiple linear regression

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Grading the professor

Many college courses conclude by giving students the opportunity to evaluate the course and the instructor anonymously. However, the use of these student evaluations as an indicator of course quality and teaching effectiveness is often criticized because these measures may reflect the influence of non-teaching related characteristics, such as the physical appearance of the instructor. The article titled, “Beauty in the classroom: instructors’ pulchritude and putative pedagogical productivity” (Hamermesh and Parker, 2005) found that instructors who are viewed to be better looking receive higher instructional ratings. (Daniel S. Hamermesh, Amy Parker, Beauty in the classroom: instructors pulchritude and putative pedagogical productivity, *Economics of Education Review*, Volume 24, Issue 4, August 2005, Pages 369-376, ISSN 0272-7757, 10.1016/j.econedurev.2004.07.013. <http://www.sciencedirect.com/science/article/pii/S0272775704001165>.)

In this lab we will analyze the data from this study in order to learn what goes into a positive professor evaluation.

The data

The data were gathered from end of semester student evaluations for a large sample of professors from the University of Texas at Austin. In addition, six students rated the professors’ physical appearance. (This is a slightly modified version of the original data set that was released as part of the replication data for *Data Analysis Using Regression and Multilevel/Hierarchical Models* (Gelman and Hill, 2007).) The result is a data frame where each row contains a different course and columns represent variables about the courses and professors.

```
load("more/evals.RData")
```

variable	description
score	average professor evaluation score: (1) very unsatisfactory - (5) excellent.
rank	rank of professor: teaching, tenure track, tenured.
ethnicity	ethnicity of professor: not minority, minority.
gender	gender of professor: female, male.
language	language of school where professor received education: english or non-english.
age	age of professor.
cls_perc_eval	percent of students in class who completed evaluation.
cls_did_eval	number of students in class who completed evaluation.

variable	description
<code>cls_students</code>	total number of students in class.
<code>cls_level</code>	class level: lower, upper.
<code>cls_profs</code>	number of professors teaching sections in course in sample: single, multiple.
<code>cls_credits</code>	number of credits of class: one credit (lab, PE, etc.), multi credit.
<code>bty_f1lower</code>	beauty rating of professor from lower level female: (1) lowest - (10) highest.
<code>bty_f1upper</code>	beauty rating of professor from upper level female: (1) lowest - (10) highest.
<code>bty_f2upper</code>	beauty rating of professor from second upper level female: (1) lowest - (10) highest.
<code>bty_m1lower</code>	beauty rating of professor from lower level male: (1) lowest - (10) highest.
<code>bty_m1upper</code>	beauty rating of professor from upper level male: (1) lowest - (10) highest.
<code>bty_m2upper</code>	beauty rating of professor from second upper level male: (1) lowest - (10) highest.
<code>bty_avg</code>	average beauty rating of professor.
<code>pic_outfit</code>	outfit of professor in picture: not formal, formal.
<code>pic_color</code>	color of professor's picture: color, black & white.

Exploring the data

1. Is this an observational study or an experiment? The original research question posed in the paper is whether beauty leads directly to the differences in course evaluations. Given the study design, is it possible to answer this question as it is phrased? If not, rephrase the question.

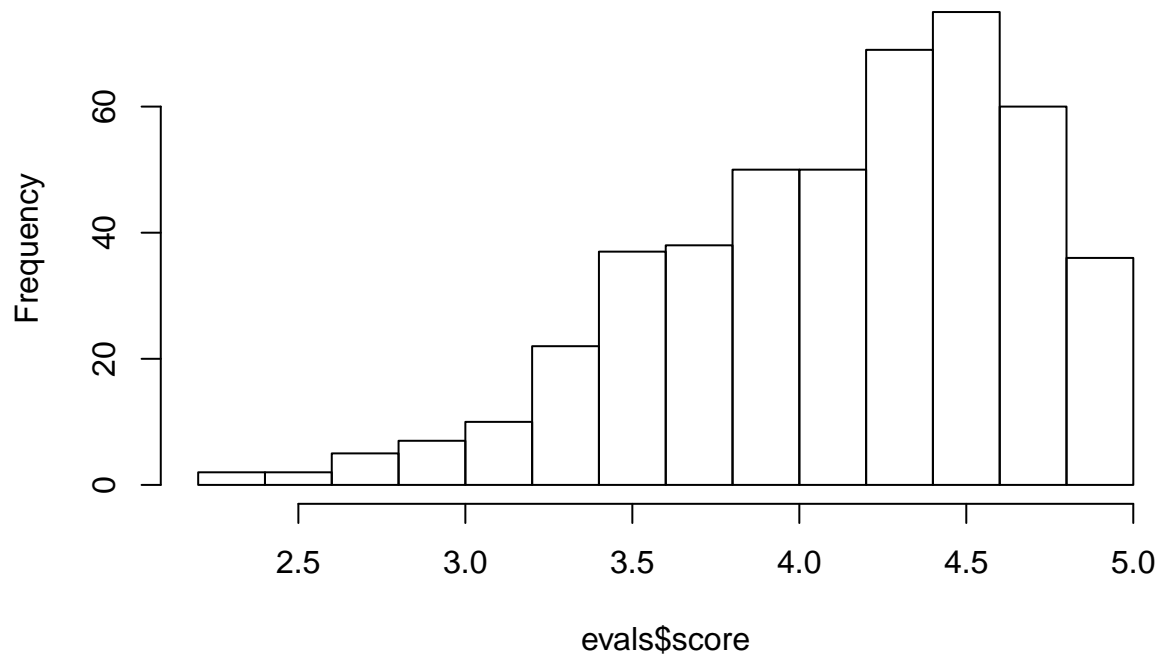
Answer: It is an observational study that doesn't have any controlled or experimented group in this study. Because of it is an observational study, we cannot determine the causation relationship in this study, but instead we can see correlation between the dependent and independent variables.

2. Describe the distribution of `score`. Is the distribution skewed? What does that tell you about how students rate courses? Is this what you expected to see? Why, or why not?

Answer:

```
hist(evals$score)
```

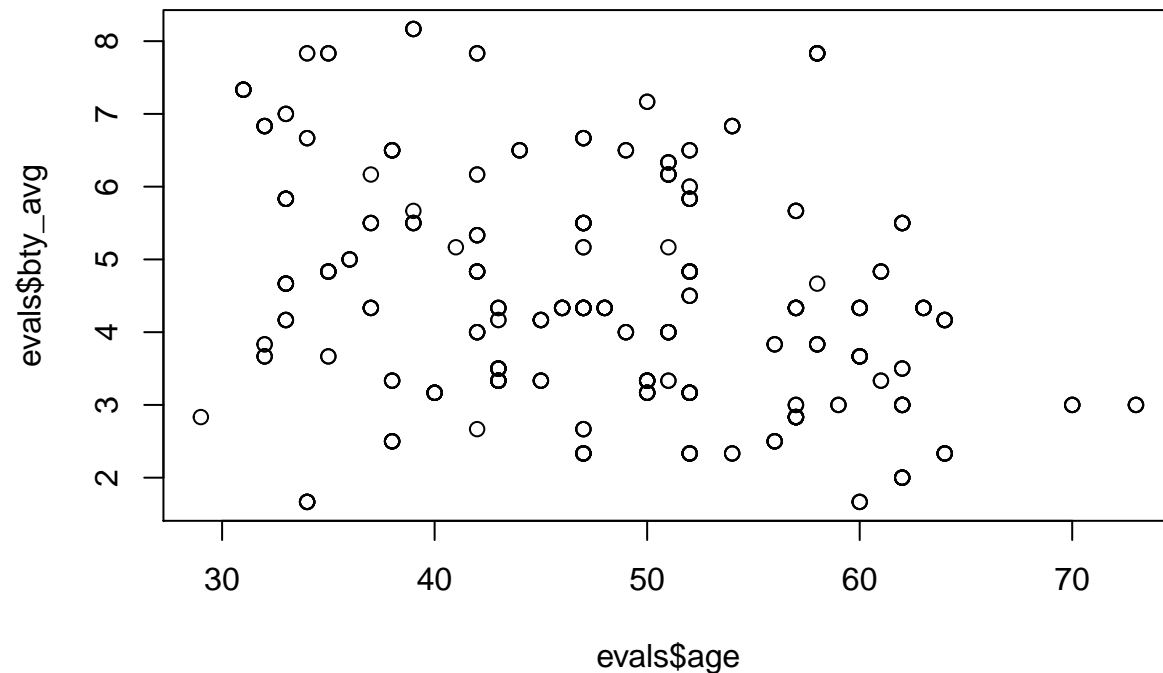
Histogram of evals\$score



It is left skewed according to the plot which means most of the students rated the class with high score. I expected exactly what the histogram shows because I believe most professors are trying their best for their classes and students and manage the class efficiently.

3. Excluding `score`, select two other variables and describe their relationship using an appropriate visualization (scatterplot, side-by-side boxplots, or mosaic plot).

```
plot(eval$btty_avg ~ eval$age)
```



An-

swer: According to the scatterplot, I don't see any relationship between age and beauty score.

Simple linear regression

The fundamental phenomenon suggested by the study is that better looking teachers are evaluated more favorably. Let's create a scatterplot to see if this appears to be the case:

```
plot(evals$score ~ evals$bty_avg)
```

Before we draw conclusions about the trend, compare the number of observations in the data frame with the approximate number of points on the scatterplot. Is anything awry?

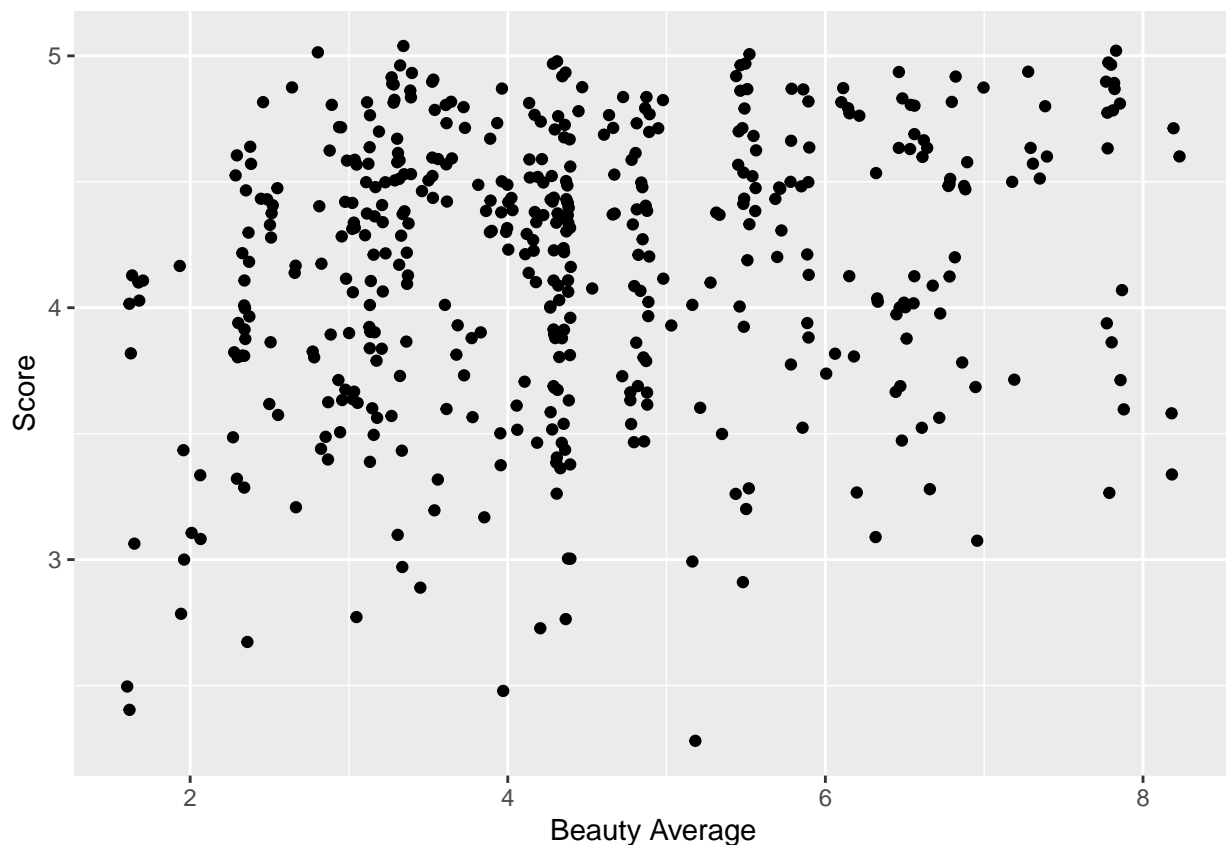
```
nrow(evals)
```

```
## [1] 463
```

Answer: It seems there are more observations in the data compared to the plot above.

4. Replot the scatterplot, but this time use the function `jitter()` on the y - or the x -coordinate. (Use `?jitter` to learn more.) What was misleading about the initial scatterplot?

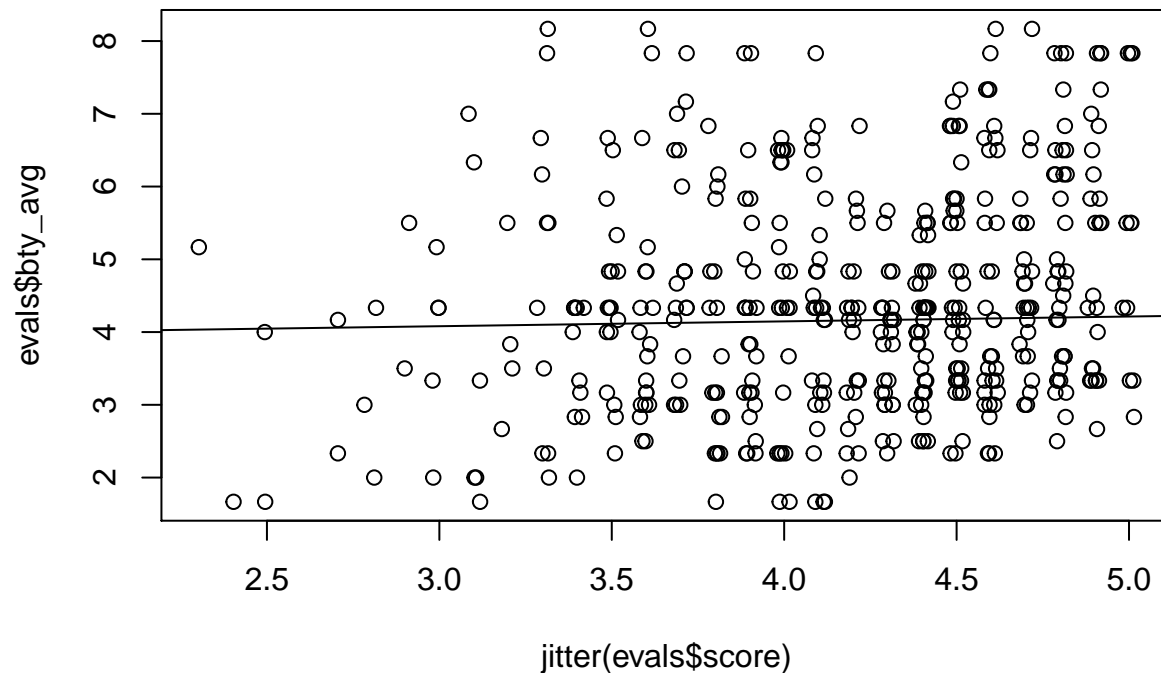
```
library(ggplot2)
g<-ggplot(evals,aes(x=bty_avg,y=score))
g<-g+geom_point(position = position_jitter())
g<-g+xlab("Beauty Average")+ylab("Score")
g
```



Answer: Eventhough use jitter for the scatter plot, still doesn't show relationship between Beauty Average and Score, because jitter helps to see the overlapped points, but seems doesn't help to see the relationships.

- Let's see if the apparent trend in the plot is something more than natural variation. Fit a linear model called `m_bty` to predict average professor score by average beauty rating and add the line to your plot using `abline(m_bty)`. Write out the equation for the linear model and interpret the slope. Is average beauty score a statistically significant predictor? Does it appear to be a practically significant predictor?

```
m_bty <- lm(evals$score~evals$bty_avg)
plot(evals$bty_avg~jitter(evals$score))
abline(m_bty)
```



```
summary(m_bty)
```

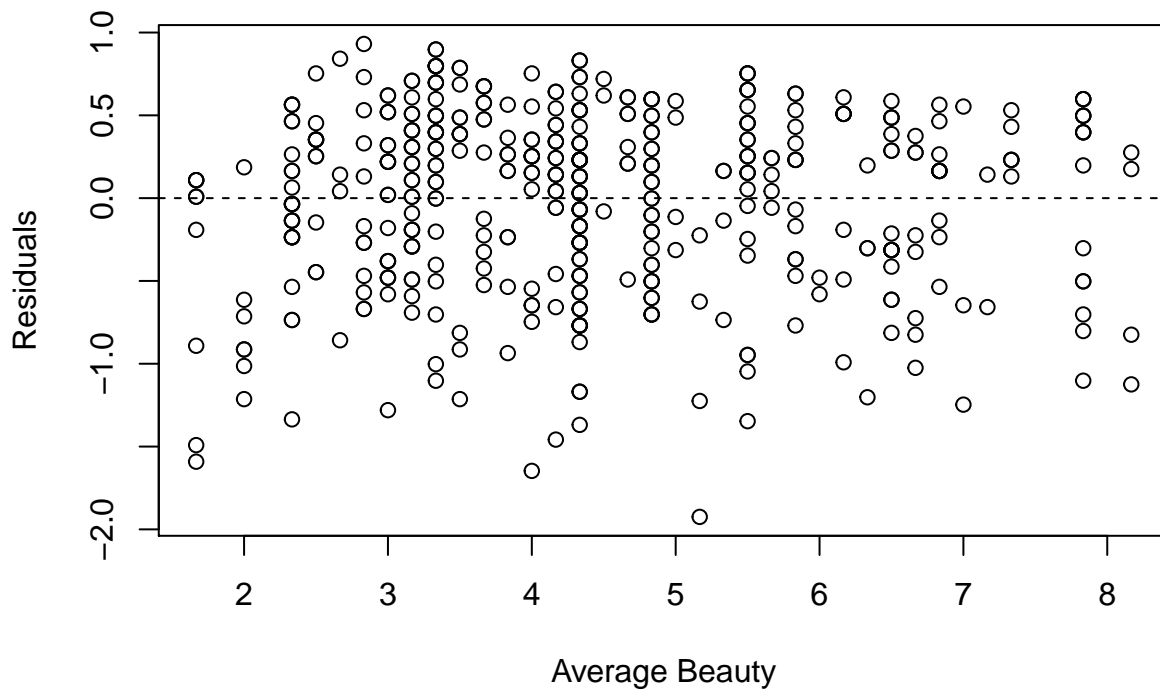
```
##
## Call:
## lm(formula = evals$score ~ evals$bty_avg)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.9246 -0.3690  0.1420  0.3977  0.9309
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   3.88034    0.07614   50.96 < 2e-16 ***
## evals$bty_avg  0.06664    0.01629    4.09 5.08e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5348 on 461 degrees of freedom
```

```
## Multiple R-squared:  0.03502,    Adjusted R-squared:  0.03293
## F-statistic: 16.73 on 1 and 461 DF,  p-value: 5.083e-05
```

Answer: The summary shows it is statistically significant predictor.

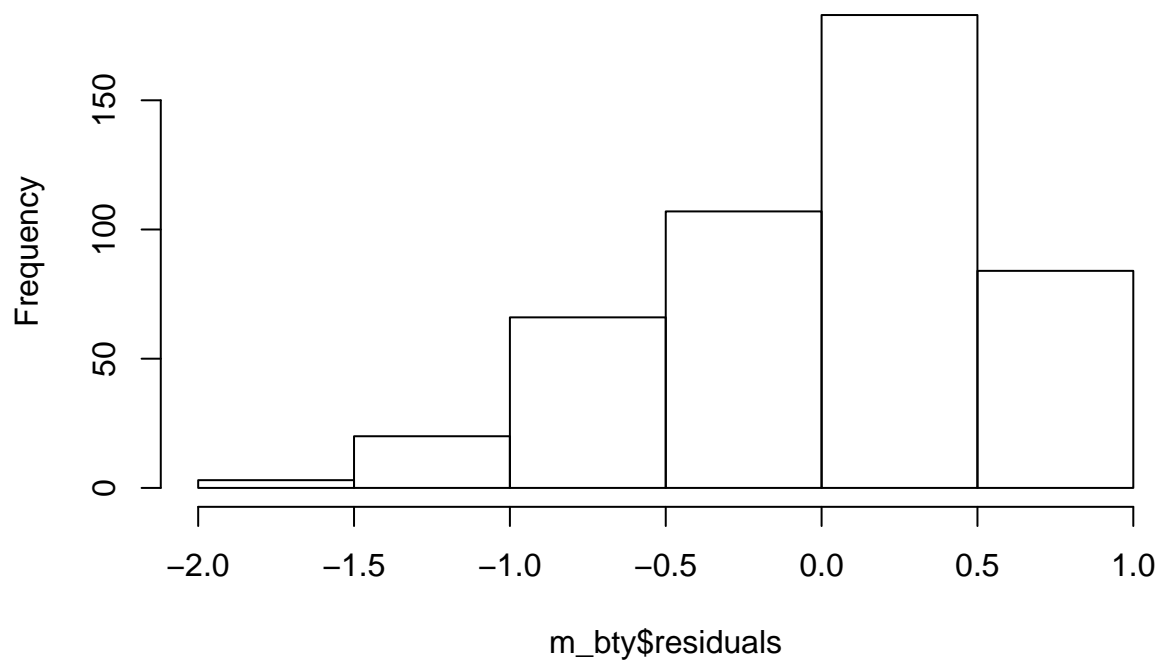
6. Use residual plots to evaluate whether the conditions of least squares regression are reasonable. Provide plots and comments for each one (see the Simple Regression Lab for a reminder of how to make these).

```
plot(m_bty$residuals ~ evals$bty_avg, ylab="Residuals", xlab="Average Beauty")
abline(h = 0, lty = 2)
```



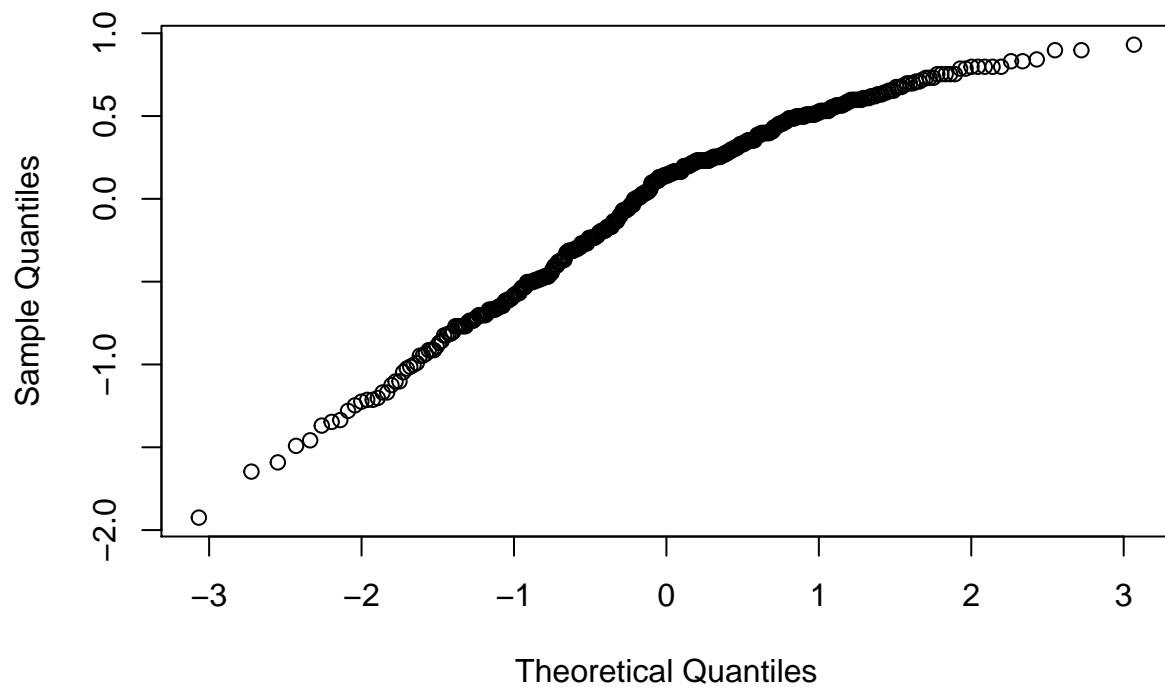
```
hist(m_bty$residuals)
```

Histogram of m_bty\$residuals



```
qqnorm(m_bty$residuals)
```

Normal Q-Q Plot



Answer: residual models are usually normally distributed, but the histogram shows it is slightly skewed left.

Multiple linear regression

The data set contains several variables on the beauty score of the professor: individual ratings from each of the six students who were asked to score the physical appearance of the professors and the average of these six scores. Let's take a look at the relationship between one of these scores and the average beauty score.

```
plot(evals$btty_avg ~ evals$btty_follower)
cor(evals$btty_avg, evals$btty_follower)
```

As expected the relationship is quite strong - after all, the average score is calculated using the individual scores. We can actually take a look at the relationships between all beauty variables (columns 13 through 19) using the following command:

```
plot(evals[,13:19])
```

These variables are collinear (correlated), and adding more than one of these variables to the model would not add much value to the model. In this application and with these highly-correlated predictors, it is reasonable to use the average beauty score as the single representative of these variables.

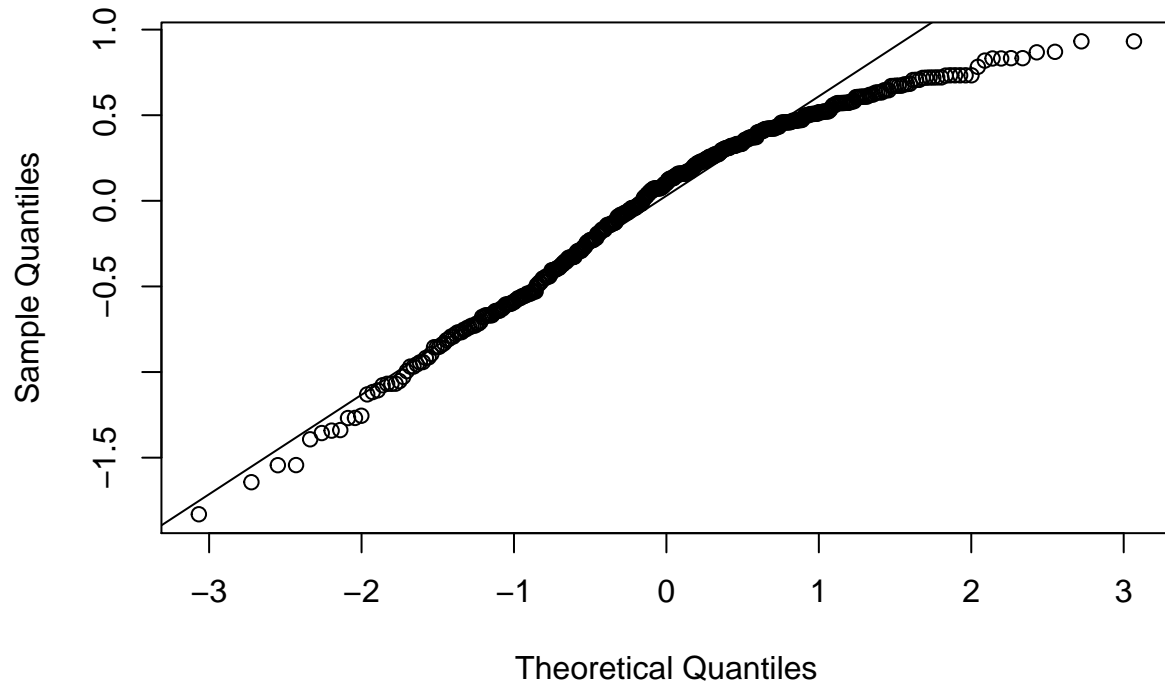
In order to see if beauty is still a significant predictor of professor score after we've accounted for the gender of the professor, we can add the gender term into the model.

```
m_bty_gen <- lm(score ~ btty_avg + gender, data = evals)
summary(m_bty_gen)
```

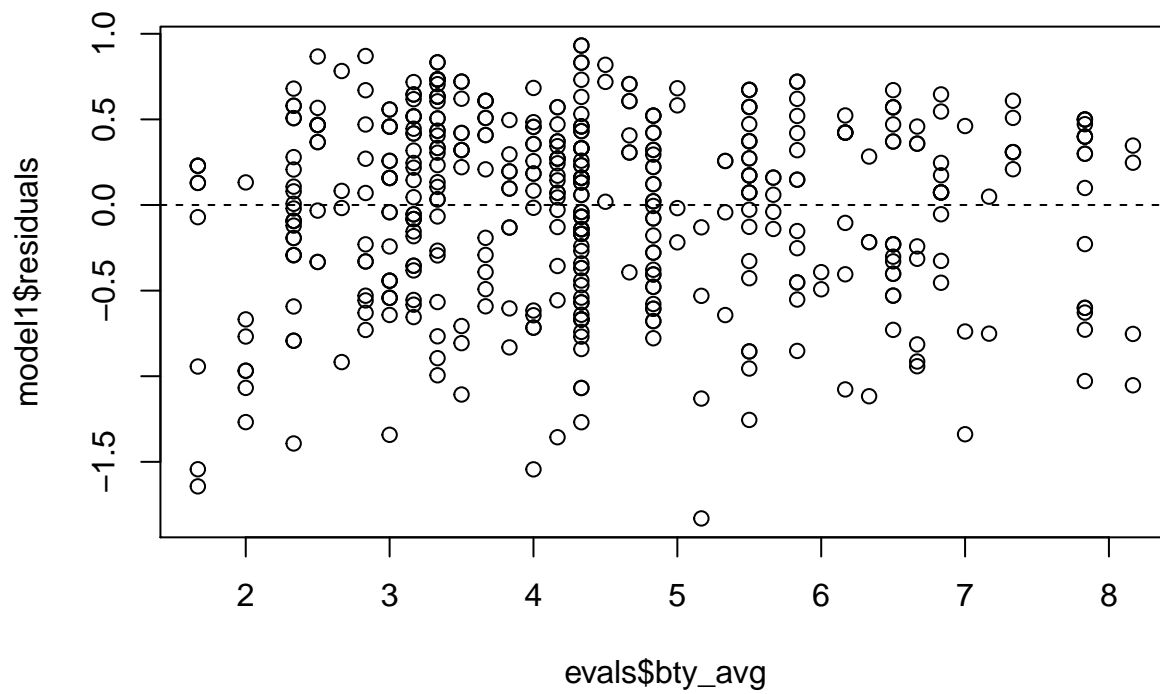
7. P-values and parameter estimates should only be trusted if the conditions for the regression are reasonable. Verify that the conditions for this model are reasonable using diagnostic plots.

```
library(ggplot2)
model1 <- lm(score ~ btty_avg + gender, data = evals)
qqnorm(model1$residuals)
qqline(model1$residuals)
```


Normal Q-Q Plot



```
plot(model1$residuals ~ evals$bty_avg)
abline(h = 0, lty = 2)
```

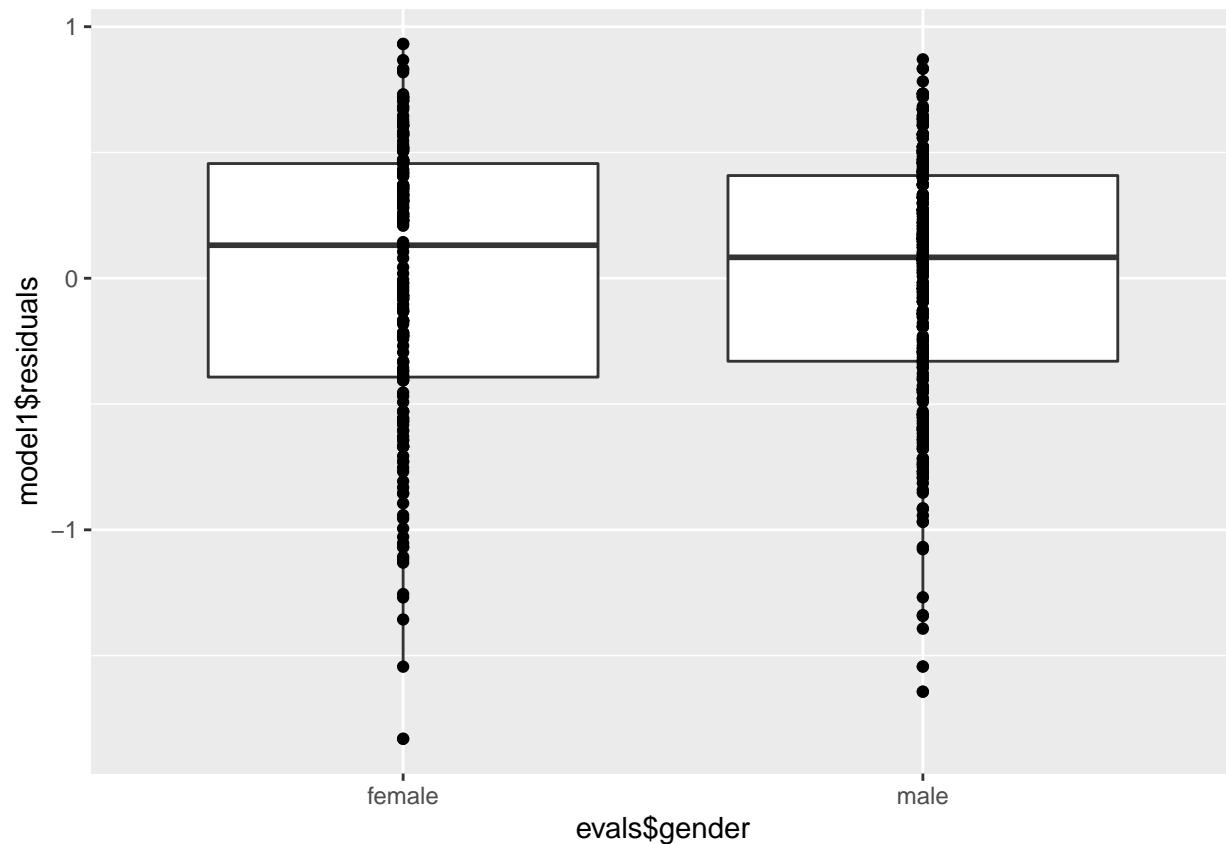


```
g<-ggplot(evals,aes(y=model1$residuals,x=evals$gender))
g<-g+geom_boxplot()
g<-g+geom_point()
```

```
g
```

```
## Warning: Use of `evals$gender` is discouraged. Use `gender` instead.
```

```
## Warning: Use of `evals$gender` is discouraged. Use `gender` instead.
```



8. Is `bty_avg` still a significant predictor of `score`? Has the addition of `gender` to the model changed the parameter estimate for `bty_avg`?

```
summary(model1)
```

```
##
## Call:
## lm(formula = score ~ bty_avg + gender, data = evals)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.8305 -0.3625  0.1055  0.4213  0.9314
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   3.74734    0.08466  44.266  < 2e-16 ***
## bty_avg        0.07416    0.01625   4.563 6.48e-06 ***
## gendermale     0.17239    0.05022   3.433 0.000652 ***
```

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5287 on 460 degrees of freedom
## Multiple R-squared:  0.05912,    Adjusted R-squared:  0.05503
## F-statistic: 14.45 on 2 and 460 DF,  p-value: 8.177e-07
```

```
summary(m_bty)
```

```
##
## Call:
## lm(formula = evals$score ~ evals$bty_avg)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.9246 -0.3690  0.1420  0.3977  0.9309
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   3.88034    0.07614   50.96 < 2e-16 ***
## evals$bty_avg  0.06664    0.01629    4.09 5.08e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5348 on 461 degrees of freedom
## Multiple R-squared:  0.03502,    Adjusted R-squared:  0.03293
## F-statistic: 16.73 on 1 and 461 DF,  p-value: 5.083e-05
```

Answer: yes, it is still significant, and slope and intercept increased.

Note that the estimate for **gender** is now called **gendermale**. You'll see this name change whenever you introduce a categorical variable. The reason is that R recodes **gender** from having the values of **female** and **male** to being an indicator variable called **gendermale** that takes a value of 0 for females and a value of 1 for males. (Such variables are often referred to as “dummy” variables.)

As a result, for females, the parameter estimate is multiplied by zero, leaving the intercept and slope form familiar from simple regression.

$$\begin{aligned}\widehat{score} &= \hat{\beta}_0 + \hat{\beta}_1 \times bty_avg + \hat{\beta}_2 \times (0) \\ &= \hat{\beta}_0 + \hat{\beta}_1 \times bty_avg\end{aligned}$$

We can plot this line and the line corresponding to males with the following custom function.

```
multiLines(m_bty_gen)
```

9. What is the equation of the line corresponding to males? (*Hint:* For males, the parameter estimate is multiplied by 1.) For two professors who received the same beauty rating, which gender tends to have the higher course evaluation score?

Answer: $score = 3.74734 + 0.17239 \times \text{beauty average}$ males tend to have a higher rating according to the blue line.

The decision to call the indicator variable **gendermale** instead of **genderfemale** has no deeper meaning. R simply codes the category that comes first alphabetically as a 0. (You can change the reference level of a

categorical variable, which is the level that is coded as a 0, using the `relevel` function. Use `?relevel` to learn more.)

10. Create a new model called `m_bty_rank` with `gender` removed and `rank` added in. How does R appear to handle categorical variables that have more than two levels? Note that the `rank` variable has three levels: `teaching`, `tenure track`, `tenured`.

```
m_bty_rank <- lm(score ~ bty_avg + rank, data = evals)
summary(m_bty_rank)

##
## Call:
## lm(formula = score ~ bty_avg + rank, data = evals)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.8713 -0.3642  0.1489  0.4103  0.9525
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    3.98155    0.09078  43.860 < 2e-16 ***
## bty_avg         0.06783    0.01655   4.098 4.92e-05 ***
## ranktenure track -0.16070    0.07395  -2.173  0.0303 *
## ranktenured     -0.12623    0.06266  -2.014  0.0445 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.5328 on 459 degrees of freedom
## Multiple R-squared:  0.04652,    Adjusted R-squared:  0.04029
## F-statistic: 7.465 on 3 and 459 DF,  p-value: 6.88e-05
```

The interpretation of the coefficients in multiple regression is slightly different from that of simple regression. The estimate for `bty_avg` reflects how much higher a group of professors is expected to score if they have a beauty rating that is one point higher *while holding all other variables constant*. In this case, that translates into considering only professors of the same rank with `bty_avg` scores that are one point apart.

The search for the best model

We will start with a full model that predicts professor score based on rank, ethnicity, gender, language of the university where they got their degree, age, proportion of students that filled out evaluations, class size, course level, number of professors, number of credits, average beauty rating, outfit, and picture color.

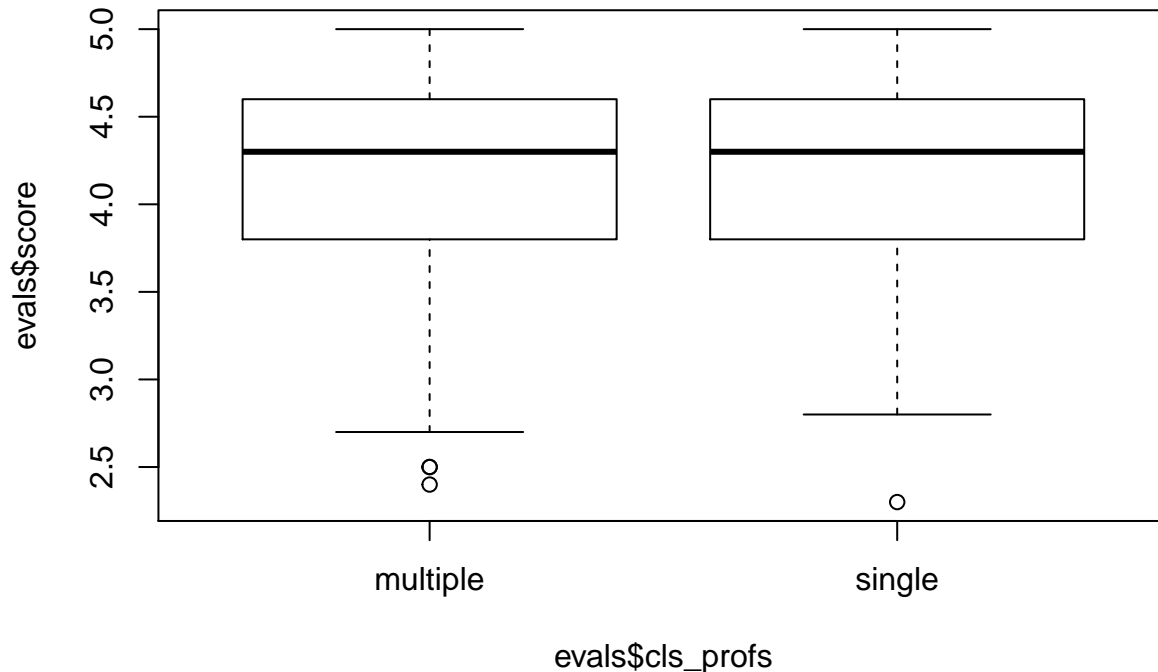
11. Which variable would you expect to have the highest p-value in this model? Why? *Hint:* Think about which variable would you expect to not have any association with the professor score.

Answer: I guess `cls_profs` as the variable to have the weakest association with the evaluation score. Let's run the model...

```
m_full <- lm(score ~ rank + ethnicity + gender + language + age + cls_perc_eval
             + cls_students + cls_level + cls_profs + cls_credits + bty_avg
             + pic_outfit + pic_color, data = evals)
summary(m_full)
```

12. Check your suspicions from the previous exercise. Include the model output in your response. Answer: the `cls_profs` has the least association to “scores” since it has the highest p-value in the model.

```
plot(evals$score ~ evals$cls_profs)
```



13. Interpret the coefficient associated with the ethnicity variable.

Answer: The p-value of about 0.11 of ethnicity means that it has a weak relationship to scores.

14. Drop the variable with the highest p-value and re-fit the model. Did the coefficients and significance of the other explanatory variables change? (One of the things that makes multiple regression interesting is that coefficient estimates depend on the other variables that are included in the model.) If not, what does this say about whether or not the dropped variable was collinear with the other explanatory variables?

```
model_drop <- lm(score ~ rank + ethnicity + gender + language + age + cls_perc_eval +  
  cls_students + cls_level + cls_credits + bty_avg + pic_outfit + pic_color,  
  data = evals)  
summary(model_drop)
```

```
##  
## Call:  
## lm(formula = score ~ rank + ethnicity + gender + language + age +  
##     cls_perc_eval + cls_students + cls_level + cls_credits +  
##     bty_avg + pic_outfit + pic_color, data = evals)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max   
## -1.7836 -0.3257  0.0859  0.3513  0.9551   
##
```

```
## Coefficients:
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept)      4.0872523   0.2888562   14.150 < 2e-16 ***
## ranktenure track  -0.1476746   0.0819824   -1.801 0.072327 .
## ranktenured       -0.0973829   0.0662614   -1.470 0.142349
## ethnicitynot minority 0.1274458   0.0772887    1.649 0.099856 .
## gendermale        0.2101231   0.0516873    4.065 5.66e-05 ***
## languagenon-english -0.2282894   0.1111305   -2.054 0.040530 *
## age               -0.0089992   0.0031326   -2.873 0.004262 **
## cls_perc_eval      0.0052888   0.0015317    3.453 0.000607 ***
## cls_students       0.0004687   0.0003737    1.254 0.210384
## cls_levelupper     0.0606374   0.0575010    1.055 0.292200
## cls_creditsone credit 0.5061196   0.1149163    4.404 1.33e-05 ***
## bty_avg           0.0398629   0.0174780    2.281 0.023032 *
## pic_outfitnot formal -0.1083227   0.0721711   -1.501 0.134080
## pic_colorcolor     -0.2190527   0.0711469   -3.079 0.002205 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4974 on 449 degrees of freedom
## Multiple R-squared:  0.187, Adjusted R-squared:  0.1634
## F-statistic: 7.943 on 13 and 449 DF, p-value: 2.336e-14
```

Answer: after cls_profs removed from the model, a slight change in the coefficients and significance.

15. Using backward-selection and p-value as the selection criterion, determine the best model. You do not need to show all steps in your answer, just the output for the final model. Also, write out the linear model for predicting score based on the final model you settle on.

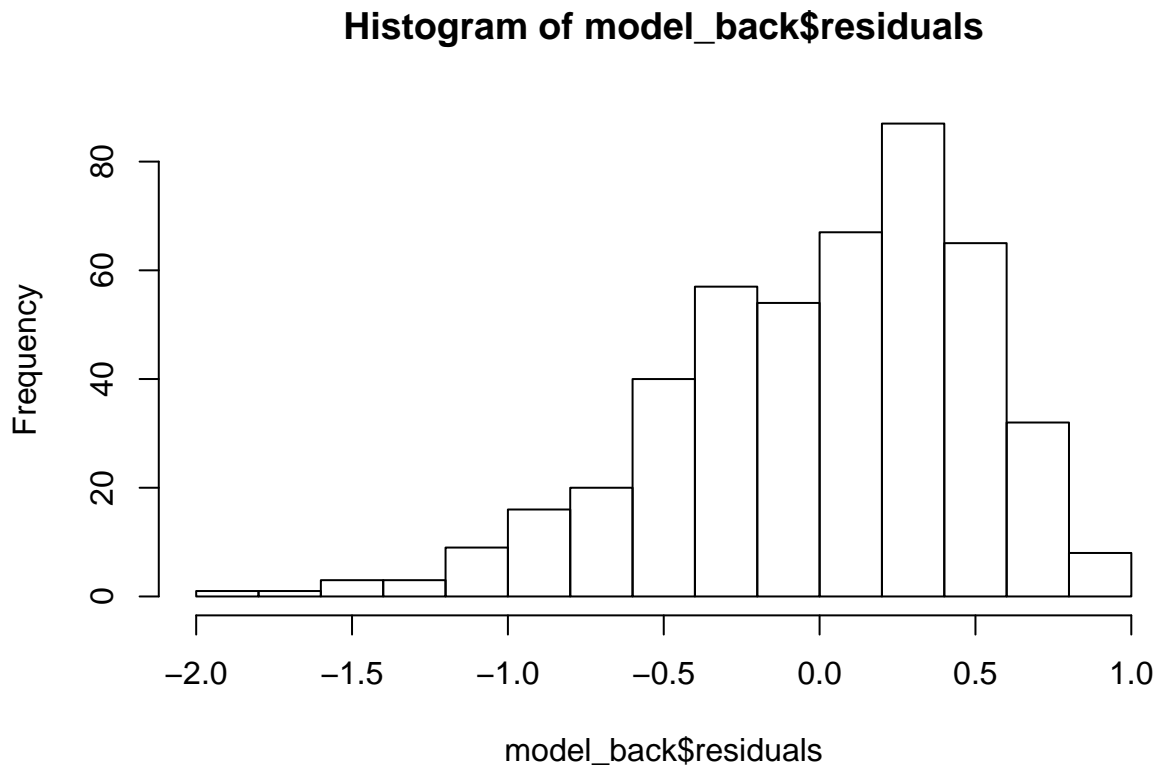
```
model_back <- lm(score ~ ethnicity + gender + language + age + cls_perc_eval +
  cls_credits + bty_avg + pic_color, data = evals)
summary(model_back)
```

```
##
## Call:
## lm(formula = score ~ ethnicity + gender + language + age + cls_perc_eval +
##     cls_credits + bty_avg + pic_color, data = evals)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.85320 -0.32394  0.09984  0.37930  0.93610
##
## Coefficients:
##               Estimate Std. Error t value Pr(>|t|)
## (Intercept)      3.771922   0.232053  16.255 < 2e-16 ***
## ethnicitynot minority 0.167872   0.075275   2.230 0.02623 *
## gendermale        0.207112   0.050135   4.131 4.30e-05 ***
## languagenon-english -0.206178   0.103639   -1.989 0.04726 *
## age               -0.006046   0.002612   -2.315 0.02108 *
## cls_perc_eval      0.004656   0.001435    3.244 0.00127 **
## cls_creditsone credit 0.505306   0.104119    4.853 1.67e-06 ***
## bty_avg           0.051069   0.016934    3.016 0.00271 **
## pic_colorcolor     -0.190579   0.067351   -2.830 0.00487 **
```

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.4992 on 454 degrees of freedom
## Multiple R-squared:  0.1722, Adjusted R-squared:  0.1576
## F-statistic: 11.8 on 8 and 454 DF, p-value: 2.58e-15
```

16. Verify that the conditions for this model are reasonable using diagnostic plots.

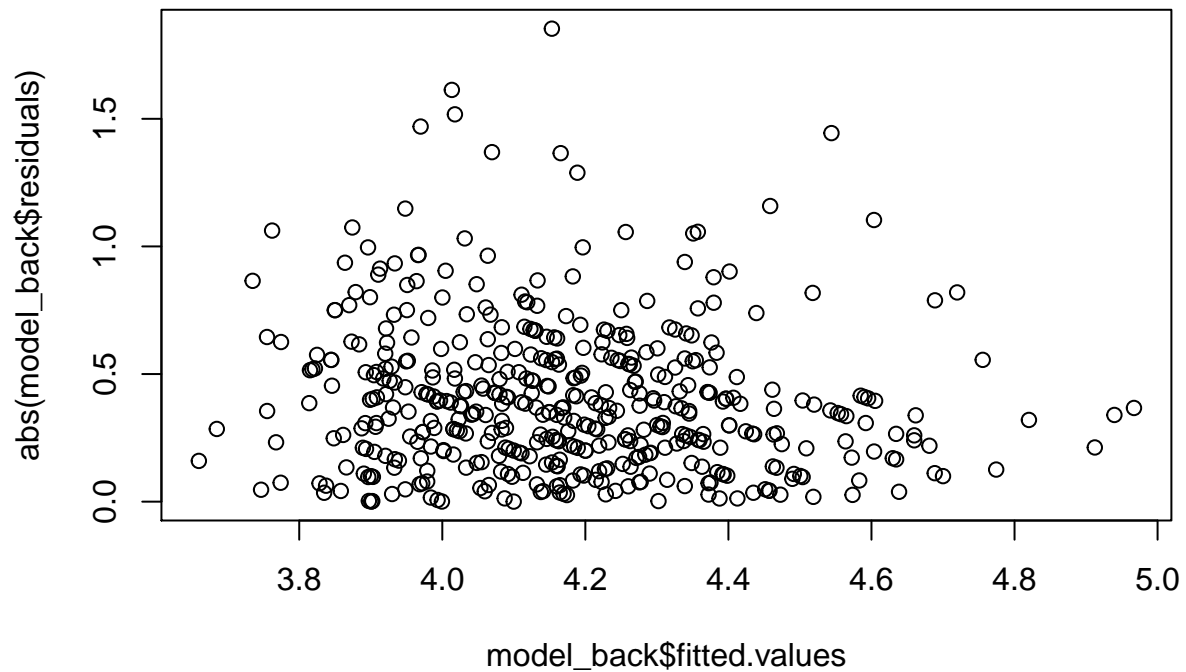
```
hist(model_back$residuals)
```



residuals of model_back is not normal.

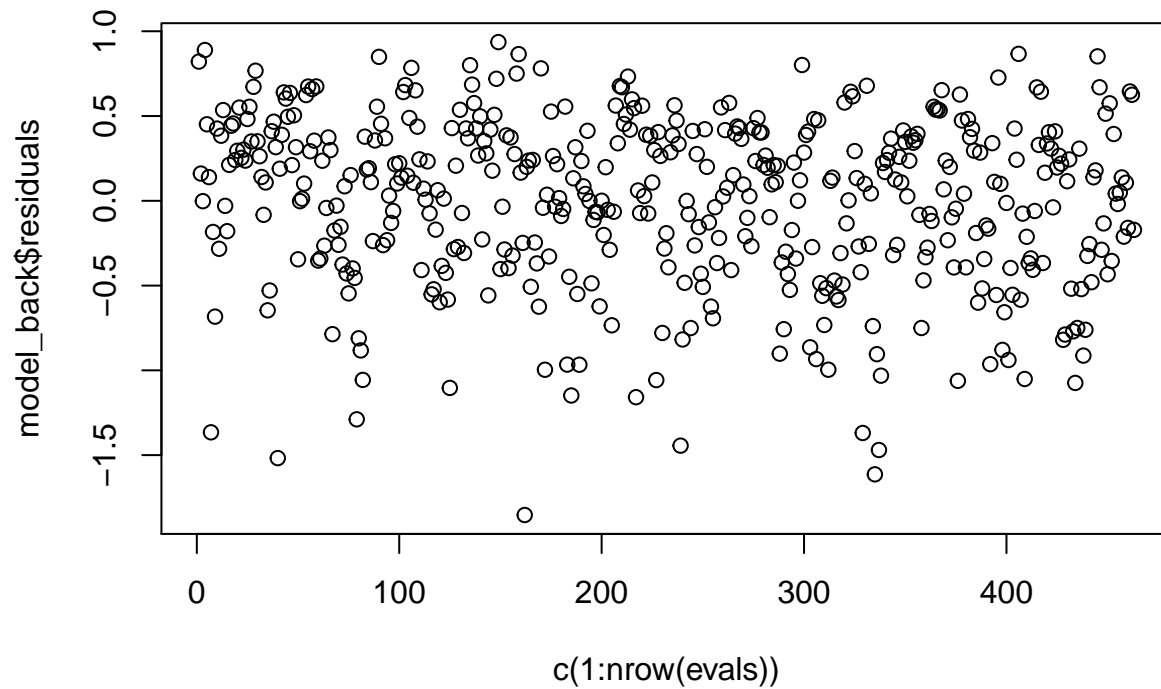
The

```
plot(abs(model_back$residuals) ~ model_back$fitted.values)
```



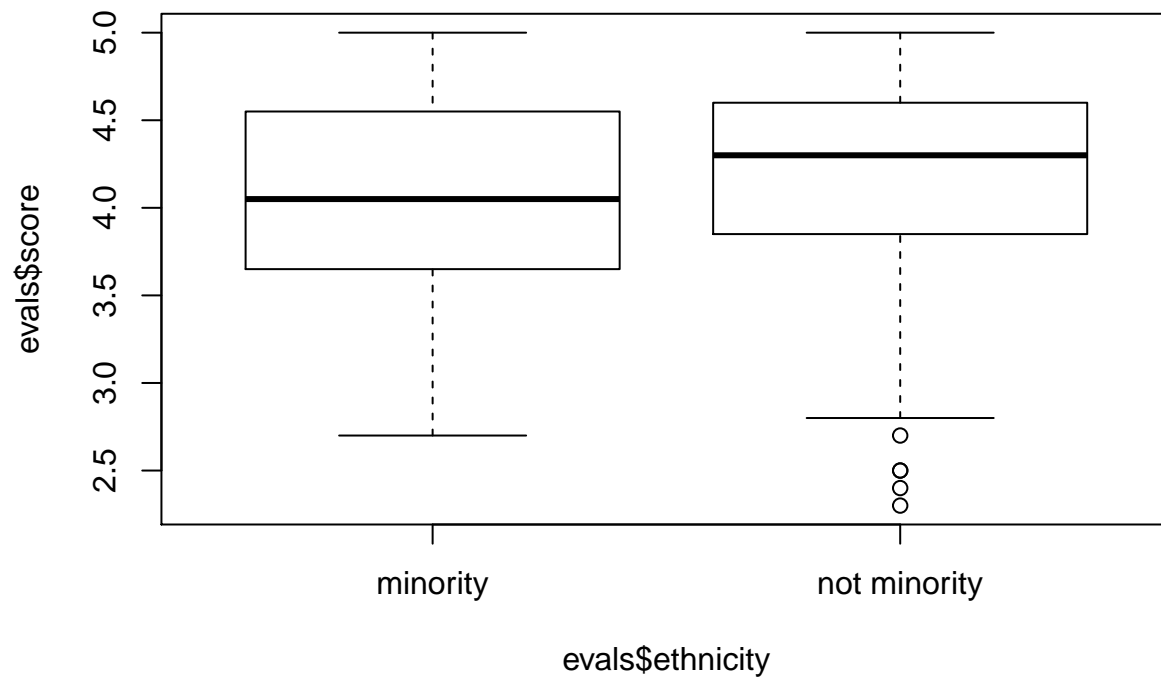
Most of the residual values are close to the fitted values.

```
plot(model_back$residuals ~ c(1:nrow(evals)))
```

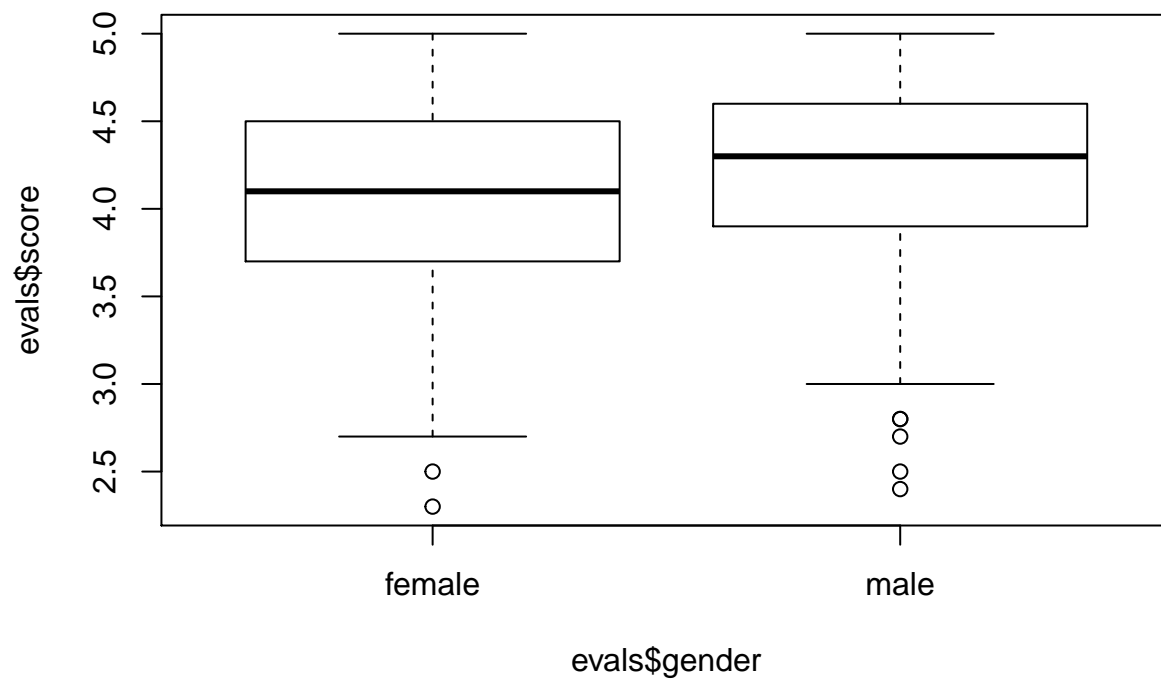


The residuals shows that they were randomly gathered.

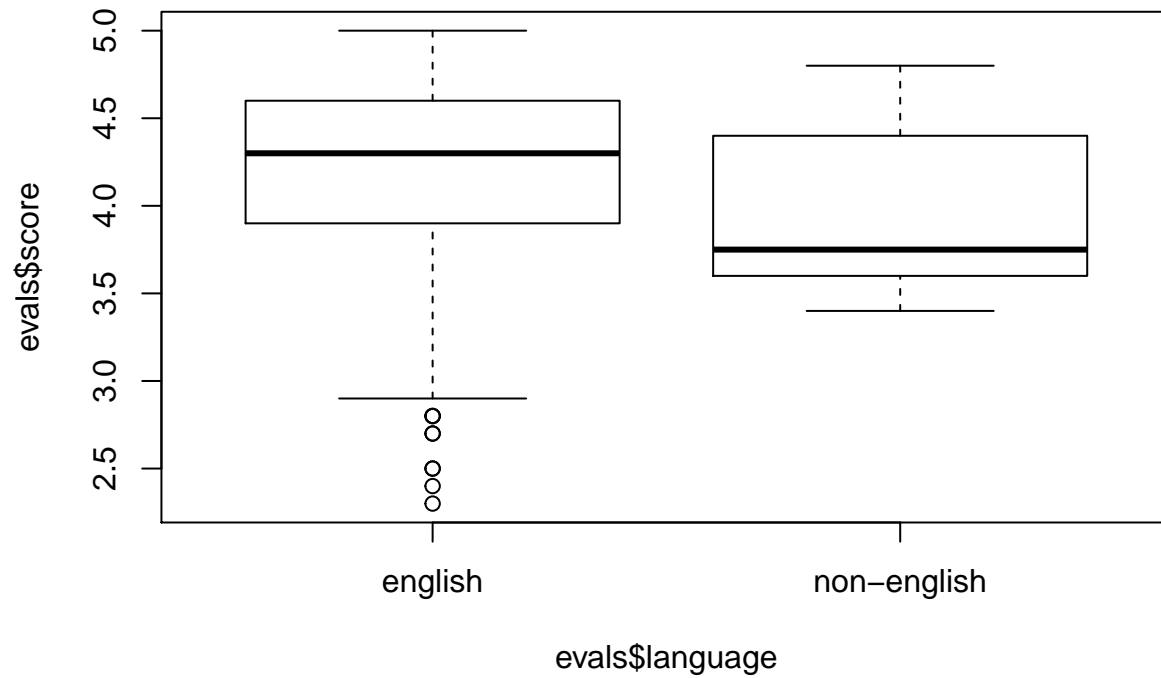
```
plot(evals$score ~ evals$ethnicity)
```

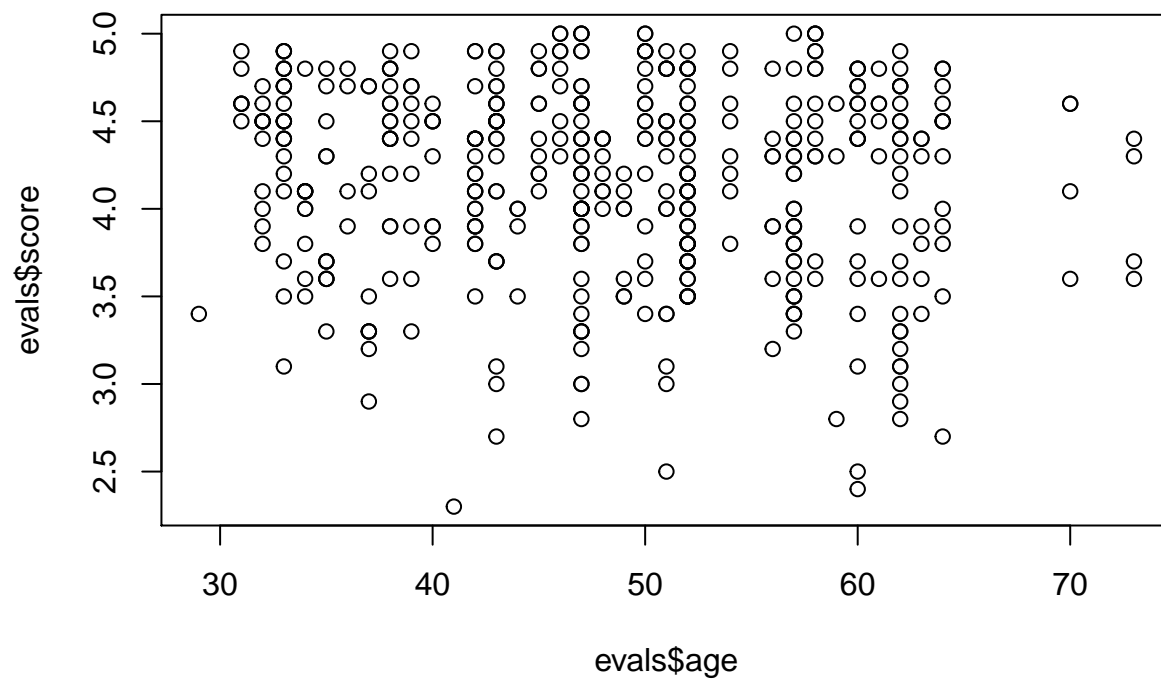
```
plot(evals$score ~ evals$gender)
```



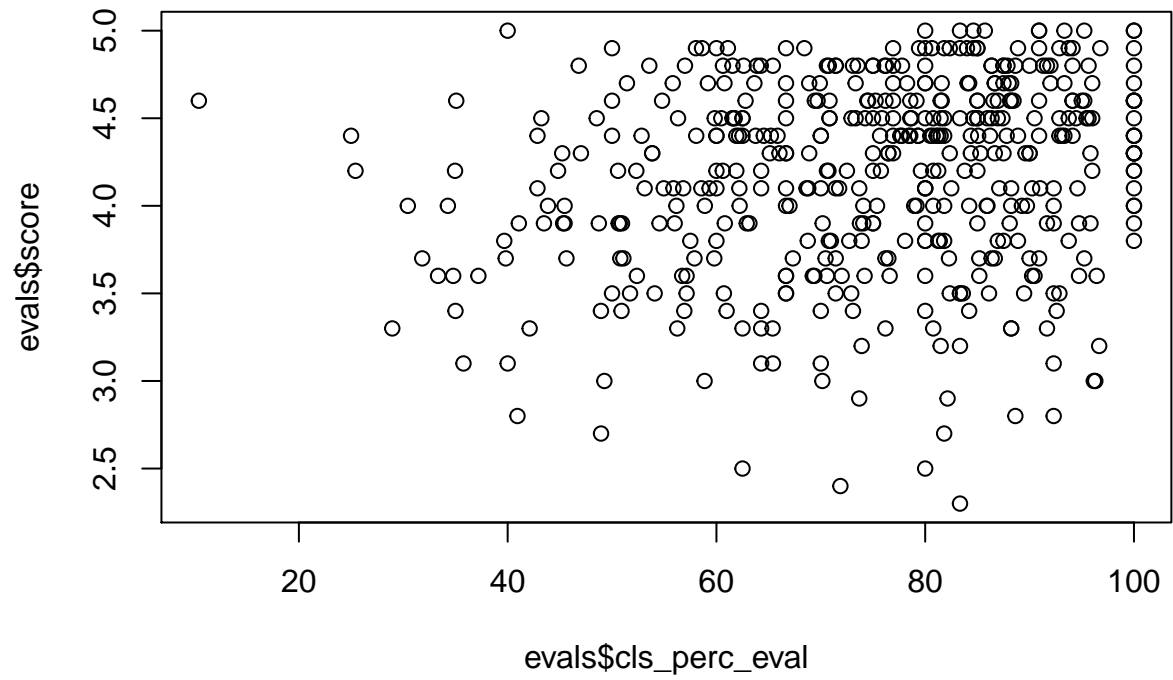
```
plot(evals$score ~ evals$language)
```



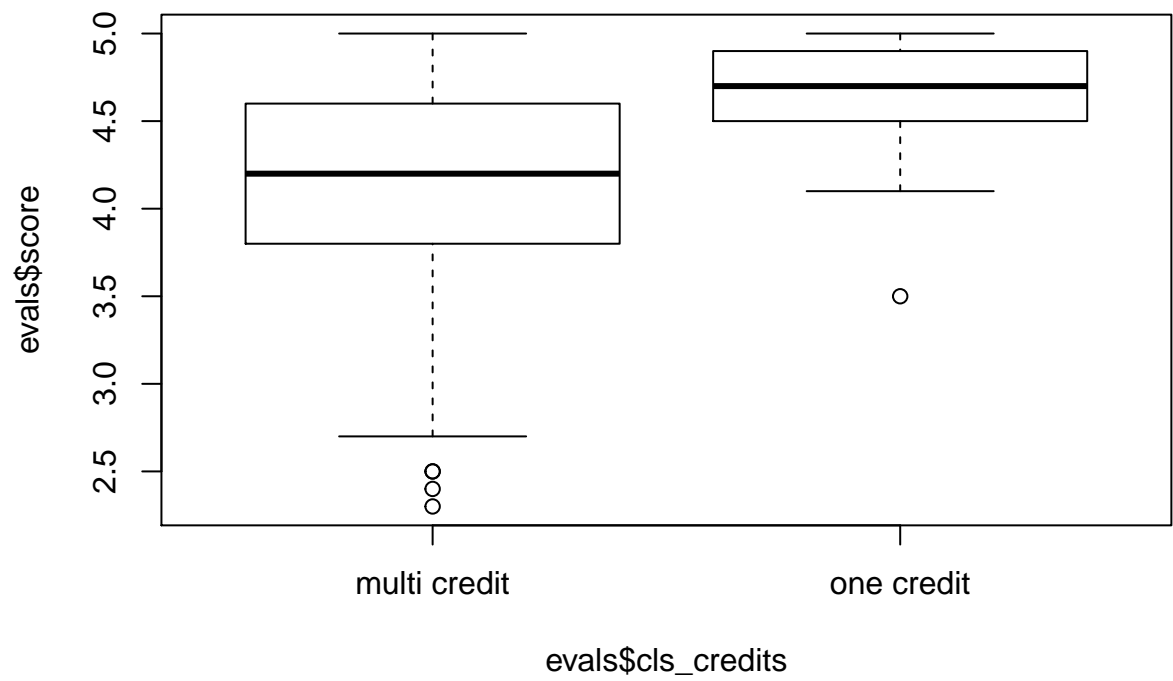
```
plot(evals$score ~ evals$age)
```



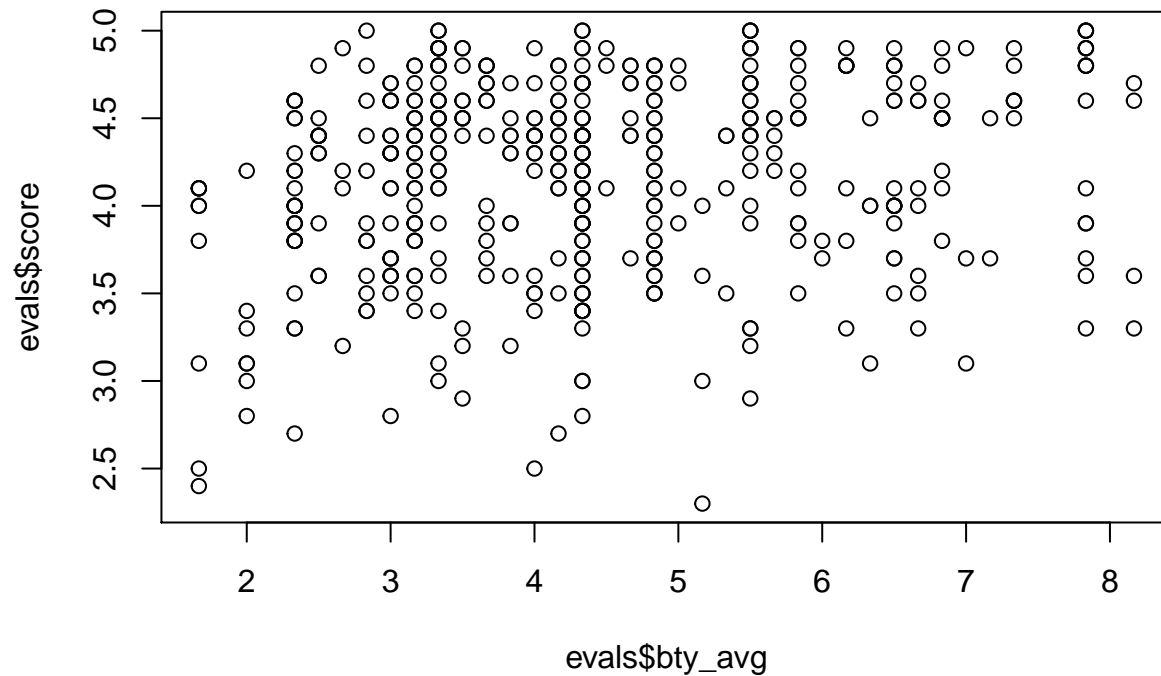
```
plot(evals$score ~ evals$cls_perc_eval)
```



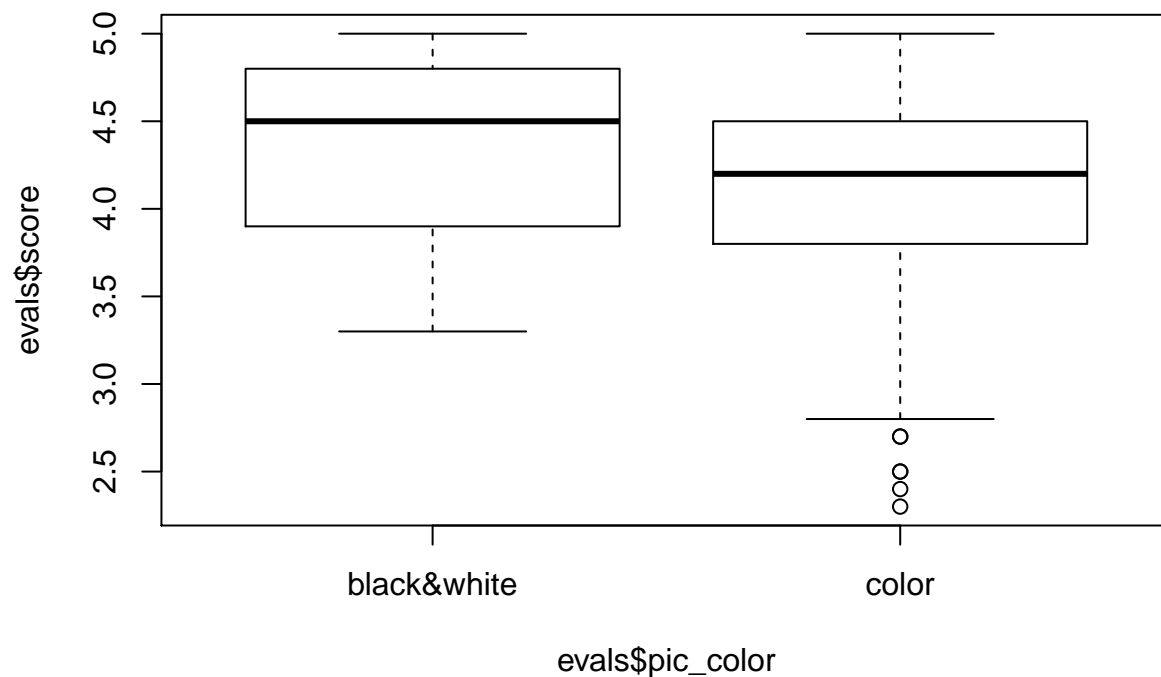
```
plot(evals$score ~ evals$cls_credits)
```



```
plot(evals$score ~ evals$bty_avg)
```



```
plot(evals$score ~ evals$pic_color)
```



17. The original paper describes how these data were gathered by taking a sample of professors from the University of Texas at Austin and including all courses that they have taught. Considering that each row represents a course, could this new information have an impact on any of the conditions of linear regression?

Answer: I don't think so. Since the classes are independent, the evaluation should be independent as well.

18. Based on your final model, describe the characteristics of a professor and course at University of Texas at Austin that would be associated with a high evaluation score. Answer: A male and non-minority professor who has graduated from an American school and teaches a one credit course, has a high beauty average score from the students and in the young age.
19. Would you be comfortable generalizing your conclusions to apply to professors generally (at any university)? Why or why not? Answer: No. because the sample size is too small.