

Analysis of the High-Dimensional Spatial Autoregressive Models with Bayesian Regularization



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Introduction

High Dimensional



That is: $p > n$

- Over parametrization
- Complexity

Spatial Auto Regressive



Bayesian Model Averaging (Lesage and Pace (2007))

Bayesian Approach



Prior Information

Bayesian Hierarchical

Posterior

$$\pi(\beta, \sigma^2 | Y) = \frac{\pi(Y | \beta, \sigma^2) \pi(\beta, \sigma^2)}{\int \pi(Y | \beta, \sigma^2) \pi(\beta, \sigma^2) d(\beta, \sigma^2)}$$

$Y | \beta, \sigma^2$

Data

Posterior

Posterior Log

Log

$$\log \pi(\beta, \sigma^2 | Y, X) \propto \log \pi(Y | X \beta, \sigma^2) + \log \pi(\beta | \sigma^2, \lambda) + \log \pi(\sigma^2) + \log \pi(\lambda)$$

Hyperparameters
Controls Complexity

Maximum a
Posterior
Estimator

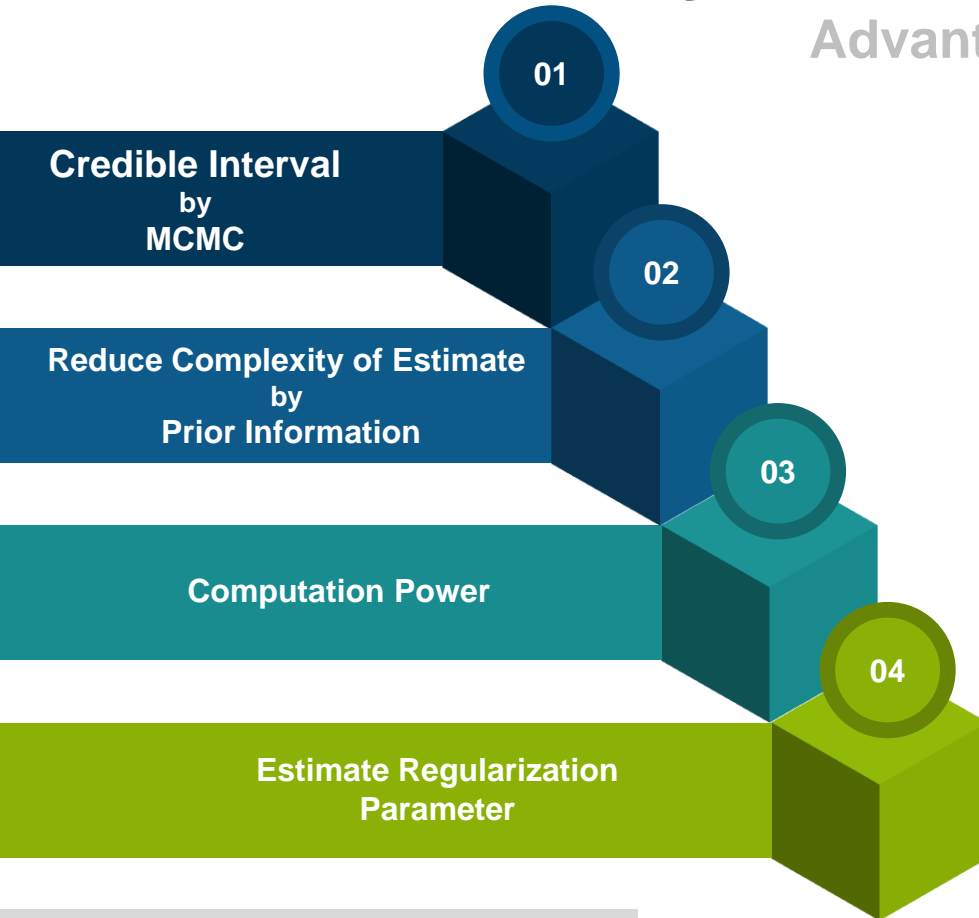
$\log \pi(\beta | \sigma^2, \lambda)$

Shrinkage Priors

Spike and Slab
Priors

Bayesian Hierarchical

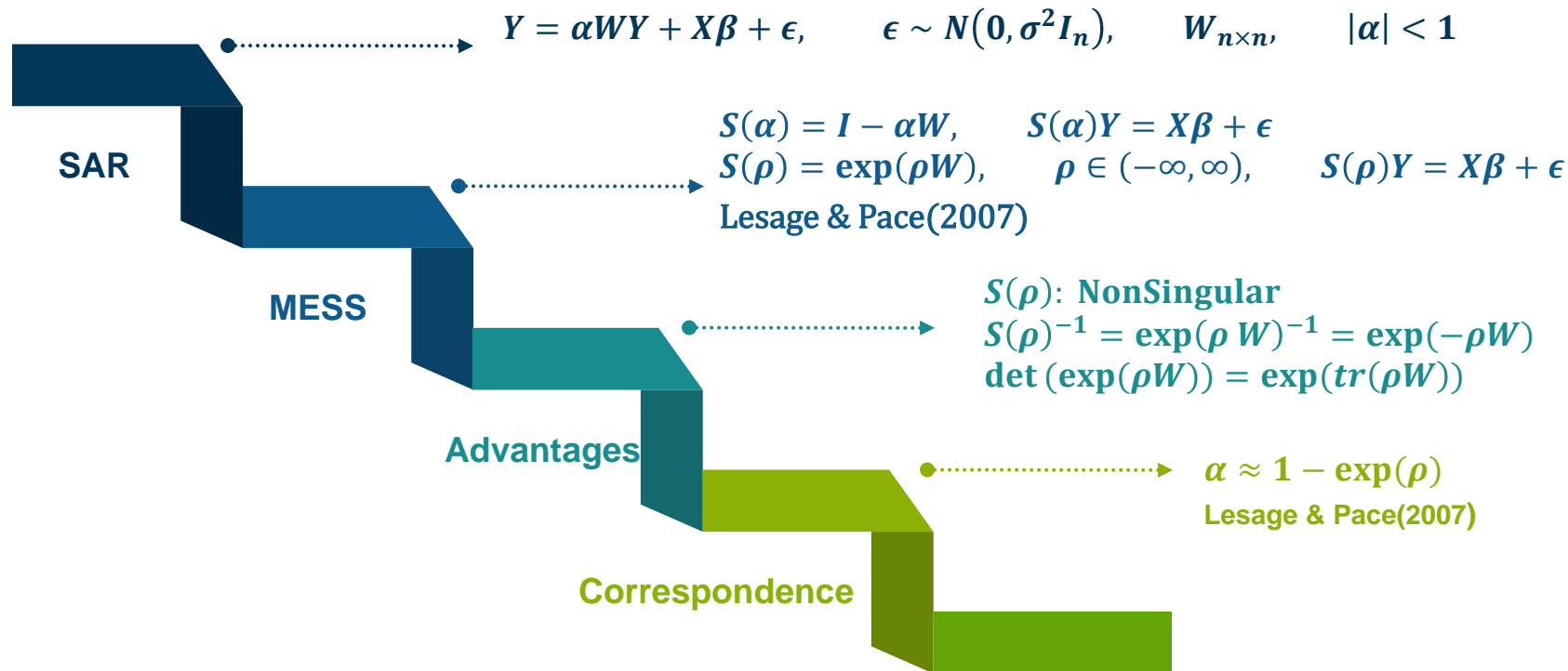
Advantages



- 01 Gibbs Variable Selection
- 02 Bayesian Model Averaging
- 03 Search Stochastic Variable Selection

Spatial Autoregressive Model

Matrix Exponential Spatial Specification



Global-Local Shrinkage Priors

✓ $S(\rho)Y = X\beta + \epsilon, \quad \epsilon \sim N(\mathbf{0}, \Omega)$ ✓ $f(Y|X\beta, \sigma^2, \rho) = (2\pi)^{-\frac{n}{2}} \det(\Omega)^{-\frac{1}{2}} \exp\left(-\frac{\mathbf{e}^T \Omega^{-1} \mathbf{e}}{2}\right)$

✓ $\beta_j | \psi_j \sim N(\mathbf{0}, \psi_j)$

$$\psi_j \sim D$$

Griffin & Brown(2010)

✓ More hierarchical layers:
More flexibility and adaptive shrinkage
Pfarrhofer & Piribauer (2019)

None

$$\underline{\beta} = 0, \underline{\Sigma} = 1000 I_p$$
$$\rho \sim N(0, c), \quad c = 10$$



$$\beta | \cdot \sim N(\bar{\beta}, \bar{\Sigma}),$$

$$\bar{\Sigma} = (\underline{\Sigma}^{-1} + X^T \Omega^{-1} X)^{-1},$$

$$\bar{\beta} = \bar{\Sigma}(\underline{\Sigma}^{-1} \underline{\beta} + X^T \Omega^{-1} S(\rho) Y)$$

$$\rho | \cdot \propto \exp\left(-\frac{e^T \Omega^{-1} e}{2}\right) \pi(\rho)$$

$$\min[1, \frac{g(\rho^* | \cdot)}{g(\rho_{t-1} | \cdot)}]$$

$$\rho^* \sim N(\rho_{t-1}, \tau)$$

$$\rho_t = \rho^*$$

Posterior

Normal-Gamma Shrinkage

Griffin & Brown (2010)

$$\beta_j | \psi_j \sim N(0, 2\lambda^{-2}\psi_j),$$

$$\psi_j \sim G(\theta, \theta),$$

$$\lambda^2 \sim G(d_0, d_1)$$



$$\psi_j | \lambda, \beta_j \sim GIG(\theta - \frac{1}{2}, \beta_j^2, \theta \lambda^2)$$

$$\lambda^2 | \psi \sim G(d_0 + \theta P, d_1 + 2^{-1}\theta \sum \psi_j)$$

$$\text{diag}(\underline{\Sigma}) = \psi$$

Dirichlet-Laplace Shrinkage

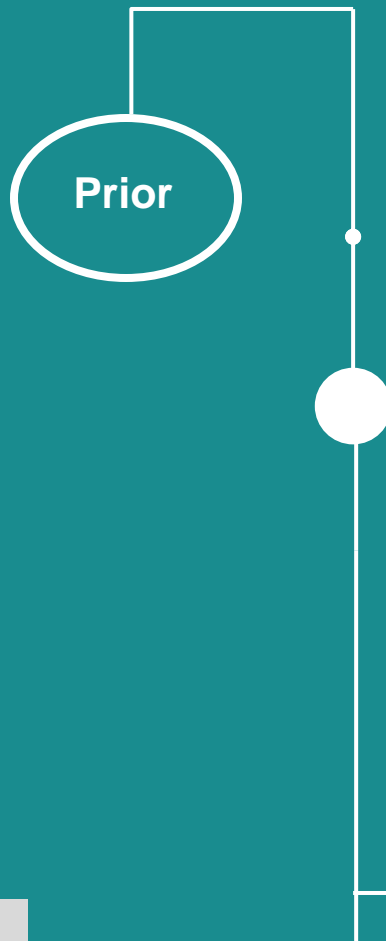
Bhattacharya & et al (2015)

$$\beta_j \sim N(0, \varphi_j \phi_j^2 \tau^2),$$

$$\varphi_j \sim \exp\left(\frac{1}{2}\right),$$

$$\phi \sim \text{Dir}(a, \dots, a),$$

$$\tau \sim G(na, \frac{1}{2})$$



$$\mu_j = \frac{\phi_j \tau}{|\beta_j|},$$

$$\tilde{\varphi}_j | \phi, \beta \sim \text{GIG}\left(-\frac{1}{2}, 1, \mu_j^{-2}\right), \varphi_j = \frac{1}{\tilde{\varphi}_j}$$

$$\tau | \phi, \beta \sim \text{GIG}(1 - p, 2 \sum_{j=1}^p \frac{|\beta_j|}{\phi_j}, 1)$$

$$T_j \sim \text{GIG}(a - 1, 2|\beta_j|, 1), \phi_j = \frac{T_j}{\sum_{j=1}^p T_j}$$

$$\text{diag}(\underline{\Sigma}) = (\varphi_1 \phi_1^2 \tau^2, \dots, \varphi_p \phi_p^2 \tau^2)^T$$

Simulation

✓ $n = 100$
 $p = 200$

✓ $X_{n \times p} \sim N_p(\mathbf{0}, \mathbf{1})$
 W : 5-Nearest Neighbors
 $\sigma^2 = 1$

✓ $\beta_{p \times 1}$: $\beta_{1:5} \sim N(\mathbf{0}, 5), \beta_{6:10} \sim N(\mathbf{0}, 1), \beta_{11:p} = \mathbf{0}$
 $\rho \sim N(\mathbf{0}, 3)$
 $Y = \exp(-\rho W)(X\beta + \epsilon), \epsilon \sim N(\mathbf{0}, \sigma^2 I_n)$

✓
$$\text{RMSE}(\hat{\theta}) = \left(\sum_{j=1}^M (\hat{\theta}_j - \tilde{\theta}_j)^2 / M \right)^{1/2}$$

✓ **Hyperparameters**
 $\theta = 0.1, \quad d_0 = d_1 = 0.01$
 $a = 1/p$

✓ **MCMC Algorithm**

- Input initial value: θ_0
- Generate random value: full conditional
- Number iteration= 2000
- Burn-in= 1000
- ACF

$N=100, P=200$

Dirichlet-Laplace

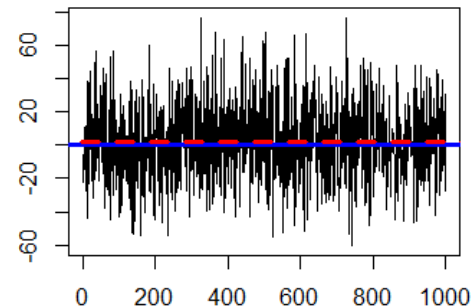
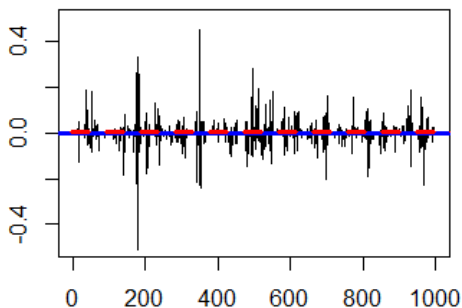
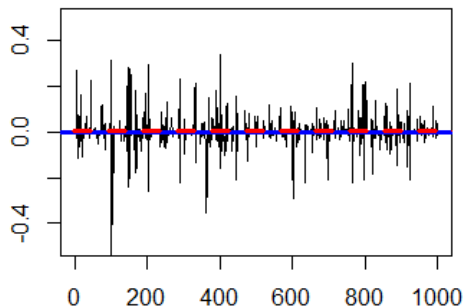
Normal-Gamma

None

DL

NG

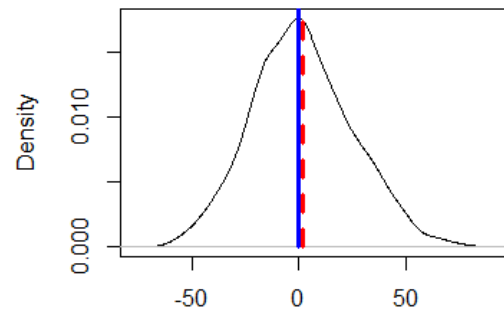
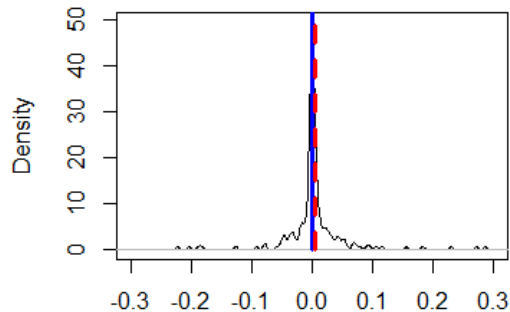
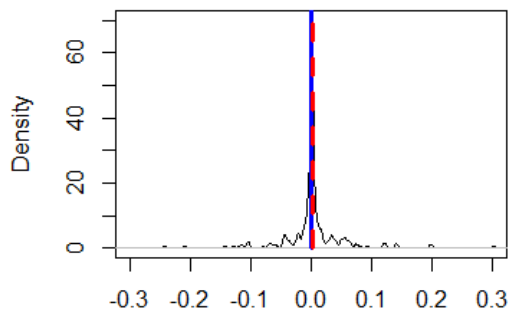
None



Iterations

Iterations

Iterations



Simulation Results

	P=200		
	RMSE_β	RMSE_ρ	Time(min.)
Dirichlet-Laplace	0.22	0.66	31.29
Normal-Gamma	0.29	0.66	31.78
None	0.39	0.94	30.59

Disadvantages and Future Works

- Estimate hyperparameters by Monte Carlo EM algorithm
- Define another shrinkage prior and compare results
- Estimate posterior density
- Reduce costs (time)

References

- Bhattacharya, A., Patti, D., Pillai, N. S., and Dunson, D. B. (2015), Dirichlet–Laplace Priors for Optimal Shrinkage, *Journal of the American Statistical Association*, 110, 1479-1490.
- Buhlman, P., and Van De Geer, S. (2011), *Statistics for High-Dimensional Data: Methods, theory and Applications*. Springer Science and Business Media.
- Buhlman, P., Kalisch, M., and Meier, L. (2014), *High-Dimensional Statistics with a View Toward Applications in Biology*.
- Griffin, J. E., and Brown, P. J. (2010), Inference with Normal-Gamma Prior Distributions in Regression Problems, *Bayesian Analysis*, 5, 171-188.
- Giraud, C. (2014), *Introduction to High-dimensional Statistics (Vol. 138)*, CRC Press.
- Himel Mallick, N. Y. (2013), Bayesian Methods for High Dimensional Linear Models, *Biostatistics*, 1, 005.
- Pfarrhofer, M., and Piribauer, P. (2019), Flexible Shrinkage in High-Dimensional Bayesian Spatial Autoregressive Models, *Spatial Statistics*, 29, 109-128.
- Park, T., and Casella, G. (2008), The Bayesian Lasso, *Journal of the American Statistical Association*, 103, 681-686.

THANKS
FOR
YOUR ATTENTION



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