

CEE 260/MIE 273: Probability and Statistics in Civil Engineering

Lecture M4a: Point Estimates, Sampling Variability and Central Limit Theorem

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Outline

- ① Statistical inference
- ② Point estimation
- ③ Method of moments
- ④ Variability and CLT
- ⑤ Outlook

Statistical inference

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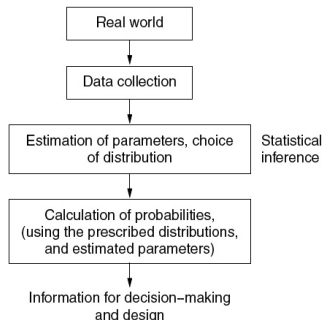
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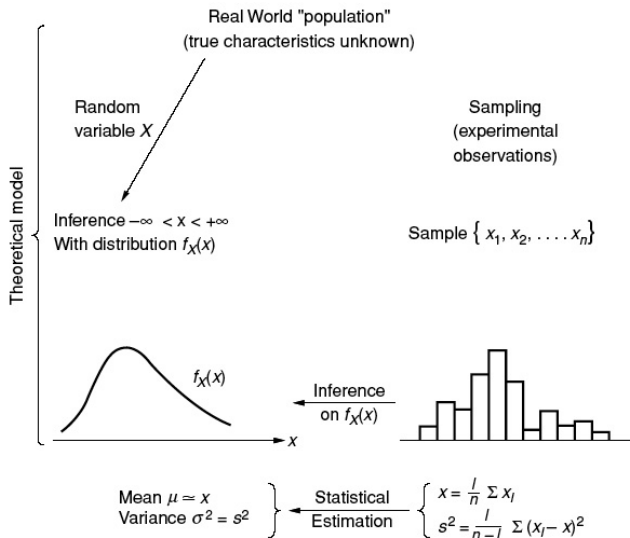
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Role of sampling in statistical inference



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- Hypothesis testing for a proportion (M4c)

Point estimates

Definition

A **point estimate** of a parameter θ (e.g. proportion p , or mean value μ) is a single number that can be regarded as a sensible value for θ and is obtained by computing the value of a suitable statistic (e.g. sample mean, sample standard deviation, etc) from given sample data.

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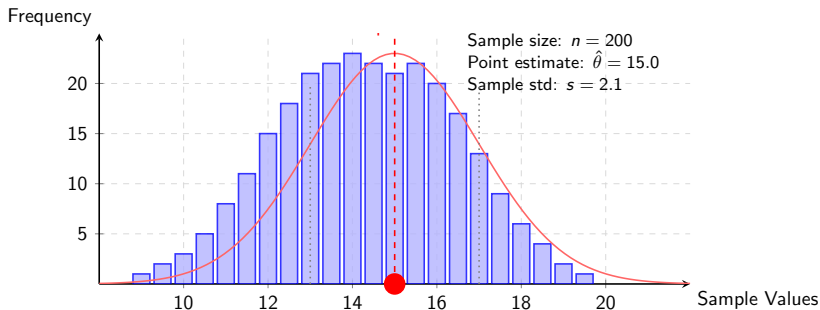


Figure: Sample histogram with point estimate $\hat{\theta}$ showing the center of the distribution

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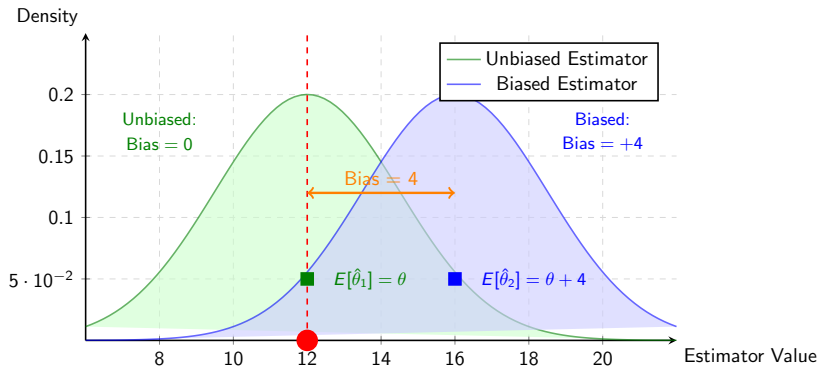
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An estimator is *consistent* if $\hat{\theta} \rightarrow \theta$ as $n \rightarrow \infty$, i.e. the estimation error should decrease with increasing sample size.

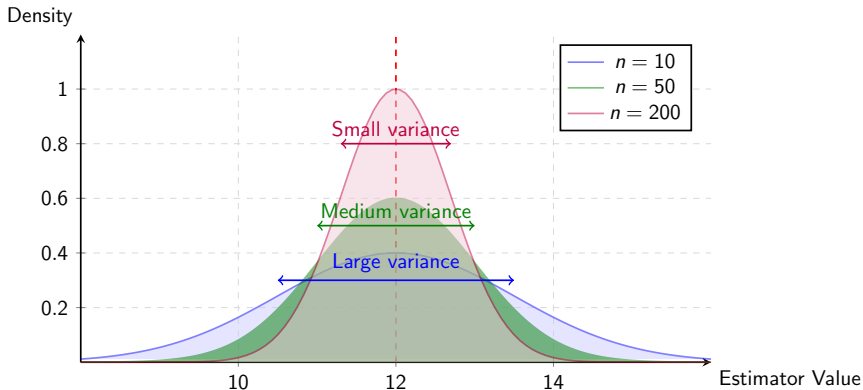


Figure: As sample size increases, the sampling distribution becomes more concentrated around the true parameter

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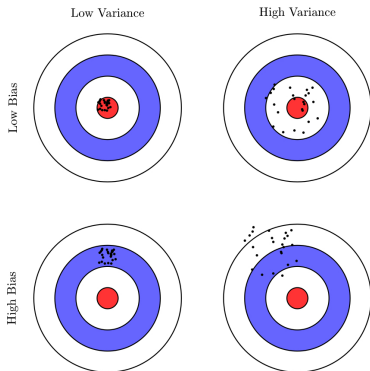


Image source: <https://tex.stackexchange.com/a/307285/2269>

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$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 \quad (5)$$

Unbiasedness of s^2

From Equation (5), you can show (as an exercise) that:

$$s^2 = \frac{1}{n-1} \left[\sum_{i=1}^n x_i^2 - n\bar{x}^2 \right] \quad (6)$$

You may be wondering why the sample variance is not just the average of the sum of squared deviations from the sample mean.

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The second estimator is biased and underestimates σ^2 by $-\frac{\sigma^2}{n}$.

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- (b) Estimate the population variance using the estimator s^2 (sample variance)
- (c) Now, estimate the variance replacing the denominator $(n - 1)$ with n in the estimator s^2 . What do you notice?

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$\hat{\sigma}^2$ underestimates σ^2 by 0.031 squared units.

Variability of a point estimate

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Example 2: Solar energy expansion

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- (b) Sample $n = 1000$ entries without replacement
- (c) Plot the histogram of the sampling distribution of \hat{p}
- (d) Compute the sample mean $x_{\hat{p}}$
- (e) Compute the standard deviation $s_{\hat{p}}$ (called the **standard error** $SE_{\hat{p}}$).
- (f) Investigate what happens as n increases.

The Central Limit Theorem (CLT)

Let X_1, X_2, \dots, X_n be a random sample from a distribution with mean μ and variance σ^2 . If n is sufficiently large, then the sample mean \bar{X} has approximately a **normal distribution** with

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- The product of large number of random components approaches the lognormal distribution

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Sample proportion and the CLT

If the observations in a given sample are a Bernoulli sequence with a constant proportion (or probability) p , then if n is large, the sample proportion \hat{p} follows a normal distribution (according to the CLT):

$$\hat{p} \sim \mathcal{N}(\mu_{\hat{p}}, SE_{\hat{p}}^2) = \mathcal{N}\left(p, \frac{p(1-p)}{n}\right) \quad (17)$$

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$$\text{Sampling error/standard error of } \hat{p}: SE_{\hat{p}} = \sqrt{\frac{p(1-p)}{n}} \approx \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

One rule of thumb for determining whether n is large enough is to check that both np and $n(1-p)$ are ≥ 10 (also known as the success-failure condition).

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- (a) According to the CLT, what is the distribution of \hat{p} ?
- (b) According to the CLT, what are $\mu_{\hat{p}}$ and $SE_{\hat{p}}$, respectively?

CLT application: sample proportion (cont.)

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- (a) First, we note that the response of each American adult in the entire population is part of a Bernoulli sequence with $p = 0.88$. According to the CLT, the distribution of \hat{p} (sample proportion) is normal/Gaussian. We can denote this as:

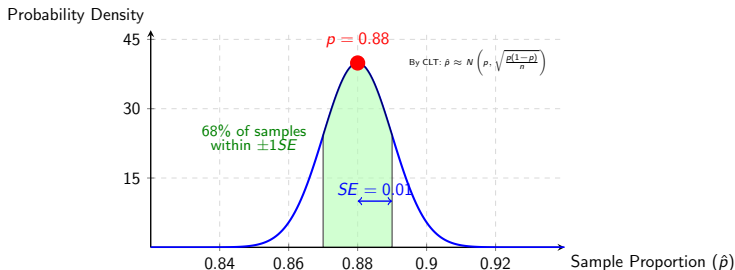
$$\hat{p} \sim \mathcal{N}\left(p, \frac{\sigma^2}{n}\right) \text{ OR } \mathcal{N}\left(\mu_p, \frac{\sigma_p^2}{n}\right) \quad (20)$$

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(b) $\mu_{\hat{p}}$ denotes the mean estimate of p , which is 0.88 (according to the CLT, the mean of the sample is the population mean if n is large).

$SE_{\hat{p}}$ denotes the sampling error, which is the the square root of the variance of the sample mean: $\sqrt{\sigma^2/n}$. Given that the sample is governed by the Binomial distribution with $\sigma^2 = p(1 - p)$. Thus:

$$SE_{\hat{p}}^2 = \frac{\sigma^2}{n} = \frac{p(1 - p)}{n} = \frac{0.88(0.12)}{1000}$$

$$SE_{\hat{p}} = \sqrt{\frac{0.88(0.12)}{1000}} = \boxed{0.01}$$

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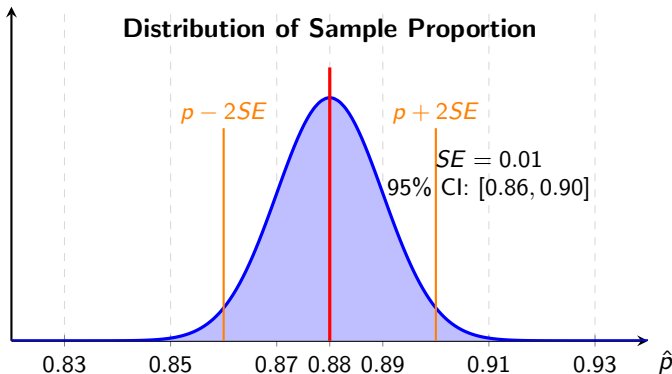


Figure: Sample proportion distribution: most samples fall within $\pm 2SE$ of the true proportion

Another application of the CLT

Example 4: Mean batch weight

A certain brand of cement is shipped in batches of 40 bags. Previous records indicate the weight of a randomly selected bag of this brand has a mean of 2.5 kg and an SD of 0.1 kg. The exact distribution is unknown.

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- (b) If the shipping company charges an overweight fee if a batch exceeds the mean batch weight by more than 1 kg, what is the probability that a batch will be charged?

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Summary

- Desired properties of point estimates: unbiasedness and efficiency
- Distribution of sample proportions (or other parameters) is called a sampling distribution
- When n is sufficiently large and observations are independent, the sample proportion (or other parameter) follows a normal distribution
- The success-failure condition can be used to determine if n is large enough for the CLT to hold (for a sample proportion)