

Final Project:

2.1: Don't worry about SMC, just ABC from class.

3.3: Model flu outbreaks

q_c : community transmission

q_h : household "

data in a Supplemental file - File006

w_{ij} give prob. of each cell in that table

then check if simulated table is close to observed data
if so keep it or reject and test new q_c, q_h

Notes on Paper:

Rather than just using ABC to find Θ , we also want to choose a model m . Global: calculate $P(m|D_0)$

↑ marginal posterior distribution of m

2 ways to find this:

Joint Space Based Approach

Define a joint space of model indicators $m = 1, 2, \dots, |\mathcal{M}|$
& model parameters Θ

Obtain joint posterior distribution over combined space
 $P(\Theta, m | D_0)$

Marginalize over parameters to obtain $P(m | D_0)$

Marginal Likelihood Based Approach

Estimate marginal likelihoods $P(D_0|m)$ & use "Evidence"

$$P(m|D_0) = \frac{P(D_0|m) P(m)}{\sum_{m'} P(D_0|m') P(m')}$$

Both are good but computationally costly so paper uses SMC (we won't). Paper prefers Joint Space Approach.

Model Selection based on ABC rejection:

particles: (m, θ)

(1) Draw m^* from the prior $P(m)$

(2) Sample θ^* from the prior $P(\theta|m^*)$

(3) Simulate a dataset $D^* \sim f(D|\theta^*, m^*)$
↑ what is f ?

(4) Compute distance

$$\begin{cases} d(D_0, D^*) \leq \varepsilon, & \text{accept } (m^*, \theta^*) \\ \text{else} & \text{reject} \end{cases}$$

(5) repeat until N particles are accepted

(6) Approximate marginal posterior distribution by:

$$P(m=m' | D_0) \approx \frac{\# \text{ accepted particles of the form } (m', \cdot)}{N}$$

Modifications made for SMC (we won't do this!)

(MS1) Initialize $\epsilon_1 > \dots > \epsilon_T$

Set population indicator $t=1$

(MS2.0) Set particle indicator $i=1$

(MS2.1) If $t=1$, sample (m^{**}, θ^{**}) from prior $P(m, \theta)$

• If $t>1$, sample m^* with $P_{t-1}(m^*)$ & draw

$$m^{**} \sim K M_t(m|m^*)$$

?

Sample θ^* from previous population $\{\theta(m^{**})_{t-1}\}$

with weights w_{t-1} & draw $\theta^{**} \sim \text{KP}_{t, m^{**}}(\theta | \theta)$

$\left\{ \begin{array}{l} P(m^{**}, \theta^{**}) = 0, \text{ return to MS2.1} \\ \text{else: simulate } D^* \sim f(D | \theta^{**}, m^{**}) \end{array} \right.$

$\left\{ \begin{array}{l} d(D_0, D^*) > \epsilon_t, \\ \text{else, continue} \end{array} \right.$

(MS2.2) Set $(m_t^i, \theta_t^i) = (m^{**}, \theta^{**})$ &

calculate weight of particle via some formula

If $i < N$, set $i=i+1$ & go to (MS2.1)

(MS3) Normalize weights w_t ← how?

Obtain marginal model probabilities by

$$P_t(m_t = m) = \sum w_t^{(i)} (m_t^{(i)}, \theta_t^{(i)})$$

If $t < T$ set $t=t+1$ & go to MS2.0

The goal of this algorithm is to obtain (m_T, θ_T) &
 $P_T(m_T)$

In 3.1 choose uniform prior distributions
(discussion of priors at the end of 2.1)

Section 3 Notes

3.3 - Goal: replicate figures 3a & 3c

q_{fc} := probability a susceptible person doesn't
get infected from community

q_{hn} := " household

w_{js} := probability that j of s susceptibles in a
row ¹ _{col} house get infected

Paper aims to infer q_c & q_n from data on w_{js}

& answer whether the 2 data sets have same q_c & q_n

Consider 2 models:

~~X~~ (1) has $q_{n1}, q_{c1}, q_{n2}, q_{c2}$ — surmises different
parameters

~~X~~ (2) has q_n, q_c — assumes same parameters

Prior distributions all assumed to be uniform on $[0, 1]$

Distance fxn given on pg. 107