### **Definition 3.11**

Poisson probability distribution of random variable *Y*:

$$p(y) = \frac{\lambda^y}{y!} e^{-\lambda}$$

#### Theorem 3.11

Expected value of Poisson distribution:

$$E(Y) = \mu = \lambda$$

Variance of Poisson distribution:

$$V(Y) = \sigma^2 = \lambda$$

## Theorem 3.14

Tchebysheff's Theorem for random variable Y, mean  $\mu$ , and variance  $\sigma^2$ . For any constant k > 0:

$$P(|Y - \mu| < k\sigma) \ge 1 - \frac{1}{k^2}$$
 OR  $P(|Y - \mu| \ge k\sigma) \le \frac{1}{k^2}$ 

### **Definition 4.1**

Distribution function of *Y* (any random variable):

$$F(y) = P(Y \le y)$$
 for all  $-\infty < y < \infty$ 

# Theorem 4.1

Properties of a distribution function:

1. 
$$F(-\infty) \equiv \lim_{y \to -\infty} F(y) = 0$$
  
2.  $F(\infty) \equiv \lim_{y \to \infty} F(y) = 1$ 

2. 
$$F(\infty) \equiv \lim_{y \to \infty} F(y) = 1$$

3. F(y) is a nondecreasing function of y.

# **Definition 4.2**

If F(y) is continuous on  $-\infty < y < \infty$ , Y is a continuous random variable.

### **Definition 4.3**

Probability density function of Y:

$$f(y) = \frac{dF(y)}{dy} = F'(y)$$

## Theorem 4.2

Properties of a density function f(y):

- 1.  $f(y) \ge 0$  for all  $-\infty < y < \infty$
- $2. \int_{-\infty}^{\infty} f(y) dy = 1$

### Theorem 4.3

If Y has a density function f(y) and bounds a < b, the probability that Y is on interval [a, b]:

$$P(a \le Y \le b) = \int_a^b f(y)dy$$

#### **Definition 4.5**

Expected value of continuous random variable Y:

$$E(Y) = \int_{-\infty}^{\infty} y f(y) dy$$

### Theorem 4.4

Expected value of continuous random variable Y:

$$E[g(Y)] = \int_{-\infty}^{\infty} g(y)f(y)dy$$

### Theorem 4.5

Let c be a constant, g(Y),  $g_1(Y)$  ...  $g_k(Y)$  be functions of Y.

- 1. E(c) = c
- 2. E(cg(Y) = cE[g(Y)]
- 3.  $E(g_1(Y) + g_2(Y) + \dots + g_k(Y)) = E[g_1(Y)] + E[g_2(Y)] + \dots + E[g_k(Y)]$

### **Definition 4.6**

Y has a continuous uniform probability distribution if its density function is:

$$f(y) = \begin{cases} \frac{1}{\theta_2 - \theta_1}, & \theta_1 \le y \le \theta_2 \\ 0, & elsewhere \end{cases}$$

#### Theorem 4.6

If  $\theta_1 < \theta_2$  and Y is uniformly distributed on  $(\theta_1, \theta_2)$ , then:

1. 
$$E(Y) = \mu = \frac{\theta_1 + \theta_2}{2}$$

2. 
$$V(Y) = \sigma^2 = \frac{(\theta_2 - \theta_1)^2}{12}$$

## **Definition 4.8**

Y has normal probability distribution if its density function is:

$$f(y) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(y-\mu)^2/(2\sigma^2)} \text{ when } \sigma > 0, \ -\infty < \mu < \infty, \text{ and } -\infty < y < \infty$$

# Theorem 4.7

Expected value of normally distributed random variable Y:

$$E(Y) = \mu$$

Variance of normally distributed random variable Y:

$$V(Y) = \sigma^2$$

# **Definition 4.9**

Y has a gamma distribution with parameters  $\alpha > 0$  and  $\beta > 0$  if the density function is:

$$f(y) = \begin{cases} \frac{y^{\alpha - 1}e^{-y/\beta}}{\beta^{\alpha}\Gamma(\alpha)}, & 0 \le y < \infty \\ 0, & elsewhere \end{cases}, \text{ where } \Gamma(\alpha) \int_0^{\infty} y^{\alpha - 1}e^{-y}dy$$

### Theorem 4.8

Expected value of normally distributed random variable Y:

$$E(Y) = \mu = \alpha \beta$$

Variance of normally distributed random variable Y:

$$V(Y) = \sigma^2 = \alpha \beta^2$$

### **Definition 4.11**

Y has an exponential distribution with parameter  $\beta > 0$  if the density function is:

$$f(y) = \begin{cases} \frac{1}{\beta} e^{-y/\beta}, & 0 \le y < \infty \\ 0, & elsewhere \end{cases}$$

# Theorem 4.10

Expected value of exponentially distributed random variable Y:

$$E(Y) = \mu = \beta$$

Variance of exponentially distributed random variable Y:

$$V(Y) = \sigma^2 = \beta^2$$

### **Definition 5.1**

Joint probability function of  $Y_1$  and  $Y_2$ :

$$p(y_{\mathrm{1}},y_{\mathrm{2}}) = p(Y_{\mathrm{1}} = y_{\mathrm{1}},Y_{\mathrm{2}} = y_{\mathrm{2}})$$
 when  $-\infty < y_{\mathrm{1}},y_{\mathrm{2}} < \infty$ 

# Theorem 5.1

If  $Y_1$  and  $Y_2$  have joint probability function, then:

1. 
$$p(y_1, y_2) \ge 0$$

2. 
$$\sum_{y_1,y_2} p(y_1,y_2) = 1$$

### **Definition 5.2**

Joint distribution function of  $Y_1$  and  $Y_2$ :

$$p(y_1, y_2) = P(Y_1 \le y_1, Y_2 \le y_2)$$
 when  $-\infty < y_1, y_2 < \infty$ 

#### **Definition 5.2**

Jointly continuous random variables  $Y_1$  and  $Y_2$ :

$$F(y_1, y_2) = \int_{-\infty}^{y_1} \int_{-\infty}^{y_2} f(t_1, t_2) dt_2 dt_1$$

#### Theorem 5.2

If  $Y_1$  and  $Y_2$  have joint distribution function, then:

1. 
$$F(-\infty, -\infty) = F(-\infty, y_2) = F(y_1, -\infty) = 0$$

2. 
$$F(\infty, \infty) = 1$$

3. 
$$f(y1, y2) \ge 0$$

4. 
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(y_1, y_2) dy_1 dy_2$$

### **Definition 5.4**

Marginal probability function of  $Y_1$  and  $Y_2$ :

$$p_1(y_1) = \sum_{all \ y_2} p(y_1, y_2)$$
 AND  $p_2(y_2) = \sum_{all \ y_1} p(y_1, y_2)$ 

Marginal density function of  $Y_1$  and  $Y_2$ :

$$f_1(y_1) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_2$$
 AND  $f_2(y_2) = \int_{-\infty}^{\infty} f(y_1, y_2) dy_1$ 

# **Definition 5.5**

Conditional discrete probability function of  $Y_1$  and  $Y_2$ :

$$p(y_1|y_2) = P(Y_1 = y_1|Y_2 = y_2) = \frac{P(Y_1 = y_1|Y_2 = y_2)}{P(Y_2 = y_2)} = \frac{p(y_1,y_2)}{p_2(y_2)} \text{ when } p2(y2) > 0$$

### **Definition 5.6**

Conditional distribution function of  $Y_1$  and  $Y_2$ :

$$p(y_1|y_2) = P(Y_1 = y_1|Y_2 = y_2)$$

### **Definition 5.7**

Conditional density of  $Y_1$ , given  $Y_2 = y_2$ 

$$f(y_1|y_2) = \frac{f(y_1,y_2)}{f_2(y_2)}$$

Conditional density of  $Y_2$ , given  $Y_1 = y_1$ 

$$f(y_2|y_1) = \frac{f(y_1,y_2)}{f_1(y_1)}$$

### **Definition 5.8**

Joint distribution variables  $Y_1$  and  $Y_2$  are independent if and only if

$$F(y_1, y_2) = F_1(y_1)F_2(y_2)$$
 for every  $(y_1, y_2)$ .

Otherwise,  $Y_1$  and  $Y_2$  are dependent.

# Theorem 5.4

Marginal probability variables  $Y_1$  and  $Y_2$  are independent if and only if

$$p(y_1, y_2) = p_1(y_1)p_2(y_2)$$
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Otherwise,  $Y_1$  and  $Y_2$  are dependent.