

Exam 2

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```
library(R2jags)
library(dplyr)
library(brms)
library(loo)
library(rstan)
```

For the first set of questions we will be using the data file **influent.dat**. You should have received that data file with the exam. Water flows into the Mississippi river from a number of streams and rivers. These sources carry nitrogen into the river. In the data set, there are six sources of the nitrogen, which are a random sample of the many hundreds of streams and rivers that flow into the Mississippi. These sources have been classified by type. The three types are: (1) no farm land in watershed, (2) less than 50% farm land in watershed, and (3) more than 50% farmland in watershed. These three types are the only types that we seek to analyze. In the data file, the three columns are river source, nitrogen, and type.

1. First read in the data and print out the first six rows of the data.

```
influent <- read.table("influent.dat")
colnames(influent) <- c("source", "nitrogen", "type")
```

2. Ignoring the source, use brm to write a model to find the differences in type. Use priors of normal(0,100) for the type effects, and gamma(2,.1) for σ_{error} . Remember that the kind of variable you are working with will make a difference. What is the level of nitrogen estimated for type 3.

Level of nitrogen estimated for type 3 is 36.36

```
source = influent$source
nitrogen = influent$nitrogen
type = influent$type
influent$vtype = as.factor(influent$type)

fit1 <- brm(formula = nitrogen ~ -1 + vtype, data = influent,
            family = "gaussian",
            prior = c(set_prior("normal(0,100)", class = "b"),
                     set_prior("gamma(2,.1)", class = "sigma")),
            warmup = 1000, iter = 5000, chains = 4,
```

```
#control = list(adapt_delta = 0.98),
save_pars = save_pars(all = TRUE))
```

```
summary(fit1)
```

```
Family: gaussian
Links: mu = identity; sigma = identity
Formula: nitrogen ~ -1 + vtype
Data: influent (Number of observations: 37)
Samples: 4 chains, each with iter = 5000; warmup = 1000; thin = 1;
         total post-warmup samples = 16000
```

Population-Level Effects:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
vtype1	15.61	2.39	10.93	20.43	1.00	16978	11131
vtype2	19.92	1.59	16.82	23.08	1.00	16060	11019
vtype3	36.32	3.33	29.83	42.94	1.00	16383	11428

Family Specific Parameters:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma	7.46	0.95	5.86	9.57	1.00	15101	12282

Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

```
chains <- as.matrix(fit1)
dim(chains)
head(chains)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)
```

3. What is the looic for this model?

The looic is 255.6

```
loo1 <- loo(fit1)
print(loo1)
```

Computed from 16000 by 37 log-likelihood matrix

	Estimate	SE
elpd_loo	-127.7	4.0
p_loo	3.3	0.8
looic	255.5	8.0

Monte Carlo SE of elpd_loo is 0.0.

All Pareto k estimates are good ($k < 0.5$).
See `help('pareto-k-diagnostic')` for details.

4. Now run the same model using JAGS. Use the same priors for the type parameters (remember, normal priors in JAGS use precision), but use a $\text{gamma}(5,1)$ for σ_{error}^2 . What is the level of nitrogen estimated for type 3?

```
mdl <- "

model {

  for(i in 1:37){
    nitrogen[i] ~ dnorm(mu[i], 1/s2error)
    mu[i] <- beta[type[i]]
  }

  # Priors
  for(i in 1:3){
    beta[i] ~ dnorm(0, 0.0001)
  }

  s2error ~ dgamma(5, 0.1)

}
"
```

```
writeLines(mdl, 'fit2.txt')

source = influent$source
nitrogen = influent$nitrogen
type = influent$type

data.jags <- c('nitrogen', 'type')
parms <- c('beta', 's2error')

fit2 <- jags(data= data.jags, parameters.to.save = parms,
             model.file = 'fit2.txt', inits = NULL,
             n.iter = 20000, n.thin = 5, n.burnin = 2000,
             n.chains = 5)

module glm loaded

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 37
  Unobserved stochastic nodes: 4
  Total graph size: 84
```

Initializing model

fit2

```
Inference for Bugs model at "fit2.txt", fit using jags,
  5 chains, each with 20000 iterations (first 2000 discarded), n.thin = 5
  n.sims = 18000 iterations saved
```

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat	n.eff
beta[1]	15.570	2.305	11.022	14.053	15.560	17.118	20.080	1.001	18000
beta[2]	19.883	1.538	16.861	18.879	19.866	20.891	22.955	1.001	18000
beta[3]	36.352	3.217	29.982	34.208	36.346	38.468	42.730	1.001	18000
s2error	52.802	11.465	34.535	44.578	51.389	59.398	78.966	1.001	18000
deviance	251.400	2.736	247.987	249.403	250.754	252.749	258.306	1.001	14000

For each parameter, n.eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor (at convergence, Rhat=1).

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

$pD = 3.7$ and $DIC = 255.1$

DIC is an estimate of expected predictive error (lower deviance is better).

```
sims <- as.mcmc(fit2)
chains <- as.matrix(sims)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)
```

5. What is the DIC of the above model?

DIC is 255.3

6. Now we want to account for the variance in sources to make inference. Redo the model in brm, but now put source in the model appropriately. Use a $\text{gamma}(2,1)$ prior for σ_{source} . The other priors can stay the same. Now what is the estimate for the type 3 mean?

Estimate for type 3 mean is 12.24. This plus whatever the effect for whatever source it is.

```
fit3 <- brm(formula = nitrogen ~ -1 + vtype + (1|source), data = influent,
  family = "gaussian",
  prior = c(set_prior("normal(0,10)", class = "b"),
    set_prior("gamma(2,.1)", class = "sigma"),
    set_prior("gamma(2,.1)", class = "sd")),
  warmup = 1000, iter = 20000, chains = 4, thin = 10,
  #control = list(adapt_delta = 0.98),
  save_pars = save_pars(all = TRUE))
```

```
summary(fit3)
```

```
Family: gaussian
Links: mu = identity; sigma = identity
Formula: nitrogen ~ -1 + vtype + (1 | source)
Data: influent (Number of observations: 37)
Samples: 4 chains, each with iter = 20000; warmup = 1000; thin = 10;
         total post-warmup samples = 7600
```

Group-Level Effects:

~source (Number of levels: 6)

	Estimate	Est.Error	1-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sd(Intercept)	16.13	7.57	4.57	34.13	1.00	6961	6295

Population-Level Effects:

	Estimate	Est.Error	1-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
vtype1	7.34	7.91	-9.51	21.51	1.00	7534	7448
vtype2	11.38	7.71	-5.39	24.20	1.00	6803	6943
vtype3	12.01	10.71	-9.27	32.01	1.00	7091	6080

Family Specific Parameters:

	Estimate	Est.Error	1-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma	6.85	0.92	5.32	8.90	1.00	6895	6900

Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

```
fit3$fit
```

Inference for Stan model: 68305c8400f0c8b7efc70543152eb5c9.

4 chains, each with iter=20000; warmup=1000; thin=10;

post-warmup draws per chain=1900, total post-warmup draws=7600.

	mean	se_mean	sd	2.5%	25%	50%	75%
b_vtype1	7.34	0.09	7.91	-9.51	2.34	7.90	12.89
b_vtype2	11.38	0.09	7.71	-5.39	6.57	12.13	16.99
b_vtype3	12.01	0.13	10.71	-9.27	4.70	11.95	19.57
sd_source__Intercept	16.13	0.09	7.57	4.57	10.86	15.03	20.22
sigma	6.85	0.01	0.92	5.32	6.20	6.75	7.40
r_source[1,Intercept]	9.93	0.10	7.97	-3.52	4.23	9.14	14.85
r_source[2,Intercept]	2.52	0.09	7.77	-10.97	-2.95	1.73	7.36
r_source[3,Intercept]	9.00	0.10	8.29	-5.58	3.10	8.51	14.21
r_source[4,Intercept]	12.63	0.10	8.20	-1.17	6.70	11.97	17.73
r_source[5,Intercept]	6.77	0.09	8.20	-7.98	0.92	6.13	12.11
r_source[6,Intercept]	23.18	0.13	11.28	1.95	15.33	23.08	31.09
lp__	-145.89	0.04	3.43	-153.63	-147.89	-145.43	-143.40
z_1[1,1]	0.63	0.01	0.48	-0.32	0.33	0.62	0.93
z_1[1,2]	0.06	0.01	0.55	-1.26	-0.24	0.12	0.42
z_1[1,3]	0.57	0.01	0.54	-0.49	0.23	0.57	0.91

z_1[1,4]	0.83	0.01	0.50	-0.10	0.50	0.80	1.14
z_1[1,5]	0.41	0.01	0.54	-0.72	0.07	0.41	0.75
z_1[1,6]	1.51	0.01	0.66	0.25	1.08	1.47	1.91
	97.5%	n_eff	Rhat				
b_vtype1	21.51	7481	1				
b_vtype2	24.20	6609	1				
b_vtype3	32.01	7035	1				
sd_source__Intercept	34.13	6854	1				
sigma	8.90	6962	1				
r_source[1,Intercept]	27.34	6688	1				
r_source[2,Intercept]	19.38	6703	1				
r_source[3,Intercept]	26.63	7372	1				
r_source[4,Intercept]	30.16	6859	1				
r_source[5,Intercept]	24.21	7533	1				
r_source[6,Intercept]	45.46	7070	1				
lp__	-140.49	6972	1				
z_1[1,1]	1.58	6914	1				
z_1[1,2]	0.98	7031	1				
z_1[1,3]	1.64	7635	1				
z_1[1,4]	1.87	7469	1				
z_1[1,5]	1.45	7832	1				
z_1[1,6]	2.91	7097	1				

Samples were drawn using NUTS(diag_e) at Mon Mar 22 15:21:30 2021.
 For each parameter, n_eff is a crude measure of effective sample size,
 and Rhat is the potential scale reduction factor on split chains (at
 convergence, Rhat=1).

Diagnostics look good

```
chains <- as.matrix(fit3)
dim(chains)
```

```
[1] 7600 18
```

```
sims <- as.mcmc(chains)
raftery.diag(sims)
```

Quantile (q) = 0.025
 Accuracy (r) = +/- 0.005
 Probability (s) = 0.95

	Burn-in (M)	Total (N)	Lower bound (Nmin)	Dependence factor (I)
b_vtype1	2	3924	3746	1.050
b_vtype2	4	8206	3746	2.190
b_vtype3	3	4188	3746	1.120
sd_source__Intercept	2	3759	3746	1.000
sigma	6	8690	3746	2.320
r_source[1,Intercept]	2	3759	3746	1.000

r_source[2,Intercept]	2	3840	3746	1.030
r_source[3,Intercept]	2	3600	3746	0.961
r_source[4,Intercept]	2	3718	3746	0.993
r_source[5,Intercept]	2	3718	3746	0.993
r_source[6,Intercept]	2	3718	3746	0.993
lp__	2	3639	3746	0.971
z_1[1,1]	2	3882	3746	1.040
z_1[1,2]	2	3924	3746	1.050
z_1[1,3]	2	3639	3746	0.971
z_1[1,4]	2	3718	3746	0.993
z_1[1,5]	2	3600	3746	0.961
z_1[1,6]	2	3718	3746	0.993

`effectiveSize(sims)`

	b_vtype1	b_vtype2	b_vtype3
	7600.000	6858.278	7175.409
sd_source__Intercept		sigma	r_source[1,Intercept]
	6811.934	7600.000	7093.651
r_source[2,Intercept]	r_source[3,Intercept]	r_source[4,Intercept]	
	6905.563	7600.000	6970.358
r_source[5,Intercept]	r_source[6,Intercept]	lp__	
	7600.000	7200.955	6966.401
z_1[1,1]	z_1[1,2]	z_1[1,3]	
	6668.253	7143.009	7600.000
z_1[1,4]	z_1[1,5]	z_1[1,6]	
	7600.000	7220.784	

7. What is the looic for this model?

looic is 252.7

```
loo3 <- loo(fit3)
```

8. Which model (problem 2 or problem 6 is better? Why?

The model for problem 6 is better. It has a looic of 252.7 compared to 255.3. Clearly accounting for source is important.

9. Now do the model with the variance for sources accounted for in JAGS. Use a $\text{gamma}(5,1)$ for the variance component σ_{source}^2 . What is the estimate for the type 3 mean?

Estimate for type 3 mean is 36.24.

```
mdl <- "
model {
  for(i in 1:37){
    nitrogen[i] ~ dnorm(mu[i], 1/s2error)
```

```

    mu[i] <- beta[type[i]] + u[source[i]]
  }

# Priors
for(i in 1:3){
  beta[i] ~ dnorm(0, 0.0001)
}
for(i in 1:6){
  u[i] ~ dnorm(0,1/s_source)
}

s2error ~ dgamma(5, 0.1)
s_source ~ dgamma(5,0.1)

}
"

writeLines(mdl, 'fit4.txt')

source = influent$source
nitrogen = influent$nitrogen
type = influent$type

data.jags <- c('nitrogen', 'type', 'source')
parms <- c('beta', 'u', 's2error')

fit4 <- jags(data= data.jags, parameters.to.save = parms,
             model.file = 'fit4.txt', inits = NULL,
             n.iter = 20000, n.thin = 5, n.burnin = 2000,
             n.chains = 5)

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 37
  Unobserved stochastic nodes: 11
  Total graph size: 135

Initializing model

(fit4)

Inference for Bugs model at "fit4.txt", fit using jags,
  5 chains, each with 20000 iterations (first 2000 discarded), n.thin = 5
  n.sims = 18000 iterations saved

```

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat	n.eff
beta[1]	15.593	5.122	5.419	12.300	15.621	18.843	25.856	1.001	18000
beta[2]	19.917	4.056	11.835	17.324	19.906	22.505	27.959	1.001	18000
beta[3]	36.243	7.261	21.861	31.563	36.218	40.889	50.659	1.001	18000

s2error	45.742	10.521	29.092	38.265	44.339	51.735	70.205	1.001	18000
u[1]	1.460	4.176	-6.915	-1.212	1.456	4.164	9.708	1.001	18000
u[2]	-5.191	4.260	-13.857	-7.930	-5.108	-2.388	3.021	1.001	18000
u[3]	0.961	5.040	-8.933	-2.290	0.941	4.204	11.108	1.001	18000
u[4]	3.790	4.291	-4.691	1.000	3.767	6.552	12.443	1.001	13000
u[5]	-0.971	5.042	-11.102	-4.198	-0.978	2.284	8.979	1.001	18000
u[6]	0.106	6.582	-13.033	-3.980	0.121	4.253	13.152	1.001	18000
deviance	244.616	3.832	239.069	241.797	244.001	246.727	253.897	1.001	18000

For each parameter, n.eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor (at convergence, Rhat=1).

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

$pD = 7.3$ and $DIC = 252.0$

DIC is an estimate of expected predictive error (lower deviance is better).

```
sims <- as.mcmc(fit4)
chains <- as.matrix(sims)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)
```

10. What is the DIC for this model?

DIC is 252.0

11. Which model (problem 4 or problem 9) is better? Why?

DIC for model 9 is better, it is lower. Therefore there is value in accounting for the which source it comes from. We can now make inference for sources.

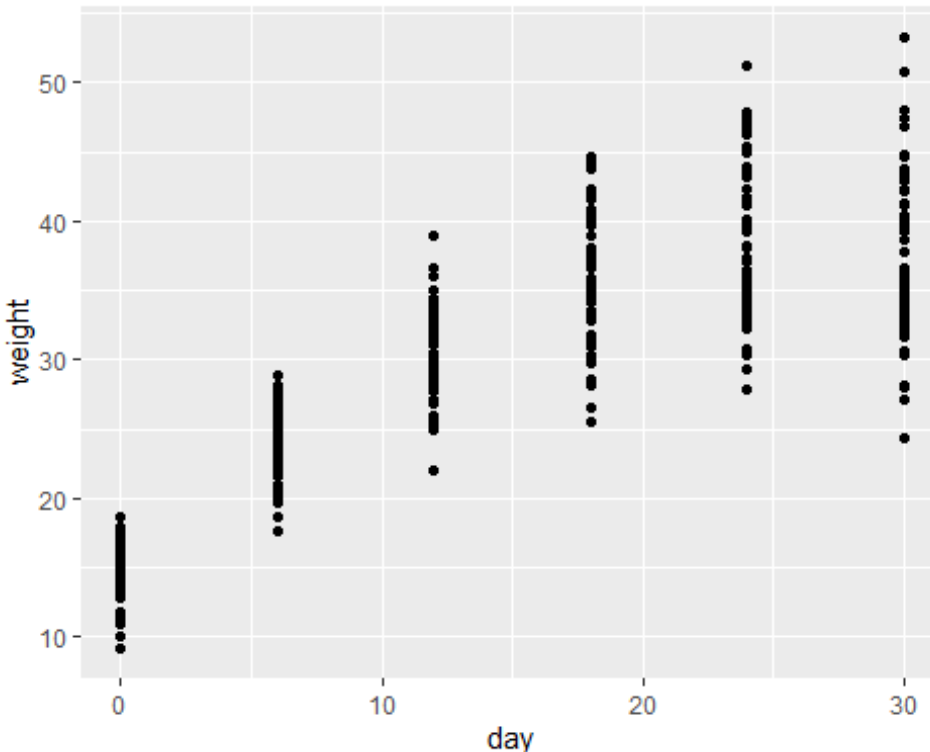
We will be using the data file **pig.dat** for the the next set of questions. You should have received that data file. The data set is from a feeding experiment comparing 3 treatments (basically how much of a particular additive was included in the food ration) on young hogs over the first 30 days of life. The measurements were taken every 6 days. The pig number identifies the pig within a treatment. That is, pig 1 in trt 1 is not the same pig as pig 1 in trt 2.

12. Plot the weight on the y-axis and the day on the x-axis, ignoring the trt and the pig.

```
pig <- read.table("pig.dat", header = TRUE)
head(pig)
```

	Obs	trt	day	pig	weight
1	1	1	0	1	14.0
2	2	1	6	1	22.1
3	3	1	12	1	27.7
4	4	1	18	1	31.8
5	5	1	24	1	35.3
6	6	1	30	1	32.6

```
ggplot(data = pig, mapping = aes(x = day, y = weight)) +  
  geom_point()
```



13. You will note from your plot that weight gain is fairly linear over the first 15 days or so, and then tails off. To account for this, we are going to add a quadratic term for day. That is, include a column for day^2 in the data set. Print the first 6 rows of the new data set.

```
pig$day2 = (pig$day)^2  
head(pig)
```

	Obs	trt	day	pig	weight	day2
1	1	1	0	1	14.0	0
2	2	1	6	1	22.1	36
3	3	1	12	1	27.7	144
4	4	1	18	1	31.8	324
5	5	1	24	1	35.3	576
6	6	1	30	1	32.6	900

14. Using JAGS, run a model with an intercept, linear term, and quadratic term for the overall data. You are essentially fitting the data you plotted, with no concern about treatments or the multiple observations on each pig. Use $\text{dnorm}(0, 0.0001)$ priors for the coefficients of the model (the β 's), and a $\text{dgamma}(1.1, 1)$ as the prior for σ_{error}^2 . Print a summary of your model. What is the DIC of the model?

DIC is 2073

```

mdl <- "
  model {

    for (i in 1:360){
      weight[i] ~ dnorm(mu[i], 1/vv)
      mu[i] <- b0 + b1*day[i] + b2*day2[i]
    }

    b0 ~ dnorm(0, 0.0001)
    b1 ~ dnorm(0, 0.0001)
    b2 ~ dnorm(0, 0.0001)
    vv ~ dgamma(1.1, 0.1)
  }

"
day = pig$day
day2 = pig$day2
weight = pig$weight

writeLines(mdl, 'fit5.txt')

data.jags <- c('weight', 'day', 'day2')
parms <- c('b0', 'b1', 'b2', 'vv')

fit5 <- jags(data= data.jags, parameters.to.save = parms,
             model.file = 'fit5.txt', inits = NULL,
             n.iter = 12000, n.thin = 5, n.burnin = 2000,
             n.chains = 5)

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 360
  Unobserved stochastic nodes: 4
  Total graph size: 1108

Initializing model

(fit5)

Inference for Bugs model at "fit5.txt", fit using jags,
  5 chains, each with 12000 iterations (first 2000 discarded), n.thin = 5
  n.sims = 10000 iterations saved

```

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat
b0	14.738	0.503	13.771	14.393	14.740	15.080	15.719	1.002
b1	1.719	0.078	1.564	1.665	1.718	1.771	1.872	1.001
b2	-0.031	0.002	-0.036	-0.033	-0.031	-0.030	-0.027	1.001
vv	18.331	1.385	15.860	17.366	18.254	19.222	21.276	1.001

```

deviance 2068.906    2.874 2065.356 2066.790 2068.267 2070.282 2076.240 1.001
      n.eff
b0      2600
b1      6000
b2     10000
vv     10000
deviance 10000

```

For each parameter, n.eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor (at convergence, Rhat=1).

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

$pD = 4.1$ and $DIC = 2073.0$

DIC is an estimate of expected predictive error (lower deviance is better).

```

sims <- as.mcmc(fit5)
chains <- as.matrix(sims)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)

```

15. Run the same model using brm. Use the same priors for the β 's (which will now be normal(0,100) since brm works in standard deviations), and a gamma(1.1,.5) prior for the standard deviation (σ_{error}). Print a summary of your model. What is the looic of the model?

looic is 2073.6

```

fit6 <- brm(formula = weight ~ day + day2, data = pig,
  family = "gaussian",
  prior = c(set_prior("normal(0,100)", class = "b"),
    #set_prior("normal(0,100)", class = "a"),
    set_prior("gamma(1.1,0.5)", class = "sigma")),
  warmup = 1000, iter = 5000, chains = 4,
  #control = list(adapt_delta = 0.98),
  save_pars = save_pars(all = TRUE), silent = TRUE)

```

(fit6)

```

Family: gaussian
Links: mu = identity; sigma = identity
Formula: weight ~ day + day2
Data: pig (Number of observations: 360)
Samples: 4 chains, each with iter = 5000; warmup = 1000; thin = 1;
total post-warmup samples = 16000

```

Population-Level Effects:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
Intercept	14.74	0.50	13.76	15.71	1.00	10405	9927
day	1.72	0.08	1.57	1.87	1.00	8871	8980
day2	-0.03	0.00	-0.04	-0.03	1.00	9148	9265

Family Specific Parameters:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma	4.28	0.16	3.98	4.61	1.00	10447	9294

Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

Convergence and ESS are good

```
chains <- as.matrix(fit6)
```

```
dim(chains)
```

```
sims <- as.mcmc(chains)
```

```
raftery.diag(sims)
```

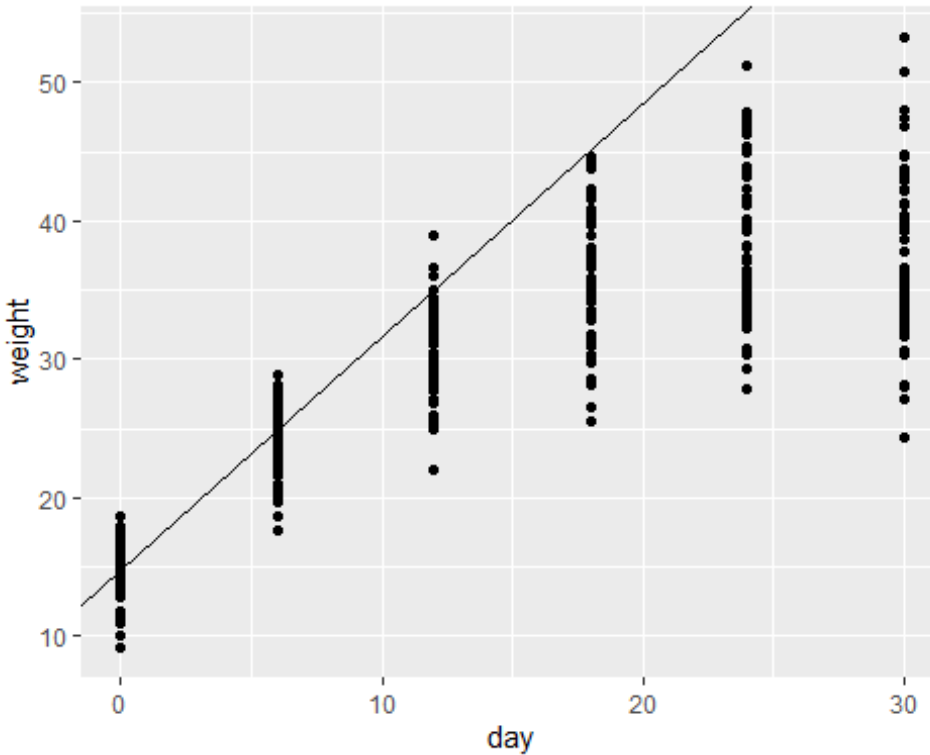
```
effectiveSize(sims)
```

looic

```
loo(fit6)
```

16. Using the estimated coefficients from either the JAGS or the brm model, add a best fit line to the plot you made in problem 12.

```
ggplot(data = pig, mapping = aes(x = day, y = weight)) +
  geom_point() +
  geom_abline(intercept = 14.73, slope = 1.69)
```



17. Using JAGS, run a model with an intercept, linear slope, and quadratic slope for each of the trt's. That is, you will be computing 9 β 's, 3 for each trt. Assume all the data are independent (that is, we are not worried about the different pigs at this point). Use $\text{dnorm}(0, 0.0001)$ priors for the coefficients of the model (the β 's), and a $\text{dgamma}(1.1, 1)$ as the prior for σ_{error}^2 . While you have now accounted for the different treatments, you are still ignoring that there are multiple measurements on each animal.

What is the DIC for this model?

DIC is 1977.4

```
mdl <- "
model {
  for (i in 1:360){
    weight[i] ~ dnorm(mu[i], 1/vv)
    mu[i] <- b0[trt[i]] + b1[trt[i]]*day[i] + b2[trt[i]]*day2[i]
  }

  for(i in 1:3){
    b0[i] ~ dnorm(0, 0.0001)
    b1[i] ~ dnorm(0, 0.0001)
    b2[i] ~ dnorm(0, 0.0001)
  }

  vv ~ dgamma(1.1, 0.1)
}
```



```

b1[1]    10000
b1[2]     5400
b1[3]    10000
b2[1]    10000
b2[2]     8800
b2[3]    10000
vv        9300
deviance  6700

```

For each parameter, `n.eff` is a crude measure of effective sample size, and `Rhat` is the potential scale reduction factor (at convergence, `Rhat=1`).

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

`pD = 10.2` and `DIC = 1977.3`

DIC is an estimate of expected predictive error (lower deviance is better).

Diagnostics are good

```

sims <- as.mcmc(fit7)
chains <- as.matrix(sims)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)

```

18. Repeat this analysis using `brm`. You will need to create a factor variable from `trt` (that is, let a variable be `as.factor(pig$trt)`). Use the same priors for the β 's and a `gamma(1.1,.5)` as the prior for σ_{error}^2 . Print a summary of the model. What is the looic?

looic is 1978.6

```

pig$trtf = as.factor(pig$trt)
fit8 <- brm(formula = weight ~ -1 + trtf + trtf:day + trtf:day2 , data =
pig,
            family = "gaussian",
            prior = c(set_prior("normal(0,100)", class = "b"),
                      set_prior("gamma(1.1,0.5)", class = "sigma")),
            warmup = 1000, iter = 5000, chains = 4,
            #control = list(adapt_delta = 0.98),
            save_pars = save_pars(all = TRUE))

```

```
summary(fit8)
```

```

Family: gaussian
Links: mu = identity; sigma = identity
Formula: weight ~ -1 + trtf + trtf:day + trtf:day2
Data: pig (Number of observations: 360)
Samples: 4 chains, each with iter = 5000; warmup = 1000; thin = 1;
         total post-warmup samples = 16000

```


Population-Level Effects:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
trtf1	14.91	0.75	13.46	16.38	1.00	9090	10205
trtf2	14.97	0.76	13.48	16.45	1.00	9245	10474
trtf3	14.36	0.75	12.88	15.84	1.00	8936	10113
trtf1:day	1.72	0.12	1.49	1.95	1.00	8112	9099
trtf2:day	2.01	0.12	1.78	2.24	1.00	7920	9231
trtf3:day	1.42	0.12	1.19	1.66	1.00	7623	8821
trtf1:day2	-0.03	0.00	-0.04	-0.02	1.00	8999	9997
trtf2:day2	-0.04	0.00	-0.05	-0.03	1.00	8704	10182
trtf3:day2	-0.03	0.00	-0.03	-0.02	1.00	8326	9337

Family Specific Parameters:

	Estimate	Est.Error	l-95% CI	u-95% CI	Rhat	Bulk_ESS	Tail_ESS
sigma	3.72	0.14	3.45	4.00	1.00	13182	10252

Samples were drawn using sampling(NUTS). For each parameter, Bulk_ESS and Tail_ESS are effective sample size measures, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat = 1).

Convergence and ESS are good

```
chains <- as.matrix(fit8)
dim(chains)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)
```

looic

```
loo(fit8)
```

Computed from 16000 by 360 log-likelihood matrix

	Estimate	SE
elpd_loo	-989.4	15.5
p_loo	11.2	1.3
looic	1978.7	30.9

Monte Carlo SE of elpd_loo is 0.0.

All Pareto k estimates are good ($k < 0.5$).
See help('pareto-k-diagnostic') for details.

19. Now we are going to take into account that there are multiple measures on each animal to properly account for this information when we make inference. Using JAGS, create a term for a random deviation from the intercept for the intercepts only (this is sometimes referred to as a random coefficients approach). That is, we are assuming each pig is a random draw from the population of all pigs. You will need to account for the numbering of the pigs in the data set, since pig 1 in trt1 is not the same as pig 1 in trt 2. In this model you will be estimating 60 intercepts (20 pigs in each of the three

treatments). Use a $\text{gamma}(1.1, 1)$ prior for $\sigma_{intercepts}^2$. You will also need a prior for all the terms that are deviations from the overall intercept. Use a normal with a mean of 0, and a precision that is $1/\sigma_{intercepts}^2$. What is the DIC?

DIC is 1696.2

```
pig$newpig <- rep(1:60, each = 6)

mdl <- "
  model {

    for (i in 1:360){
      weight[i] ~ dnorm(mu[i], 1/vv)
      mu[i] <- b0[trt[i]] + b1[trt[i]]*day[i] + b2[trt[i]]*day2[i] +
u0[newpig[i]]
    }

    for(i in 1:3){
      b0[i] ~ dnorm(0, 0.0001)
      b1[i] ~ dnorm(0, 0.0001)
      b2[i] ~ dnorm(0, 0.0001)
    }

    for(i in 1:60){
      u0[i] ~ dnorm(0, 1/vvint)
    }

    vvint ~ dgamma(1.1, 0.1)
    vv ~ dgamma(1.1, 0.1)
  }
"

newpig <- pig$newpig
trt = pig$trt
day = pig$day
day2 = pig$day2
weight = pig$weight

writeLines(mdl, 'fit9.txt')

data.jags <- c('weight', 'day', 'day2', 'trt', 'newpig')
parms <- c('b0', 'b1', 'b2', 'u0', 'vv', 'vvint')

fit9 <- jags(data= data.jags, parameters.to.save = parms,
             model.file = 'fit9.txt', inits = NULL,
             n.iter = 8000, n.thin = 5, n.burnin = 2000,
             n.chains = 5)
```

```

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 360
  Unobserved stochastic nodes: 71
  Total graph size: 2274

```

```

Initializing model

```

```

fit9

```

```

Inference for Bugs model at "fit9.txt", fit using jags,
  5 chains, each with 8000 iterations (first 2000 discarded), n.thin = 5
  n.sims = 6000 iterations saved

```

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat
b0[1]	14.898	0.821	13.278	14.353	14.894	15.461	16.519	1.001
b0[2]	14.968	0.820	13.363	14.414	14.981	15.518	16.566	1.001
b0[3]	14.365	0.801	12.787	13.815	14.386	14.914	15.910	1.001
b1[1]	1.724	0.073	1.583	1.674	1.725	1.774	1.868	1.001
b1[2]	2.011	0.073	1.871	1.960	2.011	2.061	2.152	1.002
b1[3]	1.423	0.073	1.283	1.374	1.425	1.472	1.566	1.001
b2[1]	-0.029	0.002	-0.034	-0.031	-0.029	-0.027	-0.024	1.001
b2[2]	-0.040	0.002	-0.045	-0.042	-0.040	-0.039	-0.036	1.002
b2[3]	-0.025	0.002	-0.030	-0.027	-0.025	-0.024	-0.021	1.001
u0[1]	-3.576	1.094	-5.708	-4.315	-3.566	-2.838	-1.444	1.001
u0[2]	1.733	1.117	-0.469	0.995	1.723	2.480	3.952	1.001
u0[3]	-4.090	1.099	-6.218	-4.817	-4.104	-3.375	-1.949	1.001
u0[4]	1.692	1.099	-0.488	0.953	1.692	2.421	3.790	1.001
u0[5]	5.016	1.114	2.817	4.260	5.027	5.788	7.197	1.002
u0[6]	0.967	1.105	-1.159	0.217	0.963	1.722	3.106	1.002
u0[7]	0.298	1.110	-1.901	-0.435	0.302	1.038	2.510	1.001
u0[8]	-1.753	1.108	-3.952	-2.504	-1.748	-1.012	0.420	1.001
u0[9]	-2.993	1.088	-5.200	-3.713	-2.989	-2.245	-0.930	1.001
u0[10]	-1.328	1.095	-3.429	-2.069	-1.333	-0.586	0.817	1.001
u0[11]	-1.979	1.103	-4.109	-2.729	-1.981	-1.241	0.209	1.001
u0[12]	1.527	1.110	-0.632	0.786	1.516	2.254	3.743	1.001
u0[13]	1.276	1.110	-0.869	0.526	1.284	2.014	3.463	1.001
u0[14]	-2.243	1.108	-4.467	-2.962	-2.243	-1.502	-0.110	1.001
u0[15]	4.099	1.095	1.977	3.362	4.091	4.828	6.264	1.001
u0[16]	6.148	1.126	4.001	5.367	6.127	6.898	8.371	1.001
u0[17]	-0.767	1.117	-2.986	-1.525	-0.748	-0.006	1.332	1.001
u0[18]	-2.547	1.107	-4.738	-3.285	-2.562	-1.801	-0.338	1.001
u0[19]	0.237	1.109	-1.922	-0.502	0.224	0.962	2.432	1.001
u0[20]	-1.726	1.095	-3.841	-2.458	-1.747	-0.980	0.390	1.001
u0[21]	1.593	1.082	-0.560	0.867	1.592	2.312	3.711	1.001
u0[22]	-2.463	1.099	-4.634	-3.184	-2.460	-1.739	-0.300	1.001
u0[23]	-3.352	1.108	-5.514	-4.092	-3.354	-2.590	-1.177	1.001
u0[24]	0.643	1.094	-1.507	-0.096	0.653	1.364	2.779	1.001
u0[25]	6.219	1.108	4.049	5.461	6.208	6.982	8.377	1.001

u0[26]	-0.426	1.095	-2.610	-1.149	-0.433	0.317	1.713	1.001
u0[27]	4.053	1.109	1.890	3.303	4.048	4.784	6.236	1.001
u0[28]	1.416	1.091	-0.717	0.681	1.406	2.156	3.570	1.001
u0[29]	-3.591	1.107	-5.729	-4.337	-3.586	-2.845	-1.388	1.001
u0[30]	-2.642	1.107	-4.757	-3.404	-2.646	-1.883	-0.482	1.001
u0[31]	1.797	1.101	-0.317	1.042	1.782	2.539	3.971	1.001
u0[32]	3.324	1.121	1.166	2.563	3.327	4.076	5.525	1.001
u0[33]	1.965	1.093	-0.117	1.208	1.957	2.705	4.130	1.001
u0[34]	-2.076	1.105	-4.229	-2.826	-2.058	-1.321	0.034	1.002
u0[35]	2.674	1.106	0.512	1.946	2.662	3.429	4.853	1.001
u0[36]	-3.705	1.100	-5.841	-4.456	-3.700	-2.968	-1.527	1.001
u0[37]	-1.880	1.099	-3.996	-2.623	-1.882	-1.150	0.276	1.001
u0[38]	-3.518	1.102	-5.682	-4.275	-3.512	-2.767	-1.397	1.001
u0[39]	2.672	1.093	0.508	1.948	2.660	3.397	4.800	1.001
u0[40]	-2.796	1.107	-4.960	-3.545	-2.812	-2.050	-0.596	1.001
u0[41]	0.647	1.092	-1.473	-0.082	0.654	1.376	2.834	1.001
u0[42]	-1.767	1.096	-3.932	-2.495	-1.758	-1.016	0.369	1.001
u0[43]	-4.827	1.096	-6.974	-5.587	-4.825	-4.064	-2.713	1.001
u0[44]	-3.760	1.093	-5.862	-4.511	-3.788	-3.029	-1.576	1.001
u0[45]	0.597	1.097	-1.587	-0.132	0.594	1.337	2.715	1.001
u0[46]	5.018	1.090	2.897	4.276	4.999	5.737	7.147	1.001
u0[47]	0.688	1.090	-1.455	-0.041	0.698	1.445	2.791	1.001
u0[48]	-2.065	1.073	-4.181	-2.774	-2.056	-1.353	0.013	1.002
u0[49]	0.957	1.089	-1.181	0.232	0.945	1.691	3.093	1.001
u0[50]	1.305	1.096	-0.859	0.566	1.313	2.052	3.454	1.001
u0[51]	-3.020	1.109	-5.163	-3.784	-3.017	-2.255	-0.813	1.001
u0[52]	1.603	1.096	-0.546	0.869	1.584	2.347	3.787	1.001
u0[53]	-0.017	1.089	-2.155	-0.764	-0.014	0.723	2.063	1.001
u0[54]	3.511	1.078	1.354	2.777	3.526	4.227	5.649	1.001
u0[55]	2.233	1.085	0.147	1.496	2.227	2.963	4.388	1.001
u0[56]	-2.155	1.101	-4.302	-2.895	-2.143	-1.403	-0.011	1.001
u0[57]	-2.569	1.097	-4.739	-3.285	-2.581	-1.836	-0.401	1.001
u0[58]	-0.649	1.104	-2.804	-1.406	-0.637	0.099	1.515	1.001
u0[59]	4.303	1.096	2.187	3.553	4.284	5.059	6.481	1.001
u0[60]	-0.296	1.105	-2.398	-1.066	-0.301	0.442	1.887	1.001
vv	5.243	0.437	4.444	4.935	5.221	5.527	6.147	1.001
vvint	9.174	1.907	6.139	7.827	8.947	10.220	13.607	1.001
deviance	1615.463	12.810	1592.802	1606.477	1614.755	1623.356	1642.846	1.001
n.eff								
b0[1]	6000							
b0[2]	5700							
b0[3]	6000							
b1[1]	6000							
b1[2]	2500							
b1[3]	6000							
b2[1]	6000							
b2[2]	2800							
b2[3]	6000							
u0[1]	3300							
u0[2]	4300							

u0[3]	6000
u0[4]	6000
u0[5]	1700
u0[6]	2900
u0[7]	6000
u0[8]	5800
u0[9]	5500
u0[10]	3500
u0[11]	4500
u0[12]	6000
u0[13]	3300
u0[14]	6000
u0[15]	6000
u0[16]	6000
u0[17]	6000
u0[18]	6000
u0[19]	6000
u0[20]	6000
u0[21]	4800
u0[22]	6000
u0[23]	6000
u0[24]	4900
u0[25]	5500
u0[26]	6000
u0[27]	5900
u0[28]	6000
u0[29]	6000
u0[30]	6000
u0[31]	6000
u0[32]	4200
u0[33]	6000
u0[34]	2800
u0[35]	6000
u0[36]	3800
u0[37]	6000
u0[38]	6000
u0[39]	6000
u0[40]	6000
u0[41]	4900
u0[42]	6000
u0[43]	6000
u0[44]	4400
u0[45]	6000
u0[46]	6000
u0[47]	6000
u0[48]	2200
u0[49]	6000
u0[50]	6000
u0[51]	3300
u0[52]	6000

```

u0[53]    5300
u0[54]    6000
u0[55]    4100
u0[56]    6000
u0[57]    4200
u0[58]    6000
u0[59]    6000
u0[60]    5800
vv        6000
vvint     6000
deviance  5500

```

For each parameter, n.eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor (at convergence, Rhat=1).

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

$pD = 82.0$ and $DIC = 1697.5$

DIC is an estimate of expected predictive error (lower deviance is better).

Diagnostics are good

```

sims <- as.mcmc(fit9)
chains <- as.matrix(sims)
sims <- as.mcmc(chains)
raftery.diag(sims)
effectiveSize(sims)

```

20. Continuing in JAGS, add a random coefficient term for the deviations from the linear term in the model. You will now have random coefficients for both the intercepts and the linear term (slopes). You can use a $\text{gamma}(1.1, 1)$ prior for σ_{linear}^2 . You will also need a prior for all the terms that are deviations from the overall linear term (slopes). Use a normal with a mean of 0, and a precision that is $1/\sigma_{linear}^2$. What is the DIC?

DIC is 1451.0

```

mdl <- "
model {

  for (i in 1:360){
    weight[i] ~ dnorm(mu[i], 1/vv)
    mu[i] <- b0[trt[i]] + b1[trt[i]]*day[i] + b2[trt[i]]*day2[i] +
    u0[newpig[i]] + u1[newpig[i]]*day[i]
  }

  for(i in 1:3){
    b0[i] ~ dnorm(0, 0.0001)
    b1[i] ~ dnorm(0, 0.0001)
    b2[i] ~ dnorm(0, 0.0001)
  }

  for(i in 1:60){

```

```

    u0[i] ~ dnorm(0,1/vvint)
    u1[i] ~ dnorm(0, 1/vvslp)
  }

  vvslp ~ dgamma(1.1, 0.1)
  vvint ~ dgamma(1.1, 0.1)
  vv ~ dgamma(1.1, 0.1)
}

"
newpig <- pig$newpig
trt = pig$trt
day = pig$day
day2 = pig$day2
weight = pig$weight

writeLines(mdl, 'fit10.txt')

data.jags <- c('weight', 'day', 'day2', 'trt', 'newpig')
parms <- c('b0' , 'b1', 'b2','u0','u1', 'vv', 'vvint', 'vvslp')

fit10 <- jags(data= data.jags, parameters.to.save = parms,
              model.file = 'fit10.txt', inits = NULL,
              n.iter = 20000, n.thin = 5, n.burnin = 2000,
              n.chains = 5)

```

```

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 360
  Unobserved stochastic nodes: 132
  Total graph size: 2696

```

Initializing model

fit10

Inference for Bugs model at "fit10.txt", fit using jags,
 5 chains, each with 20000 iterations (first 2000 discarded), n.thin = 5
 n.sims = 18000 iterations saved

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat
b0[1]	14.899	0.386	14.148	14.639	14.898	15.156	15.677	1.001
b0[2]	14.954	0.385	14.205	14.698	14.952	15.211	15.711	1.001
b0[3]	14.356	0.387	13.602	14.098	14.355	14.616	15.123	1.001
b1[1]	1.725	0.061	1.603	1.684	1.724	1.766	1.844	1.001
b1[2]	2.010	0.061	1.889	1.969	2.011	2.051	2.131	1.001
b1[3]	1.424	0.062	1.303	1.383	1.424	1.465	1.546	1.001
b2[1]	-0.029	0.001	-0.032	-0.030	-0.029	-0.028	-0.026	1.001

b2[2]	-0.040	0.002	-0.043	-0.041	-0.040	-0.039	-0.037	1.001
b2[3]	-0.025	0.002	-0.028	-0.026	-0.025	-0.024	-0.022	1.001
u0[1]	-0.624	0.773	-2.200	-1.134	-0.610	-0.100	0.851	1.001
u0[2]	-0.037	0.763	-1.529	-0.544	-0.033	0.474	1.471	1.001
u0[3]	-0.712	0.772	-2.245	-1.223	-0.707	-0.193	0.788	1.001
u0[4]	0.967	0.790	-0.545	0.422	0.954	1.494	2.568	1.001
u0[5]	0.948	0.781	-0.538	0.413	0.936	1.462	2.516	1.001
u0[6]	0.398	0.768	-1.108	-0.118	0.396	0.912	1.929	1.001
u0[7]	-1.181	0.799	-2.798	-1.713	-1.166	-0.629	0.325	1.001
u0[8]	0.168	0.757	-1.294	-0.337	0.167	0.668	1.654	1.001
u0[9]	0.066	0.766	-1.443	-0.441	0.071	0.568	1.584	1.001
u0[10]	-0.773	0.777	-2.334	-1.285	-0.764	-0.249	0.717	1.001
u0[11]	0.094	0.764	-1.421	-0.417	0.092	0.601	1.596	1.001
u0[12]	-0.175	0.761	-1.693	-0.682	-0.160	0.333	1.305	1.001
u0[13]	0.287	0.764	-1.185	-0.237	0.283	0.794	1.805	1.001
u0[14]	-0.183	0.763	-1.702	-0.687	-0.177	0.324	1.325	1.001
u0[15]	1.270	0.799	-0.227	0.717	1.248	1.798	2.899	1.001
u0[16]	1.245	0.800	-0.273	0.698	1.225	1.773	2.870	1.001
u0[17]	-0.279	0.761	-1.775	-0.782	-0.275	0.225	1.213	1.001
u0[18]	-0.357	0.763	-1.893	-0.860	-0.350	0.149	1.137	1.001
u0[19]	0.073	0.766	-1.432	-0.434	0.074	0.580	1.599	1.001
u0[20]	-1.206	0.801	-2.822	-1.733	-1.190	-0.660	0.345	1.001
u0[21]	0.364	0.762	-1.127	-0.148	0.363	0.868	1.859	1.001
u0[22]	0.400	0.765	-1.082	-0.106	0.385	0.904	1.932	1.001
u0[23]	-0.592	0.770	-2.131	-1.101	-0.586	-0.065	0.877	1.001
u0[24]	0.029	0.761	-1.481	-0.475	0.027	0.531	1.534	1.001
u0[25]	1.336	0.807	-0.187	0.783	1.319	1.862	2.981	1.001
u0[26]	-0.716	0.773	-2.278	-1.227	-0.707	-0.196	0.767	1.001
u0[27]	0.511	0.771	-0.996	-0.007	0.503	1.021	2.042	1.001
u0[28]	0.340	0.763	-1.140	-0.171	0.329	0.838	1.870	1.001
u0[29]	-0.880	0.774	-2.452	-1.395	-0.867	-0.351	0.598	1.001
u0[30]	-0.271	0.763	-1.787	-0.777	-0.271	0.235	1.219	1.001
u0[31]	-0.358	0.767	-1.888	-0.863	-0.357	0.151	1.159	1.001
u0[32]	0.238	0.760	-1.245	-0.273	0.232	0.744	1.734	1.001
u0[33]	0.671	0.772	-0.807	0.152	0.663	1.178	2.212	1.001
u0[34]	-0.377	0.759	-1.894	-0.883	-0.365	0.128	1.104	1.001
u0[35]	0.584	0.767	-0.891	0.069	0.567	1.092	2.126	1.001
u0[36]	-0.902	0.780	-2.493	-1.402	-0.887	-0.374	0.589	1.001
u0[37]	0.092	0.768	-1.395	-0.422	0.087	0.605	1.620	1.001
u0[38]	-0.820	0.779	-2.397	-1.333	-0.801	-0.294	0.671	1.001
u0[39]	0.729	0.777	-0.751	0.197	0.712	1.248	2.291	1.001
u0[40]	-0.355	0.762	-1.865	-0.856	-0.341	0.157	1.116	1.001
u0[41]	1.141	0.789	-0.362	0.601	1.120	1.667	2.738	1.001
u0[42]	-0.702	0.775	-2.275	-1.207	-0.689	-0.175	0.781	1.001
u0[43]	-1.075	0.794	-2.696	-1.597	-1.055	-0.532	0.424	1.001
u0[44]	-1.759	0.836	-3.466	-2.307	-1.741	-1.182	-0.180	1.001
u0[45]	1.211	0.795	-0.309	0.672	1.190	1.730	2.841	1.001
u0[46]	0.922	0.781	-0.559	0.388	0.903	1.436	2.511	1.001
u0[47]	0.569	0.773	-0.924	0.052	0.555	1.076	2.118	1.001
u0[48]	-0.067	0.764	-1.573	-0.571	-0.063	0.441	1.427	1.001

u0[49]	1.041	0.795	-0.477	0.492	1.025	1.564	2.649	1.001
u0[50]	0.364	0.757	-1.109	-0.141	0.362	0.859	1.878	1.001
u0[51]	-1.124	0.791	-2.717	-1.647	-1.095	-0.587	0.376	1.001
u0[52]	0.226	0.768	-1.276	-0.290	0.221	0.734	1.758	1.001
u0[53]	-0.128	0.761	-1.611	-0.629	-0.132	0.376	1.349	1.001
u0[54]	0.490	0.774	-0.994	-0.030	0.478	0.993	2.041	1.001
u0[55]	-0.421	0.767	-1.964	-0.926	-0.416	0.091	1.091	1.001
u0[56]	-1.617	0.822	-3.289	-2.165	-1.591	-1.050	-0.084	1.001
u0[57]	0.045	0.765	-1.463	-0.455	0.045	0.555	1.545	1.001
u0[58]	-0.551	0.769	-2.078	-1.062	-0.547	-0.035	0.936	1.001
u0[59]	0.971	0.784	-0.525	0.437	0.956	1.487	2.558	1.001
u0[60]	0.426	0.768	-1.064	-0.091	0.414	0.927	1.982	1.001
u1[1]	-0.210	0.060	-0.327	-0.250	-0.209	-0.170	-0.093	1.001
u1[2]	0.129	0.060	0.012	0.088	0.129	0.170	0.246	1.001
u1[3]	-0.237	0.060	-0.354	-0.278	-0.237	-0.198	-0.118	1.001
u1[4]	0.036	0.060	-0.082	-0.004	0.037	0.077	0.153	1.001
u1[5]	0.284	0.060	0.164	0.243	0.284	0.324	0.402	1.001
u1[6]	0.037	0.060	-0.082	-0.004	0.036	0.078	0.154	1.001
u1[7]	0.125	0.060	0.006	0.084	0.124	0.165	0.243	1.001
u1[8]	-0.144	0.059	-0.261	-0.184	-0.143	-0.104	-0.027	1.001
u1[9]	-0.225	0.060	-0.342	-0.265	-0.224	-0.184	-0.108	1.001
u1[10]	-0.030	0.060	-0.147	-0.071	-0.030	0.010	0.086	1.001
u1[11]	-0.153	0.060	-0.269	-0.193	-0.153	-0.112	-0.036	1.001
u1[12]	0.126	0.059	0.012	0.086	0.126	0.166	0.242	1.001
u1[13]	0.068	0.060	-0.047	0.028	0.068	0.108	0.186	1.001
u1[14]	-0.147	0.060	-0.264	-0.187	-0.147	-0.107	-0.030	1.001
u1[15]	0.190	0.060	0.070	0.149	0.190	0.230	0.307	1.001
u1[16]	0.341	0.061	0.222	0.300	0.341	0.381	0.462	1.001
u1[17]	-0.031	0.059	-0.147	-0.071	-0.032	0.009	0.085	1.001
u1[18]	-0.154	0.060	-0.271	-0.194	-0.154	-0.114	-0.037	1.001
u1[19]	0.009	0.060	-0.108	-0.031	0.009	0.049	0.126	1.001
u1[20]	-0.020	0.061	-0.139	-0.061	-0.020	0.022	0.100	1.001
u1[21]	0.084	0.060	-0.032	0.045	0.084	0.124	0.201	1.001
u1[22]	-0.214	0.060	-0.331	-0.254	-0.214	-0.173	-0.097	1.001
u1[23]	-0.194	0.059	-0.311	-0.234	-0.195	-0.155	-0.076	1.001
u1[24]	0.046	0.059	-0.070	0.006	0.046	0.086	0.162	1.001
u1[25]	0.337	0.061	0.218	0.295	0.337	0.377	0.458	1.001
u1[26]	0.031	0.060	-0.084	-0.009	0.031	0.072	0.149	1.001
u1[27]	0.253	0.060	0.137	0.212	0.252	0.293	0.370	1.001
u1[28]	0.077	0.060	-0.039	0.036	0.077	0.117	0.194	1.001
u1[29]	-0.187	0.060	-0.303	-0.227	-0.187	-0.147	-0.069	1.001
u1[30]	-0.171	0.060	-0.288	-0.211	-0.171	-0.130	-0.053	1.001
u1[31]	0.164	0.060	0.046	0.124	0.164	0.203	0.281	1.001
u1[32]	0.223	0.059	0.107	0.184	0.223	0.263	0.339	1.001
u1[33]	0.086	0.060	-0.031	0.046	0.086	0.126	0.205	1.001
u1[34]	-0.119	0.059	-0.237	-0.159	-0.119	-0.079	-0.004	1.001
u1[35]	0.146	0.060	0.030	0.106	0.146	0.186	0.263	1.001
u1[36]	-0.193	0.060	-0.309	-0.233	-0.193	-0.152	-0.076	1.001
u1[37]	-0.145	0.060	-0.261	-0.185	-0.145	-0.105	-0.029	1.001
u1[38]	-0.185	0.060	-0.302	-0.226	-0.185	-0.144	-0.068	1.001

u1[39]	0.131	0.059	0.016	0.091	0.131	0.171	0.248	1.001
u1[40]	-0.175	0.059	-0.291	-0.215	-0.175	-0.135	-0.059	1.001
u1[41]	-0.051	0.060	-0.171	-0.093	-0.051	-0.011	0.066	1.001
u1[42]	-0.066	0.060	-0.186	-0.106	-0.066	-0.026	0.052	1.001
u1[43]	-0.259	0.061	-0.378	-0.300	-0.259	-0.218	-0.140	1.001
u1[44]	-0.121	0.062	-0.242	-0.162	-0.121	-0.079	0.000	1.001
u1[45]	-0.061	0.061	-0.179	-0.103	-0.061	-0.021	0.057	1.001
u1[46]	0.289	0.060	0.170	0.248	0.289	0.329	0.408	1.001
u1[47]	0.000	0.060	-0.120	-0.039	0.000	0.041	0.120	1.001
u1[48]	-0.145	0.060	-0.262	-0.185	-0.145	-0.105	-0.027	1.001
u1[49]	-0.021	0.061	-0.140	-0.062	-0.022	0.019	0.098	1.001
u1[50]	0.064	0.060	-0.054	0.024	0.064	0.104	0.182	1.001
u1[51]	-0.122	0.060	-0.238	-0.162	-0.122	-0.082	-0.004	1.001
u1[52]	0.098	0.060	-0.018	0.058	0.098	0.138	0.216	1.001
u1[53]	0.010	0.060	-0.107	-0.030	0.009	0.050	0.128	1.001
u1[54]	0.215	0.060	0.101	0.175	0.215	0.255	0.334	1.001
u1[55]	0.202	0.060	0.085	0.161	0.202	0.242	0.320	1.001
u1[56]	-0.016	0.061	-0.134	-0.058	-0.016	0.025	0.108	1.001
u1[57]	-0.190	0.060	-0.307	-0.230	-0.191	-0.150	-0.073	1.001
u1[58]	0.002	0.060	-0.115	-0.038	0.001	0.042	0.121	1.001
u1[59]	0.230	0.061	0.113	0.188	0.229	0.271	0.349	1.001
u1[60]	-0.059	0.060	-0.177	-0.099	-0.059	-0.020	0.058	1.001
vv	2.154	0.193	1.808	2.016	2.144	2.276	2.563	1.001
vvint	1.233	0.460	0.491	0.907	1.180	1.498	2.278	1.001
vvslp	0.031	0.007	0.021	0.027	0.031	0.035	0.047	1.001
deviance	1294.889	17.733	1262.084	1282.569	1294.252	1306.464	1331.381	1.001

DIC info (using the rule, $pD = \text{var}(\text{deviance})/2$)

$pD = 157.2$ and $DIC = 1452.1$

DIC is an estimate of expected predictive error (lower deviance is better).

Diagnostics are good

`sims <- as.mcmc(fit10)`

`chains <- as.matrix(sims)`

`sims <- as.mcmc(chains)`

`raftery.diag(sims)`

`effectiveSize(sims)`

21. Continuing in JAGS, add a random coefficient term for the deviations from the quadratic term in the model. You will now have random coefficients for the intercepts, the linear term (slopes), and the quadratic term. You can use a $\text{gamma}(1.1, 1)$ prior for $\sigma_{\text{quadratic}}^2$. You will also need a prior for all the terms that are deviations from the overall quadratic term. Use a normal with a mean of 0, and a precision that is $1/\sigma_{\text{quadratic}}^2$. What is the DIC?

DIC is 1476.5

```

mdl <- "
  model {

    for (i in 1:360){
      weight[i] ~ dnorm(mu[i], 1/vv)
      mu[i] <- b0[trt[i]] + b1[trt[i]]*day[i] + b2[trt[i]]*day2[i] +
u0[newpig[i]] + u1[newpig[i]]*day[i] + u2[newpig[i]]*day2[i]
    }

    for(i in 1:3){
      b0[i] ~ dnorm(0, 0.0001)
      b1[i] ~ dnorm(0, 0.0001)
      b2[i] ~ dnorm(0, 0.0001)
    }

    for(i in 1:60){
      u0[i] ~ dnorm(0,1/vvint)
      u1[i] ~ dnorm(0, 1/vvslp)
      u2[i] ~ dnorm(0, 1/vvslp2)
    }

    vvslp2 ~ dgamma(1.1, 0.1)
    vvslp ~ dgamma(1.1, 0.1)
    vvint ~ dgamma(1.1, 0.1)
    vv ~ dgamma(1.1, 0.1)
  }
"

newpig <- pig$newpig
trt = pig$trt
day = pig$day
day2 = pig$day2
weight = pig$weight

writeLines(mdl, 'fit11.txt')

data.jags <- c('weight', 'day', 'day2', 'trt', 'newpig')
parms <- c('b0', 'b1', 'b2', 'u0', 'u1', 'vv', 'vvint', 'vvslp')

fit11 <- jags(data= data.jags, parameters.to.save = parms,
              model.file = 'fit11.txt', inits = NULL,
              n.iter = 30000, n.thin = 5, n.burnin = 2000,
              n.chains = 5)

Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:

```

Observed stochastic nodes: 360
Unobserved stochastic nodes: 193
Total graph size: 3118

Initializing model

fit11

Inference for Bugs model at "fit11.txt", fit using jags,
5 chains, each with 30000 iterations (first 2000 discarded), n.thin = 5
n.sims = 28000 iterations saved

	mu.vect	sd.vect	2.5%	25%	50%	75%	97.5%	Rhat
b0[1]	14.896	0.398	14.115	14.632	14.893	15.162	15.673	1.001
b0[2]	14.959	0.400	14.172	14.695	14.956	15.225	15.747	1.001
b0[3]	14.351	0.397	13.582	14.084	14.353	14.617	15.129	1.001
b1[1]	1.724	0.058	1.611	1.685	1.724	1.763	1.838	1.001
b1[2]	2.010	0.058	1.898	1.971	2.010	2.049	2.122	1.001
b1[3]	1.424	0.058	1.311	1.386	1.424	1.463	1.537	1.001
b2[1]	-0.029	0.002	-0.032	-0.030	-0.029	-0.028	-0.026	1.001
b2[2]	-0.040	0.002	-0.043	-0.041	-0.040	-0.039	-0.037	1.001
b2[3]	-0.025	0.002	-0.028	-0.026	-0.025	-0.024	-0.022	1.001
u0[1]	-0.794	0.809	-2.411	-1.336	-0.785	-0.252	0.783	1.001
u0[2]	0.025	0.791	-1.523	-0.506	0.027	0.555	1.583	1.001
u0[3]	-0.876	0.816	-2.503	-1.414	-0.860	-0.322	0.674	1.001
u0[4]	1.103	0.815	-0.448	0.545	1.089	1.641	2.739	1.001
u0[5]	1.195	0.831	-0.377	0.628	1.177	1.737	2.892	1.001
u0[6]	0.426	0.797	-1.104	-0.115	0.417	0.954	2.034	1.001
u0[7]	-1.233	0.814	-2.863	-1.778	-1.218	-0.676	0.329	1.001
u0[8]	0.100	0.793	-1.458	-0.434	0.097	0.629	1.668	1.001
u0[9]	-0.003	0.799	-1.583	-0.537	-0.001	0.530	1.561	1.001
u0[10]	-0.878	0.797	-2.462	-1.407	-0.868	-0.336	0.653	1.001
u0[11]	0.041	0.798	-1.540	-0.488	0.043	0.573	1.610	1.001
u0[12]	-0.108	0.797	-1.677	-0.637	-0.099	0.426	1.466	1.001
u0[13]	0.381	0.792	-1.163	-0.155	0.378	0.906	1.954	1.001
u0[14]	-0.327	0.797	-1.926	-0.855	-0.319	0.210	1.222	1.001
u0[15]	1.505	0.840	-0.099	0.920	1.492	2.068	3.171	1.001
u0[16]	1.534	0.856	-0.074	0.944	1.510	2.101	3.255	1.001
u0[17]	-0.307	0.795	-1.882	-0.833	-0.298	0.230	1.239	1.001
u0[18]	-0.473	0.802	-2.062	-1.006	-0.469	0.067	1.080	1.001
u0[19]	0.089	0.798	-1.459	-0.443	0.079	0.620	1.656	1.001
u0[20]	-1.365	0.822	-3.034	-1.906	-1.349	-0.806	0.198	1.001
u0[21]	0.439	0.798	-1.104	-0.094	0.429	0.967	2.037	1.001
u0[22]	0.292	0.799	-1.292	-0.240	0.289	0.821	1.870	1.001
u0[23]	-0.731	0.816	-2.367	-1.276	-0.715	-0.174	0.834	1.001
u0[24]	0.041	0.793	-1.507	-0.488	0.038	0.564	1.612	1.001
u0[25]	1.672	0.863	0.034	1.087	1.649	2.246	3.411	1.001
u0[26]	-0.771	0.803	-2.373	-1.303	-0.753	-0.233	0.766	1.001
u0[27]	0.707	0.808	-0.832	0.152	0.703	1.241	2.343	1.001
u0[28]	0.384	0.800	-1.176	-0.152	0.385	0.925	1.954	1.001
u0[29]	-1.078	0.821	-2.726	-1.628	-1.060	-0.515	0.484	1.001

u0[30]	-0.360	0.798	-1.950	-0.893	-0.347	0.172	1.200	1.001
u0[31]	-0.260	0.795	-1.832	-0.793	-0.259	0.275	1.298	1.001
u0[32]	0.405	0.803	-1.147	-0.140	0.395	0.935	1.995	1.001
u0[33]	0.750	0.809	-0.819	0.202	0.742	1.282	2.353	1.001
u0[34]	-0.482	0.791	-2.040	-1.014	-0.478	0.054	1.064	1.001
u0[35]	0.712	0.811	-0.857	0.158	0.700	1.256	2.359	1.001
u0[36]	-1.096	0.818	-2.722	-1.640	-1.085	-0.532	0.465	1.001
u0[37]	0.034	0.796	-1.533	-0.504	0.034	0.564	1.609	1.001
u0[38]	-1.007	0.819	-2.632	-1.560	-0.995	-0.447	0.568	1.001
u0[39]	0.879	0.811	-0.687	0.329	0.867	1.416	2.495	1.001
u0[40]	-0.486	0.798	-2.055	-1.016	-0.475	0.054	1.066	1.001
u0[41]	1.236	0.819	-0.334	0.676	1.218	1.770	2.886	1.001
u0[42]	-0.792	0.803	-2.393	-1.323	-0.782	-0.255	0.751	1.001
u0[43]	-1.276	0.830	-2.962	-1.821	-1.261	-0.710	0.313	1.001
u0[44]	-2.022	0.864	-3.760	-2.596	-2.003	-1.432	-0.385	1.001
u0[45]	1.322	0.817	-0.245	0.764	1.318	1.865	2.952	1.001
u0[46]	1.176	0.830	-0.399	0.613	1.156	1.720	2.852	1.001
u0[47]	0.563	0.807	-0.997	0.010	0.554	1.100	2.179	1.001
u0[48]	-0.174	0.798	-1.735	-0.707	-0.177	0.374	1.391	1.001
u0[49]	1.167	0.809	-0.370	0.615	1.156	1.697	2.812	1.001
u0[50]	0.428	0.794	-1.105	-0.113	0.424	0.962	2.001	1.001
u0[51]	-1.309	0.827	-2.970	-1.857	-1.296	-0.743	0.272	1.001
u0[52]	0.390	0.799	-1.142	-0.153	0.381	0.919	1.985	1.001
u0[53]	-0.089	0.788	-1.636	-0.615	-0.089	0.443	1.463	1.001
u0[54]	0.635	0.812	-0.931	0.082	0.628	1.167	2.267	1.001
u0[55]	-0.380	0.791	-1.929	-0.906	-0.377	0.153	1.173	1.001
u0[56]	-1.812	0.842	-3.505	-2.372	-1.791	-1.229	-0.220	1.001
u0[57]	-0.070	0.800	-1.653	-0.603	-0.064	0.465	1.497	1.001
u0[58]	-0.588	0.805	-2.190	-1.120	-0.576	-0.049	0.973	1.001
u0[59]	1.187	0.824	-0.377	0.624	1.166	1.734	2.851	1.001
u0[60]	0.437	0.796	-1.124	-0.103	0.437	0.969	1.997	1.001
u1[1]	-0.171	0.081	-0.325	-0.226	-0.173	-0.118	-0.005	1.001
u1[2]	0.105	0.078	-0.051	0.053	0.106	0.158	0.254	1.001
u1[3]	-0.206	0.081	-0.361	-0.261	-0.207	-0.153	-0.043	1.001
u1[4]	0.022	0.078	-0.132	-0.030	0.023	0.075	0.173	1.001
u1[5]	0.225	0.086	0.049	0.169	0.227	0.285	0.388	1.001
u1[6]	0.038	0.077	-0.114	-0.014	0.038	0.089	0.189	1.001
u1[7]	0.097	0.080	-0.062	0.043	0.097	0.151	0.252	1.001
u1[8]	-0.108	0.079	-0.260	-0.161	-0.109	-0.056	0.051	1.001
u1[9]	-0.198	0.080	-0.351	-0.251	-0.199	-0.145	-0.035	1.001
u1[10]	-0.019	0.078	-0.171	-0.071	-0.019	0.032	0.135	1.001
u1[11]	-0.124	0.079	-0.276	-0.178	-0.126	-0.072	0.033	1.001
u1[12]	0.089	0.080	-0.071	0.037	0.092	0.143	0.244	1.001
u1[13]	0.046	0.078	-0.112	-0.005	0.047	0.098	0.198	1.001
u1[14]	-0.083	0.084	-0.241	-0.141	-0.085	-0.028	0.089	1.001
u1[15]	0.150	0.082	-0.017	0.096	0.151	0.205	0.305	1.001
u1[16]	0.277	0.090	0.092	0.218	0.281	0.339	0.442	1.002
u1[17]	-0.036	0.077	-0.190	-0.087	-0.036	0.015	0.114	1.001
u1[18]	-0.117	0.081	-0.271	-0.171	-0.118	-0.064	0.045	1.001
u1[19]	0.014	0.077	-0.139	-0.038	0.014	0.065	0.165	1.001

u1[20]	-0.002	0.078	-0.153	-0.055	-0.002	0.050	0.155	1.001
u1[21]	0.070	0.078	-0.082	0.018	0.070	0.123	0.222	1.001
u1[22]	-0.145	0.085	-0.306	-0.204	-0.148	-0.089	0.028	1.002
u1[23]	-0.173	0.080	-0.328	-0.226	-0.174	-0.119	-0.012	1.001
u1[24]	0.045	0.077	-0.108	-0.006	0.044	0.095	0.198	1.001
u1[25]	0.261	0.091	0.073	0.201	0.266	0.324	0.430	1.001
u1[26]	0.026	0.077	-0.126	-0.025	0.026	0.078	0.178	1.001
u1[27]	0.192	0.085	0.019	0.137	0.196	0.250	0.353	1.001
u1[28]	0.083	0.077	-0.066	0.031	0.082	0.135	0.237	1.001
u1[29]	-0.131	0.084	-0.290	-0.188	-0.133	-0.075	0.039	1.001
u1[30]	-0.160	0.079	-0.312	-0.212	-0.161	-0.107	-0.002	1.001
u1[31]	0.100	0.084	-0.072	0.046	0.103	0.157	0.257	1.001
u1[32]	0.157	0.086	-0.020	0.101	0.161	0.217	0.320	1.001
u1[33]	0.088	0.078	-0.063	0.036	0.088	0.140	0.242	1.001
u1[34]	-0.088	0.078	-0.239	-0.141	-0.089	-0.037	0.069	1.001
u1[35]	0.122	0.079	-0.037	0.070	0.122	0.175	0.275	1.001
u1[36]	-0.148	0.082	-0.304	-0.203	-0.150	-0.094	0.019	1.001
u1[37]	-0.120	0.079	-0.271	-0.172	-0.120	-0.067	0.040	1.001
u1[38]	-0.145	0.082	-0.301	-0.200	-0.146	-0.091	0.017	1.001
u1[39]	0.105	0.079	-0.053	0.052	0.106	0.158	0.257	1.001
u1[40]	-0.142	0.080	-0.297	-0.197	-0.143	-0.089	0.018	1.001
u1[41]	-0.043	0.078	-0.195	-0.095	-0.044	0.009	0.109	1.001
u1[42]	-0.058	0.078	-0.209	-0.109	-0.058	-0.006	0.097	1.001
u1[43]	-0.227	0.083	-0.385	-0.283	-0.229	-0.173	-0.060	1.001
u1[44]	-0.078	0.082	-0.235	-0.134	-0.080	-0.023	0.090	1.001
u1[45]	-0.054	0.078	-0.207	-0.105	-0.054	-0.002	0.100	1.001
u1[46]	0.222	0.088	0.041	0.163	0.226	0.284	0.387	1.001
u1[47]	0.043	0.079	-0.108	-0.011	0.040	0.095	0.205	1.001
u1[48]	-0.105	0.080	-0.259	-0.159	-0.106	-0.053	0.058	1.001
u1[49]	-0.030	0.078	-0.184	-0.080	-0.030	0.022	0.125	1.001
u1[50]	0.057	0.078	-0.097	0.005	0.057	0.109	0.210	1.001
u1[51]	-0.095	0.079	-0.248	-0.148	-0.096	-0.042	0.064	1.001
u1[52]	0.023	0.085	-0.152	-0.033	0.026	0.082	0.182	1.001
u1[53]	-0.019	0.078	-0.178	-0.071	-0.018	0.034	0.132	1.001
u1[54]	0.188	0.081	0.027	0.135	0.188	0.241	0.344	1.001
u1[55]	0.180	0.079	0.022	0.128	0.180	0.232	0.336	1.001
u1[56]	0.000	0.079	-0.152	-0.054	-0.001	0.052	0.157	1.001
u1[57]	-0.135	0.084	-0.293	-0.192	-0.137	-0.080	0.038	1.001
u1[58]	-0.013	0.078	-0.169	-0.065	-0.013	0.039	0.138	1.001
u1[59]	0.188	0.083	0.020	0.133	0.189	0.244	0.346	1.001
u1[60]	-0.044	0.077	-0.194	-0.096	-0.045	0.007	0.108	1.001
vv	2.075	0.195	1.731	1.939	2.063	2.198	2.490	1.001
vvint	1.503	0.540	0.633	1.125	1.438	1.808	2.738	1.001
vvslp	0.024	0.007	0.011	0.019	0.024	0.029	0.041	1.002
deviance	1281.125	19.963	1243.238	1267.563	1280.745	1294.475	1321.159	1.001

For each parameter, n.eff is a crude measure of effective sample size,
and Rhat is the potential scale reduction factor (at convergence, Rhat=1).

```
DIC info (using the rule,  $pD = \text{var}(\text{deviance})/2$ )  
pD = 199.2 and DIC = 1480.3  
DIC is an estimate of expected predictive error (lower deviance is better).
```

```
sims <- as.mcmc(fit10)  
chains <- as.matrix(sims)  
sims <- as.mcmc(chains)  
raftery.diag(sims)  
effectiveSize(sims)
```

22. Using DIC, which of the models that allow inference over all pigs would you choose?

Using DIC I would choose the model from question 20. This model just has random coefficients for the intercept and the linear day term. Adding a random effect for pig on the quadratic day term does not improve the model.

Problems 23 and 24 are not required. If completed they will be worth bonus points. 10 points for problem 23, and 5 points for problem 24.

23. Since brm does not do well with these types of models, now you will do the work in Stan. Create a hierarchical model with hierarchical terms for both the intercepts and linear terms (slopes), but not for the quadratic term. This is a hierarchical model that is much like the random coefficients model in number 20. Use normal priors with means of 0 and standard deviations of 100 where appropriate, and $\text{gamma}(1.1, .5)$ priors for the variance terms. What are means of the chains for the overall intercept in treatment 1, the overall linear term coefficient (slope) in treatment 1, and the overall quadratic term coefficient in treatment 1?

Intercept for treatment 1 is 18.5, Linear term for treatment 1 is 0.85

```
model <- "  
  
data {  
  int <lower = 1> N;  
  int q;  
  real weight[N];  
  int day[N];  
  int trt[N];  
}  
  
parameters {  
  real alpha[q];  
  real beta[q];  
  real mu_alpha;  
  real mu_beta;  
  real <lower = 0> serr;  
  real <lower = 0> salpha;  
  real <lower = 0> sbeta;  
}
```

```

model {
  mu_alpha ~ normal(0, 100);
  mu_beta ~ normal(0, 100);
  alpha ~ normal(mu_alpha, salpha);
  beta ~ normal(mu_beta, sbeta);
  serr ~ gamma(1.1, 0.5);
  salpha ~ gamma(1.1, 0.5);
  sbeta ~ gamma(1.1, .5);
  for(i in 1:N){
    weight[i] ~ normal(alpha[trt[i]] + beta[trt[i]] * day[i], serr);
  }
}

generated quantities {
  vector[N] log_lik;
  real s2error;
  real s2int;
  real s2slope;
  s2error = serr*serr;
  s2int = salpha*salpha;
  s2slope = sbeta*sbeta;
  for (i in 1:N) log_lik[i] = normal_lpdf(weight[i] | alpha[trt[i]] +
beta[trt[i]]*day[i], serr);
}

"

writeLines(model, 'hier.stan')
trt = pig$trt
day = pig$day
weight = pig$weight
N <- 360
q <- 3
hier_dat <- list(N=N, weight = weight, trt = trt, day = day, q=q)
fit12 <- stan(file = "hier.stan", data = hier_dat, iter = 11000,
             warmup = 1000, chains = 4, thin = 2)

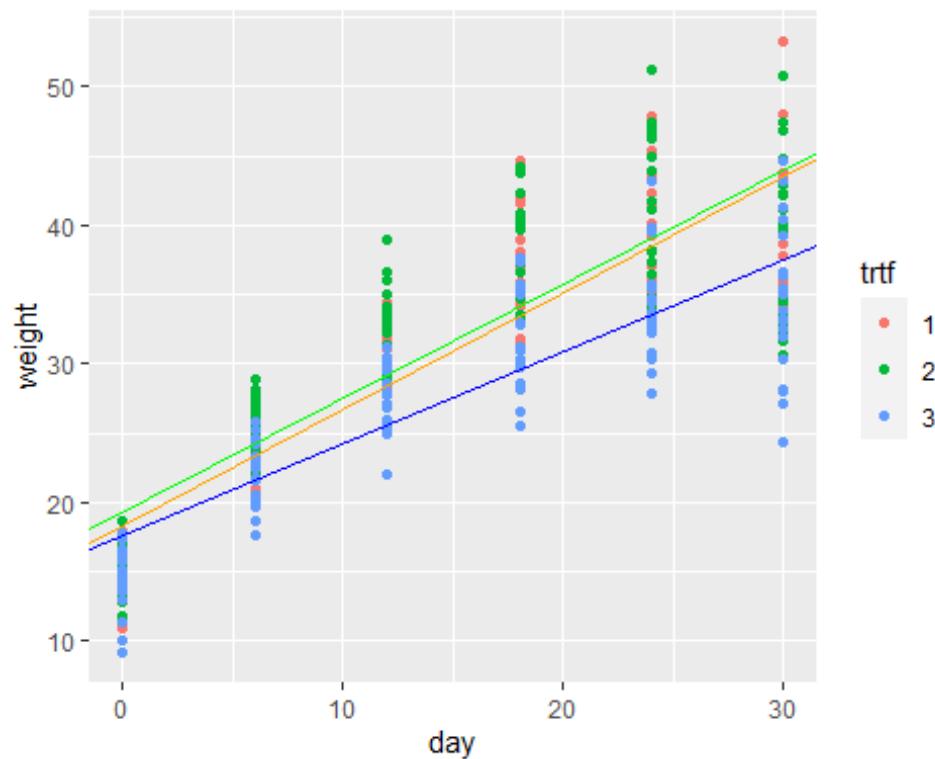
chains <- as.matrix(fit12)
dim(chains)

[1] 20000 375
sims <- as.mcmc(chains)

```



```
ggplot(data = pig, mapping = aes(x = day, y = weight, col = trtf)) +
  geom_point() +
  geom_abline(intercept = 18.4, slope = 0.84, col = "orange") +
  geom_abline(intercept = 19.38, slope = 0.82, col = "green") +
  geom_abline(intercept = 17.664, slope = 0.663, col = "blue")
```



24. Compute the waic for the model in problem 23.

waic is 2145.3

```
llfit <- extract_log_lik(fit12, parameter_name = "log_lik",
  merge_chains = TRUE)
waic(llfit)
```

Computed from 20000 by 360 log-likelihood matrix

	Estimate	SE
elpd_waic	-1072.7	14.0
p_waic	7.7	0.8
waic	2145.4	28.0