

# Table of Common Distributions

## Discrete Distributions

### Bernoulli( $p$ )

pmf	$P(X = x p) = p^x(1 - p)^{1-x}; \quad x = 0, 1; \quad 0 \leq p \leq 1$
mean and variance	$EX = p, \quad \text{Var } X = p(1 - p)$
mgf	$M_X(t) = (1 - p) + pe^t$

### Binomial( $n, p$ )

pmf	$P(X = x n, p) = \binom{n}{x}p^x(1 - p)^{n-x}; \quad x = 0, 1, 2, \dots, n; \quad 0 \leq p \leq 1$
mean and variance	$EX = np, \quad \text{Var } X = np(1 - p)$
mgf	$M_X(t) = [pe^t + (1 - p)]^n$
notes	Related to Binomial Theorem (Theorem 3.2.2). The <i>multinomial distribution</i> (Definition 4.6.2) is a multivariate version of the binomial distribution.

### Discrete uniform

pmf	$P(X = x N) = \frac{1}{N}; \quad x = 1, 2, \dots, N; \quad N = 1, 2, \dots$
mean and variance	$EX = \frac{N+1}{2}, \quad \text{Var } X = \frac{(N+1)(N-1)}{12}$
mgf	$M_X(t) = \frac{1}{N} \sum_{i=1}^N e^{it}$

### Geometric( $p$ )

pmf	$P(X = x p) = p(1 - p)^{x-1}; \quad x = 1, 2, \dots; \quad 0 \leq p \leq 1$
mean and variance	$EX = \frac{1}{p}, \quad \text{Var } X = \frac{1-p}{p^2}$

*mgf*  $M_X(t) = \frac{pe^t}{1-(1-p)e^t}$ ,  $t < -\log(1-p)$

*notes*  $Y = X - 1$  is negative binomial(1,  $p$ ). The distribution is memoryless:  
 $P(X > s | X > t) = P(X > s - t)$ .

### Hypergeometric

*pdf*  $P(X = x | N, M, K) = \frac{\binom{M}{x} \binom{N-M}{K-x}}{\binom{N}{K}}$ ;  $x = 0, 1, 2, \dots, K$ ;

$$M - (N - K) \leq x \leq M; \quad N, M, K \geq 0$$

*mean and variance*  $EX = \frac{KM}{N}$ ,  $\text{Var } X = \frac{KM}{N} \frac{(N-M)(N-K)}{N(N-1)}$

*notes* If  $K \ll M$  and  $N$ , the range  $x = 0, 1, 2, \dots, K$  will be appropriate.

### Negative binomial( $r, p$ )

*pmf*  $P(X = x | r, p) = \binom{r+x-1}{x} p^r (1-p)^x$ ;  $x = 0, 1, \dots$ ;  $0 \leq p \leq 1$

*mean and variance*  $EX = \frac{r(1-p)}{p}$ ,  $\text{Var } X = \frac{r(1-p)}{p^2}$

*mgf*  $M_X(t) = \left( \frac{p}{1-(1-p)e^t} \right)^r$ ,  $t < -\log(1-p)$

*notes* An alternate form of the pmf is given by  $P(Y = y | r, p) = \binom{y-1}{r-1} p^r (1-p)^{y-r}$ ,  $y = r, r+1, \dots$ . The random variable  $Y = X + r$ . The negative binomial can be derived as a gamma mixture of Poissons. (See Exercise 4.34.)

### Poisson( $\lambda$ )

*pmf*  $P(X = x | \lambda) = \frac{e^{-\lambda} \lambda^x}{x!}$ ;  $x = 0, 1, \dots$ ;  $0 \leq \lambda < \infty$

*mean and variance*  $EX = \lambda$ ,  $\text{Var } X = \lambda$

*mgf*  $M_X(t) = e^{\lambda(e^t-1)}$

### Continuous Distributions

#### Beta( $\alpha, \beta$ )

*pdf*  $f(x | \alpha, \beta) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1}$ ,  $0 \leq x \leq 1$ ,  $\alpha > 0$ ,  $\beta > 0$

*mean and variance*  $EX = \frac{\alpha}{\alpha+\beta}$ ,  $\text{Var } X = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$

*mgf*  $M_X(t) = 1 + \sum_{k=1}^{\infty} \left( \prod_{l=0}^{k-1} \frac{\alpha+l}{\alpha+\beta+l} \right) \frac{t^k}{k!}$

*notes* The constant in the beta pdf can be defined in terms of gamma function  $B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$ . Equation (3.2.18) gives a general expression for the moments.

#### Cauchy( $\theta, \sigma$ )

*pdf*  $f(x | \theta, \sigma) = \frac{1}{\pi\sigma} \frac{1}{1 + \left(\frac{x-\theta}{\sigma}\right)^2}$ ,  $-\infty < x < \infty$ ;  $-\infty < \theta < \infty$ ,  $\sigma > 0$

*mean and variance* do not exist

*mgf* does not exist

*notes* Special case of Student's  $t$ , when degrees of freedom = 1. Also, if  $X$  and  $Y$  are independent  $n(0, 1)$ ,  $X/Y$  is Cauchy.

#### Chi squared( $p$ )

*pdf*  $f(x | p) = \frac{1}{\Gamma(p/2) 2^{p/2}} x^{p/2-1} e^{-x/2}$ ;  $0 \leq x < \infty$ ;  $p = 1, 2, \dots$

*mean and variance*  $EX = p$ ,  $\text{Var } X = 2p$

*mgf*  $M_X(t) = \left( \frac{1}{1-2t} \right)^{p/2}$ ,  $t < \frac{1}{2}$

*notes* Special case of the gamma distribution.

#### Double exponential( $\mu, \sigma$ )

*pdf*  $f(x | \mu, \sigma) = \frac{1}{2\sigma} e^{-|x-\mu|/\sigma}$ ,  $-\infty < x < \infty$ ,  $-\infty < \mu < \infty$ ,  $\sigma > 0$

*mean and variance*  $EX = \mu$ ,  $\text{Var } X = 2\sigma^2$

*mgf*  $M_X(t) = \frac{e^{\mu t}}{1-(\sigma t)^2}$ ,  $|t| < \frac{1}{\sigma}$

*notes* Also known as the Laplace distribution.

*Exponential*( $\beta$ )

pdf  $f(x|\beta) = \frac{1}{\beta}e^{-x/\beta}, \quad 0 \leq x < \infty, \quad \beta > 0$

mean and variance  $EX = \beta, \quad \text{Var } X = \beta^2$

mgf  $M_X(t) = \frac{1}{1-\beta t}, \quad t < \frac{1}{\beta}$

notes Special case of the gamma distribution. Has the memoryless property. Has many special cases:  $Y = X^{1/\gamma}$  is Weibull,  $Y = \sqrt{2X/\beta}$  is Rayleigh,  $Y = \alpha - \gamma \log(X/\beta)$  is Gumbel.

*F*

pdf  $f(x|\nu_1, \nu_2) = \frac{\Gamma(\frac{\nu_1+\nu_2}{2})}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \left(\frac{\nu_1}{\nu_2}\right)^{\nu_1/2} \frac{x^{\nu_1-2}}{(1+(\frac{\nu_1}{\nu_2})x)^{(\nu_1+\nu_2)/2}},$   
 $0 \leq x < \infty; \quad \nu_1, \nu_2 = 1, \dots$

mean and variance  $EX = \frac{\nu_2}{\nu_2-2}, \quad \nu_2 > 2,$   
 $\text{Var } X = 2 \left(\frac{\nu_2}{\nu_2-2}\right)^2 \frac{(\nu_1+\nu_2-2)}{\nu_1(\nu_2-4)}, \quad \nu_2 > 4$

moments (mgf does not exist)  $EX^n = \frac{\Gamma(\frac{\nu_1+2n}{2})\Gamma(\frac{\nu_2-2n}{2})}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \left(\frac{\nu_2}{\nu_1}\right)^n, \quad n < \frac{\nu_2}{2}$

notes Related to chi squared ( $F_{\nu_1, \nu_2} = \left(\frac{\chi_{\nu_1}^2}{\nu_1}\right) / \left(\frac{\chi_{\nu_2}^2}{\nu_2}\right)$ ), where the  $\chi^2$ s are independent) and  $t$  ( $F_{1, \nu} = t_\nu^2$ ).

*Gamma*( $\alpha, \beta$ )

pdf  $f(x|\alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-x/\beta}, \quad 0 \leq x < \infty, \quad \alpha, \beta > 0$

mean and variance  $EX = \alpha\beta, \quad \text{Var } X = \alpha\beta^2$

mgf  $M_X(t) = \left(\frac{1}{1-\beta t}\right)^\alpha, \quad t < \frac{1}{\beta}$

notes Some special cases are exponential ( $\alpha = 1$ ) and chi squared ( $\alpha = p/2, \beta = 2$ ). If  $\alpha = \frac{3}{2}, Y = \sqrt{X/\beta}$  is Maxwell.  $Y = 1/X$  has the inverted gamma distribution. Can also be related to the Poisson (Example 3.2.1).

*Logistic*( $\mu, \beta$ )

pdf  $f(x|\mu, \beta) = \frac{1}{\beta} \frac{e^{-(x-\mu)/\beta}}{[1+e^{-(x-\mu)/\beta}]^2}, \quad -\infty < x < \infty, \quad -\infty < \mu < \infty, \quad \beta > 0$

mean and variance  $EX = \mu, \quad \text{Var } X = \frac{\pi^2\beta^2}{3}$

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mgf  $M_X(t) = e^{\mu t} \Gamma(1-\beta t) \Gamma(1+\beta t), \quad |t| < \frac{1}{\beta}$   
 notes The cdf is given by  $F(x|\mu, \beta) = \frac{1}{1+e^{-(x-\mu)/\beta}}$ .

*Lognormal*( $\mu, \sigma^2$ )

pdf  $f(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \frac{e^{-(\log x - \mu)^2/(2\sigma^2)}}{x}, \quad 0 \leq x < \infty, \quad -\infty < \mu < \infty, \quad \sigma > 0$

mean and variance  $EX = e^{\mu + (\sigma^2/2)}, \quad \text{Var } X = e^{2(\mu + \sigma^2)} - e^{2\mu + \sigma^2}$

moments (mgf does not exist)  $EX^n = e^{n\mu + n^2\sigma^2/2}$

notes Example 2.3.5 gives another distribution with the same moments.

*Normal*( $\mu, \sigma^2$ )

pdf  $f(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} e^{-(x-\mu)^2/(2\sigma^2)}, \quad -\infty < x < \infty, \quad -\infty < \mu < \infty, \quad \sigma > 0$

mean and variance  $EX = \mu, \quad \text{Var } X = \sigma^2$

mgf  $M_X(t) = e^{\mu t + \sigma^2 t^2/2}$

notes Sometimes called the Gaussian distribution.

*Pareto*( $\alpha, \beta$ )

pdf  $f(x|\alpha, \beta) = \frac{\beta\alpha^\beta}{x^{\beta+1}}, \quad a < x < \infty, \quad \alpha > 0, \quad \beta > 0$

mean and variance  $EX = \frac{\beta\alpha}{\beta-1}, \quad \beta > 1, \quad \text{Var } X = \frac{\beta\alpha^2}{(\beta-1)^2(\beta-2)}, \quad \beta > 2$

mgf does not exist

$t$

pdf  $f(x|\nu) = \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\frac{\nu}{2})} \frac{1}{\sqrt{\nu\pi}} \frac{1}{(1+(\frac{x^2}{\nu}))^{(\nu+1)/2}}, \quad -\infty < x < \infty, \quad \nu = 1, \dots$

mean and variance  $EX = 0, \quad \nu > 1, \quad \text{Var } X = \frac{\nu}{\nu-2}, \quad \nu > 2$

moments (mgf does not exist)  $EX^n = \frac{\Gamma(\frac{\nu+1}{2})\Gamma(\frac{\nu-n}{2})}{\sqrt{\pi}\Gamma(\frac{\nu}{2})} \nu^{n/2}$  if  $n < \nu$  and even,  
 $EX^n = 0$  if  $n < \nu$  and odd.

notes Related to  $F$  ( $F_{1, \nu} = t_\nu^2$ ).

