

# HW2

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## Problem 1

(a)

$$\hat{\beta}x_i = \hat{\alpha}z_i \quad \hat{\beta} = c\hat{\alpha}$$

$$\hat{\alpha}_j = \frac{\hat{\beta}_j}{c}$$

(b)

(c)

(d)

I would say this is incorrect because the student forgot the  $\epsilon$ .

(e)

False. You can make the training model whatever size you want.

(f)

Unbiased means that if you did the experiment a large amount of times and calculate  $\hat{\beta}$  for each of those times, the average of all the estimates would exactly equal  $\beta$ .

## Problem 2

```
#install.packages("ISLR2") #you only need to do this one time.  
library(ISLR2) #you will need to do this every time you open a new R session.
```

```
## Warning: package 'ISLR2' was built under R version 4.5.2
```

(a)

```
head(Boston)
```

```
##      crim zn  indus chas   nox    rm   age    dis rad tax ptratio lstat medv  
## 1 0.00632 18   2.31    0 0.538 6.575 65.2 4.0900   1 296    15.3   4.98 24.0  
## 2 0.02731  0   7.07    0 0.469 6.421 78.9 4.9671   2 242    17.8   9.14 21.6  
## 3 0.02729  0   7.07    0 0.469 7.185 61.1 4.9671   2 242    17.8   4.03 34.7  
## 4 0.03237  0   2.18    0 0.458 6.998 45.8 6.0622   3 222    18.7   2.94 33.4  
## 5 0.06905  0   2.18    0 0.458 7.147 54.2 6.0622   3 222    18.7   5.33 36.2  
## 6 0.02985  0   2.18    0 0.458 6.430 58.7 6.0622   3 222    18.7   5.21 28.7
```

```

set.seed(1)

n <- nrow(Boston)

train_index <- sample(1:n, floor(n / 2), replace = FALSE)

train_boston <- Boston[train_index, ]
test_boston <- Boston[-train_index, ]

m1 <- lm(crim ~ ., data = train_boston)

summary(m1)

##
## Call:
## lm(formula = crim ~ ., data = train_boston)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -10.574  -2.723  -0.566   1.351   57.279
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  18.665277   10.693890   1.745  0.08219 .
## zn           0.046105    0.028198   1.635  0.10336
## indus       -0.127032    0.116665  -1.089  0.27730
## chas        -0.916885    1.769171  -0.518  0.60476
## nox        -11.606805    7.924234  -1.465  0.14431
## rm           0.738859    0.913675   0.809  0.41951
## age        -0.010585    0.026291  -0.403  0.68761
## dis        -1.184115    0.427288  -2.771  0.00602 **
## rad          0.671788    0.130702   5.140  5.7e-07 ***
## tax        -0.004607    0.007552  -0.610  0.54237
## ptratio    -0.515160    0.284824  -1.809  0.07175 .
## lstat       0.296310    0.114591   2.586  0.01031 *
## medv       -0.249594    0.097588  -2.558  0.01115 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 6.693 on 240 degrees of freedom
## Multiple R-squared:  0.5021, Adjusted R-squared:  0.4772
## F-statistic: 20.17 on 12 and 240 DF,  p-value: < 2.2e-16

trainMSE <- sum(m1$residuals^2)/nrow(train_boston)

trainMSE

## [1] 42.49345

test_predict <- predict(m1, newdata = test_boston)

test_MSE <- mean((test_predict - test_boston$crim)^2)

test_MSE

```

```
## [1] 41.19923
```

(b)

```
set.seed(1)

n <- nrow(Boston)

train_index <- sample(1:n, floor(n / 2), replace = FALSE)

train_boston <- Boston[train_index, ]
test_boston <- Boston[-train_index, ]

m1 <- lm(crim ~ zn + indus + nox + dis + rad + ptratio + medv, data = train_boston)

summary(m1)
```

```
##
## Call:
## lm(formula = crim ~ zn + indus + nox + dis + rad + ptratio +
##     medv, data = train_boston)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -9.145 -2.542 -0.645  1.175  57.028
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  29.29068    8.84292   3.312 0.001065 **
## zn           0.04940    0.02779   1.777 0.076733 .
## indus       -0.13410    0.10956  -1.224 0.222122
## nox        -12.25307    7.66941  -1.598 0.111408
## dis         -1.36389    0.39946  -3.414 0.000748 ***
## rad          0.63065    0.07015   8.990 < 2e-16 ***
## ptratio     -0.57530    0.28209  -2.039 0.042483 *
## medv        -0.35150    0.06458  -5.443 1.27e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 6.739 on 245 degrees of freedom
## Multiple R-squared:  0.4847, Adjusted R-squared:  0.47
## F-statistic: 32.93 on 7 and 245 DF, p-value: < 2.2e-16
```

```
trainMSE <- sum(m1$residuals^2)/nrow(train_boston)
```

```
trainMSE
```

```
## [1] 43.97466
```

```
test_predict <- predict(m1, newdata = test_boston)
```

```
test_MSE <- mean((test_predict - test_boston$crim)^2)
```

```
test_MSE
```

```
## [1] 39.62763
```

The training MSE collected from part A was slightly smaller than the one from part B. The test MSE collected from part A was slightly larger than the one from part B.

(c)

I expected part b to have a larger MSE because it is being trained on less variables so it's not likely to be as good as part a that was trained on every variable.

(d)

### Problem 3

(a)

$$\beta_0 = 2 \quad \beta_1 = 3 \quad \beta_2 = 5$$

(b)

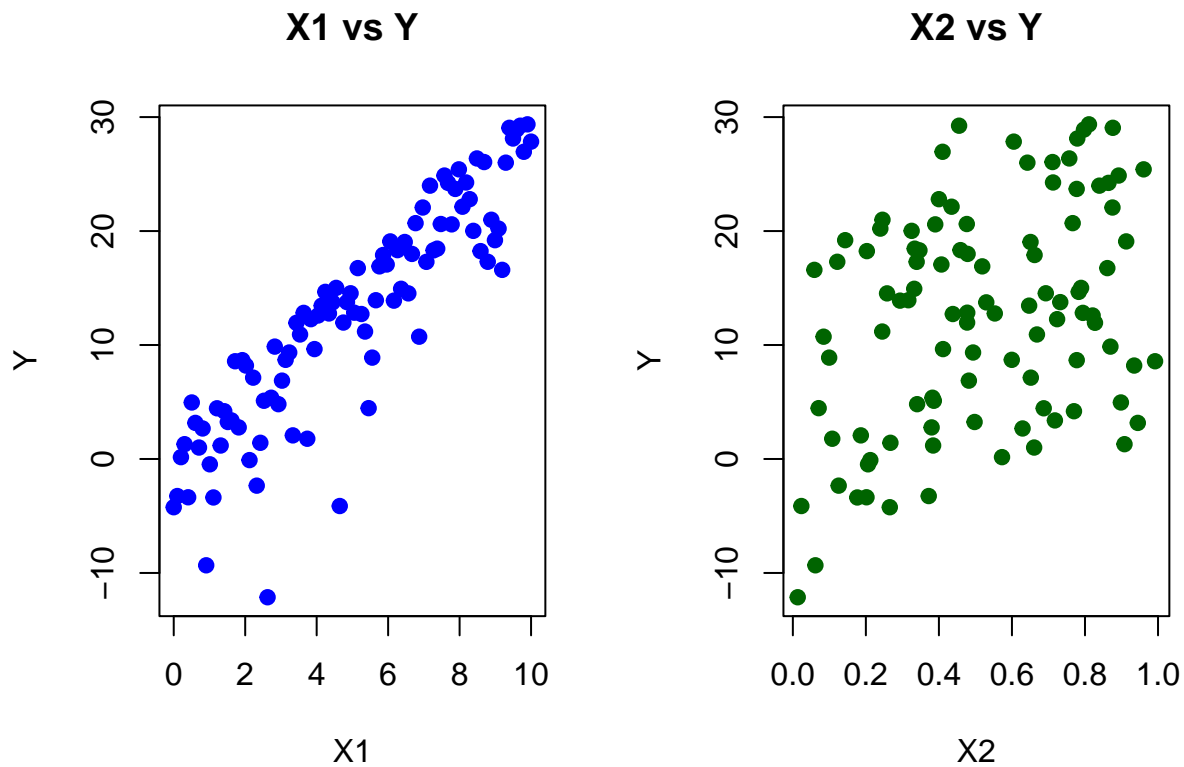
```
set.seed(1) # For reproducibility
X1 = seq(0, 10, length.out = 100)
X2 = runif(100)
epsilon = rnorm(100, mean = 0, sd = 1)
```

```
Y = 2 + 3*X1 + 5*log(X2) + epsilon
Y
```

```
## [1] -4.23243317 -3.25163817 0.16155277 1.29831740 -3.36017219
## [6] 4.95979463 3.16639013 1.00554051 2.67674841 -9.32815238
## [11] -0.46809191 -3.37647122 4.44916434 1.18318278 4.19129737
## [16] 3.24545027 3.38444034 8.57643582 2.77034153 8.67147697
## [21] 8.19849529 -0.09879487 7.13633701 -2.33945369 1.42069144
## [26] 5.10909193 -12.13061519 5.37632558 9.86110241 4.80943921
## [31] 6.87401554 8.70101323 9.34431309 2.07223625 11.94948013
## [36] 10.92516778 12.82034189 1.77720644 12.26835426 9.64280773
## [41] 12.59220416 13.45553046 14.66413305 12.76885860 13.74312926
## [46] 15.01216225 -4.12701656 11.97037589 13.76311067 14.53952062
## [51] 12.83644531 16.74957378 12.72008067 11.18201057 4.46102116
## [56] 8.89426498 13.93063644 16.90011981 17.89753266 17.06416690
## [61] 19.09030520 13.89557628 18.32735160 14.93305155 19.03933549
## [66] 14.53050849 17.99498458 20.69307892 10.73023226 22.06593930
## [71] 17.29846356 23.98309105 18.30688716 18.45521373 20.61605275
## [76] 24.86960493 24.22779022 20.58752240 23.69519179 25.41423139
## [81] 22.13662294 24.26178623 22.79845699 20.01888537 26.36971877
## [86] 18.24079376 26.05506809 17.30404313 20.99204629 19.19887827
## [91] 20.21505958 16.59568737 26.00022507 29.05787821 28.11965954
## [96] 28.90449384 29.24380140 26.95436970 29.36243300 27.84620901
```

(c)

```
par(mfrow=c(1,2))
plot(X1, Y, main="X1 vs Y", col="blue", pch=19)
plot(X2, Y, main="X2 vs Y", col="darkgreen", pch=19)
```



(d)

```
n_sim <- 10000
beta1_hats <- rep(NA, n_sim)

for(i in 1:n_sim) {
  # We generate new noise each time to simulate a new sample
  eps_sim <- rnorm(100, 0, 1)
  Y_sim <- 2 + 3*X1 + 5*log(X2) + eps_sim

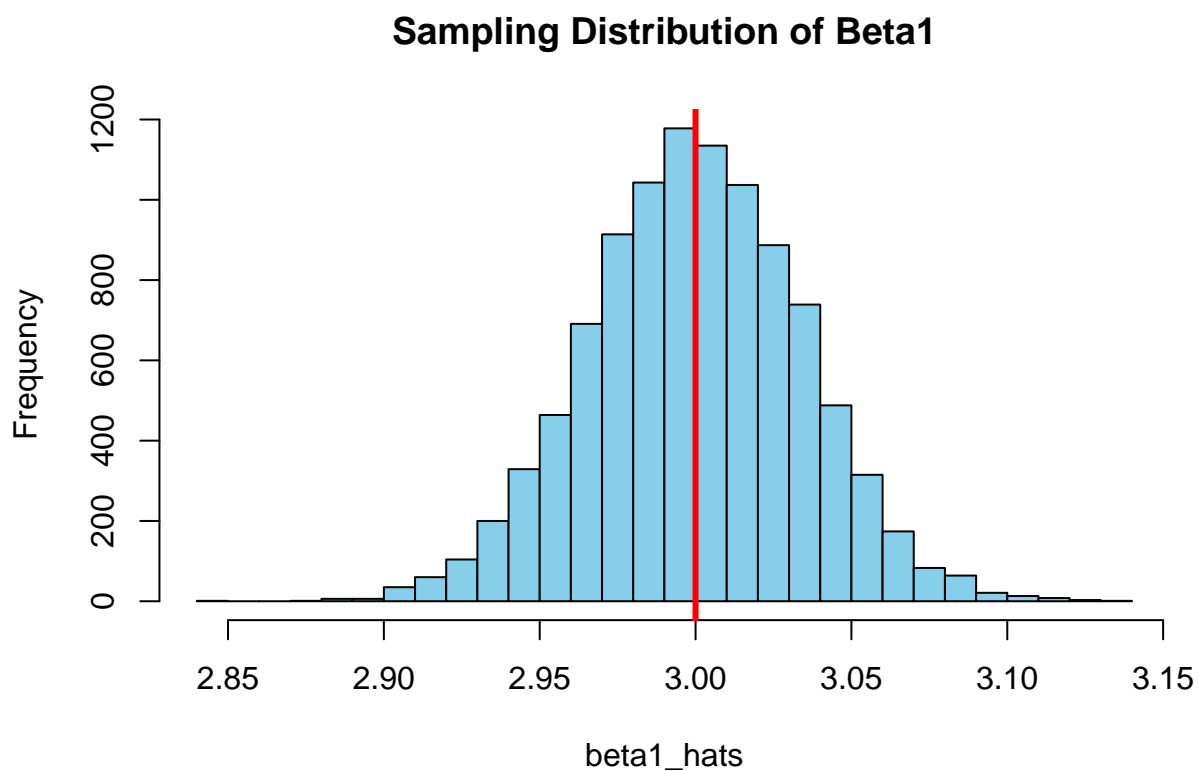
  # Fit the model (using log(X2) since that is the true relationship)
  fit <- lm(Y_sim ~ X1 + log(X2))
  beta1_hats[i] <- coef(fit)[2]
}

mean(beta1_hats)
```

## [1] 2.999792

(e)

```
hist(beta1_hats, breaks=30, col="skyblue", main="Sampling Distribution of Beta1")
abline(v = 3, col = "red", lwd = 3) # The true Beta1
```



## (f)

```
beta2_hats <- rep(NA, n_sim)

for(i in 1:n_sim) {
  eps_sim <- rnorm(100, 0, 1)
  Y_sim <- 2 + 3*X1 + 5*log(X2) + eps_sim
  fit <- lm(Y_sim ~ X1 + log(X2))
  beta2_hats[i] <- coef(fit)[3]
}

mean(beta2_hats)
```

## [1] 4.999968

(g)

```
hist(beta2_hats, breaks=30, col="lightgreen", main="Sampling Distribution of Beta2")
abline(v = 5, col = "red", lwd = 3) # The true Beta2
```

## Sampling Distribution of Beta2

