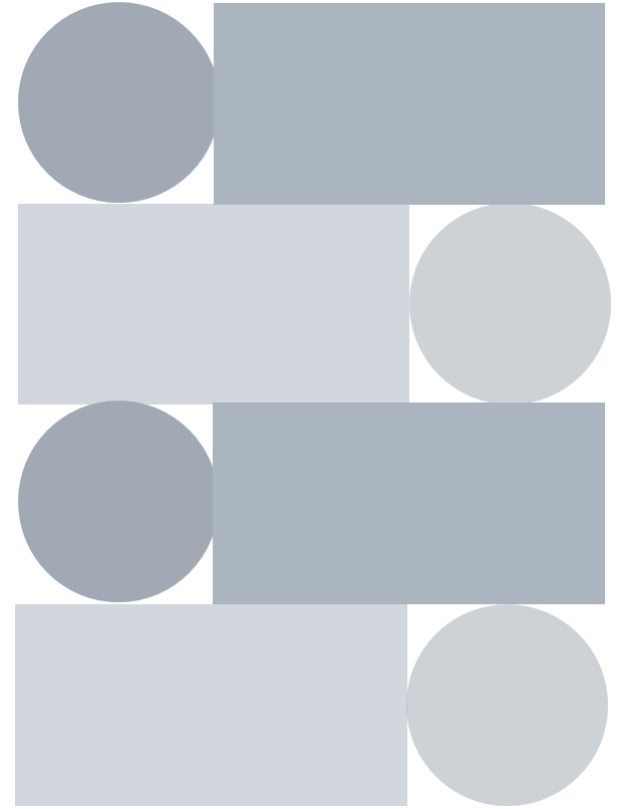
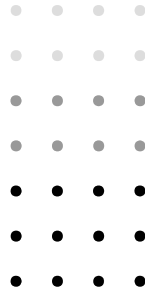


STATISTICAL COMPUTATION

WEEK 13 – LINEAR REGRESSION (1/2)

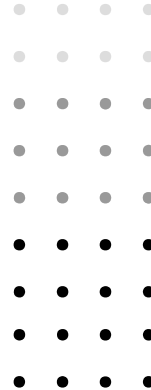
Annisa Auliya
I Melda Puspita



GET TO KNOW US

ANNISA AULIYA R.

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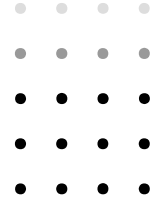


I MELDA PUSPITA L.

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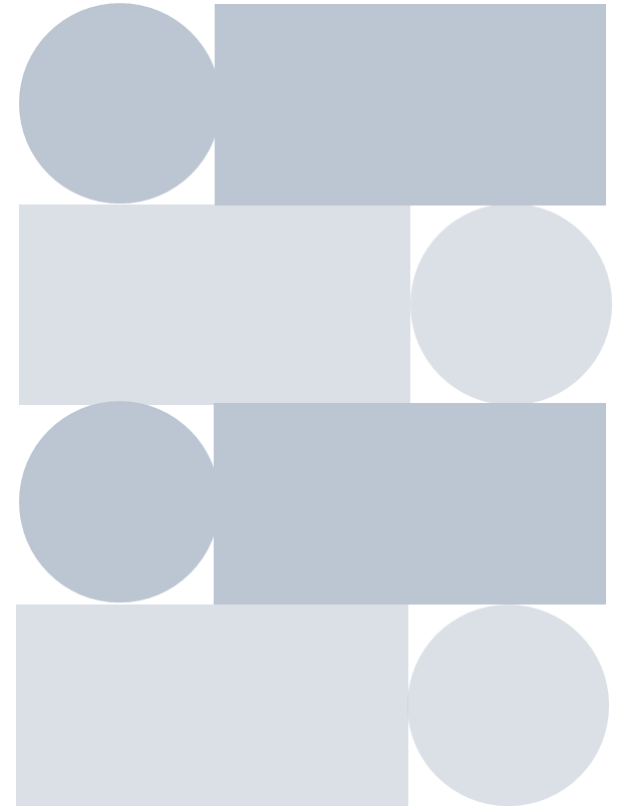
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MATERIALS



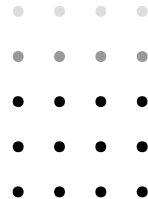
Simple Linear Regression

- Parameter Estimation
- Significance Test (Simultan & Partial)



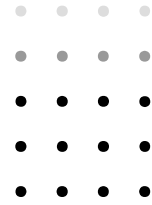
01

PARAMETER ESTIMATION





ESTIMATION



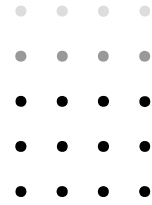
The **method of least squares** is used to estimate β_0 and β_1 . That is, we estimate β_0 and β_1 so that the sum of the squares of the differences between the observations y_i and the straight line is a minimum. From Eq. (2.1) we may write

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, \quad i = 1, 2, \dots, n \quad (2.3)$$

Equation (2.1) maybe viewed as a **population regression model** while Eq. (2.3) is a **sample regression model**, written in terms of the n pairs of data (y_i, x_i) ($i = 1, 2, \dots, n$). Thus, the least-squares criterion is

$$S(\beta_0, \beta_1) = \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 \quad (2.4)$$

ESTIMATION



The least-squares estimators of β_0 and β_1 , say $\hat{\beta}_0$ and $\hat{\beta}_1$, must satisfy

$$\left. \frac{\partial S}{\partial \beta_0} \right|_{\hat{\beta}_0, \hat{\beta}_1} = -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0$$

and

$$\left. \frac{\partial S}{\partial \beta_1} \right|_{\hat{\beta}_0, \hat{\beta}_1} = -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) x_i = 0$$

ESTIMATION

$$\begin{aligned} n\hat{\beta}_0 + \hat{\beta}_1 \sum_{i=1}^n x_i &= \sum_{i=1}^n y_i \\ \hat{\beta}_0 \sum_{i=1}^n x_i + \hat{\beta}_1 \sum_{i=1}^n x_i^2 &= \sum_{i=1}^n y_i x_i \end{aligned} \quad (2.5)$$

Equations (2.5) are called the **least-squares normal equations**. The solution to the normal equations is

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x} \quad (2.6)$$

and

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n y_i x_i - \frac{\left(\sum_{i=1}^n y_i\right)\left(\sum_{i=1}^n x_i\right)}{n}}{\sum_{i=1}^n x_i^2 - \frac{\left(\sum_{i=1}^n x_i\right)^2}{n}} \quad (2.7)$$



ESTIMATION

Since the denominator of Eq. (2.7) is the corrected sum of squares of the x_i and the numerator is the corrected sum of cross products of x_i and y_i , we may write these quantities in a more compact notation as

$$S_{xx} = \sum_{i=1}^n x_i^2 - \frac{\left(\sum_{i=1}^n x_i\right)^2}{n} = \sum_{i=1}^n (x_i - \bar{x})^2 \quad (2.9)$$

and

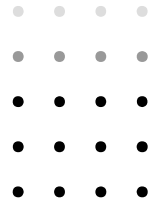
$$S_{xy} = \sum_{i=1}^n y_i x_i - \frac{\left(\sum_{i=1}^n y_i\right)\left(\sum_{i=1}^n x_i\right)}{n} = \sum_{i=1}^n y_i (x_i - \bar{x}) \quad (2.10)$$

Thus, a convenient way to write Eq. (2.7) is

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} \quad (2.11)$$

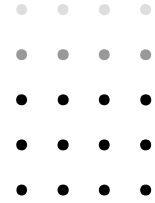
02

SIGNIFICANCE TEST





SIGNIFICANCE TEST

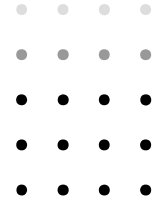


Source of Variation	Sum of Squares	Degrees of Freedom	Mean Square	F_0
Regression	$SS_R = \hat{\beta}_1 S_{xy}$	1	MS_R	MS_R / MS_{Res}
Residual	$SS_{Res} = SS_T - \hat{\beta}_1 S_{xy}$	$n - 2$	MS_{Res}	
Total	SS_T	$n - 1$		

$$SS_{Res} = \sum_{i=1}^n e_i^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2 \quad \sum_{i=1}^n y_i^2 - n\bar{y}^2 = \sum_{i=1}^n (y_i - \bar{y})^2 \equiv SS_T \quad \hat{\sigma}^2 = \frac{SS_{Res}}{n-2} = MS_{Res}$$

$$SS_{Res} = \sum_{i=1}^n y_i^2 - n\bar{y}^2 - \hat{\beta}_1 S_{xy} \quad SS_{Res} = SS_T - \hat{\beta}_1 S_{xy}$$

SIGNIFICANCE TEST



The test procedure for $H_0: \beta_1 = 0$ may be developed from two approaches. The first approach simply makes use of the t statistic in Eq. (2.27) with $\beta_{10} = 0$, or

$$t_0 = \frac{\hat{\beta}_1}{\text{se}(\hat{\beta}_1)}$$

The null hypothesis of significance of regression would be rejected if $|t_0| > t_{\alpha/2, n-2}$.

$$\text{se}(\hat{\beta}_1) = \sqrt{\frac{MS_{\text{Res}}}{S_{xx}}}$$

SIGNIFICANCE TEST

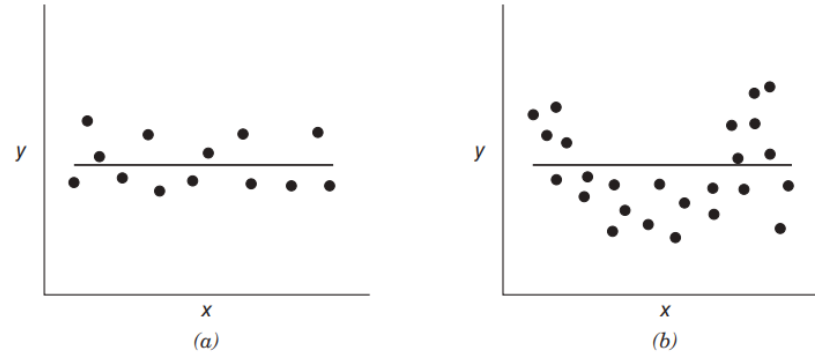


Figure 2.2 Situations where the hypothesis $H_0: \beta_1 = 0$ is not rejected.

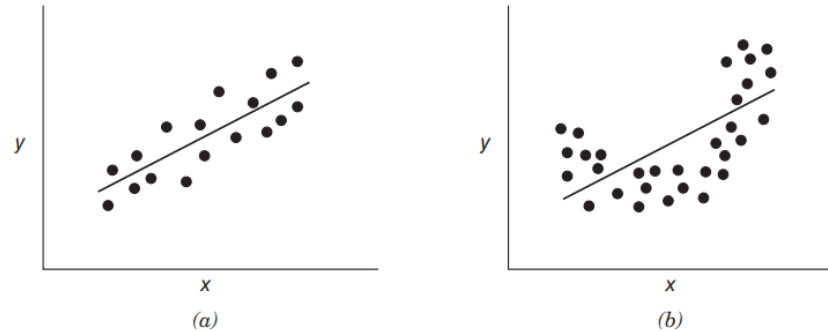


Figure 2.3 Situations where the hypothesis $H_0: \beta_1 = 0$ is rejected.



THANKS

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