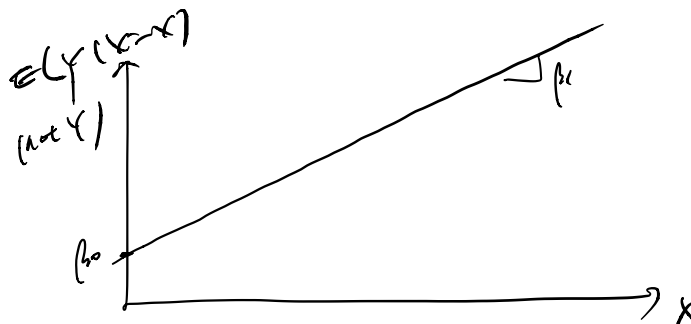


Simple linear regression. $\xrightarrow{\text{dataset}} (x_1, y_1) \dots (x_n, y_n)$
 \Rightarrow # of pred. var = 1

- there is a prob. distⁿ of Y for each level of X (= for given $X = x_i$)
- means of these prob distⁿ vary in some fashion wrt X .

$$\Rightarrow y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, \quad i=1 \dots n, \quad \varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

$$\Rightarrow E(y_i | X = x_i) = \underline{\beta_0} + \underline{\beta_1} x_i \sim N(\beta_0 + \beta_1 x_i, \sigma^2)$$



LS method. Use of $\beta_0, \beta_1 \Rightarrow \begin{matrix} \hat{\beta}_0 \\ \vdots \\ \hat{\beta}_1 \end{matrix} = \begin{matrix} b_0 \\ \vdots \\ b_1 \end{matrix}$

b_0 which minimizes $S(\beta_0, \beta_1) = \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2$

(i.e) $b_0 (= \hat{\beta}_0) = \arg \min_{\beta_0} S(\beta_0, \beta_1)$

$b_1 (= \hat{\beta}_1) = \arg \min_{\beta_1} S(\beta_0, \beta_1)$

$$\Rightarrow \frac{\partial S}{\partial \beta_0} = 0, \quad \frac{\partial S}{\partial \beta_1} = 0 \quad \left(\frac{\partial S}{\partial \beta} > 0 \right)$$

$$\textcircled{1} \sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \quad \textcircled{2} \sum x_i (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0.$$

$$\Downarrow \quad \Downarrow \hat{\beta}_1$$

$$\sum e_i = 0$$

$$\checkmark \quad \hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}, \quad \checkmark \quad \hat{\beta}_1 = \frac{s_{xy}}{s_{xx}} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$

$$s_{xy} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} = \frac{\sum (x_i - \bar{x}) y_i - \bar{y} \sum (x_i - \bar{x})}{\sum (x_i - \bar{x})^2}$$

$$= \frac{\sum (x_i - \bar{x}) y_i}{\sum (x_i - \bar{x})^2} = \frac{\sum (y_i - \bar{y}) x_i}{\sum (x_i - \bar{x})^2}$$

$$s_{xx} = \sum (x_i - \bar{x})^2$$

$$b_1 = \hat{\beta}_1 = \frac{\sum (x_i - \bar{x}) y_i}{s_{xx}} \sim N \left(\beta_1, \frac{\sigma^2}{s_{xx}} \right)$$

\therefore for $y_i \sim N(\mu, \sigma^2)$

$$b_0 = \bar{y} - \hat{\beta}_1 \bar{x} \sim N \left(\beta_0, \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{s_{xx}} \right) \right)$$

$$\text{Cov}(\bar{y}, \hat{\beta}_1) = 0.$$

$$V(b_0) = V(\bar{y}) + V(\hat{\beta}_1 \bar{x}) = \frac{\sigma^2}{n} + \bar{x}^2 \text{Var}(\hat{\beta}_1)$$

$$\sigma^2 \leftarrow \left(\hat{\sigma}^2 = \frac{\sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2}{n-2} \right) \quad \sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

$$\sigma^2 \leftarrow \hat{\sigma}^2 = \frac{\sum (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2}{n-2}$$

$$\frac{(n-2)\hat{\sigma}^2}{\sigma^2} \sim \chi_{n-2}^2 \quad \text{if } b_0, b_1$$

by def. of t -dist

$$z = \frac{\hat{\beta}_1 - \beta_1}{\frac{\hat{\sigma}}{\sqrt{s_{xx}}}} \sim N(0,1) \quad \frac{(n-2)\hat{\sigma}^2}{\sigma^2} \sim \chi_{n-2}^2$$

$$\sqrt{\frac{\chi^2}{n-2}} \quad (X)$$

$$\Rightarrow \frac{Z}{\sqrt{\chi^2/n-2}} \sim t_{n-2}$$

$$\hookrightarrow \frac{\hat{\beta}_1 - \beta_1}{\sqrt{\hat{\sigma}^2/n-2}} \sim t_{n-2}$$

$$\left[\begin{array}{c} \frac{\hat{\beta}_1 - \beta_1}{\sqrt{\hat{\sigma}^2/n-2}} \\ \frac{\hat{\beta}_0 - \beta_0}{\sqrt{\frac{1}{n} + \frac{\bar{x}^2}{S_{xx}}}} \end{array} \right] \sim t_{n-2}$$

$H_0: (\text{goal goal})^c$
 $H_a: \text{your goal}$
 \Rightarrow

#6 $Z = \begin{bmatrix} X \\ Y \end{bmatrix} \sim N_2 \left(\underline{\mu} = \begin{pmatrix} \mu_x \\ \mu_y \end{pmatrix}, \Sigma \right)$ $|\Sigma| = \det(\Sigma)$ $\Sigma = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$

(a) $f_Z(Z) = (2\pi)^{-\frac{2}{2}} |\Sigma|^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} \underbrace{(Z - \underline{\mu})^T}_{1 \times 2} \underbrace{\Sigma^{-1}}_{2 \times 2} \underbrace{(Z - \underline{\mu})}_{2 \times 1} \right\}$

$Z = \begin{pmatrix} X \\ Y \end{pmatrix}, \Sigma^{-1}$ $|\Sigma| = ad - bc.$ $\Rightarrow |X| \in \mathbb{R}^1.$

$f(x,y) = \left[\right]$ $\left(\begin{array}{l} v = \frac{x - \mu_x}{\sigma_x} \\ w = \frac{y - \mu_y}{\sigma_y} \end{array} \right)$

(b), Show $X \sim N(\mu_x, \sigma_x^2),$

$$f(x) = \int f(x,y) dy$$

$\underline{Z} = \begin{bmatrix} X \\ Y \end{bmatrix} \sim N_2(\underline{\mu}, \Sigma) \Rightarrow \underline{CZ} \sim N_2(C\underline{\mu}, C\Sigma C^T)$

\underline{C} : some constant vector

(c) $f_{X|Y} = \frac{f(X, Y)}{f(Y)}$

$\underline{Z} = \frac{\exp\left(-\frac{(X - \hat{\beta})^2}{2\sigma^2_{Y|X}}\right)}{\sqrt{\sigma^2_{Y|X}}} \quad C = \begin{bmatrix} 1 \\ 1 \end{bmatrix}$

#5 (b): $H_0: \beta_1 = 0$ vs $H_1: \beta_1 \neq 0$

$t^* = \frac{\hat{\beta}_1 - \beta_1}{\sqrt{\sigma^2_{Y|X} / S_{XX}}} \sim t_{n-2}$

rej. region $C = \{X: \left| \frac{\hat{\beta}_1 - 0}{\sqrt{\sigma^2_{Y|X} / S_{XX}}} \right| > t_{n-2, \alpha/2}\}$

cool! 100%

C.I. for β_1 : $\Pr(-t_{n-2, \alpha/2} < t^* < t_{n-2, \alpha/2}) = 1 - \alpha$

$\Rightarrow 1 - \alpha = 1 - \sqrt{\sigma^2_{Y|X} / S_{XX}}$

power

$$r(\beta) = \Pr[\text{rej } H_0 \mid H_0 \text{ is not true } (\beta_1 \neq 0)]$$

$\alpha = 0.05$
 $\delta = 2.5$
 $\beta_1 = \pm 0.5$

$$= \Pr \left[\left| \frac{\hat{\beta}_1 - 0}{\sqrt{\sigma^2 / 544}} \right| \geq z_{\alpha/2} \mid \beta_1 \neq 0 \right]$$

$\text{S.E.}(\hat{\beta}_1) = \text{se.}$

$$= \Pr \left[\frac{\hat{\beta}_1}{\text{se}} \geq z_{\alpha/2} \text{ or } \frac{\hat{\beta}_1}{\text{se}} < -z_{\alpha/2} \mid \beta_1 \neq 0 \right]$$

$$\frac{\hat{\beta}_1 - \beta_1}{\text{se}} \sim N(0, 1)$$

$$= \Pr \left[\frac{\hat{\beta}_1 - \beta_1}{\text{se}} > z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right] + \Pr \left[\frac{\hat{\beta}_1 - \beta_1}{\text{se}} < -z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right]$$

$\Phi(x) = \Pr(Z \leq x) \text{ with } Z \sim N(0, 1)$

$$= 1 - \Pr \left(Z < z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right) + \Pr \left(Z < -z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right)$$

$$= 1 - \Phi \left(z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right) + \Phi \left(-z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right)$$

$$r(\beta) = 1 - \Phi \left(z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right) + \Phi \left(-z_{\alpha/2} - \frac{\beta_1}{\text{se}} \right)$$

$$= r(-\beta_1)$$

$$\Phi(x) = 1 - \Phi(-x)$$

