




Single-Source Capture-Recapture Models With **singleRcapture**

Piotr Chlebicki
Stockholm University

Maciej Beręsewicz 
Poznań University of Economics
and Business

Abstract

Estimating population size is an important issue in official statistics, social sciences and natural sciences. One way to approach this problem is to use capture-recapture methods, which can be classified according to the number of sources used, the main distinction being between methods based on one source and those based on two or more sources. In this presentation we will introduce the **singleRcapture** R package for fitting SSCR models. The package implements state-of-the-art models as well as some new models proposed by the authors (e.g. extensions of zero-truncated one-inflated and one-inflated zero-truncated models). The software is intended for users interested in estimating the size of populations, particularly those that are difficult to reach or for which information is available from only one source and dual/multiple system estimation cannot be used.

Keywords: population size estimation, truncated distributions, count regression models, R.

1. Introduction

1.1. Literature review

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The subject of this workshop is the **singleRcapture** package and its lightweight extension that allows for integration with other R packages called **singleRcaptureExtra**.

The package is available on CRAN: [CRAN.R-project.org/package=singleRcapture](https://cran.r-project.org/package=singleRcapture) while the extension is available on: <https://github.com/ncn-foreigners/singleRcaptureExtra>.

The **singleRcapture** package is an R language package that focuses on implementing state of the art methods for frequentist point and interval estimation of size of closed populations in single-source capture-recapture (SSCR) setting (e.g. estimation of the population size of irregular migrants at set time point in a given area).

1.2. How do we estimate population size with only one register

1.3. Example with R code

Installation:

```
R> install.packages("singleRcapture")
R> remotes::install_github("https://github.com/ncn-foreigners/singleRcaptureExtra")

R> library(singleRcapture)
R>
R> head(netherlandsimmigrant)
```

	capture	gender	age	reason	nation
1	1	male	<40yrs	Other reason	North Africa
2	1	male	<40yrs	Other reason	North Africa
3	1	male	<40yrs	Other reason	North Africa
4	1	male	<40yrs	Other reason	Asia
5	1	male	<40yrs	Other reason	Asia
6	2	male	<40yrs	Other reason	North Africa

```
R> basicModel <- estimatePopsizel(
+   formula = capture ~ gender + age + nation,
+   model    = ztpoisson(),
+   data     = netherlandsimmigrant
+ )
R>
R> summary(basicModel)
```

Call:

```
estimatePopsizel.default(formula = capture ~ gender + age + nation,
  data = netherlandsimmigrant, model = ztpoisson())
```

Pearson Residuals:

	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
	-0.486442	-0.486442	-0.298080	0.002093	-0.209444	13.910844

Coefficients:

For linear predictors associated with: lambda

	Estimate	Std. Error	z value	P(> z)	
(Intercept)	-1.3411	0.2149	-6.241	4.35e-10	***
gendermale	0.3972	0.1630	2.436	0.014832	*
age>40yrs	-0.9746	0.4082	-2.387	0.016972	*
nationAsia	-1.0926	0.3016	-3.622	0.000292	***
nationNorth Africa	0.1900	0.1940	0.979	0.327398	
nationRest of Africa	-0.9106	0.3008	-3.027	0.002468	**
nationSurinam	-2.3364	1.0136	-2.305	0.021159	*
nationTurkey	-1.6754	0.6028	-2.779	0.005445	**

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

AIC: 1712.901

BIC: 1757.213

Residual deviance: 1128.553

Log-likelihood: -848.4504 on 1872 Degrees of freedom

Number of iterations: 8

Population size estimation results:

Point estimate 12690.35

Observed proportion: 14.8% (N obs = 1880)

Std. Error 2808.169

95% CI for the population size:

	lowerBound	upperBound
normal	7186.444	18194.26
logNormal	8431.275	19718.32

95% CI for the share of observed population:

	lowerBound	upperBound
normal	10.332927	26.16037
logNormal	9.534281	22.29793

2. Detailed information

2.1. Fitting method

2.2. Available models

The full list of implemented models in **singleRcapture** along with the expressions for probability density functions and point estimates is found in the collective help file for all family functions:

```
R> ?ztpoisson
```

Here we limit ourselves to just listing the family functions:

- Zero-truncated and zero-one-truncated Poisson, geometric, NB type II regression where the untruncated distribution is parameterized as:

$$\mathbb{P}[Y = y|\lambda, \alpha] = \frac{\Gamma(y + \alpha^{-1})}{\Gamma(\alpha^{-1}) y!} \left(\frac{\alpha^{-1}}{\alpha^{-1} + \lambda} \right)^{\alpha^{-1}} \left(\frac{\lambda}{\lambda + \alpha^{-1}} \right)^y.$$

- Zero-truncated one-inflated (ztoi) modifications distributions where the new probability \mathbb{P}^* measure is defined in terms of count data measure \mathbb{P} with support on $\mathbb{N} \cup \{0\}$ as:

$$\mathbb{P}^*[Y = y] = \begin{cases} \mathbb{P}[Y = 0] & y = 0, \\ \omega(1 - \mathbb{P}[Y = 0]) + (1 - \omega)\mathbb{P}[Y = 1] & y = 1, \\ (1 - \omega)\mathbb{P}[Y = y] & y > 1, \end{cases}$$

$$\mathbb{P}^*[Y = y|Y > 0] = \omega 1_{\{1\}}(y) + (1 - \omega)\mathbb{P}[Y = y|Y > 0].$$

- One-inflated zero-truncated (oizt) modifications distributions where the new probability \mathbb{P}^* measure is defined as:

$$\mathbb{P}^*[Y = y] = \omega 1_{\{1\}}(y) + (1 - \omega)\mathbb{P}[Y = y],$$

$$\mathbb{P}^*[Y = y|Y > 0] = \omega \frac{1_{\{1\}}(y)}{1 - (1 - \omega)\mathbb{P}[Y = 0]} + (1 - \omega) \frac{\mathbb{P}[Y = y]}{1 - (1 - \omega)\mathbb{P}[Y = 0]}.$$

- Generalized Chao's and Zelterman's estimators via logistic regression on variable Z defined as $Z = 1$ if $Y = 2$ and $Z = 0$ if $Y = 1$ with $Z \sim b(p)$ where $\text{logit}(p) = \ln(\lambda/2)$ for poisson parameter λ ,

$$\hat{N} = N_{obs} + \sum_{k=1}^{f_1+f_2} \left(2 \exp(\mathbf{x}_k \hat{\beta}) + 2 \exp(2\mathbf{x}_k \hat{\beta}) \right)^{-1}, \quad (\text{Chao's estimator})$$

$$\hat{N} = \sum_{k=1}^{N_{obs}} \left(1 - \exp(-2 \exp(\mathbf{x}_k \hat{\beta})) \right)^{-1}. \quad (\text{Zelterman's estimator})$$

- Alternative approaches to modelling one-inflation that mimic hurdle models where the first type zero truncated hurdle model (ztHurdle) is defined as:

$$\mathbb{P}^*[Y = y] = \begin{cases} \frac{\mathbb{P}[Y=0]}{1-\mathbb{P}[Y=1]} & y = 0, \\ \pi(1 - \mathbb{P}[Y = 1]) & y = 1, \\ (1 - \pi) \frac{\mathbb{P}[Y=y]}{1-\mathbb{P}[Y=1]} & y > 1, \end{cases}$$

$$\mathbb{P}^*[Y = y|Y > 0] = \pi 1_{\{1\}}(y) + (1 - \pi) 1_{\mathbb{N} \setminus \{1\}}(y) \frac{\mathbb{P}[Y = y]}{1 - \mathbb{P}[Y = 0] - \mathbb{P}[Y = 1]}$$

- The Hurdle zero truncated (Hurdlezt) is defined as:

$$\mathbb{P}^*[Y = y] = \begin{cases} \pi & y = 1, \\ (1 - \pi) \frac{\mathbb{P}[Y=y]}{1-\mathbb{P}[Y=1]} & y \neq 1, \end{cases} \quad \mathbb{P}^*[Y = y|Y > 0] = \begin{cases} \pi \frac{1-\mathbb{P}[Y=1]}{1-\mathbb{P}[Y=0]-\mathbb{P}[Y=1]} & y = 1, \\ (1 - \pi) \frac{\mathbb{P}[Y=y]}{1-\mathbb{P}[Y=0]-\mathbb{P}[Y=1]} & y > 1. \end{cases}$$

*Structure of a family function***2.3. Marginal frequencies**

A popular method of testing the model fit in single source capture-recapture studies is comparing the fitted marginal frequencies $\sum_{j=1}^{N_{obs}} \hat{\mathbb{P}}[Y_j = k | \mathbf{x}_j, Y_k > 0]$ with the observed marginal frequencies $\sum_{j=1}^N 1(Y_k = k) = \sum_{j=1}^{N_{obs}} 1(Y_k = k)$ for $k \geq 1$.

If a fitted model bears sufficient resemblance to the real data collection process these quantities should be quite close.

Affiliation:

Piotr Chlebicki
 Stockholm University
 Matematiska institutionen
 106 91 Stockholm, Albano hus 1
 E-mail: piotr.chlebicki@math.su.se
 URL: <https://github.com/Kertoo>

Maciej Beręsewicz
 Poznań University of Economics
 and Business
 Poznań University of Economics and Business
 Department of Statistics
 Institute of Informatics and Quantitative Economics
 Al. Niepodległości 10
 61-875 Poznań, Poland
 E-mail: maciej.beresewicz@ue.poznan.pl