# Unit 1 Simple Linear Regression

#### Prof. Phil Schniter



ECE 5307: Introduction to Machine Learning, Sp23

#### Learning objectives

- Understand how to load datasets in Python using pandas
- Visualize data using a scatter plot
- Understand sample mean, variance, and covariance
- Describe a linear model for data
  - Identify the target variable and predictor
- Compute the least-squares fit using the regression formula
- Compute the  $R^2$  measure-of-fit
- Visually assess goodness-of-fit and identify causes of poor fit

#### Outline

- Motivating Example: Predicting Automobile MPG
- Linear Model
- Least-Squares Fit: Problem Definition
- Sample Mean, Variance, Standard Deviation, Covariance
- Least-Squares Fit: The Solution
- Assessing Goodness-of-Fit

## Example: Predicting automobile mpg

- We will consider the task of predicting the fuel efficiency (in mpg) of a car from features like # cylinders, horsepower, weight, etc.
- Demo in Jupyter notebook: demo01\_auto\_mpg.ipynb
  - Learn from demo, and test your understanding in the lab
- Uses data from UCI library: https://archive.ics.uci.edu/ml/
- Data is loaded using Python's pandas library
  - Pandas routines have many options!
  - Learn from examples; Google is your friend!
  - Pandas read\_csv command creates a dataframe object with 3 main components:
    - df.values: numerical values in Numpy format
    - df.columns: column labels
    - df.index: row labels

## Result of pandas' read\_csv

4 17.0

**5** 15.0

8

8

302.0

429.0

```
import pandas as pd
 import numpy as np
         names = ['mpg', 'cylinders', 'displacement', 'horsepower',
In [3]:
                     'weight', 'acceleration', 'model year', 'origin', 'car name']
         df = pd.read csv('https://archive.ics.uci.edu/ml/machine-learning-databases/'+
In [4]:
                              'auto-mpg/auto-mpg.data'.
                              header=None, delim whitespace=True, names=names, na values='?')
         df.head(6)
Out[4]:
                  cylinders
                           displacement horsepower
                                                   weight acceleration model year origin
                                                                                                  car name
          0 18.0
                        8
                                  307.0
                                             130.0
                                                   3504.0
                                                                 12.0
                                                                             70
                                                                                    1 chevrolet chevelle malibu
             15.0
                         8
                                  350.0
                                             165.0
                                                   3693.0
                                                                 11.5
                                                                             70
                                                                                            buick skylark 320
          2 18.0
                        8
                                  318 0
                                             150.0 3436.0
                                                                 11.0
                                                                             70
                                                                                            plymouth satellite
             16.0
                        8
                                  304.0
                                             150.0
                                                   3433.0
                                                                 12.0
                                                                             70
                                                                                                amc rebel sst
```

10.5

10.0

70

70

ford torino

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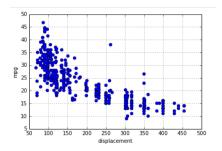
ford galaxie 500

140.0 3449.0

198.0 4341.0

#### Visualizing the data

- When possible, visualize the data before working with it.
- Python has MATLAB-like plotting features in the matplotlib module.



#### Visualizing the Data

We load the matplotlib module to plot the MATLAB.

```
In [15]: import matplotlib
import matplotlib.pyplot as plt
%matplotlib inline
```

First, let's extract some data columns and co

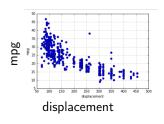
```
In [16]: xstr = 'displacement'
x = np.array(df[xstr])
y = np.array(df['mpg'])
```

Now we can create a scatter plot:

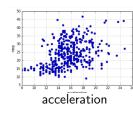
```
In [17]: plt.plot(x,y,'o')
   plt.xlabel(xstr)
   plt.ylabel('mpg')
   plt.grid(True)
```

## Postulating a model

- What relationships do you see?
- Is there a mathematical model relating the variables?
- How well can you predict mpg from these variables?







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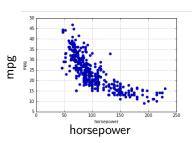
#### Outline

Motivating Example: Predicting Automobile MPG

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## Describing and visualizing data

- y: the variable we are trying to predict
  - Many names: target, label, regressand, response variable, dependent variable
- x: the variable we are using for prediction
  - Many names: feature, predictor, regressor, attribute, independent variable
  - For now we consider a single variable. Next unit we'll consider multiple
- Data: the set of pairs  $\{(x_i, y_i)\}_{i=1}^n$ 
  - Each pair is called a sample
  - We will use n for the number of samples
- A scatter plot is used to visualize the data pairs



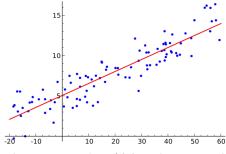
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## Single-variable linear model

Assume a linear relationship:

$$y \approx \beta_0 + \beta_1 x \triangleq \widehat{y}$$

- β<sub>1</sub>: slope
- lacksquare  $\beta_0$ : intercept
- $\widehat{y}$ : prediction
- $\beta \triangleq [\beta_0, \beta_1]^T$  are the parameters of the model. Also called coefficients or weights
- Why this model?
  - easy to interpret & analyze
  - simple to optimize parameters (as we will see)
- When is this a good model?



The regression line  $\widehat{y}(x)$  is shown in red

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#### Linear model: The residual

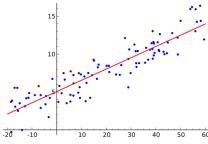
■ Note: *x* does not *exactly* predict *y*:

$$y \approx \beta_0 + \beta_1 x$$

Let's model this behavior explicitly using a residual,  $\epsilon$ :

$$y = \beta_0 + \beta_1 x + \epsilon$$

- For the *i*th sample, we then have:
  - **predicted value:**  $\hat{y}_i = \beta_0 + \beta_1 x_i$
  - residual:  $\epsilon_i = y_i \widehat{y}_i$



The residual  $\epsilon_i$  is the vertical deviation from  $y_i$  to the regression line

## Least-squares fit

- How do we choose the model parameters  $\beta = [\beta_0, \beta_1]^T$ ?
- Minimize the residual sum of squares (RSS):

$$RSS(\beta_0, \beta_1) \triangleq \sum_{i=1}^{n} \epsilon_i^2 = \sum_{i=1}^{n} (y_i - \widehat{y}_i)^2$$

https://en.wikipedia.org/wiki/Residual\_sum\_of\_squares

- Also called the sum of squared errors (SSE) and sum of square residuals (SSR)
- Note that  $\epsilon_i$  and  $\widehat{y}_i$  are implicitly functions of  $[\beta_0, \beta_1]$
- The value of  $\beta$  that minimizes RSS is the least-squares fit.
  - Geometrically: minimizes sum of squared distances to the regression line.

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## The optimization approach to ML: A general recipe

#### General ML problem

- Assume a model with some parameters
- Get training data
- Choose a loss function
- Find parameters that minimize loss

#### Simple Linear Regression

- $\rightarrow$  Linear model:  $\hat{y} = \beta_0 + \beta_1 x$
- $\rightarrow$  Data:  $\{(x_i, y_i)\}_{i=1}^n$
- $\rightarrow \operatorname{RSS}(\beta_0, \beta_1) \triangleq \sum_{i=1}^n (y_i \widehat{y}_i)^2$
- $\rightarrow$  Find  $(\beta_0, \beta_1)$  that minimizes  $RSS(\beta_0, \beta_1)$

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## Sample mean, variance, & standard deviation

If we summarize the data in the right way, we can easily design the optimal  $\beta!$ 

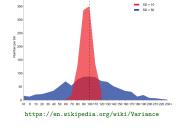
■ Sample mean:

$$\overline{x} \triangleq \frac{1}{n} \sum_{i=1}^{n} x_i, \quad \overline{y} \triangleq \frac{1}{n} \sum_{i=1}^{n} y_i$$

Sample variance:

$$s_x^2 \triangleq \frac{1}{n} \sum_{i=1}^n (x_i - \overline{x})^2, \quad s_y^2 \triangleq \frac{1}{n} \sum_{i=1}^n (y_i - \overline{y})^2$$

Note: some authors use  $\frac{1}{n-1}$  to give an "unbiased estimate." We'll explain this later, in Unit 3.

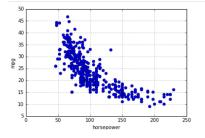


- Sample standard deviation (SD):  $s_x$ ,  $s_y$ 
  - Simply the square-root of the sample variance

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## Visualizing sample mean & SD on scatter plot

- Sample means  $\overline{x}$  and  $\overline{y}$ :
  - The center of mass in each axis
- Standard deviations  $s_x$  and  $s_y$ :
  - The "spread" in each axis about the mean
  - If the data was Gaussian distributed...
    - $\blacksquare$  68% of points <1 SD from mean
    - $\blacksquare$  95% of points < 2 SDs from mean
    - 99.7% of points < 3 SDs from mean



■ What are your estimates of  $\overline{x}, \overline{y}, s_x, s_y$  from the above scatter plot?

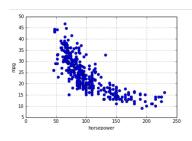
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## Computing the sample mean & SD in Python

• We can exactly compute  $\overline{x}, \overline{y}, s_x, s_y$  using the Numpy package in Python

```
In [27]: xm = np.mean(x)
ym = np.mean(y)
sxx = np.mean((x-xm)**2)
syy = np.mean((y-ym)**2)

xm = 104.47, ym= 23.45
sgrt(sxx)= 38.44, sgrt(syy)= 7.80
```



#### Sample covariance

The sample covariance is

$$s_{xy} \triangleq \frac{1}{n} \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y})$$

which indicates how "related"  $\{x_i\}$  and  $\{y_i\}$  are.

■ The value  $s_{xy}$  is easier to interpret after normalization. This motivates the sample (Pearson) correlation coefficient:

$$\rho_{xy} \triangleq \frac{s_{xy}}{s_x s_y} \in [-1, 1]$$

- lacktriangle The property  $ho_{xy} \in [-1,1]$  is a consequence of the Cauchy-Schwarz inequality.
- **Example** scatterplots for datasets with various  $\rho_{xy}$ :



https://en.wikipedia.org/wiki/Pearson\_correlation\_coefficient

## Alternative expressions for sample variance & covariance

- Recall that the sample variance was defined as  $s_x^2 \triangleq \frac{1}{n} \sum_{i=1}^n (x_i \overline{x})^2$
- A very useful alternative formula can be found by expanding the square:

$$s_x^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - 2\overline{x} \frac{1}{n} \sum_{i=1}^n x_i + \overline{x}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2 - \overline{x}^2 \quad \Rightarrow \quad \frac{1}{n} \sum_{i=1}^n x_i^2 = s_x^2 + \overline{x}^2$$

- Similarly, for the sample covariance we had  $s_{xy} \triangleq \frac{1}{n} \sum_{i=1}^{n} (x_i \overline{x})(y_i \overline{y})$
- A useful alternative expression for that is

$$s_{xy} = \frac{1}{n} \sum_{i=1}^{n} x_i y_i - \overline{x} \overline{y} \quad \Rightarrow \quad \frac{1}{n} \sum_{i=1}^{n} x_i y_i = s_{xy} + \overline{x} \overline{y}$$

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#### Notation

- We will use the following notation in this class (we'll try to be consistent)
- Note: some books/authors use different notations

Statistic	Notation	Formula	Python
sample mean	$\overline{x}$	$\frac{1}{n} \sum_{i=1}^{n} x_i$	xm
sample variance	$s_x^2 = s_{xx}$	$\frac{1}{n} \sum_{i=1}^{n} (x_i - \overline{x})^2$	sxx
sample standard deviation	$s_x = \sqrt{s_{xx}}$	$\sqrt{\frac{1}{n}\sum_{i=1}^{n}(x_i-\overline{x})^2}$	sx
sample covariance	$s_{xy}$	$ \frac{1}{n} \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y}) $	sxy
sample correlation coefficient	$ ho_{xy}$	$rac{s_{xy}}{s_x s_y}$	rhoxy

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## Minimizing RSS

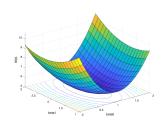
■ To minimize  $RSS(\beta_0, \beta_1)$ , we find the  $\beta_0$  and  $\beta_1$  that zero the gradient, i.e., the vector of partial derivatives:

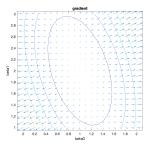
$$\left[\frac{\partial \operatorname{RSS}(\beta_0, \beta_1)}{\partial \beta_0}, \frac{\partial \operatorname{RSS}(\beta_0, \beta_1)}{\partial \beta_1}\right]^\mathsf{T} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}$$

Because RSS is quadratic with a PSD Hessian, the zero-gradient point is guaranteed to be an RSS minimum (more on this later)

 After some manipulations (see next page), we obtain the optimal values

$$\beta_1 = \frac{s_{xy}}{s_{xx}} = \frac{\rho_{xy}s_y}{s_x}, \qquad \beta_0 = \overline{y} - \beta_1 \overline{x}$$





## Minimizing RSS: Derivation

The minimum RSS is achieved by values of  $(\beta_0, \beta_1)$  that zero the gradient, i.e.,

$$0 = \frac{\partial \operatorname{RSS}(\beta_0, \beta_1)}{\partial \beta_0} = \frac{\partial}{\partial \beta_0} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 = -2 \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)$$
 (1)

$$0 = \frac{\partial \operatorname{RSS}(\beta_0, \beta_1)}{\partial \beta_1} = \frac{\partial}{\partial \beta_1} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 = -2 \sum_{i=1}^n x_i (y_i - \beta_0 - \beta_1 x_i)$$
 (2)

Starting with (1), we can multiply both sides by  $-\frac{1}{2n}$  to give

$$0 = \frac{1}{n} \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 x_i) = \underbrace{\frac{1}{n} \sum_{i=1}^{n} y_i - \beta_0 - \beta_1}_{\overline{y}} \underbrace{\frac{1}{n} \sum_{i=1}^{n} x_i}_{\overline{x}} \quad \Leftrightarrow \quad \boxed{\beta_0 = \overline{y} - \beta_1 \overline{x}}$$
(3)

Doing the same with (2) gives

$$0 = \frac{1}{n} \sum_{i=1}^{n} x_i (y_i - \beta_0 - \beta_1 x_i) = \frac{1}{n} \sum_{i=1}^{n} x_i y_i - \overline{x} \beta_0 - \beta_1 \frac{1}{n} \sum_{i=1}^{n} x_i^2$$
 (4)

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## Minimizing RSS: Derivation (continued)

Plugging in (3) into (4) gives

$$0 = \underbrace{\frac{1}{n} \sum_{i=1}^{n} x_i y_i - \overline{x} \, \overline{y}}_{S_{xy}} - \beta_1 \left( \underbrace{\frac{1}{n} \sum_{i=1}^{n} x_i^2 - \overline{x}^2}_{S_{xx}} \right) \quad \Leftrightarrow \quad \boxed{\beta_1 = \frac{s_{xy}}{s_{xx}}}$$

To find the minimum RSS, we plug the optimal value of  $\beta_0$  into the RSS definition to get

$$RSS(\beta_0, \beta_1) = \sum_{i} \left( y_i - \beta_0 - \beta_1 x_i \right)^2 = \sum_{i} \left( \left( y_i - \overline{y} \right) - \beta_1 (x_i - \overline{x}) \right)^2$$

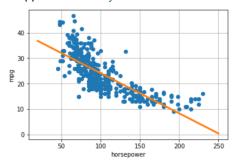
$$= \underbrace{\sum_{i} (y_i - \overline{y})^2 - 2\beta_1}_{ns_{yy}} \underbrace{\sum_{i} (y_i - \overline{y})(x_i - \overline{x})}_{ns_{xy}} + \beta_1^2 \underbrace{\sum_{i} (x_i - \overline{x})^2}_{ns_{xx}}$$
(5)

and then we plug the optimal value of  $\beta_1$  into (5) to get

$$RSS(\beta_0, \beta_1) = n \left( s_{yy} - 2 \frac{s_{xy}^2}{s_{xx}} + \frac{s_{xy}^2}{s_{xx}} \right) = n \left( s_{yy} - \frac{s_{xy}^2}{s_{xx}} \right)$$
$$= n \left( 1 - \frac{s_{xy}^2}{s_{xx}s_{yy}} \right) s_{yy} = n (1 - \rho_{xy}^2) s_{yy}$$

#### Automobile demo

Applied to our Python demo. . .



```
xm = np.mean(x)
ym = np.mean(y)
sxx = np.mean((x-xm)**2)
syy = np.mean((y-ym)**2)
syx = np.mean((y-ym)*(x-xm))
beta1 = syx/sxx
beta0 = ym - beta1*xm
Rsq = syx**2/sxx/syy
```

beta0=39.94, beta1=-0.16 Rsq=0.61

$$\widehat{\mathsf{mpg}} = \beta_0 + \beta_1 \times \mathsf{horsepower}$$

Another good simple-linear-regression demo can be found at https://stattrek.com/regression/regression-example.aspx?Tutorial=AP

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#### $R^2$ Goodness-of-fit

- Question: How to judge whether a predictor is doing "well" on a dataset?
- Answer: Use a normalized version of RSS:
  - $\blacksquare$  RSS includes contributions from n training samples. By considering RSS /n, we remove the dependence on n.
  - $\blacksquare$  RSS /n depends on  $s_y^2$ , the variance-of-y (i.e., if  $s_y^2$  doubles then RSS /n doubles). By considering  $\frac{RSS/n}{s_x^2}$ , we remove the dependence on  $s_y^2$ .
- More commonly, we report

$$1 - \frac{\text{RSS}/n}{s_y^2} \triangleq R^2,$$

 $1 - \frac{\text{RSS}/n}{s_y^2} \triangleq R^2$ , known as the "coefficient of determination"

https://en.wikipedia.org/wiki/Coefficient\_of\_determination

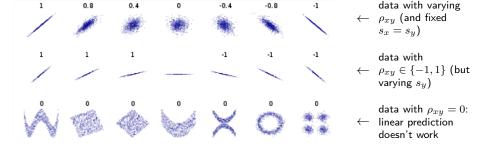
- $Arr R^2 = 1$  implies that the predictor is perfect (i.e.,  $\hat{y}_i = y_i$ )
- $Arr R^2 = 0$  implies the predictor is no better than the trivial one (i.e.,  $\hat{y}_i = \overline{y}$ )
- $\blacksquare R^2 < 0$  implies worse than trivial!
- With RSS-minimizing  $\beta_0$  and  $\beta_1$ , we know

$$RSS(\beta_0, \beta_1) = n(1 - \rho_{xy}^2)s_y^2 \quad \Rightarrow \quad \boxed{R^2 = \rho_{xy}^2}$$

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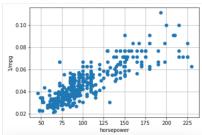
## Visualizing $\rho_{xy}^2$ from data $\{(x_i, y_i)\}_{i=1}^n$

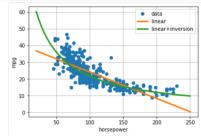
- If  $\rho_{xy}^2 \approx 1$ , then LS linear model gives a very good fit
- If  $\rho_{xy}^2 \approx 0$ , then LS linear model gives a very poor fit
- Since LS coef  $\beta_1 = \frac{\rho_{xy}s_y}{s_x}$ , we have that  $\mathrm{sgn}(\beta_1) = \mathrm{sgn}(\rho_{xy})$



## A better model for the automobile example?

- What if we predicted the inverse:  $\frac{1}{\text{mpg}} \approx \beta_0 + \beta_1 \times \text{horsepower}$
- This uses a nonlinear data transformation followed by linear regression
- Will explore this idea later in the course





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