

# CS215 Assignment 5

## Estimation Transformations

Tathagat Verma 180050111  
Neel Aryan Gupta 180050067

Lucky	
PAGE NO.	
DATE	/ /

### Question - 4

Pareto distribution  $P(\theta) \propto \left(\frac{\theta_m}{\theta}\right)^\alpha$  for  $\theta \geq \theta_m$   
0 otherwise

First we find the normalization constant

$$\therefore \int_{-\infty}^{\infty} P(\theta | \theta_m, \alpha) d\theta = 1$$

$$\Rightarrow k \int_{-\infty}^{\theta_m} 0 d\theta + k \int_{\theta_m}^{\infty} \left(\frac{\theta_m}{\theta}\right)^\alpha d\theta = 1$$

$$\Rightarrow k \cdot \theta_m^\alpha \int_{\theta_m}^{\infty} \frac{d\theta}{\theta^\alpha} = 1$$

$$\Rightarrow k \cdot \theta_m^\alpha \left[ \frac{\theta^{1-\alpha}}{1-\alpha} \right]_{\theta_m}^{\infty} = 1$$

$$\Rightarrow k \cdot \theta_m^\alpha \cdot \frac{\theta_m^{1-\alpha}}{\alpha-1} = 1$$

$$\Rightarrow \boxed{k = \frac{\alpha-1}{\theta_m}}$$

Now, we find the mean

$$\therefore \int_{-\infty}^{\infty} \theta \cdot P(\theta | \theta_m, \alpha) d\theta \equiv \text{Mean}$$

$$\Rightarrow \text{Mean} = \int_{\theta_m}^{\infty} \theta \cdot \frac{\alpha-1}{\theta_m} \cdot \left(\frac{\theta_m}{\theta}\right)^\alpha d\theta = (\alpha-1) \theta_m^{\alpha-1} \int_{\theta_m}^{\infty} \frac{d\theta}{\theta^{\alpha-1}}$$

$$= (\alpha-1) \theta_m^{\alpha-1} \left[ \frac{\theta^{2-\alpha}}{2-\alpha} \right]_{\theta_m}^{\infty}$$

$\therefore$  If  $\alpha < 2$  then mean =  $\infty$  (absurd)

Assume  $\alpha > 2$

$$\begin{aligned}
 \Rightarrow \text{Mean} &= (d-1) \theta_m^{d-1} \left[ \frac{\theta_m^{2-d}}{2-d} \right]_{\theta_m}^{\infty} \\
 &= (d-1) \theta_m^{d-1} \cdot \frac{\theta_m^{2-d}}{d-2} \\
 &= \left( \frac{d-1}{d-2} \right) \theta_m \quad \text{where } d > 2
 \end{aligned}$$

Let the dataset be  $\{x_i\}_{i=1}^N$   
 (i) Likelihood:  $\prod_{i=1}^N \frac{1}{\theta} = \frac{1}{\theta^N}$

$\Rightarrow \theta$  should be minimised

But  $(\theta, \theta)$  should still contain all the data  $\{x_i\}_{i=1}^N$

$\Rightarrow$  Lower bound for  $\theta = \max \{x_i\}_{i=1}^N$

Since  $\theta$  has to be minimised

$$\Rightarrow \theta = \max \{x_i\}_{i=1}^N$$

$$\text{let } \tilde{\theta} = \max \{x_i\}_{i=1}^N$$

$$\therefore \hat{\theta}_{MLE} = \tilde{\theta}$$

$$\text{Posterior} \propto \prod_{i=1}^N \frac{1}{\theta} \cdot \text{Prior} = \frac{1}{\theta^N} \cdot \left( \frac{\theta_m}{\theta} \right)^d \quad \text{for } \theta \geq \theta_m$$

$\therefore \theta$  should be minimised

$\theta$  still has to contain the dataset but

Prior should also be maximised

$$\Rightarrow \theta = \max(\theta_m, \tilde{\theta})$$

If  $\theta_m < \tilde{\theta}$  then likelihood would become zero at datapoints  $> \theta_m$  thereby halting the maximality of Posterior. That's why  $\theta = \max(\theta_m, \tilde{\theta})$  and NOT  $\theta = \tilde{\theta}$ .

$$\Rightarrow \hat{\theta}^{MAP} = \max(\theta_m, \tilde{\theta})$$

$$\& \tilde{\theta}^{MLE} = \tilde{\theta}$$

(ii) Note that if  $\theta_m < \tilde{\theta}$  then  $\hat{\theta}^{MAP} = \hat{\theta}^{MLE}$  always.  
For large sample size, we see that  
 $P(\theta_m > \tilde{\theta}) \approx 0$

$$\therefore P(\theta_m > \tilde{\theta}) = P(\theta_m > x_1, \theta_m > x_2, \dots, \theta_m > x_N)$$

$$= \frac{\theta_m}{\tilde{\theta}} \cdot \frac{\theta_m}{\tilde{\theta}} \dots$$

$$= \left(\frac{\theta_m}{\tilde{\theta}}\right)^N$$



But here  $\theta_m < \theta \Rightarrow \theta_m/\theta < 1$

$$\Rightarrow \left(\frac{\theta_m}{\theta}\right)^N \approx 0 \text{ for large } N$$

So  $\hat{\theta}^{MAP}$  will tend to  $\hat{\theta}^{MLE}$  as  $N$  increases.

This is desirable because we know that MLE estimators are the most efficient estimators for large sample sizes, using CRLB.

(ii) Note that posterior is zero if  $\theta < \max(\tilde{\theta}, \theta_m)$   
and posterior  $\propto \frac{\theta_m^{\alpha'}}{\theta^{\alpha'+N}} = \frac{1}{\theta^N}$

$$\Rightarrow \text{Posterior} \propto \frac{\theta_m^{\alpha'}}{\theta^{\alpha'+N}} \text{ for } \theta \geq \max(\tilde{\theta}, \theta_m) = \theta_m^*$$

0 otherwise

This is also a Pareto distribution which has  $\alpha' = \alpha + N$  &  $\theta_m^* = \max(\tilde{\theta}, \theta_m)$

We know the Pareto mean which is

$$\left(\frac{\alpha'-1}{\alpha'-2}\right) \theta_m$$



So we get

$$\text{Posterior mean} = \left( \frac{\alpha + N - 1}{\alpha + N - 2} \right) \cdot \max(\tilde{\theta}, \theta_m)$$

$$\Rightarrow \hat{\theta}^{\text{Posterior mean}} = \left( \frac{\alpha + N - 1}{\alpha + N - 2} \right) \cdot \theta_m^* \text{ where } \theta_m^* = \max(\tilde{\theta}, \theta_m)$$

(iv) As  $N \rightarrow \infty$

$$\lim_{N \rightarrow \infty} \left( \frac{\alpha + N - 1}{\alpha + N - 2} \right) \theta_m^* = \theta_m^* = \hat{\theta}^{\text{MAP}}$$

We know that  $\hat{\theta}^{\text{MAP}}$  tends to  $\hat{\theta}^{\text{MLE}}$  for large  $N$   
 $\Rightarrow \hat{\theta}^{\text{Posterior mean}} \rightarrow \hat{\theta}^{\text{MLE}}$  as  $N \rightarrow \infty$

This is desirable due to the same reason as mentioned in (ii) that  $\hat{\theta}^{\text{MLE}}$  is the most efficient estimator for  $N \rightarrow \infty$ .

This shows Bayes estimator gives correct results even for large  $N$  (in this case).