

Recall $\text{Var}(X) = \sigma^2 = E[(X - \mu_1)^2]$, where $\mu_1 = E(X)$.

3.6. Covariance

Ex 1. Let $E(X) = \mu_1$, $E(Y) = \mu_2$.

a) Confirm

$$\boxed{\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2E[(X-\mu_1)(Y-\mu_2)]} \quad (1)$$

Answer. $E(X+Y) = \mu_1 + \mu_2$

$$X+Y - E(X+Y) = X+Y - \mu_1 - \mu_2 = (X-\mu_1) + (Y-\mu_2)$$

$$\text{Var}(X+Y) = E[(X-\mu_1) + (Y-\mu_2)]^2 = E[(X-\mu_1)^2] + E[(Y-\mu_2)^2]$$

$$+ 2E[(X-\mu_1)(Y-\mu_2)] = \text{Var}(X) + \text{Var}(Y) + 2E[(X-\mu_1)(Y-\mu_2)].$$

b) Find $\text{Var}(X+Y)$, assuming X, Y are independent.

Answer. $\boxed{\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)}$, because

$$E[(X-\mu_1)(Y-\mu_2)] = E(X-\mu_1)E(Y-\mu_2) = 0$$

X, Y are indep.

Def. Covariance of X, Y is the number

$$\text{Cov}(X, Y) = E[(X-\mu_1)(Y-\mu_2)], \text{ where } \mu_1 = E(X), \mu_2 = E(Y)$$

Note 1. $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$

2. $\text{Cov}(X, Y) = 0$ if X, Y are independent. In that case,

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

$$3. \text{Cor}(X, Y) = \text{Cor}(Y, X), \quad \text{Cor}(X, X) = \text{Var}(X)$$

$$4. \text{Cor}\left(\sum_{i=1}^n a_i X_i, Y\right) = \sum_{i=1}^n a_i \text{Cor}(X_i, Y)$$

Why 4? Let $E(Y) = \mu$, $E(X_i) = \mu_i$. Then $E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i \mu_i$

$$\sum_{i=1}^n a_i X_i - E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i (X_i - \mu_i)$$

$$\text{Cor}\left(\sum_{i=1}^n a_i X_i, Y\right) = E\left[\sum_{i=1}^n a_i (X_i - \mu_i)(Y - \mu)\right] = \sum_{i=1}^n a_i \text{Cor}(X_i, Y)$$

5. (Consequence of 4)

$$\text{Cor}\left(\sum_{i=1}^n a_i X_i, \sum_{j=1}^m b_j Y_j\right) = \sum_{i,j} a_i b_j \text{Cor}(X_i, Y_j)$$

Variance covariance expansion

$$\begin{aligned} \text{Var}\left(\sum_{i=1}^n a_i X_i\right) &= \text{Cor}\left(\sum_{i=1}^n a_i X_i, \sum_{j=1}^n a_j X_j\right) \\ &= \sum_{i,j=1}^n a_i a_j \text{Cor}(X_i, X_j) = \sum_{i=1}^n a_i^2 \text{Var}(X_i) + \sum_{i \neq j} a_i a_j \text{Cor}(X_i, X_j) \\ &= \sum_{i=1}^n a_i^2 + 2 \sum_{i < j} a_i a_j \text{Cor}(X_i, X_j) \end{aligned}$$

In particular, $\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab \text{Cor}(X, Y)$.

$$\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y) \quad \text{if } X, Y \text{ are indep.}$$

Def. X, Y are uncorrelated if $\text{Cor}(X, Y) = 0$.

If X, Y are independent, then they are uncorrelated.

Meaning of $\text{Cor}(X, Y) = E[(X - \mu_1)(Y - \mu_2)]$ ($\mu_1 = E(X)$, $\mu_2 = E(Y)$):

$\text{Cor}(X, Y) > 0$ if $X - \mu_1$ and $Y - \mu_2$ are predominantly of the same sign.

$\text{Cor}(X, Y) < 0$ if " " of different sign

$\text{Cor}(X, Y) = 0$: X, Y uncorrelated

Def. Let $\sigma_1 = \sqrt{\text{Var}(X)}$, $\sigma_2 = \sqrt{\text{Var}(Y)}$.

Correlation (correlation coefficient) of X, Y is the number

$$\rho = \rho_{X, Y} = \frac{\text{Cor}(X, Y)}{\sigma_1 \sigma_2}$$

Basic properties 1. $\text{Cor}(X, Y) = \rho \sigma_1 \sigma_2$

2. Always $-1 \leq \rho \leq 1$, equivalently, $|\rho| \leq 1$

Ex 1. Let $\text{Var}(X) = \sigma_1^2$, $\text{Var}(Y) = \sigma_2^2$ with correlation $\rho = \rho(X, Y)$. Show that

a) $Y = \rho \frac{\sigma_2}{\sigma_1} X + Z$ with $\text{Cor}(X, Z) = 0$.

Such a representation is unique: if $Y = aX + \tilde{Z}$ with $\text{Cor}(X, \tilde{Z}) = 0$, then $\tilde{Z} = Z$, $a = \rho \frac{\sigma_2}{\sigma_1}$.

b) $\text{Var}(Z) = \sigma_2^2(1 - \rho^2)$

Answer a) Since $Z = Y - \rho \frac{\sigma_2}{\sigma_1} X$,

$$\begin{aligned} \text{Cor}(X, Z) &= \text{Cor}\left(X, Y - \rho \frac{\sigma_2}{\sigma_1} X\right) = \text{Cor}(X, Y) - \rho \frac{\sigma_2}{\sigma_1} \text{Var}(X) = \\ &= \rho \sigma_1 \sigma_2 - \rho \frac{\sigma_2}{\sigma_1} \sigma_1^2 = 0. \end{aligned}$$

$$b) \text{Var}(Z) = \text{Var}(Y) + \cancel{\rho^2 \frac{\sigma_2^2}{\sigma_1^2} \sigma_1^2} - 2\rho \frac{\sigma_2}{\sigma_1} \rho \sigma_1 \sigma_2 = \sigma_2^2(1 - \rho^2).$$

Comments on Ex 1:

1. $\text{Var}(Z) = \sigma_2^2(1 - \rho^2) \geq 0 \Rightarrow \rho^2 \leq 1$, equivalently, $|\rho| \leq 1$.

2. "Extreme correlation" is when $\rho^2 = 1$, $\rho = \pm 1$.

In this case, $\text{Var}(Z) = \sigma_z^2(1-\rho^2) = 0 \Rightarrow Z = c$ (constant), and

$Y = \pm \frac{\sigma_z}{\sigma_1} X + c$: Y is linear function of X .

3. The roles of X and Y can be reversed:

$X = \rho \frac{\sigma_1}{\sigma_2} Y + V$ with $\text{Cor}(V, Y) = 0$, and $\text{Var}(V) = \sigma_1^2(1-\rho^2)$.