Regression I

COMP9417, 23T2

- Intro
- 2 Linear Regression
- Multiple Linear Regression
- **5** Question 2 (a \rightarrow h)

Who am I?

Intro ○●

Who am I? Who are you?

Intro

Who am I? Who are you?

What you'll get from this course:

- Understand the basis of machine learning
- ML algorithms and the math behind them
- Ability to implement these ideas in Python

How to do well:

- Fully understand tut questions from week to week (they pile up)
- Don't be afraid of notation, break it all down
- Keep researching

Section 2

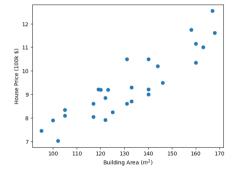
Linear Regression

Linear Regression

Say we're given a task to explain the relationship of the prices of homes based on their size in square meters.

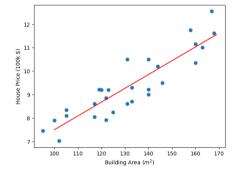
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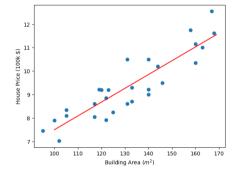


Let's try fitting a line of best fit:

Intro 00



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How do we know that this is the line of best fit?

$$E = e_1 + e_2 + e_3 + \dots + e_n$$
$$= \sum_{i=1}^{n} e_i$$

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We can generalise this to a function in nicer form:

$$L(\hat{y}) = \sum_{i=1}^{n} (y_i - \hat{y_i})$$

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$$L(\hat{y}) = \sum_{i=1}^{n} (y_i - \hat{y_i})$$

Something is wrong here.

Formally, we define our error/loss function as:

$$L(\hat{y})=\frac{1}{n}\sum_{i=1}^n(y_i-\hat{y_i})^2$$
 a.k.a MSE
$$L(w_0,w_1)=\frac{1}{n}\sum_{i=1}^n(y_i-w_0-w_1x_i)^2$$
 by definition

The minimum of our loss function w.r.t w_0 and w_1 will be their optimal values respectively.

Section 3

Question 1 (a \rightarrow c)

Derive the least-squares estimates for the univariate linear regression model.

i.e Solve:

$$\underset{w_0, w_1}{\operatorname{arg \, min}} \quad L(w_0, w_1) \\
\underset{w_0, w_1}{\operatorname{arg \, min}} \quad \frac{1}{n} \sum_{i=1}^{n} (y_i - w_0 - w_1 x_i)^2$$

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$$\frac{\partial L(w_0, w_1)}{\partial w_0} = -\frac{2}{n} \sum_{i=1}^n (y_i - w_0 - w_1 x_i)$$
$$= -\frac{2}{n} \left(\sum_{i=1}^n y_i - nw_0 - w_1 \sum_{i=1}^n x_i \right)$$

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For the minimum, $\frac{\partial L(w_0,w_1)}{\partial w_0}=0$,

$$-\frac{2}{n}\left(\sum_{i=1}^{n} y_i - nw_0 - w_1 \sum_{i=1}^{n} x_i\right) = 0$$

To find w_1 , we follow a similar process and use simple simultaneous equations to solve for the final solution.

So,

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$$\frac{\partial L(w_0, w_1)}{\partial w_1} = -\frac{2}{n} \sum_{i=1}^n x_i (y_i - w_0 - w_1 x_i)$$
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$$= -\frac{2}{n} \left(\sum_{i=1}^n x_i y_i - w_0 \sum_{i=1}^n x_i - w_1 \sum_{i=1}^n x_i^2 \right)$$

$$\frac{\partial L(w_0, w_1)}{\partial w_1} = 0,$$

$$\frac{1}{n} \left(\sum_{i=1}^{n} x_i y_i - w_0 \sum_{i=1}^{n} x_i - w_1 \sum_{i=1}^{n} x_i^2 \right) = 0$$

$$\overline{xy} - w_0 \overline{x} - w_1 \overline{x^2} = 0$$

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$$w_1 = \frac{\overline{xy} - w_0 \overline{x}}{\overline{x^2}}$$
(2)

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Sub (1) into (2):

(2)

$$\overline{xy} - w_0 \overline{x} - w_1 \overline{x^2} = 0$$

$$w_1 = \frac{\overline{xy} - w_0 \overline{x}}{\overline{x^2}}$$

Sub (1) into (2):

$$w_1 = \frac{\overline{x}\overline{y} - (\overline{y} - w_1 \overline{x})\overline{x}}{\overline{x}^2}$$

$$w_1 = \frac{\overline{x}\overline{y} - \overline{x}\overline{y} + w_1 \overline{x}^2}{\overline{x}^2}$$

$$v_1(\frac{\overline{x^2} - \overline{x}^2}{\overline{x}^2}) = \frac{\overline{x}\overline{y} - \overline{x}\overline{y} + w_1 \overline{x}^2}{\overline{x}^2}$$

$$w_1 = \frac{\overline{x}\overline{y} - \overline{x}\overline{y}}{\overline{x}^2 - \overline{x}^2}$$

Finally, we have

$$w_1=rac{\overline{x}\overline{y}-ar{x}ar{y}}{\overline{x^2}-ar{x}^2}$$
 and $w_0=ar{y}-w_1ar{x}$

Problem: Prove (\bar{x}, \bar{y}) is on the line.

From 1(a), the equation of our line $(\hat{y} = w_0 + w_1 x)$ becomes:

$$\hat{y} = \bar{y} - \bar{x} \frac{\overline{xy} - \bar{x}\bar{y}}{\overline{x^2} - \bar{x}^2} + \frac{\overline{xy} - \bar{x}\bar{y}}{\overline{x^2} - \bar{x}^2} x$$

Sub $x = \bar{x}$.

$$\hat{y} = \bar{y} - \bar{x} \frac{\overline{xy} - \bar{x}\bar{y}}{\overline{x^2} - \bar{x}^2} + \frac{\overline{xy} - \bar{x}\bar{y}}{\overline{x^2} - \bar{x}^2} \bar{x}$$

$$\hat{y} = \bar{y}$$

 (\bar{x}, \bar{y}) is on the line

Multiple Linear Regression

1c

Similar to 1a, though take care with the partial derivatives:

$$\frac{\partial L(w_0, w_1)}{\partial w_0} = -\frac{2}{n} \sum_{i=1}^n (y_i - w_0 - w_1 x_i)$$

$$\frac{\partial L(w_0, w_1)}{\partial w_1} = -\frac{2}{n} \sum_{i=1}^{n} x_i (y_i - w_0 - w_1 x_i) + 2\lambda w_1$$

Final result is:

$$w_0 = \bar{y} - w_1 \bar{x}$$

$$w_1 = \frac{\bar{x}\bar{y} - \bar{x}\bar{y}}{\bar{x}^2 - \bar{x}^2 + \lambda}$$

Notice how the coefficients have an inverse relationship with λ .

Recall the previous problem where we were tasked with finding price patterns of homes using the size of the home.

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Simple, just add another parameter:

$$\hat{y} = w_0 + w_1 x_1 + w_2 x_2$$

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$$\hat{y} = w_0 + w_1 x_1 + w_2 x_2$$

What if we're given the year the house was built and the coordinates? Let's say d more features?

Let's vectorise our model, say:

$$x_i = \begin{bmatrix} 1 \\ x_{i1} \end{bmatrix}$$
 to represent our input & the bias (w_0)

$$y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$
 to represent the target variable

$$w = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$
 to represent the parameters

$$X = \begin{bmatrix} 1 & x_{11} \\ 1 & x_{21} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$$

Multiple Linear Regression

$$X = \begin{bmatrix} 1 & x_{11} \\ 1 & x_{21} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$$

So,

$$Xw = \begin{bmatrix} w_0 + w_1 x_{11} \\ w_0 + w_1 x_{21} \\ \vdots \\ w_0 + w_1 x_{n1} \end{bmatrix}$$

$$\hat{y} = Xw$$

$$\mathcal{L}(w) = \frac{1}{n} \sum_{i=1}^{n} (y_i - [Xw]_i)^2$$

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Formally,

$$\mathcal{L}(w) = \frac{1}{n} ||y - Xw||_2^2$$

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Squared 2-Norm Identity

For a vector v,

$$||v||_2^2 = v^T v$$

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Vector Calculus

If $u \in \mathbb{R}^n$ and $v \in \mathbb{R}^n$ is a fixed vector, then

$$\frac{\partial v^T u}{\partial u} = v$$

If $A \in \mathbb{R}^{n \times n}$ is a fixed matrix,

$$\frac{\partial (u^T A u)}{\partial u} = A u + A^T u$$

Section 5

Question 2 (a \rightarrow h)

2a

Problem: Show that $\mathcal{L}(w) = \frac{1}{n} \|y - Xw\|_2^2$ has critical point $\hat{w} = (X^T X)^{-1} X^T y$.

To find optimal w, solve $\frac{\partial \mathcal{L}(w)}{\partial w} = 0$

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To find optimal w, solve $\frac{\partial \mathcal{L}(w)}{\partial w} = 0$

$$\mathcal{L}(w) = \frac{1}{n} (y - Xw)^T (y - Xw)$$

$$= \frac{1}{n} \left(y^T y - y^T Xw - w^T X^T y + w^T X^T Xw \right)$$

$$= \frac{1}{n} \left(y^T y - 2y^T Xw + w^T X^T Xw \right)$$

Let's find the derivative w.r.t w,

$$\frac{\partial \mathcal{L}(w)}{\partial w} = -\frac{1}{n}(-2X^Ty + 2X^TXw)$$

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To solve for \hat{w} ,

$$-2X^{T}y + 2X^{T}X\hat{w} = 0$$
$$\hat{w} = (X^{T}X)^{-1}X^{T}y$$

2b

Problem: Prove $\hat{w} = (X^TX)^{-1}X^Ty$ is a global minimum.

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$$\nabla_w^2 \mathcal{L}(w) = \nabla_w (\nabla_w \mathcal{L}(w))$$
$$= \nabla_w (-2X^T y + 2X^T X w)$$
$$= 2X^T X$$

Problem: Prove $\hat{w} = (X^T X)^{-1} X^T y$ is a global minimum.

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So, for a vector $u \in \mathbb{R}^p$.

$$u^{T}(2X^{T}X)u = 2(u^{T}X^{T})(Xu)$$

$$= 2(Xu)^{T}(Xu)$$

$$= 2||Xu||_{2}^{2} \ge 0$$

Therefore, ${\cal L}$ is convex and \hat{w} is the unique global minimum.

$$x_i = egin{bmatrix} 1 \\ x_{i1} \end{bmatrix}$$
 to represent our input & the bias (w_0)

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$$w = \begin{bmatrix} w_0 \\ w_1 \end{bmatrix}$$
 to represent the parameters

$$X = \begin{bmatrix} 1 & x_{11} \\ 1 & x_{21} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$$

$$X^T y = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_{11} & x_{21} & \cdots & x_{n1} \end{bmatrix} \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

$$X = \begin{bmatrix} 1 & x_{11} \\ 1 & x_{21} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$$

$$X^{T}y = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_{11} & x_{21} & \cdots & x_{n1} \end{bmatrix} \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$
$$= \begin{bmatrix} n\bar{y} \end{bmatrix}$$

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 $X^T X = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_{11} & x_{21} & \cdots & x_{n1} \end{bmatrix} \begin{bmatrix} 1 & x_{11} \\ 1 & x_{11} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$

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$$= \begin{bmatrix} n\overline{y} \\ n\overline{xy} \end{bmatrix}$$

 $X^{T}X = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_{11} & x_{21} & \cdots & x_{n1} \end{bmatrix} \begin{bmatrix} 1 & x_{11} \\ 1 & x_{11} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$ $= \begin{bmatrix} n & \sum_{i=1}^{n} x_{i} \\ \sum_{i=1}^{n} x_{i} & \sum_{i=1}^{n} x_{i}^{2} \end{bmatrix}$

$$X = \begin{bmatrix} 1 & x_{11} \\ 1 & x_{21} \\ \vdots & \vdots \\ 1 & x_{n1} \end{bmatrix}$$

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We have:

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Recall the inverse of a matrix
$$A=\begin{bmatrix} a & b \\ c & d \end{bmatrix}$$
 is $A^{-1}=\frac{1}{ad-bc}\begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$.

Multiple Linear Regression

$$X^T X = \begin{bmatrix} n & n\bar{x} \\ n\bar{x} & n\bar{x}^2 \end{bmatrix}$$

Recall the inverse of a matrix $A = \begin{bmatrix} a & b \\ c & d \end{bmatrix}$ is $A^{-1} = \frac{1}{ad-bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$.

$$(X^T X)^{-1} = \frac{1}{n^2 \overline{x^2} - n^2 \overline{x}^2} \begin{bmatrix} n \overline{x^2} & -n \overline{x} \\ -n \overline{x} & n \end{bmatrix}$$
$$= \frac{1}{n(\overline{x^2} - \overline{x}^2)} \begin{bmatrix} \overline{x^2} & -\overline{x} \\ -\overline{x} & 1 \end{bmatrix}$$

$$(X^T X)^{-1} X^T y = \frac{1}{n(\overline{x^2} - \bar{x}^2)} \begin{bmatrix} \overline{x^2} & -\bar{x} \\ -\bar{x} & 1 \end{bmatrix} \begin{bmatrix} n\bar{y} \\ n\overline{xy} \end{bmatrix}$$

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$$= \frac{1}{\overline{x^2} - \bar{x}^2} \begin{bmatrix} \overline{x^2} \bar{y} - \bar{x} \overline{xy} \\ \overline{xy} - \bar{x} \bar{y} \end{bmatrix}$$

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$$= \frac{1}{\overline{x^2} - \bar{x}^2} \begin{bmatrix} \overline{x^2} \bar{y} - \bar{x} \overline{x} \overline{y} \\ \overline{xy} - \bar{x} \overline{y} \end{bmatrix}$$
$$= \begin{bmatrix} \bar{y} - \hat{w}_1 \bar{x} \\ \frac{\overline{xy} - \bar{x} \bar{y}}{x^2 - \bar{x}^2} \end{bmatrix}$$

2e - Lab

Given $x_1,\cdots,x_5=3,6,7,8,11$ and $y_1,\cdots,y_5=13,8,11,2,6$ compute the least squares solution by hand and using Python. Check your results with the sklearn implementation.

$$\mathsf{MSE}(w) = \operatorname*{arg\,min}_{w} \frac{1}{n} \|y - Xw\|_2^2 \text{ and } \mathsf{SSE}(w) = \operatorname*{arg\,min}_{w} \|y - Xw\|_2^2$$

- i) Is the minimiser of MSE and SSE the same?
- ii) Is the minimum value of MSE and SSE the same?