

(1) Consider a two-class classification problem:

$$p(x|w_1) = N(x|\mu_1, \Sigma) \quad (\text{class 1})$$

$$p(x|w_2) = N(x|\mu_2, \Sigma) \quad (\text{class 2})$$

with $p(w_1) = p(w_2) = \frac{1}{2}$, $\Sigma = \begin{pmatrix} \Sigma_a & 0 \\ 0 & \Sigma_b \end{pmatrix}$, and

$$\Sigma_a = \text{diag}(1, 2, \dots, D), (\Sigma_b)_{ij} = \frac{\sqrt{\epsilon_{ij}}}{2|i-j|} \in \mathbb{R}^{D \times D},$$

μ_a and μ_b are vectors, both in the form $(\underbrace{1, \dots, 1}_{\text{first } k \text{ components}}, 0, \dots, 0)$.

Find the Bayes error L^* .

The decision boundary is given whether the posteriors $p(w_1|x) > p(w_2|x)$ or $p(w_1|x) \leq p(w_2|x)$.

Since $p(w_1) = p(w_2)$ this gives

$$\log p(x|w_1) - \log p(x|w_2) > 0$$

$$-(x-\mu_1)^T \Sigma^{-1} (x-\mu_1) + (x-\mu_2)^T \Sigma^{-1} (x-\mu_2) > 0$$

Because of equal covariance the quadratic terms cancel out and we get

$$2\mu_1^T \Sigma^{-1} x - 2\mu_2^T \Sigma^{-1} x - \mu_1^T \Sigma^{-1} \mu_1 + \mu_2^T \Sigma^{-1} \mu_2 > 0,$$

And since $\mu_2 = -\mu_1 = -\mu_1$, we have

$$\boxed{\mu_1^T \Sigma^{-1} x > 0} \quad (\text{class 1})$$

or $x \in \text{class 2}$ otherwise.

Thus, the discriminant function is

$$g(x) = \begin{cases} 1 & \text{if } \mu^T \Sigma^{-1} x > 0 \\ 0 & \text{otherwise} \end{cases}$$

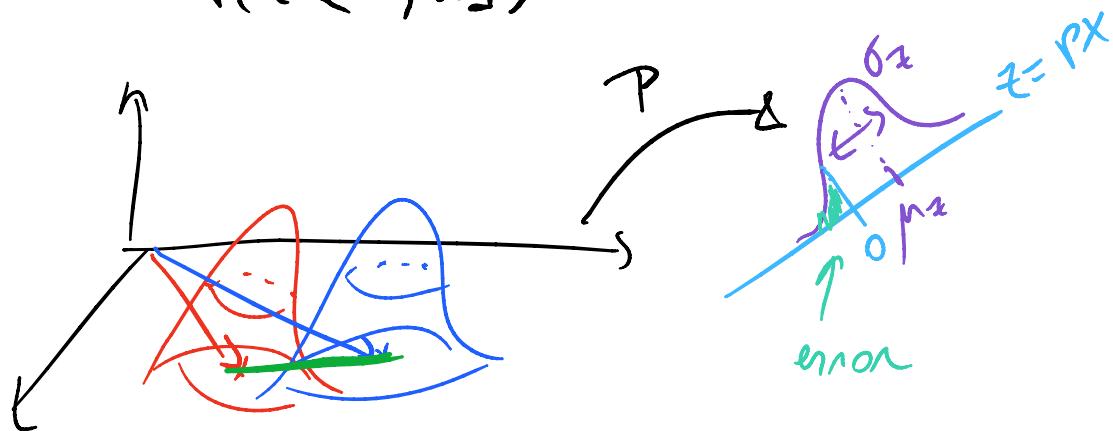
(2)

Everything is happening on the line

$$z = P^T X, \quad P \in \mu^T \Sigma^{-1}$$

The Bayes error is

$$\begin{aligned} L^* &= P(z > 0, w_2) + P(z \leq 0, w_1) \\ &= \frac{1}{2} P(z > 0 | w_2) + \frac{1}{2} P(z \leq 0 | w_1) \\ &= P(z \leq 0 | w_1) \end{aligned}$$



thus the Bayes error is the "projection" of $\mu_1(X|w_1)$ into the line $P^T X = z$, when $z \leq 0$.

z is a random variable, also normally distributed with

$$\begin{aligned} \mu_z &= E[P^T X] = P E[X] = P \mu = \mu^T \Sigma^{-1} \mu // \\ \sigma_z^2 &= E[(z - \mu_z)(z - \mu_z)^T] \\ &= E[P(X - \mu)(X - \mu)^T P^T] \\ &= P \sigma_X^2 P^T = (\mu^T \Sigma^{-1}) \Sigma (\Sigma^{-1} \mu) = \mu^T \Sigma^{-1} \mu // \end{aligned}$$

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Therefore $L^* = \int_{-\infty}^0 dz \frac{1}{\sqrt{2\pi}} \frac{1}{\sigma_z} e^{-\frac{1}{2} \frac{(x-\mu_z)^2}{\sigma_z^2}}$ (3)

$$L^* = \int_{-\infty}^{-\frac{\mu_z}{\sigma_z}} dy \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} y^2} = \int_{\frac{\mu_z}{\sigma_z}}^{\infty} dy \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} y^2}$$

Note that $\frac{\mu_z}{\sigma_z} = \frac{\mu^T \Sigma^{-1} \mu}{\sqrt{\mu^T \Sigma^{-1} \mu}} = \sqrt{\mu^T \Sigma^{-1} \mu}$, thus

$$\boxed{L^* = \int_{\sqrt{\mu^T \Sigma^{-1} \mu}}^{\infty} dy \frac{e^{-\frac{1}{2} y^2}}{\sqrt{2\pi}} = \frac{1}{2} - \frac{1}{2} \operatorname{erf}\left(\sqrt{\frac{\mu^T \Sigma^{-1} \mu}{2}}\right)}$$

$$\begin{aligned} \int_x^{\infty} dy \frac{e^{-\frac{1}{2} y^2}}{\sqrt{2\pi}} &= \int_0^{\infty} \dots - \int_0^x \dots \\ &= \frac{1}{2} - \int_0^{\frac{x}{\sqrt{2}}} \frac{\sqrt{2} dt}{\sqrt{\pi}} e^{-t^2} = \frac{1}{2} - \frac{1}{2} \operatorname{erf}\left(\frac{x}{\sqrt{2}}\right) \end{aligned}$$

$$\operatorname{erf}(x) \equiv \frac{2}{\sqrt{\pi}} \int_0^x e^{-t^2} dt$$

This is the general answer to the problem.

Just as a consistency check, if $\Sigma = I$ then $\sqrt{\mu^T \Sigma^{-1} \mu} = \sqrt{\mu^T \mu} = \|\mu\|$, and $L^* = \int_{\|\mu\|}^{\infty} dy \frac{e^{-y^2/2}}{\sqrt{\pi}}$ which is the error in Trunk's problem.

Now, all we need to do is to specify the above formula to this particular problem.

Since Σ is block diagonal we have

$$\Sigma^{-1} = \begin{pmatrix} \Sigma_a^{-1} & 0 \\ 0 & \Sigma_b^{-1} \end{pmatrix}$$

$$\text{thus } \mu^T \Sigma^{-1} \mu = (\mu_a^T \mu_b^T) \begin{pmatrix} \Sigma_a^{-1} & 0 \\ 0 & \Sigma_b^{-1} \end{pmatrix} \begin{pmatrix} \mu_a \\ \mu_b \end{pmatrix}$$

$$= \underbrace{\mu_a^T \Sigma_a^{-1} \mu_a}_1 + \underbrace{\mu_b^T \Sigma_b^{-1} \mu_b}_2$$

$$\Sigma_a = \begin{pmatrix} 1 & & & \\ & 2 & & \\ & & \ddots & \\ & & & n \end{pmatrix} \quad (\Sigma_b)_{ij} = \frac{\sqrt{ij}}{2^{l(i-j)}}$$

diagonal not diagonal, but symmetric.

Now if $(\mu_a)_i = \begin{cases} 1 & \text{for } i \leq k \\ 0 & \text{for } i > k \end{cases}$, we then have

$$\mu_a^T \Sigma_a^{-1} \mu_a = \sum_{i=1}^k \frac{1}{i} \quad (\text{divergent series})$$

Analogously, $\mu_b = \mu_a$, so

$$\mu_b^T \Sigma_b^{-1} \mu_b = \sum_{i=1}^k \sum_{j=1}^k (\Sigma_b^{-1})_{ij}$$

Numerically we can easily compute these things.

We have

$$f(K) = \mu^T \Sigma^{-1} \mu = \sum_{i=1}^k \frac{1}{i} + \sum_{i=1}^k \sum_{j=1}^k (\Sigma_b^{-1})_{ij}$$

where $(\Sigma_b^{-1})_{ij} = \frac{\sqrt{ij}}{2^{l(i-j)}}$. This matrix is invertible.

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Therefore,

$$L^*(k) = \frac{1}{2} - \frac{1}{2} \operatorname{erf}\left(\sqrt{\frac{R(k)}{2}}\right).$$

(5)

Now consider $k=5$ and any $D > k$.

- Only the first k rows and columns of Σ_b matter:

$$f(5) = \sum_{i=1}^5 \frac{1}{i} + \left(\begin{array}{cc|ccccc} \frac{4}{3} & -\frac{\sqrt{2}}{3} & 0 & 0 & 0 & 0 & \dots \\ -\frac{\sqrt{2}}{3} & \frac{5}{6} & -\frac{\sqrt{2}}{3} & 0 & 0 & 0 & \dots \\ 0 & -\frac{\sqrt{2}}{3} & \frac{5}{9} & -\frac{1}{6\sqrt{3}} & 0 & 0 & \dots \\ 0 & 0 & -\frac{1}{15} & \frac{5}{12} & -\frac{1}{3\sqrt{5}} & 0 & \dots \\ 0 & 0 & 0 & -\frac{1}{3\sqrt{5}} & \frac{1}{3} & -\frac{\sqrt{2}}{3\sqrt{15}} & \dots \\ \hline - & - & - & - & - & \frac{5}{18} & \dots \\ - & - & - & - & - & \vdots & \ddots \end{array} \right)$$

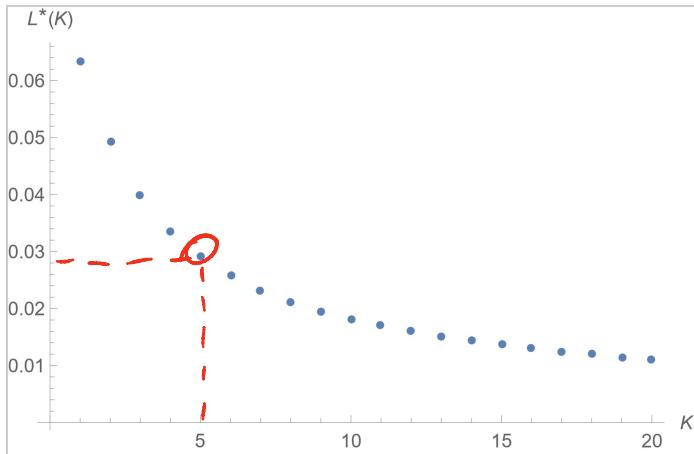
Sum of the elements in the blue block.

$$= \frac{259}{45} - \frac{2\sqrt{2}}{3\sqrt{3}} - \frac{2\sqrt{2}}{3} - \frac{2}{3\sqrt{3}} - \frac{2}{3\sqrt{5}}$$

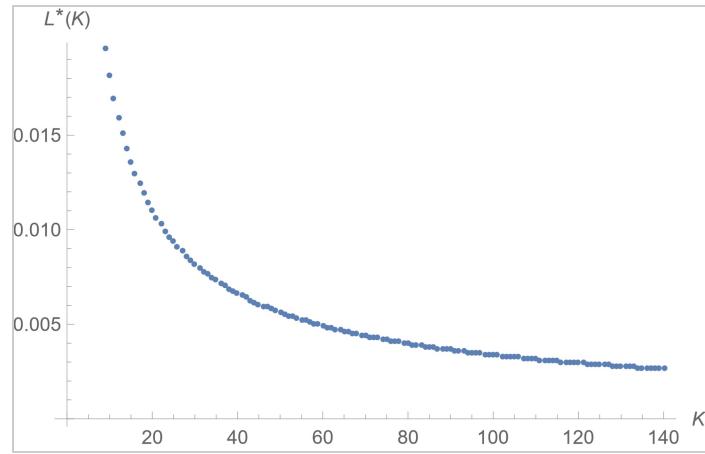
$$\text{Thus } L^* = \frac{1}{2} - \frac{1}{2} \operatorname{erf}\left(\sqrt{\frac{R(5)}{2}}\right) \approx 0.03.$$

- This does not change for any $D > k$.
- If k increases, $L^* \rightarrow 0$ since the first term in $f(k)$ is a divergent series.
- It may be possible to find $\bar{\Sigma}_b$ exactly, or at least approximately, but this would not change anything since the first term is already divergent. Thus I won't even bother doing that, but neither show the following plots.

⑥



(a)



(b)

In the above Fig. we plot $L^*(K)$ versus K , given by the above formula. we fix a large D_K , thus compute Σ_b^{-1} only once, and vary K . Notice that $L^* \rightarrow 0$ as $K \rightarrow \infty$ as shown in Fig (b).

(2) Consider the simple hypothesis test

$$H_0: f = f_A \quad \text{vs} \quad H_A: f = f_B$$

Under the rule $R_A(x)/R_0(x) > K \Rightarrow X \sim f_A$, and $R_A(x)/R_0(x) \leq K \Rightarrow X \sim f_0$. This leads to a hypothesis test with Type I error α , and Type II error β .

* Correction!

Now consider two transformations $S: X \mapsto S(X)$ and $T: X \mapsto T(X)$ such that we have

$$\frac{f_A(S(X))}{f_0(S(X))} > K_S \rightarrow \alpha_S, \beta_S$$

$$\frac{f_A(T(X))}{f_0(T(X))}$$

$$\frac{f_A(T(X))}{f_0(T(X))} > K_T \rightarrow \alpha_T, \beta_T$$

where $\alpha_S = \alpha_T$ and $\beta_S < \beta_T$. Moreover, all these tests are the most powerful.

7)

Show that, if $k_S = k_T$ then there exists priors π ($1-\pi$) such that $L_S^* < L_T^*$. What can be said if $k_S \neq k_T$?

Let $p_A = \pi$ and $p_0 = 1-\pi$ be the priors. Let's denote $X \in C_0$ if $X \sim f_0$, $X \in C_A$ if $X \sim f_A$. The decision boundary is determined when the posteriors satisfy

$$\frac{p(C_A | X)}{p(C_0 | X)} = \frac{p(C_0 | X)}{p(C_A | X)}$$

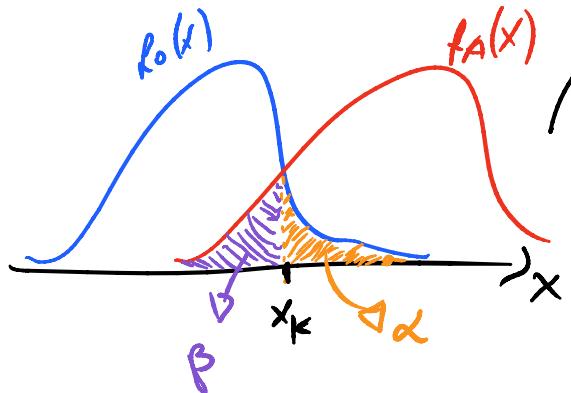
$$\frac{p(x | C_A) p_A}{p(x | C_0) p_0} = \frac{p(x | C_0)}{p(x | C_A)}$$

$$f_A(x)\pi = f_0(x)(1-\pi)$$

$$\frac{f_A(x)}{f_0(x)} = \frac{1-\pi}{\pi} = K$$

Thus $\boxed{\pi = \frac{1}{K+1}}$

Notice that



$$\boxed{L^* = \pi \beta + (1-\pi)\alpha}$$

since the test is MP.

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We thus have $\begin{cases} L_S^* = \pi_S \beta_S + (1-\pi_S) \alpha_S \\ L_T^* = \pi_T \beta_T + (1-\pi_T) \alpha_T \end{cases}$

(8)

Since $\alpha_S = \alpha_T$ we have

$$L_S^* - L_T^* = \pi_S \beta_S - \pi_T \beta_T + \alpha_S (\pi_T - \pi_S)$$

- If $k_S = k_T$ then $\pi_S = \pi_T$, therefore

$$L_S^* - L_T^* = \pi_S (\beta_S - \beta_T) < 0 //$$

since $\beta_S < \beta_T$.

- What if $k_S \neq k_T$? Let us consider the two possible cases separately.

(a) $k_S > k_T$. Then $\pi_S = \frac{1}{k_S+1} < \frac{1}{k_T+1} = \pi_T$. We still have $\beta_S < \beta_T$, thus

$$L_S^* - L_T^* = \underbrace{\pi_S \beta_S - \pi_T \beta_T}_{< 0} + \underbrace{\alpha_S (\pi_T - \pi_S)}_{> 0}$$

without further knowledge about $\beta_S, \beta_T, \alpha_S$ the result is inconclusive. Write $\beta_T = \beta_S + \varepsilon$. Then

$$\begin{aligned} L_S^* - L_T^* &= (\pi_S - \pi_T) \beta_S - \pi_T \varepsilon + \alpha_S (\pi_T - \pi_S) \\ &= \underbrace{(\pi_S - \pi_T)(\beta_S - \alpha_S)}_{< 0} - \underbrace{\pi_T \varepsilon}_{< 0} \end{aligned}$$

If $\boxed{\beta_S > \alpha_S}$ then certainly $L_S^* < L_T^*$.

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If $\boxed{\beta_S < \alpha_S}$ then $L_S^* < L_T^*$ only if

$$(\pi_S - \pi_T)(\beta_S - \alpha_S) < \pi_T \varepsilon$$

$$\beta_S - \alpha_S < \frac{\pi_T}{\pi_S - \pi_T} \varepsilon = \frac{1}{\frac{\pi_S}{\pi_T} - 1} \varepsilon$$

$$\alpha_S - \beta_S > \frac{1}{1 - \frac{\pi_S}{\pi_T}} \varepsilon = \frac{1}{1 - \frac{\pi_S}{\pi_T}} (\beta_T - \beta_S)$$

(9)

(b) $k_S < k_T$. Then $\pi_S = \frac{1}{k_S+1} > \frac{1}{k_T+1} = \pi_T$, thus

$$L_S^* - L_T^* = \underbrace{\pi_S \beta_S - \pi_T \beta_T}_{?} + \underbrace{\alpha_S (\pi_T - \pi_S)}_{< 0}$$

Again, without further assumptions it is inconclusive.

Write $\beta_T = \beta_S + \varepsilon$ (since $\beta_S < \beta_T$). Then

$$L_S^* - L_T^* = \underbrace{(\pi_S - \pi_T)(\beta_S - \alpha_S)}_{> 0} - \pi_T \varepsilon$$

If $\boxed{\beta_S \leq \alpha_S}$ then $L_S^* < L_T^*$.

If $\boxed{\beta_S > \alpha_S}$ then $L_S^* < L_T^*$ if

$$\pi_T \varepsilon > (\pi_S - \pi_T)(\beta_S - \alpha_S)$$

$$\beta_S - \alpha_S < \frac{\pi_T}{\pi_S - \pi_T} \varepsilon = \frac{\pi_T / \pi_S}{1 - \frac{\pi_T}{\pi_S}} \varepsilon$$

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Let us summarize everything.

We have transformations such that

$$\alpha_s = \alpha_T, \quad \beta_s < \beta_T \quad (\beta_T - \beta_s = \varepsilon)$$

$\nwarrow \beta$ is not power!

power = $J - \beta$ in
my convention.

Then $L_s^* < L_T^*$ provided:

$$(1) \quad k_s = k_T.$$

$$(2) \quad k_s > k_T \text{ with } \begin{cases} \beta_s > \alpha_s \text{ or} \\ \alpha_s - \beta_s > \frac{1}{1 - \frac{\pi s}{\pi T}} \varepsilon \text{ otherwise} \end{cases}$$

$$(3) \quad k_s < k_T \text{ with } \begin{cases} \beta_s \leq \alpha_s \text{ or} \\ \beta_s - \alpha_s < \frac{\pi T}{\pi s - \pi T} \varepsilon \text{ otherwise} \end{cases}$$

In any other situation $L_s^* \geq L_T^*$.