

I, Nevin Liang, with UID 705575353 have read and understood the policy on Academic Integrity

1.

Q function:  $1 - \Phi(a) = Q(a) = P(X > a)$

$$P(-1 < X < 3) = P(X > -1) - P(X > 3)$$

$$= \boxed{Q(-1) - Q(3)}$$

Note: technically  
 $P(X \geq 3)$  but  
Continuous so  
 $P(X=3) = 0$ .

2.

$$a) f(x) = \begin{cases} Ce^{-x}(1+x) & x \geq 0 \\ 0 & \text{else} \end{cases}$$

$$\int_{-\infty}^{\infty} f(x) dx = 1 \Rightarrow \int_0^{\infty} C \cdot e^{-x}(1+x) dx = 1$$

$$C \left[ \int_0^{\infty} e^{-x}(1+x) dx \right] = 1 \Rightarrow \lim_{t \rightarrow \infty} \int_0^t e^{-x}(1+x) dx = \frac{1}{C}$$

$$\lim_{t \rightarrow \infty} \left[ -e^{-x}(1+x) - \int -e^{-x} dx \right] \Big|_0^t = \lim_{t \rightarrow \infty} \left( -e^{-x}(1+x) - e^{-x} \Big|_0^t \right)$$

Integration  
by parts

$$= \lim_{t \rightarrow \infty} \left( -e^{-x}(x+2) \Big|_0^t \right) = \lim_{t \rightarrow \infty} \left( -\frac{t+2}{e^t} + \frac{2}{1} \right)$$

$$= 2 - \lim_{t \rightarrow \infty} \frac{t+2}{e^t} = 2 - \lim_{t \rightarrow \infty} \frac{1}{e^t} = \boxed{2}$$

$$2 = \frac{1}{C} \Rightarrow \boxed{C = \frac{1}{2}}$$

l'Hopital's.

$$b) E[X] = \int_0^{\infty} x \cdot f(x) dx = \int_0^{\infty} x \cdot C e^{-x}(1+x) dx = \frac{1}{2} \int_0^{\infty} x \cdot e^{-x}(1+x) dx$$

$$= \frac{1}{2} \lim_{t \rightarrow \infty} \int_0^t e^{-x}(x+x^2) dx \longrightarrow$$

Integrate by parts again

$\frac{1}{2} \lim_{t \rightarrow \infty} \int_0^t e^{-x} (x+x^2) dx$  ... let's ignore limit and bounds for now...

$$\begin{aligned} \int e^{-x} (x+x^2) dx &= -(x+x^2)e^{-x} - \int -e^{-x} (2x+1) dx \\ &= -(x+x^2)e^{-x} + \int e^{-x} (2x+1) dx \end{aligned}$$

integrate by parts again

$$\begin{aligned} &= -(x+x^2)e^{-x} - (2x+1)e^{-x} - \int -e^{-x} (2) dx \\ &= -(x+x^2+2x+1)e^{-x} + \int e^{-x} \cdot 2 dx \\ &= -(x^2+3x+1)e^{-x} + 2 \cdot (-e^{-x}) \\ &= -(x^2+3x+3)e^{-x} \end{aligned}$$

$$\frac{1}{2} \lim_{t \rightarrow \infty} \left( -(t^2+3t+3) \cdot e^{-t} + (3) \right) = \left[ 3 - \lim_{t \rightarrow \infty} \left( \frac{t^2+3t+3}{e^t} \right) \right] \cdot \frac{1}{2}$$

double l'Hopital's

$$= [3 - 0] \cdot \frac{1}{2} = \boxed{3/2}$$

3.

Discrete:

$$\Phi_x(\omega) = \sum_{x_i} e^{j\omega x_i} \cdot P(x=x_i)$$

$$= e^{j\omega \cdot 1} \cdot P(x=1) + e^{j\omega \cdot 2} \cdot P(x=2) \dots$$

$$= e^{j\omega} \cdot p + e^{j\omega \cdot 2} \cdot (1-p) \cdot p \dots$$

$$= \boxed{\frac{e^{j\omega} \cdot p}{1 - e^{j\omega} \cdot (1-p)}} \quad \text{char func.}$$

2nd deriv wrt  $\omega$ :

$$\frac{\partial}{\partial \omega} \left( \frac{\partial}{\partial \omega} \left( \frac{p}{e^{-j\omega} - (1-p)} \right) \right)$$

$$\downarrow$$

$$\frac{d}{d\omega} \left[ p (e^{-j\omega} - (1-p))^{-1} \right]$$

$$= -p (e^{-j\omega} - (1-p))^{-2} \cdot (-j \cdot e^{-j\omega})$$

$$= \frac{pj (e^{-j\omega})}{(e^{-j\omega} - (1-p))^2} = \frac{pj \cdot e^{j\omega}}{(1 - e^{j\omega}(1-p))^2}$$

$$\frac{\partial}{\partial \omega} \left( \frac{pj \cdot e^{j\omega}}{(1 - e^{j\omega}(1-p))^2} \right) = \cancel{pj \cdot j e^{j\omega} (1 - e^{j\omega}(1-p))^{-2}} \cdot e^{j\omega}$$

$$= pj \cdot \frac{j e^{j\omega} (1 - e^{j\omega}(1-p))^2 - e^{j\omega} (2(1 - e^{j\omega}(1-p))) \cdot (- (1-p) \cdot j e^{j\omega})}{(1 - e^{j\omega}(1-p))^4}$$

$$\text{when } \omega=0, \quad pj \cdot \frac{j (1 - (1-p))^2 - (2 \cdot (1 - (1-p))) \cdot (- (1-p) \cdot j)}{(1 - (1-p))^4}$$

$$= \frac{pj \cdot (jp^2 + 2p \cdot (1-p)j)}{p^4} = \frac{-p^3 - 2p^2(1-p)}{p^4}$$

$$= \frac{-2p^2 + p^3}{p^4} = \boxed{\frac{p-2}{p^2}} \quad \frac{p-2}{p^2}$$

HOWEVER

$$E[X^2] = \frac{1}{j^2} \cdot \frac{\partial^2 \Phi_x(\omega)}{\partial \omega^2} = -1 \cdot \frac{p-2}{p^2}$$

$$= \boxed{\frac{2-p}{p^2}}$$

4.

$$Y = \begin{cases} X & X > 0 \\ X^2 & X < 0 \end{cases}$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\frac{1}{4} \cdot \sigma^2 = 4$$

if  $X > 0$ :  $Y = X \Rightarrow f_Y(y) = \frac{1}{4\sqrt{2\pi}} \cdot e^{-\frac{y^2}{2 \cdot 8}}$  ✓

$$\frac{1}{4\sqrt{2\pi}} \cdot e^{-\frac{y^2}{8}}$$

if  $X < 0$ :  $Y = X^2 \Rightarrow$  chi-squared with  $k=1$ :

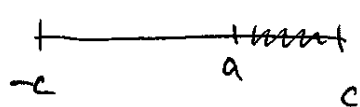
$$f_Y(y) = \frac{y^{-1/2} \cdot e^{-y/2}}{2^{1/2} \cdot \Gamma(1/2)} \quad \swarrow \text{chi-square}$$

$$= \frac{e^{-y/2}}{\sqrt{y} \cdot \sqrt{2} \cdot \sqrt{\pi}} = \boxed{\frac{e^{-y/2}}{\sqrt{2\pi y}}}$$

$$\text{PDF} = \begin{cases} \frac{1}{4\sqrt{2\pi}} \cdot e^{-\frac{y^2}{8}} + \frac{e^{-y/2}}{\sqrt{2\pi y}} & Y > 0 \\ 0 & Y < 0 \end{cases}$$

5.

(a)  $P(X > c) = \boxed{\frac{a-c}{2a}}$   $\rightarrow$   $\begin{cases} 1 & c < -a \\ \frac{a-c}{2a} & -a \leq c \leq a \\ 0 & c > a \end{cases}$



(b)  $P(|X-m| > a) \leq \frac{\sigma^2}{a^2} \dots$  cheby

uniform RV  $[-a, a]$  so  $m=0$   
 $a=c$

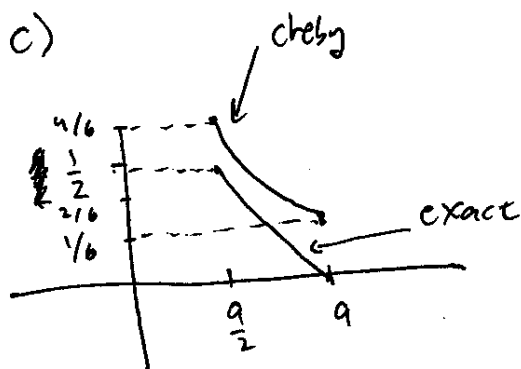
$$P(|X| \geq c) \leq \frac{\sigma^2}{c^2}$$

$$\text{VAR} = \frac{(a - (-a))^2}{12} = \frac{a^2}{3}$$

$$P(|X| \geq c) \leq \frac{a^2}{3c^2}$$

BUT we want  $P(X > c) = P(X > c) = \frac{1}{2} P(|X| > c)$

c)  $\xrightarrow{\text{continuous}} = \boxed{\frac{a^2}{6c^2}}$



$$\frac{a^2}{6 \cdot \frac{a^2}{4}} = \frac{2}{3}$$

$$\frac{a^2}{6 \cdot a^2} = \frac{1}{6}$$

6.

$$(a) \mathbb{E}[Y] = \mathbb{E}[X+1] = \mathbb{E}[X] + 1 = \boxed{\frac{1}{\lambda} + 1}$$

$$\text{VAR}[Y] = \text{VAR}[X+1] = \mathbb{E}[(X+1)^2] - \mathbb{E}[X+1]^2$$

$$= \mathbb{E}[X^2] + 2 \cdot \mathbb{E}[X] + 1 - \mathbb{E}[X+1]^2$$

$$= \cancel{\frac{1}{\lambda^2}} + 2 \cdot \frac{1}{\lambda} + 1 - \left(\frac{1}{\lambda} + 1\right)^2$$

$$\mathbb{E}[X^2]$$

$$= \mathbb{E}[X^2] + \cancel{\frac{2}{\lambda}} + \cancel{1} - \frac{1}{\lambda^2} - \cancel{\frac{2}{\lambda}} - \cancel{1} =$$

$$\mathbb{E}[X^2] - \frac{1}{\lambda^2}$$

$$= \frac{1}{\lambda^2} + \mathbb{E}[X]^2 - \frac{1}{\lambda^2} = \mathbb{E}[X]^2 = \boxed{\frac{1}{\lambda^2}}$$

makes sense.  
when shifting,  $\sigma$   
does not change

$$(b) P(Y > y) = P(X+1 > y)$$

$$= P(X > y-1) = e^{-\lambda(y-1)}$$

$$\text{CDF} = 1 - P(Y > y) = 1 - e^{-\lambda(y-1)}$$

$$\text{CDF} = \begin{cases} 1 - e^{-\lambda(y-1)} & y \geq 1 \\ 0 & \text{elsewhere.} \end{cases}$$



(c)  $z = 1 - e^{-\lambda(y-1)}$

$$e^{-\lambda(y-1)} = 1-z$$

$$-\lambda(y-1) = \ln(1-z)$$

$$y-1 = -\frac{1}{\lambda} \cdot \ln(1-z)$$

$$y = 1 - \frac{1}{\lambda} \cdot \ln(1-z)$$

we only want  $0 \leq z \leq 1$

$$\left\{ \begin{array}{ll} \text{---} & z < 0 \\ 1 - \frac{1}{\lambda} \cdot \ln(1-z) & 0 \leq z < 1 \\ \infty & z \geq 1 \quad (z=1) \end{array} \right.$$

- (d) Step 1. generate a pseudorandom # in range  $[0, 1]$ .
- 2: Use the answer from (c) to get  $F_Y^{-1}(z)$ .
- $0 \leq z < 1$  case
- 3: the output of  $F_Y^{-1}$  will have distribution of the shifted exponential  $F_Y$