EE 131A Probability and Statistics Instructor: Lara Dolecek Homework 2 Solution Monday, January 11, 2021 Due: Wednesday, January 20, 2021 before class begins levtauz@ucla.edu debarnabucla@ucla.edu

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You may type your homework or scan your handwritten version. Make sure all the work is discernible.

Reading: Chapters 2.4-2.5 & 3 of *Probability, Statistics, and Random Processes* by A. Leon-Garcia

1. Roll two fair dice independently. In terms of the possible outcomes, define the events:

$$A = \{ \text{First die is 1, 2 or 3} \}$$

 $B = \{ \text{First die is 2, 3 or 6} \}$
 $C = \{ \text{Sum of outcomes is 9} \}$

Are A, B, and C mutually independent? Hint: Three events A, B, and C are independent if all the four following constraints hold:

$$P(A \cap B) = P(A)P(B),$$

$$P(A \cap C) = P(A)P(C),$$

$$P(B \cap C) = P(B)P(C),$$

$$P(A \cap B \cap C) = P(A)P(B)P(C).$$

Solution: Sample space is $S = \{(i, j) | 1 \le i, j \le 6\}$. And,

$$A = \{(i, j) : i = 1, 2, 3 \text{ and } 1 \le j \le 6\},\$$

 $B = \{(i, j) : i = 2, 3, 6 \text{ and } 1 \le j \le 6\},\$
 $C = \{(i, j) : i + j = 9\} = \{(3, 6), (4, 5), (5, 4), (6, 3)\}.$

So, |S| = 36, |A| = 18, |B| = 18, and |C| = 4. Since the two fair dice rolls are independent, $P(A) = \frac{18}{36} = \frac{1}{2}$, $P(B) = \frac{18}{36} = \frac{1}{2}$, and $P(C) = \frac{4}{36} = \frac{1}{9}$.

$$A \cap B = \{(i, j) | i = 2, 3 \text{ and } 1 \le j \le 6\},\$$

 $A \cap C = \{(3, 6)\},\$
 $B \cap C = \{(3, 6), (6, 3)\},\$
 $A \cap B \cap C = \{(3, 6)\}.$

Hence, $P(A \cap B) = \frac{12}{36} = \frac{1}{3}$, $P(A \cap C) = \frac{1}{36}$, $P(B \cap C) = \frac{2}{36} = \frac{1}{18}$, $P(A \cap B \cap C) = \frac{1}{36}$. Three events A, B, and C are independent if all the four following constraints hold:

$$P(A \cap B) = P(A)P(B),\tag{1}$$

$$P(A \cap C) = P(A)P(C), \tag{2}$$

$$P(B \cap C) = P(B)P(C), \tag{3}$$

$$P(A \cap B \cap C) = P(A)P(B)P(C). \tag{4}$$

Observe that (1) and (2) do not hold. Therefore, A, B, and C are **not** independent.

- 2. Assume there are 5 jars numbered 1 to 5. The $i^{\rm th}$ jar contains i black balls, 6-i red balls, and 5 green balls. A jar is selected uniformly at random and a ball is selected uniformly at random from that jar. Let the events B, R, and G represent the events that a black, red, or green ball is chosen, respectively. Let J_k represent the event that the $k^{\rm th}$ jar is chosen.
 - (a) What is $P(B|J_k)$?

Solution: By the problem definition, $P(B|J_k) = \frac{k}{11}$.

(b) What is P(G), P(B), and P(R)?

Solution:Regardless of which jar is chosen, the green balls always make up 5 of the 11 available balls. Hence, $P(G) = \frac{5}{11}$.

By symmetry, P(B) = P(R). Therefore,

$$1 = P(B) + P(R) + P(G) = 2 \cdot P(B) + \frac{5}{11}$$
$$\implies P(B) = \frac{3}{11}$$

Hence, $P(B) = P(R) = \frac{3}{11}$.

(c) Given that the selected ball is black, what is the probability that the ball came from the kth jar, i.e. $P(J_k|B)$?

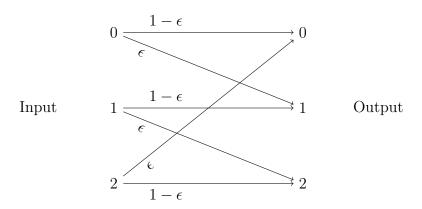
Solution:By Bayes rule, we have

$$P(J_k|B) = \frac{P(B|J_k)P(J_k)}{P(B)}$$

By using the values determined in previous parts, we get

$$P(J_k|B) = \frac{\frac{k}{11} \cdot \frac{1}{5}}{\frac{3}{11}} = \frac{k}{15}$$

3. A ternary communication channel is shown in the figure. Assume that input symbols 0, 1, and 2 are chosen for transmission with probabilities $\frac{1}{4}$, $\frac{1}{2}$, and $\frac{1}{4}$, respectively.



(a) Calculate the probability of each output.

Solution:

$$\begin{split} P(\text{Output} \ = 0) &= P(\text{Output} \ = 0 | \text{Input} \ = 0) P(\text{Input} \ = 0) \\ &+ P(\text{Output} \ = 0 | \text{Input} \ = 1) P(\text{Input} \ = 1) \\ &+ P(\text{Output} \ = 0 | \text{Input} \ = 2) P(\text{Input} \ = 2) \\ &= \frac{1}{4}(1-\epsilon) + 0 \cdot \frac{1}{2} + \frac{1}{4}\epsilon = \frac{1}{4} \end{split}$$

$$\begin{split} P(\text{Output} \ = 1) &= P(\text{Output} \ = 1|\text{Input} \ = 0)P(\text{Input} \ = 0) \\ &+ P(\text{Output} \ = 1|\text{Input} \ = 1)P(\text{Input} \ = 1) \\ &+ P(\text{Output} \ = 1|\text{Input} \ = 2)P(\text{Input} \ = 2) \\ &= \frac{1}{4}\epsilon + \frac{1}{2}(1-\epsilon) + 0 \cdot \frac{1}{4} = \frac{2-\epsilon}{4} \end{split}$$

$$\begin{split} P(\text{Output} &= 2) = P(\text{Output} = 2|\text{Input} = 0)P(\text{Input} = 0) \\ &+ P(\text{Output} = 2|\text{Input} = 1)P(\text{Input} = 1) \\ &+ P(\text{Output} = 2|\text{Input} = 2)P(\text{Input} = 2) \\ &= 0 \cdot \frac{1}{4} + \frac{1}{2}\epsilon + \frac{1}{4}(1-\epsilon) = \frac{1}{4} + \frac{1}{4}\epsilon = \frac{1+\epsilon}{4} \end{split}$$

(b) Given that the output was 1, what is the probability that the input was 0? 1? 2? Solution:

We want to figure out P(Input = k | Output = 1). By Bayes rule, we get

$$P(\text{Input} = k | \text{Output} = 1) = \frac{P(\text{Output} = 1 | \text{Input} = k)P(\text{Input} = k)}{P(\text{Output} = 1)}.$$

All the necessary terms were calculated in part (a) which gives us the result

$$\begin{split} P(\text{Input} &= 0 | \text{Output} = 1) = \frac{\frac{1}{4}\epsilon}{\frac{2-\epsilon}{4}} = \frac{\epsilon}{2-\epsilon} \\ P(\text{Input} &= 1 | \text{Output} = 1) = \frac{\frac{1}{2}(1-\epsilon)}{\frac{2-\epsilon}{4}} = \frac{2-2\epsilon}{2-\epsilon} \\ P(\text{Input} &= 2 | \text{Output} = 1) = 0 \end{split}$$

4. A family has 5 natural children and has adopted 2 girls. Each natural child has equal probability of being a girl or a boy, independent of the other children. Find the PMF of the number of girls out of the 7 children.

Solution: First, let N be the number of natural children that are girls. Let us first consider how we would get the pmf of N. Note that N can only take values in $\{0,1,2,3,4,5\}$ since there are only 5 natural children. Now, consider the probability that N=k. This means that we want to figure out the probability that k out of the 5 children were born as girls and that 5-k were born as boys. There are $\binom{5}{k}$ ways to select k out of 5 children to be girls and that the probability that those k children are girls and that the other 5-k children are boys is $(P(\text{Child is girl}))^k(P(\text{Child is boy}))^{5-k} = \frac{1}{2^k} \frac{1}{2^{5-k}} = \frac{1}{2^5}$.

Combining this together, we get the pmf of N as

$$P(N = k) = \begin{cases} \binom{5}{k} \frac{1}{2^5} & \text{if } 0 \le k \le 5\\ 0 & \text{else} \end{cases}$$

This distribution is known as a binomial distribution.

Since there are already 2 definite girls, we just need to shift this distribution. Thus, the pmf for G which is the number of girls out of the 7 children is

$$P(G = g) = \begin{cases} \binom{5}{g-2} \frac{1}{2^5} & \text{, if } 2 \le g \le 7\\ 0 & \text{else} \end{cases}$$

- 5. Throw a pair of six-sided dice. Let X_1 be the number of dots on the resulting face of the first die and let X_2 be the number of dots on the resulting face of the second die. Let $Z = X_1 + X_2$ be the sum of the two dice rolls.
 - (a) What is the pmf of Z?

Solution:

$$P(Z=z) = \begin{cases} \frac{z-1}{36} & z \in \{2, 3, 4, 5, 6, 7\} \\ \frac{13-z}{36} & z \in \{8, 9, 10, 11, 12\} \\ 0 & \text{otherwise} \end{cases}$$

(b) What is $\mathbb{E}[Z]$? Var(Z)? **Solution:** By linearity of expectation, we get

$$\mathbb{E}[Z] = \mathbb{E}[X_1 + X_2]$$

= $\mathbb{E}[X_1] + \mathbb{E}[X_2]$
= $3.5 + 3.5 = 7$

Note that X_1 and X_2 are independent, thus we get

$$Var(Z) = Var(X_1 + X_2)$$

$$= Var(X_1) + Var(X_2)$$

$$= \frac{35}{12} + \frac{35}{12} = \frac{70}{12}$$

(c) Given that Z = 10, what is the probability that $X_1 = k$ for $k \in \{1, 2, 3, 4, 5, 6\}$? Solution:

By Bayes rule,

$$P(X_1 = k | Z = 10) = \frac{P(Z = 10 | X_1 = k)P(X_1 = k)}{P(Z = 10)}.$$

From part (a), we know $P(Z=10)=\frac{3}{36}$. Additionally, we know $P(X_1=k)=\frac{1}{6}$. Due to $Z=X_1+X_2$ and the independence of X_1 and X_2 , we can write $P(Z=10|X_1=k)=P(X_2=10-X_1|X_1=k)=P(X_2=10-k)$. Note that

$$P(X_2 = 10 - k) = \begin{cases} \frac{1}{6} & k \in \{4, 5, 6, 7, 8, 9\} \\ 0 & \text{otherwise} \end{cases}.$$

Combining these results together, we get

$$P(X_1 = k | Z = 10) = \begin{cases} \frac{1}{3} & k \in \{4, 5, 6\} \\ 0 & \text{otherwise} \end{cases}$$

6. Assume that we flip a biased coin with probability of heads being p until a 2nd head is seen. Let X be the number of flips up until and including the flip that has the 2nd head. What is the pmf and expectation of X?

Solution:

Suppose X = n. Then the n^{th} flip must be a head. Of the first (n-1) flips, there must be exactly 1 head. Since there are (n-1) ways of getting exactly 1 head in (n-1) flips, the probability that X = n is

$$P(X = n) = \begin{cases} (n-1)p^{2}(1-p)^{n-2} &, n = 2, 3, 4, 5, \dots \\ 0 & else \end{cases}$$

Now, we find the expectation of X.

$$\mathbb{E}[X] = \sum_{n=2}^{\infty} n \cdot P(X = n)$$

$$= \sum_{n=2}^{\infty} n(n-1)p^2(1-p)^{n-2}$$

$$= p^2 \sum_{n=2}^{\infty} n(n-1)(1-p)^{n-2}$$

$$= p^2 \frac{d^2}{dp^2} (\sum_{n=2}^{\infty} (1-p)^n)$$
Differentiation Trick
$$= p^2 \frac{d^2}{dp^2} (\frac{(1-p)^2}{1-(1-p)})$$
Geometric Sum Formula
$$= p^2 \frac{d^2}{dp^2} (\frac{1-2p+p^2}{p})$$

$$= p^2 \frac{d^2}{dp^2} (\frac{1}{p} - 2 + p)$$

$$= p^2 (\frac{2}{p^3}) = \frac{2}{p}$$

Another way we could have arrived at this expectation is by using the linearity of expectation. Observe that the number of flips until the first head arrives is a geometric random variable regardless of when you start counting. As such, the number of flips from the first head to the second head is also a geometric random variable. Hence, X can be written as $X = Y_1 + Y_2$ where Y_1 and Y_2 are geometric random variables with parameter p.

Hence,

$$\mathbb{E}[X] = \mathbb{E}[Y_1] + \mathbb{E}[Y_2] = \frac{1}{p} + \frac{1}{p} = \frac{2}{p}$$