

Heirarchical Linear Modeling

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►

In [1]:

```
install.packages("lmerTest") # to get p-values
library(lmerTest)
library(tidyverse, quietly=T)
library("lme4", quietly=T)

# useful for computing ICC
library("merTools", quietly=T)
```

package 'lmerTest' successfully unpacked and MD5 sums checked

The downloaded binary packages are in
C:\Users\nnfon\AppData\Local\Temp\RtmpcXKKJv\downloaded_packages

Loading required package: Matrix
Loading required package: lme4

Attaching package: 'lmerTest'

The following object is masked from 'package:lme4':

lmer

The following object is masked from 'package:stats':

step

Loading tidyverse: ggplot2
Loading tidyverse: tibble
Loading tidyverse: tidyr
Loading tidyverse: readr
Loading tidyverse: purrr
Loading tidyverse: dplyr
Conflicts with tidy packages -----

expand(): tidyr, Matrix
filter(): dplyr, stats
lag(): dplyr, stats

Attaching package: 'MASS'

The following object is masked from 'package:dplyr':

select

arm (Version 1.9-3, built: 2016-11-21)

Working directory is C:/Users/nnfon/Desktop/INFO 4340 - Data Mining and Viz



In [2]:

```
dat <- read.csv("NEW_NELS_CLEAN.csv")
names(dat) <- tolower(names(dat))
head(dat)
```

stu_id	sch_id	mathscore	family_inc	urban
7898401	81	48.26	11	1
7898402	81	37.66	12	1
7898406	81	60.32	12	1
7898407	81	49.96	8	1
7898418	81	51.64	9	1
7898424	81	57.29	9	1

Introduction to Heirarchical Linear Modeling (HLM)

- In HLM, there is regression equation at level 1, which models a y variable and regression equations at level 2 which model the intercept and slope
- The full model has a predictor at level 1 and a predictor at level 2 and the intercepts (β_{0j}) and slopes(β_{1j}) are allow to vary across level 2 units.
- The intercepts has fixed components γ_{00} = mean of intercepts and a variance component $\tau_{00} = \text{var}(\beta_{0j}) = \text{var}(u_{0j})$ where u_{0j} is the random variation of the intercept of a unit
- The slope has fixed components γ_{10} = mean of slopes, and a variance component $\tau_{11} = \text{var}(\beta_{1j}) = \text{var}(u_{1j})$ where u_{1j} is the random variation of the slope of a unit.
- The covariance between slopes and intercepts is τ_{01} or $\tau_{10} = \text{cov}(\beta_{0j}, \beta_{1j})$

Tips for Formulating a Heirarchical Linear Model

- There is always a fixed intercept γ_{00} ,
- If a predictor exists at level 1, there is always a fixed slope, γ_{10} ,
- We can make the intercepts or slopes vary across level 2 units by including the random variation of the intercept u_{0j} or random variation of the slopes u_{1j}

Tips for Formulating a Heirarchical Linear Model in R

- Use 1 to indicate the intercept
- To make the intercept vary, use 1 and the group (or unit) you want the effect to vary across: (1|level2_unit)
- To make the slope vary, use the variable associated with the slope and the unit you want the effect to vary across: (variable_name|level2_unit)

Full Model

A full model has predictors at level 1 and level 2 as well as both fixed and random compoents for the slope and intercept.

Level-1 Model

$$\text{science}_{ij} = \beta_{0j} + \beta_{1j}(\text{family_inc}_{ij}) + r_{ij}$$

Level-2 Model

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(\text{schttyp}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(\text{schttyp}_j) + u_{1j}$$

Model with Predictor at Level 1

- Fixed Slopes only, Random Intercept

Level-1 Model

$$\text{science}_{ij} = \beta_{0j} + \beta_{1j}(\text{family_inc}_{ij}) + r_{ij}$$

Level-2 Model

$$\beta_{0j} = \gamma_{00} + u_{0j}$$

$$\beta_{1j} = \gamma_{10}$$



In [3]:

```
model.1 <- lmer(mathscore ~ 1 + family_inc + (1|sch_id), data=dat, REML =  
summary(model.1)
```

Linear mixed model fit by maximum likelihood t-tests use Sat
terthwaite

approximations to degrees of freedom [lmerMod]
Formula: mathscore ~ 1 + family_inc + (1 | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
59944.7	59972.8	-29968.3	59936.7	8300

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.10428	-0.73893	-0.01012	0.73980	2.91628

Random effects:

Groups	Name	Variance	Std.Dev.
sch_id	(Intercept)	15.35	3.918
Residual		72.44	8.511

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	4.069e+01	4.261e-01	5.017e+03	95.50	<2e-16

family_inc	1.060e+00	4.053e-02	7.983e+03	26.15	<2e-16

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

	(Intr)
family_inc	-0.903

Interpretation

- Random Effect: The random effect estimate for this model is the variance of the intercept across different schools and this variance is 15.40.
- Fixed Effect: The fixed effect of the intercept is 40.68

Coefficients

- We can see that the intercepts vary and the coefficients are fixed.



In [4]:

```
coef <- coef(model.1)$sch_id  
head(coef)
```

	(Intercept)	family_inc
81	42.12045	1.059821
391	39.26540	1.059821
421	36.19551	1.059821
461	39.28299	1.059821
615	37.29476	1.059821
750	49.74479	1.059821



In [5]:

```
# change column names  
names(coef) <- c("intercept", "slope_family_inc")  
head(coef)
```

	intercept	slope_family_inc
81	42.12045	1.059821
391	39.26540	1.059821
421	36.19551	1.059821
461	39.28299	1.059821
615	37.29476	1.059821
750	49.74479	1.059821

Significant Test For a Fixed Effect

Raudenbush and Bryk (2002), and therefore the HLM software, use a t -distribution to evaluate this ratio (Fotiu, 1989).

$$t = \frac{\hat{\gamma}_h}{S.E.(\hat{\gamma}_h)}$$

where γ_h is either the intercept or slope coefficient and $S.E.(\gamma_h)$ is the standard error estimate.¹ The fixed effects hypothesis tests (whether for level-1 or level-2 predictors) used by the HLM software use a degrees of freedom based on the number of level-2 units (i.e., number of groups).

$$df = N - q - 1,$$

N= number of groups

q= number of predictors

http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf
(http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf)

- Test of significance for a fixed effect is basically testing whether the mean estimate for the intercept or slope significantly differ from 0

»

In [6]:

```
intercepts = coef$intercept  
t.test(intercepts)
```

One Sample t-test

```
data: intercepts  
t = 414.03, df = 853, p-value < 2.2e-16  
alternative hypothesis: true mean is not equal to 0  
95 percent confidence interval:  
 40.49337 40.87912  
sample estimates:  
mean of x  
 40.68624
```



In [7]:

```
mean(intercepts)
```

40.6862423006384

Test of Significance for Random Effects

The chi-square test. The chi-square test used in the HLM package is based on the deviation of group means from the grand mean, given in Raudenbush and Bryk (2002, p.64) as:

$$\chi^2 = \frac{\sum_j \left(\hat{\beta}_{qj} - \hat{\gamma}_{q0} - \sum_{s=1}^{s_q} \hat{\gamma}_{qs} W_{sj} \right)^2}{\hat{V}_{qqj}} .$$

In the above formula, β is the group estimate (intercept or slope), γ is the average estimate (grand mean or average slope), and W is a predictor. The numerator in the equation represents the sum of squared deviations from the average value adjusting for the predictors in the model. The denominator, V_{qqj} , is a variance error estimate (i.e., square of the standard error). Degrees of freedom for this test are $J - S_q - 1$, where J is the number of groups and S_q is the number of predictors in the model (in Snijders & Bosker, 2012, this is $N - q - 1$). Small groups are omitted from the computations (the number omitted is noted in the HLM output).

http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf
[\(http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf\)](http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf)

The variance of a random effect is basically the deviation of the slope or intercept from the grand mean or mean of intercept, adjusting for effect of the predictor at level 2.



In [8]:

```
intercepts = coef$intercept
gamma00 = mean(intercepts)
gamma01.predictor = 0 # there is no predictor at level 2 so gamma01*pred
variance_00_error = var(intercepts)/length(intercepts)

deviations = (intercepts - gamma00 - 0)^2
chi_squared = sum(deviations)/variance_00_error
chi_squared
```

728462



In [9]:

```
# df = dree of freedom = number of groups - number of level2 predictors -
pchisq(q = 728462, df =854-0-1, lower.tail = FALSE)
```

0

Model Fit Indices

- Chi-square $-2 * \ell$
- Akaike Information Criteria

$$AIC = -2 * \ell + 2K$$
- Bayesian Informaiton Criteria

$$BIC = -2 * \ell + K * \ln(N)$$

Model comparison

When we compare two models, we find the change in deviance of the models. Normally models with smaller deviance have a better fit. We can find the change in deviance and that gives a chi-statistic with a degree of freedom being equal to the difference between the parameters in the model. We can then find if that chi-squared statistics is significant.

http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf
http://web.pdx.edu/~newsomj/mlrclass/ho_significance.pdf

Likelihood Ratio Test

This test that compares the deviance of two models is called the likelihood ratio test. The deviance of a model is calculated as deviance = $-2 * \log \text{likelihood}$ or

$$-2 \log (p(y \mid \hat{\theta}_0))$$



In [10]:

```
deviance(model.1)
```

59936.6860974382

Model with Predictor at Level 1

- Random Slopes, Random Intercept



In [11]:

```
model.2 <- lmer(mathscore ~ 1 + family_inc + (family_inc|sch_id),
               data=dat, REML = FALSE)
summary(model.2)
```

Linear mixed model fit by maximum likelihood t-tests use Sat
terthwaite

approximations to degrees of freedom [lmerMod]
Formula: mathscore ~ 1 + family_inc + (family_inc | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
59940.6	59982.7	-29964.3	59928.6	8298

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.14262	-0.74398	-0.01008	0.73675	2.90751

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
sch_id	(Intercept)	6.91135	2.6289	
	family_inc	0.01718	0.1311	1.00
Residual		72.44775	8.5116	

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	4.060e+01	4.044e-01	1.458e+03	100.4	<2e-16

family_inc	1.061e+00	4.048e-02	7.210e+03	26.2	<2e-16

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

	(Intr)
family_inc	-0.895





In [12]:

```
anova(model.1, model.2)
```

	Df	AIC	BIC	logLik	deviance	Chisq	Chi Df	Pr(>Chisq)
object	4	59944.69	59972.78	-29968.34	59936.69	NA	NA	NA
..1	6	59940.57	59982.72	-29964.28	59928.57	8.117954	2	0.01726668



In [13]:

```
# check how chi-squared was computed  
deviance(model.1) - deviance(model.2)
```

8.11795381792035



In [14]:

```
# we can also get the p-value too  
pchisq(q=452.1748, df=2, lower.tail = FALSE)
```

6.47872807522628e-99

There is a significant difference between the two models, the model with lower deviance is better, which is model 2.

Note that, we can generate a model with random slopes and random intercepts without specifying the intercepts in the model. By default, R always generates random intercepts with the lmer() function.



In [15]:

```
m <- lmer(mathscore ~ family_inc + (family_inc | sch_id),
          data=dat, REML = FALSE)
summary(m)
```

Linear mixed model fit by maximum likelihood t-tests use Sat
terthwaite

approximations to degrees of freedom [lmerMod]

Formula: mathscore ~ family_inc + (family_inc | sch_id)

Data: dat

AIC	BIC	logLik	deviance	df.resid
59940.6	59982.7	-29964.3	59928.6	8298

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.14262	-0.74398	-0.01008	0.73675	2.90751

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
sch_id	(Intercept)	6.91135	2.6289	
	family_inc	0.01718	0.1311	1.00
Residual		72.44775	8.5116	

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	4.060e+01	4.044e-01	1.458e+03	100.4	<2e-16

family_inc	1.061e+00	4.048e-02	7.210e+03	26.2	<2e-16

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

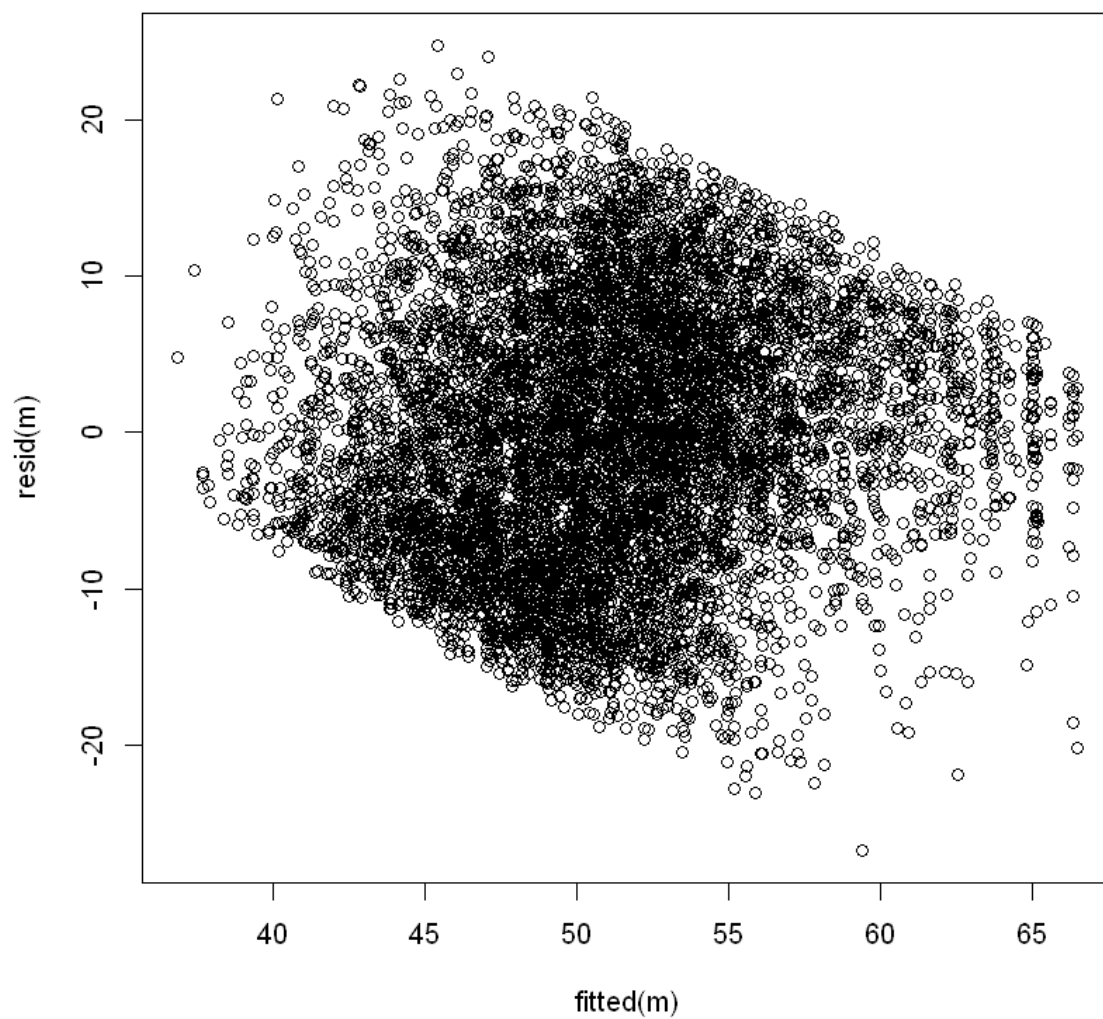
	(Intr)
family_inc	-0.895





In [16]:

```
# check residuals  
plot(resid(m) ~ fitted(m))
```



Null Model

This is a model with no predictor at level 1 or level 2. Intercepts is random. There is no slope since there is no predictor at level 1.



In [17]:

```
model.3 <- lmer(mathscore ~ 1 + (1|sch_id),  
               data=dat, REML = FALSE)  
summary(model.3)
```

summary from lme4 is returned
some computational error has occurred in lmerTest

Linear mixed model fit by maximum likelihood ['lmerMod']
Formula: mathscore ~ 1 + (1 | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
60548.8	60569.9	-30271.4	60542.8	8301

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.70061	-0.75164	-0.00107	0.70824	2.93238

Random effects:

Groups	Name	Variance	Std.Dev.
sch_id	(Intercept)	27.49	5.243
Residual		75.43	8.685

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	50.7035	0.2259	224.5

Predictor at Level 1, Random Slope, Fixed intercept

»

In [18]:

```
model.4 <- lmer(mathscore ~ 1 + (family_inc|sch_id),
                  data=dat, REML = FALSE)
summary(model.4)
```

summary from lme4 is returned
some computational error has occurred in lmerTest

Linear mixed model fit by maximum likelihood ['lmerMod']
Formula: mathscore ~ 1 + (family_inc | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
60390.7	60425.9	-30190.4	60380.7	8299

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.93050	-0.74194	-0.00765	0.71750	2.83936

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
sch_id	(Intercept)	92.867	9.637	
	family_inc	1.012	1.006	-0.89
Residual		71.726	8.469	

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	50.1727	0.2055	244.1

»

In [19]:

```
# compare model 2 and 3 to 1
anova(model.1, model.2, model.3)
```

	Df	AIC	BIC	logLik	deviance	Chisq	Chi Df	Pr(>Chisq)
..2	3	60548.84	60569.91	-30271.42	60542.84	NA	NA	NA
object	4	59944.69	59972.78	-29968.34	59936.69	606.151586	1	7.687611e-134
..1	6	59940.57	59982.72	-29964.28	59928.57	8.117954	2	1.726668e-02

Intra class correlation (ICC)

This is a measure of variability between groups. This is important to compute before even trying to apply a model that assumes that groups vary.

$$ICC = \rho = \frac{\tau_0^2}{\sigma^2 + \tau_0^2}$$

- Sigma squared is variance within the groups and tau squared is variance across the groups with respect to the dependent variable.
- So, ICC captures the proportion of variance in the dependent variable due to the groups. The variance in the dependent variable can be partitioned into variance within groups and variance across groups.

►

In [20]:

```
ICC(outcome = "mathscore", group = "sch_id", data=dat)
```

0.267488882772605

►

In [21]:

```
coef2 <- coef(model.2)$sch_id  
names(coef2) <- tolower(c("intercept", "family_inc_slope"))  
head(coef2)
```

	intercept	family_inc_slope
81	41.61890	1.1114946
391	39.65244	1.0134415
421	37.43830	0.9030384
461	39.62227	1.0119368
615	38.41604	0.9517909
750	46.57756	1.3587465

Visualize Random Intercepts and Random Slopes

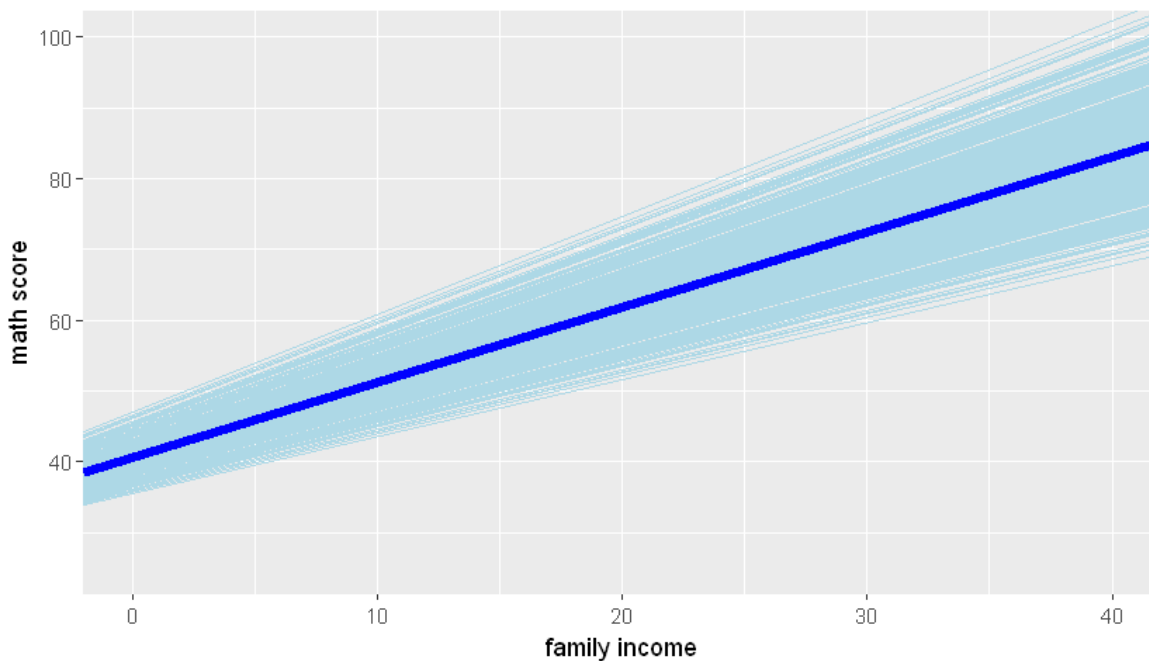
- Let's take a look at the random slopes and random intercept model where both the slope and intercept are allow to vary randomly.



In [22]:

```
options(repr.plot.height=4, repr.plot.width=7)

ggplot() +
  scale_x_continuous(name="family income", limits=c(0,40)) +
  scale_y_continuous(name="math score", limits=c(25,100)) +
  scale_linetype(name="family_inc_slope") +
  geom_abline(data=coef2,
             mapping=aes(slope=family_inc_slope,
                        intercept=intercept), color="lightblue") +
  geom_abline(slope=mean(coef2$family_inc_slope),
             intercept=mean(coef2$intercept), size=2, color="b
```



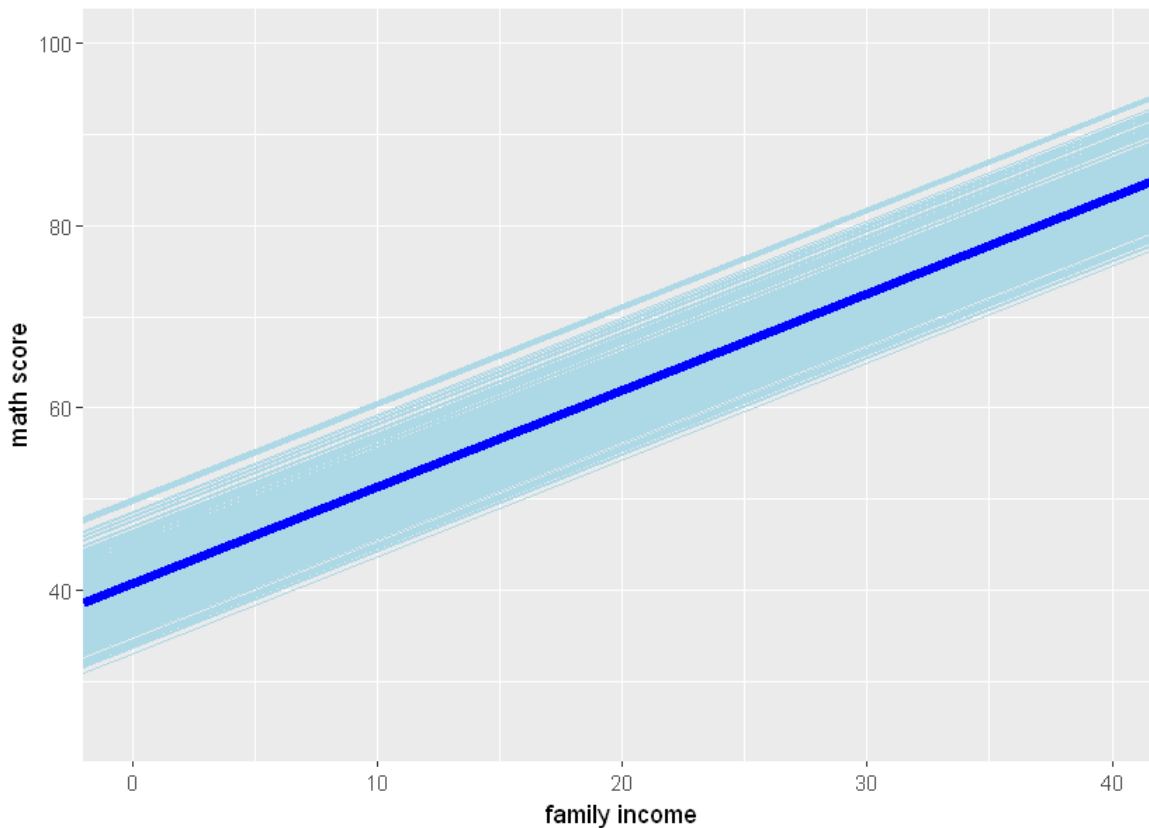
Visualize Fixed Slopes and Random Intercepts



In [23]:

```
options(repr.plot.height=5, repr.plot.width=7)

ggplot() +
  scale_x_continuous(name="family income", limits=c(0,40)) +
  scale_y_continuous(name="math score", limits=c(25,100)) +
  scale_linetype(name="family_inc_slope") +
  geom_abline(data=coef,
              mapping=aes(slope=slope_family_inc,
                          intercept=intercept), color="lightblue") +
  geom_abline(slope=mean(coef$slope_family_inc),
              intercept=mean(coef$intercept), size=2, color="bl
```



How to Obtain Means of Each Unit



In [24]:

```
1 dat %>%
2   group_by(sch_id) %>%
3   summarize(math_score_mean=mean(mathscore)) %>%
4   head()
```

sch_id	math_score_mean
81	51.75667
391	43.16000
421	44.05118
461	51.87000
615	46.70235
750	62.67000

Adding a Level 2 Predict

- The lmer() can identify which variable is a level 2 predictor because there is no within group variability for a level 2 predictor. (If you try to use a level 2 predictor as a level 1 predictor, for example if you specify a random slope for a level 2 predictor, the model will fail due to no within group variability in the level 2 predictor).
- So, the predictor at level 2 is added to the right of the formula in the lmer() function.

Predictor at Level 1 and 2: Random Intercept, Fixed Slope Model

Level-1 Model

$$\text{MATHSCOR}_{ij} = \beta_{0j} + \beta_{1j} * (\text{FAMILY_I}_{ij}) + r_{ij}$$

Level-2 Model

$$\beta_{0j} = \gamma_{00} + \gamma_{01} * (\text{URBAN}_j) + u_{0j}$$

$$\beta_{1j} = \gamma_{10}$$

Mixed Model

$$\text{MATHSCOR}_{ij} = \gamma_{00} + \gamma_{01} * \text{URBAN}_j + \gamma_{10} * \text{FAMILY_I}_{ij} + u_{0j} + r_{ij}$$



In [25]:

```
model.5 <- lmer(mathscore ~ family_inc + urban + (1|sch_id),
               data=dat, REML = FALSE)
summary(model.5)
```

Linear mixed model fit by maximum likelihood t-tests use Sat
terthwaite

approximations to degrees of freedom [lmerMod]
Formula: mathscore ~ family_inc + urban + (1 | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
59941.6	59976.7	-29965.8	59931.6	8299

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.10848	-0.74150	-0.01192	0.73562	2.92341

Random effects:

Groups	Name	Variance	Std.Dev.
sch_id	(Intercept)	15.09	3.885
	Residual	72.47	8.513

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t
)					
(Intercept)	41.11556	0.46980	4165.00000	87.517	<2e-1
6 ***					
family_inc	1.05714	0.04057	8010.00000	26.057	<2e-1
6 ***					
urban	-0.82167	0.36348	627.00000	-2.261	0.024
1 *					

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1					

Correlation of Fixed Effects:

	(Intr)	fmly_n
family_inc	-0.841	
urban	-0.424	0.055

Predictor at Level 1: Fixed Intercept, Random Slope

- We can make the intercept fixed by adding 0 to the predictor in the parenthesis.



In [26]:

```
model.6 <- lmer(mathscore ~ family_inc + (0+family_inc|sch_id),
               data=dat, REML = FALSE)
summary(model.6)
```

Linear mixed model fit by maximum likelihood t-tests use Sat
terthwaite

approximations to degrees of freedom [lmerMod]
Formula: mathscore ~ family_inc + (0 + family_inc | sch_id)
Data: dat

AIC	BIC	logLik	deviance	df.resid
59969.0	59997.1	-29980.5	59961.0	8300

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.2009	-0.7568	-0.0189	0.7329	2.8611

Random effects:

Groups	Name	Variance	Std.Dev.
sch_id	family_inc	0.1464	0.3826
Residual		73.1022	8.5500

Number of obs: 8304, groups: sch_id, 854

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t)
(Intercept)	4.014e+01	3.802e-01	8.303e+03	105.59	<2e-16

family_inc	1.095e+00	4.223e-02	5.303e+03	25.93	<2e-16

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

	(Intr)
family_inc	-0.901





In [27]:

```
coef.6 <- coef(model.6)$sch_id
names(coef.6) <- c("intercept", "slope_family_inc")
head(coef.6)
```

	intercept	slope_family_inc
81	40.14411	1.2523836
391	40.14411	0.9626563
421	40.14411	0.6038350
461	40.14411	0.9465261
615	40.14411	0.7981257
750	40.14411	1.9289007

- The coefficient data frame indicates that the intercept is fixed

Visualize the Fixed Intercept, Random Slopes Model



In [28]:

```
options(repr.plot.height=5, repr.plot.width=7)

ggplot() +
  scale_x_continuous(name="family income", limits=c(0,40)) +
  scale_y_continuous(name="math score", limits=c(25,100)) +
  scale_linetype(name="family_inc_slope") +
  geom_abline(data=coef.6,
              mapping=aes(slope=slope_family_inc,
                           intercept=intercept), color="lightblue") +
  geom_abline(slope=mean(coef.6$slope_family_inc),
              intercept=mean(coef.6$intercept), size=2, color="
```

