Partial Operator Induction with Beta Distributions

Nil Geisweiller^{1,2,3}

SingularityNET Foundation
 OpenCog Foundation
 Novamente LLC

Abstract. A specialization of Solomonoff Operator Induction considering partial operators described by second order distributions (probabilities over probabilities), is introduced. An estimate, to predict the second order distribution of new data, obtained by averaging these partial operators is derived. The problem of taking into account partial operators in that estimate is presented. This problem appears to be non-trivial. A simplistic solution with a heuristic based on estimating the Kolomogorov complexity of completions of partial operators is given.

Keywords: Solomonoff Operator Induction \cdot Beta Distribution \cdot Bayesian Averaging.

1 Introduction

Rarely natural intelligent agents attempt to construct complete models of their environment. Often time they compartmentalize their knowledge into contextual rules and make use of them without worrying about the details of the assumingly remote and irrelevant parts of the world.

This is typically how AGI Prime, aka OpenCog Prime, the AGI agent implemented over the OpenCog framework may utilize knowledge [3]. The models we are specifically targeting here are conditional probabilities, or to be more precise probability distributions over conditional probabilities, also called *second order* conditional probabilities. Maintaining second order probabilities is how OpenCog accounts for uncertainties [7] and by that properly manages integrating knowledge from heterogeneous sources, balancing exploitation and exploration and so on.

We will sometimes call these models, rules, understanding that they actually represent second order conditional probabilities. Here are some examples of rules

- 1. If the sun shines, then the temperature rises
- 2. If the sun shines and there is no wind, then the temperature rises
- 3. If the sun shines and I am in a cave, then the temperature rises

These 3 rules have different degrees of truth. The first one is often true, the second is nearly always true and the last one is rarely true. The traditional way

to quantify these degrees of truth is to assign probabilities. In practice though these probabilities are unknown, and instead one may only assign probability estimates based on limited evidence. Another possibility is to assign second order probabilities, distributions over probabilities to capture their degrees of certainty. The wider the distribution the less certain, the narrower the more certain.

Once degrees of truth are properly represented, an agent should be able to utilize these rules to predict and operate in its environment. This raises a question. How to choose between rules? Someone wanting to predict whether the temperature will rise will have to make a choice. If one is in a cave, should he/she follow the third rule? Why not the first one which is still valid, or assuming there is no wind, maybe the second?

Systematically picking the rule with the narrowest context (like being in a cave) is not always right. Indeed, the narrower the context the less evidence we have, the broader the uncertainty, the more prone to overfitting it might be.

In this paper we attempt to address this issue by adapting Solomonoff Operator Induction [8] for a special class of operators representing such rules. These operators have two particularities. First, their outcomes are second order probabilities, specifically Beta distributions. Second, they are partial, that is they are only defined over a subset of observations, the available observations encompassed by their associated contexts. An estimate of the second order probability predicting new data, obtained by averaging the second order probabilities of these partial operators, is derived. Then the problem of dealing with partial operators is presented and a simplistic heuristic consisting of estimating the Kolmogorov complexity of a perfect completion (to turn a partial program into a complete one that perfectly explains the remaining data) is offered.

In Section 2 we briefly recall what is Solomonoff Operator Induction, in Section 3 what are Beta distributions. In Section 4 we introduce our specialization of Solomonoff Operator Induction for partial operators with Beta distributions. Finally, in Section 5 we conclude and present some directions for research.

2 Solomonoff Operator Induction

Solomonoff Universal Operator Induction [8] is a general, parameter free induction method that has been shown to theoretically converge to any true computable distribution. It is a special case of Bayesian Model Averaging [5] though is universal in the sense that the models across which the averaging is taking place are Turing-complete.

Let us recall its formulation, using the same notations as in the original paper of Solomonoff (Section 3.2 of [8]). Given a sequence of n questions and answers $(Q_i, A_i)_{i \in [1,n]}$, and a countable family of operators O^j (the superscript j denotes the j^{th} operator, not the exponentiation) computing partial functions mapping pairs of question and answer to probabilities, then one may estimate

the probability of the next answer A_{n+1} given question Q_{n+1} as follows

$$\hat{P}(A_{n+1}|Q_{n+1}) = \sum_{j} a_0^j \prod_{i=1}^{n+1} O^j(A_i|Q_i)$$
(1)

where a_0^j is the prior of the j^{th} operator, its probability after zero observation, and is generally approximated by $2^{-K(O^j)}$ where K is the Kolmogorov complexity [10]. Using Hutter's convergence theorem to arbitrary alphabets [6] it can be shown that such estimate rapidly converges to the true probability.

Let us rewrite Eq. 1 by making the prediction term and the likelihood explicit

$$\hat{P}(A_{n+1}|Q_{n+1}) = \sum_{j} a_0^j l^j O^j(A_{n+1}|Q_{n+1})$$
(2)

where $l^j = \prod_{i=1}^n O^j(A_i|Q_i)$ is the likelihood, the probability of the data given the j^{th} operator.

Remark 1. In the remaining of the paper the superscript j is always used to denote the index of the j^{th} operator. Sometimes, though in a consistent manner, it is used as subscript. All other superscript notations not using j denote exponentiation.

3 Beta Distribution

Beta distributions [1] are convenient to model probability distributions over probabilities, i.e. second order probabilities. In particular, given a prior over a probability p of some event, like a coin toss to head, defined by a Beta distribution, and a sequence of experiments, like coin tosses, the posterior of p is still a Beta distribution. For that reason the Beta distribution is called a conjugate prior for the binomial distribution.

Let us recall the probability density and cumulative distribution functions of the Beta distribution as it will be useful later on.

3.1 Prior and Posterior Probability Density Function

The probability density function (pdf) of the Beta distribution with parameters α and β , is

$$f(x; \alpha, \beta) = \frac{x^{\alpha - 1} (1 - x)^{\beta - 1}}{B(\alpha, \beta)}$$
(3)

where x is a probability and $B(\alpha, \beta)$ is the beta function

$$B(\alpha, \beta) = \int_0^1 p^{\alpha - 1} (1 - p)^{\beta - 1} dp$$
 (4)

One may notice that multiplying the density by the likelihood

$$x^m(1-x)^{n-m} (5)$$

of a particular sequence of n experiments with m positive outcomes with probability x, is also a Beta distribution

$$f(x; m+\alpha, n-m+\beta) \propto x^{m+\alpha-1} (1-x)^{n-m+\beta-1}$$
(6)

3.2 Cumulative Distribution Function

The cumulative distribution function (cdf) of the Beta distribution is

$$I_x(\alpha, \beta) = \frac{B(x; \alpha, \beta)}{B(\alpha, \beta)} \tag{7}$$

where $B(x; \alpha, \beta)$ is the incomplete beta function

$$B(x; \alpha, \beta) = \int_0^x p^{\alpha - 1} (1 - p)^{\beta - 1} dp$$
 (8)

 I_x is also called the regularized incomplete beta function [12].

4 Partial Operator Induction with Beta Distributions

In this section we introduce our specialization of Solomonoff Operator Induction for partial operators describing second order distributions, and more specifically Beta distributions. An estimate of the second order conditional probability of the next data is derived, Eq. 25. Such equation contains unknown terms, the likelihoods of the unaccounted data by the partial operators, them estimated by a simplistic heuristic.

4.1 Second Order Probability Estimate

Let us first modify the Solomonoff Operator Induction probability estimate to represent a second order probability. This is crucial to maintain, and ultimately propagate to efferent cognitive processes, the uncertainty of that estimate. It directly follows from Eq. 2 of Section 2, that the cumulative distribution function of the probability estimate of observing answer A_{n+1} given question Q_{n+1} is

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) = \sum_{O^j(A_{n+1}|Q_{n+1}) \le x} a_0^j l^j$$
(9)

Due to O^j not being complete in general $cdf(A_{n+1}|Q_{n+1})(1)$ may not be equal to 1. It means that some normalization will need to take place, that is even more true since only a very small fraction of the operator space is explored in practice. We need not to worry about the continuity or the differentiability of $cdf(A_{n+1}|Q_{n+1})$. What matters is that a spread of probabilities is represented to account for the uncertainty. It is expected that the breadth would be wide at first, and progressively shrinks, fluctuating depending on the novelty of the data, as measure as more questions and answers get collected.

4.2 Continuous Parameterized Operators

Let us now extend the definition of this estimate for parameterized operators, so that each operator is a second order distribution. Let us consider a subclass of parameterized operators such that, if p is the parameter of operator O_p^j , the result of the conditional probability of A_{n+1} given Q_{n+1} is

$$O_p^j(A_{n+1}|Q_{n+1}) = p (10)$$

Doing so will enable us to consider operators as Beta distribution later on, in Section 4.3. Given that assumption, the cumulative distribution function of the estimate $c\hat{d}f(A_{n+1}|Q_{n+1})$ becomes

$$\hat{cdf}(A_{n+1}|Q_{n+1})(x) = \sum_{j} a_0^j \int_0^x f_p l_p^j dp$$
 (11)

where f_p is the prior density of p, and $l_p^j = \prod_{i=1}^n O_p^j(A_i|Q_i)$ is the likelihood of the data according to the j^{th} operator with parameter p.

Proof. Consider continuous families of parameterized operators combined with Eq. 10. Let us start by expressing Eq.9 with a discretization of O_p^j with prior $a_0^j f_p \Delta p$

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) = \sum_{O_p^j(A_{n+1}|Q_{n+1}) \le x} a_0^j f_p l_p^j \Delta p$$
(12)

where the sum runs over all j and p by steps of Δp such that $O_p^j(A_{n+1}|Q_{n+1}) \leq x$. Since a_0^j does not depends on p, it can be moved in its own sum

$$\hat{cdf}(A_{n+1}|Q_{n+1})(x) = \sum_{j} a_0^j \sum_{O_p^j(A_{n+1}|Q_{n+1}) \le x} f_p l_p^j \Delta p$$
(13)

now the second sum only runs over p. Due to Eq. 10 this can be simplified into

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) = \sum_{j} a_0^j \sum_{p \le x} f_p l_p^j \Delta p$$
 (14)

which is turns into Eq. 11 when Δp tends to 0.

Using continuous integration may seem like a departure from Solomonoff Induction. First, it does not correspond to a countable class of models. And second, the Kolmogorov complexity of p, that would in principle determine its prior, is likely chaotic and very different than how priors are typically defined over continuous parameters in Bayesian inference. In practice however integration is discretized and values are truncated up to some fixed precision. Moreover any prior can probably be made compatible with Solomonoff induction by selecting an adequate Turing machine of reference.

4.3 Operators as Beta Distributions

We have now what we need to model our rules, second order conditional probabilities, as operators.

First, we need to assume that operators are partial, that is the j^{th} operator is only defined for a subset of n^j questions, those that meet the conditions of the rule. For instance, considering the rule

- If the sun shines, then the temperature rises

questions and answers pertaining to what happens at night will be ignored.

Second, we assume that answers are Boolean, so that $A_i \in \{0,1\}$ for $i \in [1, n+1]$. In reality, OpenCog rules manipulate predicates (generally fuzzy predicates but that can be let aside), and the questions they represent are: if some instance holds property R, what are the odds that it holds property S? We simplify this by fixing predicate S so that the problem is reduced to finding R that best predicts it. Thus we assume that if $A_i = A_{n+1}$ then O_p^j models the odds of $S(Q_i)$, and if $A_i \neq A_{n+1}$, it models the odds of $\neg S(Q_i)$. The class of operators under consideration can be represented as programs of the form

$$O_p^j(A_i|Q_i) = \text{if } R^j(Q_i) \text{ then } \begin{cases} p, & \text{if } A_i = A_{n+1} \\ 1 - p, & \text{otherwise} \end{cases}$$
 (15)

where R^{j} is the condition of the rule. This allows an operator to be modeled as a Beta distribution, with cumulative distribution function

$$cdf_{O^j}(x) = I_x(m^j + \alpha, n^j - m^j + \beta)$$
(16)

where m^j is the number of times $A_i = A_{n+1}$ for the subset of n^j questions such that $R^j(Q_i)$ is true. The parameters α and β are the parameters of the prior of p, itself a Beta distribution. Eq. 16 is in fact the definition of OpenCog Truth Values as described in Chapter 4 of the PLN book [4].

4.4 Handling Partial Operators

When attempting to use such operators we still need to account for their partiality. Although Solomonoff Operator Induction does in principle encompass partial operators⁴, it does so insufficiently, in our case anyway. Indeed, if a given operator cannot compute the conditional probability of some question/answer pair, the contribution of that operator may simply be ignored in the estimate. This does not work for us since partial operators (rules over restricted contexts) might carry significant predictive power and should not go to waste.

To the best of our knowledge, the existing literature does not cover that problem. The Bayesian inference literature contains in-depth treatments about

⁴ more by necessity, since the set of partial operators is enumerable, while the set of complete ones is not.

how to properly consider missing data [11]. Unfortunately, they do not directly apply to our case because our assumptions are different. In particular, here, data omission depends on the model. However, the general principle of modeling missing data and taking into account these models in the inference process, can be applied. Let us attempt to do that by explicitly representing the portion of the likelihood over the missing data according to the j^{th} operator by a dedicated term, denoted r^j . In the rest of the paper rather than calling these data missing we prefer to refer to them as unexplained or unaccounted, which better suits our problem. Let us also define a completion of O_p^j as a program that can explain the unaccounted data.

Definition 1. A completion C of O_p^j is a program that completes O_p^j for the unaccounted data (when $R^j(Q_i)$ is false)

$$O_{p,C}^{j}(A_{i}|Q_{i}) = if R^{j}(Q_{i}) \ then \ \begin{cases} p, & if A_{i} = A_{n+1} \\ 1 - p, & otherwise \end{cases}$$
 else $C(A_{i}|Q_{i})$

The likelihood given such operator once completed would be

$$l_n^j = p^{m^j} (1 - p)^{n^j - m^j} r^j (17)$$

where the binomial term account for the likelihood of the explained data, and r^{j} is a term that accounts for the likelihood of the unexplained data, more specifically

$$r^{j} = \prod_{i \le n \ \land \ \neg R^{j}(Q_{i})} C^{j}(A_{i}|Q_{i}) \tag{18}$$

where C^j is the underlying completion of O_p^j explaining the unaccounted data. One may notice that r^j does not depend on p. Such assumption tremendously simplifies the analysis and is somewhat reasonable to make. Generally we may assume that the completion of the model is independent on its pre-existing part. By replacing the likelihood in Eq. 11 by Eq. 17 we obtain

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) = \sum_{j} a_0^j \int_0^x f_p p^{m^j} (1-p)^{n^j - m^j} r^j dp$$
 (19)

Choosing a Beta distribution as the prior of f_p simplifies the equation as the posterior remains a Beta distribution

$$f_p = f(p; \alpha, \beta) \tag{20}$$

where f is the pdf of the Beta distribution as define in Eq. 3. Usual priors are Bayes' with $\alpha=1$ and $\beta=1$, Haldane's with $\alpha=0$ and $\beta=0$ and Jeffreys' with $\alpha=\frac{1}{2}$ and $\beta=\frac{1}{2}$. The latter is probably the most accepted due to being uninformative in some sense [9]. We do not need to commit to a particular one at that point and let the parameters α and β free, giving us

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) = \sum_{j} a_0^j \int_0^x \frac{p^{\alpha-1}(1-p)^{\beta-1}}{B(\alpha,\beta)} p^{m^j} (1-p)^{n^j-m^j} r^j dp$$
 (21)

 r^{j} can be moved out of the integral and the constant $B(\alpha, \beta)$ can be ignored on the ground that our estimate will require normalization anyway

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) \propto \sum_{j} a_0^j r^j \int_0^x p^{m^j + \alpha - 1} (1-p)^{n^j - m^j + \beta - 1} dp$$
 (22)

 $\int_0^x p^{m^j+\alpha-1}(1-p)^{n^j-m^j+\beta-1}dp \text{ is the incomplete Beta function with parameters } m^j+\alpha \text{ and } n^j-m^j+\beta, \text{ thus}$

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) \propto \sum_{j} a_0^j r^j B(x; m^j + \alpha, n^j - m^j + \beta)$$
 (23)

Using the regularized incomplete beta function we obtain

$$\hat{cdf}(A_{n+1}|Q_{n+1})(x) \propto \sum_{j} a_0^j r^j I_x(m^j + \alpha, n^j - m^j + \beta) B(m^j + \alpha, n^j - m^j + \beta)$$
 (24)

As I_x is the cumulative distribution function of O^j (Eq. 16), we finally get

$$c\hat{d}f(A_{n+1}|Q_{n+1})(x) \propto \sum_{j} a_0^j r^j c df_{O^j}(x) B(m^j + \alpha, n^j - m^j + \beta)$$
 (25)

We have expressed our cumulative distribution function estimate as an averaging of the cumulative distribution functions of the operators. This gives us an estimate to predict whether S holds for a new question and how much certainty we have about that prediction.

However, we still need to address r^j , the likelihood of the unaccounted data. In theory, the right way to model r^j would be to consider all possible completions of the j^{th} operator, but that is intractable. One would be tempted to simply ignore r^j , however, as we have already observed in some preliminary experiments, this gives an unfair advantage to rules that have a lot of unexplained data, and thus make them more prone to overfitting. This is true even in spite of the fact that such rules naturally exhibit more uncertainty due to carrying less evidence.

4.5 Perfectly Explaining Unaccounted Data

Instead we attempt to consider the most prominent completions. For now we consider completions that perfectly explain the unaccounted data. Moreover, to simplify further, we assume that unaccounted answers are entirely determined by their corresponding questions. This is generally not true, the same question may relate to different answers. But under such assumptions r^j becomes 1. This may seem equivalent to ignoring r^j unless the complexity of the completion is taken into account. What that means is that we must consider, not only the complexity of the rule but also the complexity of its completion. Unfortunately

calculating that complexity is intractable. To work around that we estimate it as function of the length of the unexplained data. Specifically, we suggest as prior

$$a_0^j = 2^{-K(O^j) - v_j^{(1-c)}} (26)$$

where $K(O^j)$ is the Kolmogorov complexity of the j^{th} operator (estimated by the length of its rule), v_j is the length of its unaccounted data, and c is a compressability parameter. If c=0 then the unaccounted data are incompressible. If c=1 then the unaccounted data can be compressed to a single bit. It is a very crude heuristic and is not parameter free, but it is simple and computationally lightweight. When applied to experiments, not described here due to their early stage nature and space limitation, a value of c=0.5 was actually shown to be somewhat satisfactory.

5 Conclusion

We have introduced a specialization of Solomonoff Operator Induction over operators with the particularities of being partial and modeled by Beta distributions. A second order probability estimate to predict new data and represent the uncertainty of the prediction has been derived. While doing so we have uncovered an interesting problem, how to account for partial operators in the estimate. This problem appears to have no obvious solution, is manifestly under-addressed by the research community, and yet important in practice. Although the solution we provide is very lacking (crudely estimating the Kolmogorov complexity of a perfect completion) we hope that it provides some initial ground for experimentation and motivates further research. Even though, ultimately, it is expected that this problem might be hard enough to require some form of meta-learning [2], improvements in the heuristic by, for instance, considering completions reusing available models that do explain some unaccounted data could help.

Experiments using this estimate are currently being carried out in the context of inference control meta-learning within the OpenCog framework and will make the object of future publications.

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