Nicholas Gluze n/cq2

$$C) \int_{\Lambda} (\alpha) = \log \left(\prod_{i=1}^{n} \frac{\alpha x_{0}^{\alpha}}{x_{i}^{\alpha+1}} \right) = \sum_{i=1}^{n} \log \left(\frac{\alpha x_{0}^{\alpha}}{x_{i}^{\alpha+1}} \right) = \sum_{i=1}^{n} \left(\log \left(\alpha x_{0}^{\alpha} \right) - \log \left(x_{0}^{\alpha} \right) - \log \left(x_{0}^{\alpha} \right) \right)$$

$$\frac{dl_{i}(\alpha)}{d\alpha} = \sum_{i=1}^{n} \left(\frac{1}{\alpha} + \log x_{o} - \log x_{i} \right)$$

$$0 = \frac{1}{2} + n \log x_0 - \sum_{i=1}^{n} \log x_i$$

$$\frac{1}{2} = \sum_{i=1}^{n} \log x_i - n \log x_0 = \sum_{i=1}^{n} \log x_i - n \log x_0$$

•# triangles per node, clockwise: choose 2 of the node's clockwise abos within

ovoid double =
$$\begin{pmatrix} k \\ 2 \end{pmatrix} = \frac{k!}{2!(k-2)!} = \frac{k(k-1)}{2}$$

.# corrected triplets per node: (2k) = 2k(2k-1)

$$= \frac{3nk(k-1)}{2nk(3k-1)} = \frac{3(k-1)}{2(2k-1)} = \frac{3(k-3)}{4k-2}$$

b) + + riangles; -still nk(k-1) original -new: · can be much by correcting nodes between k+1 and 2k hops away; probthat they're corrected is = 1 n 2k . P. /n(n-1) = 2kp # triples; -still 2nk(2k-1) o(igina) - each new edge creates triples w/all of the original > however, as n > 00 this aknodes @ either of itsends (=(2k)2 triples) term ones to its so it. n possible, ~/p.cob.p. 1. egl: g. ible. =># Δ= 12 $=7(2k)^2\Lambda\rho$ - new edges also coente poiss who the new edges:
- l'new edges -> (2)= l(e-1) new triples. -expected # of new edges to a node is 21EIP = 2 (2016) => exp. new triples is nakp (akp-1) = (ak)2p2-3 uk (. k -1) =>c(G) = $\frac{\frac{1}{2}nk(k-1)\cdot 3}{\frac{1}{2}nk(ak-1)+(ak)^{2}np+\frac{1}{2}n(ak^{2})p^{2}}$ $\frac{\frac{3}{2}nk(k-1)}{nk(ak-1)+(ak)^{2}np}$ 20 k2-nk+4k2np 2(2/k2- K+4/162p) 3k-3 2(2K-1+4KP) 1= · 3 k-3 4K-2+8kp

U. a) k:(i)=M (each node connects to mothers \emptyset birth; b) $dk:(t) \sim M$ (a catio of M nodes have new connections added $dt \sim T$ from time $t \rightarrow t+1$) $k:(t)=\int_{T}^{\infty}dt=\frac{m\log t}{t} + C \rightarrow k:(i)=m=\frac{m\log i}{t} + C$ $=>C=\frac{m-m\log i}{t}$ () $k:(t)=m+n\log(\frac{t}{t}) \geq d$ $=>k:(t)=m\log t+m-m\log i$ $=>k:(t)=m\log t+m-m\log t+m-m\log t$ $==m\log t+m-m\log t+m-$

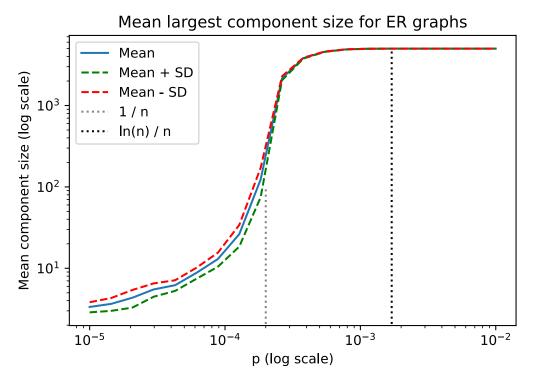
```
import networkx as nx
import numpy as np
import matplotlib.pyplot as plt
from collections import Counter, defaultdict
```

1. ER phase transitions

a) ER graph component sizes

```
In [ ]:
         def analyze er graph components(n, p range, graphs per p=20):
             means, sds = [], []
             for p in p range:
                 print('Analyzing ER graphs for p={}'.format(p))
                 comp sizes = []
                 for i in range(graphs per p):
                     G = nx.generators.fast gnp random graph(n, p, seed=i)
                     comp sizes.append(max([len(comp) for comp in nx.connected components(G)]))
                 means.append(np.mean(comp sizes))
                 sds.append(np.std(comp sizes))
             return np.array(means), np.array(sds)
In [ ]:
         n = 5000
         p range = np.logspace(-5, -2, num=20)
         means, sds = analyze er graph components(n, p range)
        Analyzing ER graphs for p=1e-05
        Analyzing ER graphs for p=1.438449888287663e-05
        Analyzing ER graphs for p=2.06913808111479e-05
        Analyzing ER graphs for p=2.9763514416313192e-05
        Analyzing ER graphs for p=4.281332398719396e-05
        Analyzing ER graphs for p=6.158482110660267e-05
        Analyzing ER graphs for p=8.858667904100833e-05
        Analyzing ER graphs for p=0.00012742749857031334
        Analyzing ER graphs for p=0.00018329807108324357
        Analyzing ER graphs for p=0.00026366508987303583
        Analyzing ER graphs for p=0.000379269019073225
        Analyzing ER graphs for p=0.0005455594781168515
```

```
Analyzing ER graphs for p=0.0007847599703514606
        Analyzing ER graphs for p=0.0011288378916846883
        Analyzing ER graphs for p=0.001623776739188721
        Analyzing ER graphs for p=0.002335721469090121
        Analyzing ER graphs for p=0.003359818286283781
        Analyzing ER graphs for p=0.004832930238571752
        Analyzing ER graphs for p=0.0069519279617756054
        Analyzing ER graphs for p=0.01
In [ ]:
         def plot with sds(x, y, y sds, n=5 000):
             plt.plot(x, y)
             plt.plot(x, y - y sds, c='green', linestyle='dashed')
             plt.plot(x, y + y sds, c='red', linestyle='dashed')
             # plot vertical lines at 1 / n and ln(n) / n
             plt.vlines(1 / n, 0, 100, linestyle='dotted', colors='gray')
             plt.vlines(np.log(n) / n, 0, 4900, linestyle='dotted', colors='black')
             plt.yscale('log')
             plt.xscale('log')
             plt.title('Mean largest component size for ER graphs')
             plt.xlabel('p (log scale)')
             plt.ylabel('Mean component size (log scale)')
             plt.legend(['Mean', 'Mean + SD', 'Mean - SD', '1 / n', 'ln(n) / n'])
In [ ]:
         plot with sds(p range, means, sds)
```



b) Interesting values of p

From class, we have that when p = lambda / n, lambda > 1 makes the largest component size converge to n, while lambda < 1 makes it converge to ln(n). Here, lambda = 1 represents a point at which the largest component size is transitioning rapidly, as it has the graph's peak slope. Also, l notice that as lambda shrinks less than 1, the size remains low (near $ln(n) \sim = 8$), while as lambda grows above 1, the size approaches n.

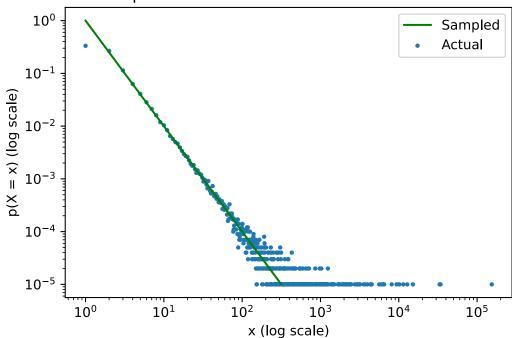
Also, we have that when p = lambda * ln(n) / n, lambda < 1 makes the graph's probability of being connected converge to 0, while lambda > 1 makes it converge to 1. Here, to the left of ln(n) / n (i.e. lambda > 1), the mean component size has not yet converged to n, but to the right, the mean is equal to n and the standard deviation is 0, meaning that, as expected, all sampled graphs converge to being connected.

2. Fitting power-law exponents

b) Pareto sampling

```
def sample pareto(alpha=1, x 0=1, size=100 000):
             \# sample uniformly in [0, 1] and apply the pareto function to each sample
             pareto_func = np.vectorize(lambda u: x_0 / (u^{**}(1 / alpha)))
             u samples = np.random.sample(size=size)
             return pareto func(u samples)
In [ ]:
         def compute pdf(samples):
             # round samples to nearest integer
             sample ints = np.rint(samples).astype(int)
             # count occurrences of each integer
             cts = Counter(sample ints)
             x = np.array(sorted(cts.keys()))
             px_x = np.array([cts[i] for i in sorted(cts.keys())]) / len(sample_ints)
             return x, px x
In [ ]:
         def compute ccdf(x, px x):
             # compute running total from right to left
             s = 0
             fbarx x = np.zeros(x.shape)
             for i in range(len(x) - 1, -1, -1):
                 s += px x[i]
                 fbarx x[i] = s
             return x, fbarx x
In [ ]:
         def plot observations(x, y, alpha=1):
             plt.scatter(x, y, s=5)
             plt.xscale('log')
             plt.yscale('log')
             y = x + 1 / x ** (alpha + 1)
             plt.plot(x[y_expected >= 1e-5], y_expected[y_expected >= 1e-5], c='green')
             plt.title('Sampled and actual PDFs for Pareto distribution')
             plt.legend(['Sampled', 'Actual'])
             plt.xlabel('x (log scale)')
             plt.ylabel('p(X = x) (log scale)')
In [ ]:
         plot observations(x, px x)
```





d) Estimating alpha

```
alpha_pred = -np.polyfit(x_log, y_log, deg=1)[0] + 1
             return alpha pred
In [ ]:
         def estimate alpha mle(samples, x 0=1):
             n = len(samples)
             # from c), MLE is alpha = n / (sum i(log(x i)) - (n * log(x 0)))
             return n / (np.log(samples).sum() - (n * np.log(x 0)))
In [ ]:
         def eval alpha estimate(estimate alpha, name, samples=100):
             preds = []
             for in range(samples):
                 samples = sample_pareto()
                 preds.append(estimate alpha(samples))
             m, sd = np.mean(preds), np.std(preds)
             print('{}:'.format(name))
             print('Mean: {}, SD: {}'.format(m, sd))
In [ ]:
         eval alpha estimate(estimate alpha pdf ls, 'Least-squares on PDF');
        Least-squares on PDF:
        Mean: 0.9587292458008777, SD: 0.048126100747554156
In [ ]:
         eval alpha estimate(estimate alpha ccdf ls, 'Least-squares on CCDF');
        Least-squares on CCDF:
        Mean: 1.9955482093481538, SD: 0.03403874599398693
In [ ]:
         eval alpha estimate(estimate alpha mle, 'Log-likelihood maximization');
```

Least-squares regression on the CCDF gives a very poor result, while least-squares on the PDF and the log-likelihood MLE both give results somewhat close to the true value. The MLE almost exactly approximates true value, while the PDF least-squares is off by a bit more; also, the MLE has near-0 standard deviation, so since it is also the least biased, it is clearly the best estimate.

note: my numbers don't line up with Dr. Segarra's for i and ii, but he looked through my code during office hours and said that what I'm doing seems right, so we're not sure why my results are different. I plotted the PDF, CCDF, and fitted lines on a log-log graph below, if that helps at all.

Out[]: <matplotlib.collections.PathCollection at 0x7f18a828ee50>

