QUY NHON UNIVERSITY DEPARTMENT OF MATHEMATICS AND STATISTICS

When Models Meet Data

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Roadmap

- (1) Data, Models, and Learning
- (2) Models as Functions: Empirical Risk Minimization
- (3) Models as Probabilistic Models: Parameter Estimation (ML and MAP)
- (4) Probabilistic Modeling and Inference
- (5) Directed Graphical Models
- (6) Model Selection



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Data, Models, and Learning

- Three major components of a machine learning system
 - **1** Data: $\{(x_1, y_1), \dots, (x_n, y_n), \dots, (x_N, y_N)\}$
 - Models: deterministic functions or probabilistic models
 - 3 Learning: Training, and prediction/inference
- Good machine learning models: Perform well for unseen (untrained) data
- Machine learning algorithm: training and prediction



Data as Vectors

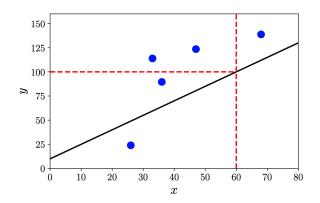
- Tabular format or not, numerical or not, good feature extraction etc.
- Assume that data is given as D-dimensional vector \mathbf{x}_n of real numbers, each called features, attributes, or covariates.
- Dataset: consisting of data points or examples $\{x_1, x_2, \dots, x_N\}$
- In supervised learning, $\{(x_1, y_1), \dots, (x_n, y_n), \dots, (x_N, y_N)\}$, where y_n is the label (or target, response variable, or annotation).
- Better representation of data as vectors
 - finding lower-dimensional approximations of the original feature vector (e.g., PCA via SVD or EVD)
 - using nonlinear higher-dimensional combinations of the original feature vector (e.g., feature map and kernel)

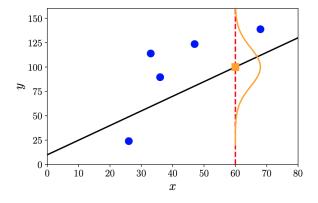
Models: Functions vs. Probabilistic Models

- Now, the business of constructing a predictor
- Models as functions
 - $f: \mathbb{R}^D \mapsto \mathbb{R}$.
 - Example. $f(\mathbf{x}) = \boldsymbol{\theta}^{\mathsf{T}} \mathbf{x} + \theta_0$, Unknown parameter: $\boldsymbol{\theta}, \theta_0$
- Models as probabilistic models
 - model our uncertainty due to the observation process and our uncertainty in the parameters of our model
 - predictors should be able to express some sort of uncertainty via probabilistic models

Mathematics for Machine Learning

 Parameters: parameters of a chosen probabilistic model (e.g., mean and variance of Gaussian)







Learning Algorithms

Three algorithmic phases

- (1) Prediction or inference: via function or probabilitic models
- (2) Training or parameters estimation
 - fixed parameter assumption (non-probabilistic) or Bayesisan approach (probabilistic)
 - non-probabilistic: e.g., empirical risk minimization
 - probabilistic: e.g., ML (Maximum Likelihood), MAP (Maximum A Posteriori)
 - cross-validation: simulation of performing for unseen data
 - regularization/prior: balancing models between training and unseen data
- (3) Hyperparameter tuning or model selection



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Empirical Risk Minimization

- Predictor as a function
- Given $\{(x_1, y_1), \dots, (x_n, y_n), \dots, (x_N, y_N)\}$, estimate a predictor $f(\cdot, \boldsymbol{\theta}) : \mathbb{R}^D \mapsto \mathbb{R}$
- Find a good parameter θ^* , such that $f(\mathbf{x}_n, \theta^*) = \hat{y}_n \approx y_n$, for all $n = 1, \dots, N$
- Example. Affine function: By adding the unit feature $x^{(0)} = 1$ and θ_0 , i.e., $\mathbf{x}_n = [1, x_n^{(1)}, \dots, x_n^{(D)}]^\mathsf{T}$, $\mathbf{\theta} = [\theta_0, \theta_1, \dots, \theta_D]^\mathsf{T}$

$$f(\mathbf{x}_n, \boldsymbol{\theta}) = \boldsymbol{\theta}^\mathsf{T} \mathbf{x}_n = \theta_0 + \sum_{d=1}^D \theta_d x_n^{(d)}$$

Example. Neural network: Complex non-linear function



Loss Function

- Training set: $\{(\boldsymbol{x}_1, y_1), \dots, (\boldsymbol{x}_N, y_n), \dots, (\boldsymbol{x}_N, y_N)\}$, an example matrix¹ $\boldsymbol{X} := [\boldsymbol{x}_1, \dots, \boldsymbol{x}_N]^\mathsf{T} \in \mathbb{R}^{N \times D}$, a label vector $\boldsymbol{y} := [y_1, \dots, y_N]^\mathsf{T}$,
- Average loss, empirical risk

$$\mathbf{R}_{\mathsf{emp}}(f, \mathbf{X}, \mathbf{y}) = \frac{1}{N} \sum_{n=1}^{N} \ell(y_n, \hat{y}_n)$$

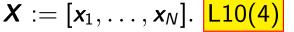
- Goal: Minimizing empirical risk
- Example. The squared loss function $\ell(y_n, \hat{y}_n) = (y_n \hat{y}_n)^2$ leads to:

$$\min_{oldsymbol{ heta} \in \mathbb{R}^D} rac{1}{oldsymbol{N}} \|oldsymbol{y} - oldsymbol{X} oldsymbol{ heta}\|^2$$

Question. Ultimgate goal: Minimizing expected risk (for unseen data)

$$R_{\mathsf{true}} = \mathbb{E}_{\boldsymbol{x},y}[\ell(y,f(\boldsymbol{x}))]?$$

¹In other chapters, we often use $D \times N$ example matrix by defining it as





Overfitting and Regularization

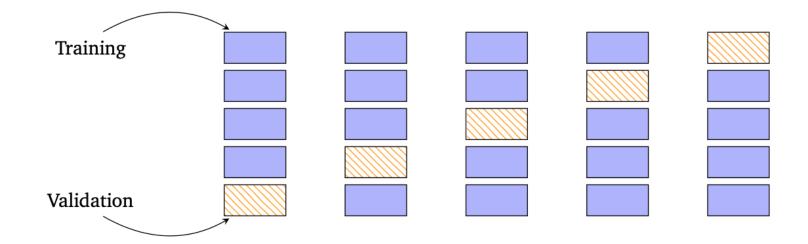
- The predictor fits too closely to the training data and does not generalize well to new data
- Need to somehow bias the search for the minimizer of empirical risk by introducing a penalty term
- Regularization: compromise between accurate solution of empirical risk minimization and the size or complexity of the solution.
- Example. Regularized Least Squares

$$\min_{\boldsymbol{\theta} \in \mathbb{R}^D} \frac{1}{N} \| \boldsymbol{y} - \boldsymbol{X}\boldsymbol{\theta} \|^2 + \lambda \| \boldsymbol{\theta} \|^2$$

• $\|\theta\|^2$: regularizer, λ : regularization parameter



Cross-Validation for Generalization Performance





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MLE (Maximum Likelihood Estimation): Concept

- Idea: define a function of the parameters called likelihood function.
- Negative log-likelihood for data x and a family of probability densities $\mathbb{P}(x \mid \theta)$ parameterized by θ :

$$\mathcal{L}_{\mathbf{x}}(\boldsymbol{\theta}) = \mathcal{L}(\boldsymbol{\theta}) := -\log \mathbb{P}(\mathbf{x} \mid \boldsymbol{\theta})$$

- $\mathcal{L}(\theta)$: how likely a particular setting of θ is for the observations x.
- MLE: Find θ such that $\mathcal{L}(\theta)$ is minimized (i.e., likelihood is maximized)



MLE: Supervised Learning

- The set of iid examples $(x_1, y_1), \ldots, (x_N, y_N)$
- $\mathcal{X} = \{ \mathbf{x}_1, \dots, \mathbf{x}_N \}$ and $\mathcal{Y} = \{ y_1, \dots, y_N \}$
- Negative log-likelihood

$$\mathcal{L}(oldsymbol{ heta}) = -\log \mathbb{P}(\mathcal{Y} \mid \mathcal{X}, oldsymbol{ heta}) = \sum_{n=1}^N \log \mathbb{P}(y_n \mid oldsymbol{x}_n, oldsymbol{ heta})$$

• Example. Assume independent Gaussian noise $\mathcal{N}(0, \sigma^2)$ and linear model $y_n = \mathbf{x}_n^\mathsf{T} \boldsymbol{\theta}$ for prediction. Then, $Y_n | (\mathbf{x}_n, \boldsymbol{\theta}) \sim \mathcal{N}(\mathbf{x}_n^\mathsf{T} \boldsymbol{\theta}, \sigma^2)$.

$$\mathcal{L}(\boldsymbol{\theta}) = -\sum_{n=1}^{N} \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right) = \frac{1}{2\sigma^2} \sum_{n=1}^{N} (y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2 - \sum_{n=1}^{N} \log \frac{1}{2\sigma^2} \exp\left(-\frac{(y_n - \boldsymbol{x}_n^\mathsf{T}\boldsymbol{\theta})^2}{2\sigma^2}\right)$$



MAP (Maximum A Posteriori)

- What if we have some prior knowledge about θ ? Then, how should we change our knowledge about θ after observing data x?
- Compute a posteriori distribution (using Bayes' Theorem) and find θ that maximizes the distribution:

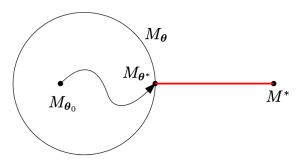
$$\max_{\boldsymbol{\theta}} \mathbb{P}(\boldsymbol{\theta} \mid \boldsymbol{x}) = \max_{\boldsymbol{\theta}} \frac{\mathbb{P}(\boldsymbol{x} \mid \boldsymbol{\theta}) \mathbb{P}(\boldsymbol{\theta})}{\mathbb{P}(\boldsymbol{x})} \Longleftrightarrow \min_{\boldsymbol{\theta}} \Big(-\log \mathbb{P}(\boldsymbol{\theta} \mid \boldsymbol{x}) \Big)$$

- In finding the optimal θ , $\mathbb{P}(x)$ can be ignored
- ML and MAP: Bridging the non-probabilistic and probabilistic worlds as it explicitly acknowledges the need for a prior distribution, yet producing a point estimate (one single parameter return).

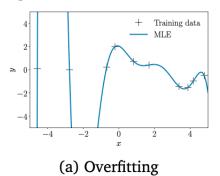


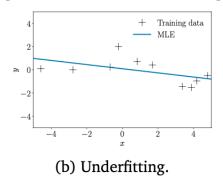
Model Fitting

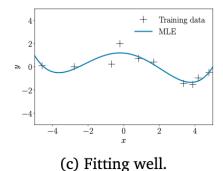
• Model class M_{θ} vs. Right model M^*



Overfitting vs. Underfitting vs. Good fitting









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Modeling Generative Process and Probabilistic Models

- Many machine learning tasks: prediction of future events and decision making
- Often build (probabilistic) models that describe the generative process that generates the observed data
- In probabilistic modeling, the joint distribution $\mathbb{P}(x,\theta)$ of the observed variables x and the hidden parameters θ encapsulate the key information
 - Given: prior $\mathbb{P}(\theta)$ and likelihood $\mathbb{P}(x|\theta)$
 - Joint dist. from prior and likelihood: $\mathbb{P}(x, \theta) = \mathbb{P}(x|\theta)\mathbb{P}(\theta)$
 - We get: marginal likelihood $\mathbb{P}(x) = \int \mathbb{P}(x, \theta) d\theta$ and posterior $\mathbb{P}(\theta|x) = \frac{\mathbb{P}(x, \theta)}{\mathbb{P}(x)}$



Fully Bayesian vs. ML/MAP

Given the data set \mathcal{X} , we want to predict A, i.e., $\mathbb{P}(A \mid \mathcal{X})$

ML: Easy (high), Exact (low)

$$\mathbb{P}(A \mid \mathcal{X}) \approx \mathbb{P}(A \mid \boldsymbol{\theta}), \quad \boldsymbol{\theta} = \arg \max \mathbb{P}(\mathcal{X} \mid \boldsymbol{\theta})$$

MAP: Easy (mid), Exact (mid)

$$\mathbb{P}(A \mid \mathcal{X}) \approx \mathbb{P}(A \mid \boldsymbol{\theta}), \quad \boldsymbol{\theta} = \arg \max \mathbb{P}(\boldsymbol{\theta} \mid \mathcal{X})$$

- Fully Bayesian: Easy (low), Exact (high)
 - predictive inference, use of posterior predictive distribution, bayesian prediction
 - remove dependence on the model parameters heta

$$\mathbb{P}(A \mid \mathcal{X}) = \int \mathbb{P}(A \mid \boldsymbol{\theta}) \mathbb{P}(\boldsymbol{\theta} \mid \mathcal{X}) d\boldsymbol{\theta}$$

- Only possible by getting the full posterior distribution $\mathbb{P}(\theta \mid \mathcal{X})$



(Fully) Bayesian Inference: Hardness

• For a data set \mathcal{X} , a parameter prior $\mathbb{P}(\theta)$, and a likelihood function, the posterior is:

$$\mathbb{P}(oldsymbol{ heta} \mid \mathcal{X}) = rac{\mathbb{P}(\mathcal{X} \mid oldsymbol{ heta})\mathbb{P}(oldsymbol{ heta})}{\mathbb{P}(\mathcal{X})}, \quad \mathbb{P}(\mathcal{X}) = \int \mathbb{P}(\mathcal{X} \mid oldsymbol{ heta})\mathbb{P}(oldsymbol{ heta}) \, \mathrm{d}oldsymbol{ heta}$$

- Implementation hardness
 - Bayesian inference requires to solve integration, which is often challenging. In particular, a conjugate prior is not chosen, the integration is not analytically tractable.
 - Approximation techniques: MCMC (Markov Chain Monte Carlo),
 Laplace approximation, variational inference, expectation propagation



Latent-Variable Models (1)

- ullet Including latent variables in the model o contributing to the interpretability of the model
- General discussions here would be applied the following examples later
 - PCA for dimensionality reduction

L10(7)

• Gaussian mixture models for density estimation

L11(3)

- In latent-variable models (LVMs)²,
 - ullet Given: prior $\mathbb{P}(z)$ and likelihood $\mathbb{P}_{ heta}(x|z)$
 - Joint dist. from prior and likelihood: $\mathbb{P}_{\theta}(x,z) = \mathbb{P}_{\theta}(x|z)\mathbb{P}(z)$
 - ullet Our interest: marginal likelihood $\mathbb{P}_{m{ heta}}(m{x})$ and posterior $\mathbb{P}_{m{ heta}}(m{z}|m{x})$

²In our note, we express the dependence on the model parameters θ using subscript notations, e.g., $\mathbb{P}_{\theta}(x|z)$ rather than $\mathbb{P}(x|z,\theta)$ to highlight the role of z = 1

LVM (2)

- Assuming we know θ , to generate a data sample from the model (i) sample z from $\mathbb{P}(z)$ and (ii) sample x from $\mathbb{P}_{\theta}(x|z)$
- Inference. computing the posterior distribution $\mathbb{P}_{\theta}(z|x)$:

$$\mathbb{P}_{\theta}(z|x) = \frac{\mathbb{P}_{\theta}(x,z)}{\mathbb{P}_{\theta}(x)} = \frac{\mathbb{P}_{\theta}(x,z)}{\int \mathbb{P}_{\theta}(x,z) dz}$$

 This requires to solve the sub-problem of computing the marginal likelihood of the observation:

$$\mathbb{P}_{m{ heta}}(m{x}) = \int \mathbb{P}_{m{ heta}}(m{x}, m{z}) dm{z}$$



LVM (3): Why the posterior distribution $\mathbb{P}_{\theta}(z|x)$?

- Explanation of the observation. Allows us to figure out which latent configurations could have plausibly generated the observation data samples.
- Learning of model parameters θ . Training LVMs to estimate θ (e.g., ML) requires $\mathbb{P}_{\theta}(z|x)$ in its inner loops

marginal likelihood $\mathbb{P}_{ heta}(extbf{ extit{x}}) \implies$ posterior distribution $\mathbb{P}_{ heta}(extbf{ extit{z}}| extbf{ extit{x}}) \implies$



LVM (4): How is $\mathbb{P}_{\theta}(z|x)$? used for θ_{ML} ?

• In ML, we need the gradient of the marginal log-likelihood. For a data sample x,

$$\nabla_{\theta} \log p_{\theta}(\mathbf{x}) = \frac{\nabla_{\theta} p_{\theta}(\mathbf{x})}{p_{\theta}(\mathbf{x})} = \frac{\int \nabla_{\theta} p_{\theta}(\mathbf{x}, \mathbf{z}) d\mathbf{z}}{p_{\theta}(\mathbf{x})} = \frac{\int p_{\theta}(\mathbf{x}, \mathbf{z}) \nabla_{\theta} \log p_{\theta}(\mathbf{x}, \mathbf{z})}{p_{\theta}(\mathbf{x})}$$
$$= \int p_{\theta}(\mathbf{z}|\mathbf{x}) \nabla_{\theta} \log p_{\theta}(\mathbf{x}, \mathbf{z}) d\mathbf{z}$$

ullet $\mathbb{P}_{ heta}(z|x)$ performs credit assignment over latent configurations



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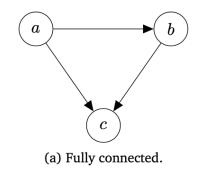
Graphical Models

- Joint distribution of a probabilistic model: key quantity of interest, but quite complicated without structural properties
- However, there exist relations of independence, conditional independence among random variables.
- (Probabilistic) graphical models: Roughly speaking, a graph of random variables.
 - Simple ways to visualize the structure of the model
 - Insights into the structural properties, e.g., conditional independence
 - Computations for inference and learning can be expressed in terms of graphical manipulations

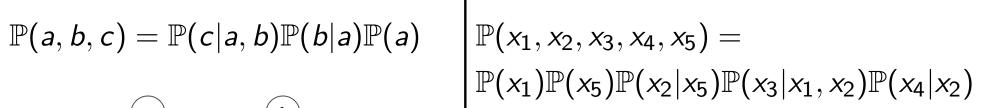


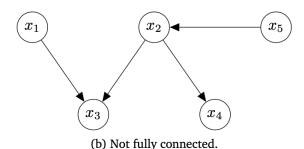
Graph Semantics

$$\mathbb{P}(a,b,c) = \mathbb{P}(c|a,b)\mathbb{P}(b|a)\mathbb{P}(a)$$



Nodes: random variables



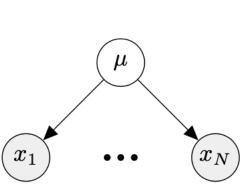


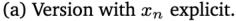
- Directed edge for direct dependence: b directly depends on a: $a \rightarrow b$
- Graph layout: factorization of the joint distribution

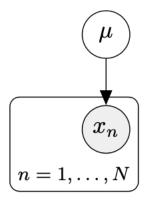
$$\mathbb{P}(x_1,\ldots,x_K)=\prod_{k=1}^K\mathbb{P}(x_k\mid\mathsf{Pa}_k),\quad\mathsf{Pa}_k$$
 are the parent nodes of $x_k.$



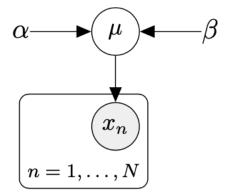
Example: N coin-flip experiments







(b) Version with plate notation.



- (c) Hyperparameters α and β on the latent μ .
- Shaded nodes: observables, μ : probability of head, a (latent) random variable
- Joint distribution

$$\mathbb{P}(x_1,\ldots,x_N\mid \mu)=\prod_{n=1}^N\mathbb{P}(x_n\mid \mu)$$



Conditional Independence and d-Separation

- Question. How can we see conditional independence in the directed graphical models? For example, $A \perp \!\!\!\perp B \mid C$?
- d-separation
 - ullet All possible trails from any node ${\mathcal A}$ to any node in ${\mathcal B}$
 - Any such path is blocked if it includes any node such that either of the following is true:
 - The arrows on the path meet either head to tail or tail to tail at the node, and the node is in $\mathcal C$
 - The arrows meet head to head at the node, and neither the node nor any of its descendants is in ${\mathcal C}$
 - If all the paths are blocked, then \mathcal{A} is d-separated from \mathcal{B} by \mathcal{C} .
 - If d-separated, $\mathcal{A} \perp \!\!\! \perp \mathcal{B} \mid \mathcal{C}$

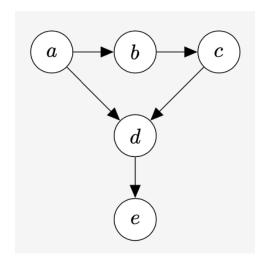




April 1, 2023

³paths that ignore the direction of the arrows

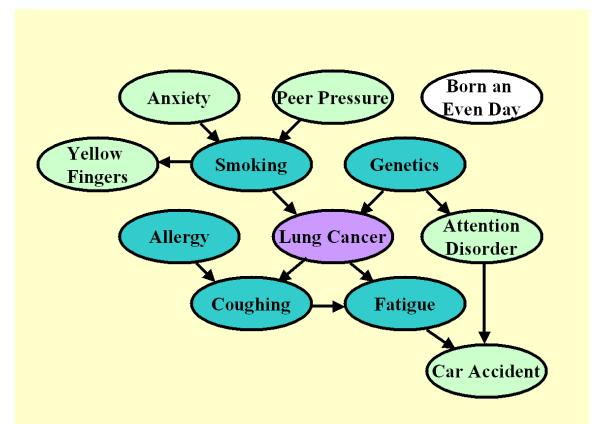
Example



- $b \perp \!\!\! \perp d \mid a, c$
- a ⊥⊥ c | b
- *b* ⊥⊥ *d* | *c*
- a ⊥ c | b, e



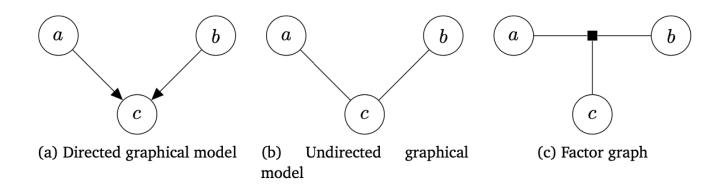
Example in Healthcare



Source: http://www.causality.inf.ethz.ch/data/LUCAS.html



Three Types of Graphical Models



- Directed graphical models (or Bayesian Networks)
- Undirected graphical models (Markov Random Fields)
- Factor graphs

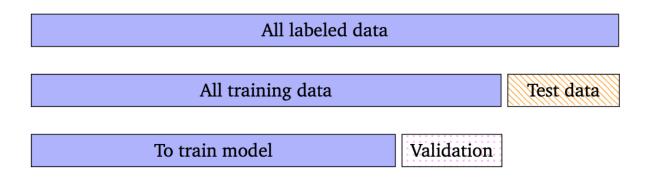


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Nested Cross-Validation



- Model selection
 - Tradeoff between model complexity and data fit
 - Occam's razor. Find the simplest model that explains the data resonably well.
- Test set: estimate the generalization performance
- Validation set: choose the best model



Bayesian Model Selection

• A set of models $M = \{M_1, \dots, M_k\}$, where each M_k has θ_k parameters. A prior $\mathbb{P}(M)$ on each model $M \in M$.

$$M_k \sim \mathbb{P}(M), \quad oldsymbol{ heta}_k \sim \mathbb{P}(oldsymbol{ heta} \mid M_k), \quad \mathcal{D} \sim \mathbb{P}(\mathcal{D} \mid oldsymbol{ heta}_k)$$

• Posterior distribution $\mathbb{P}(M_k \mid \mathcal{D}) \propto \mathbb{P}(M_k)\mathbb{P}(\mathcal{D} \mid M_k)$, where we have the following model evidence or marginal likelihood:

$$\mathbb{P}(\mathcal{D} \mid M_k) = \int \mathbb{P}(\mathcal{D} \mid \boldsymbol{\theta}_k) \mathbb{P}(\boldsymbol{\theta}_k \mid M_k) d\boldsymbol{\theta}_k \quad (***)$$

- MAP for the model: $M^* = \arg\max_{M_k} \mathbb{P}(M_k \mid \mathcal{D})$
- With the uniform model prior (i.e., $\mathbb{P}(M_k) = 1/k$), the MAP estimate equals to maximization of model evidence.

Bayes Factors for Model Comparison

• Compare two probabilistic models M_1 and M_2 :

$$(\text{Posterior odds}) = \frac{\mathbb{P}(M_1 \mid \mathcal{D})}{\mathbb{P}(M_2 \mid \mathcal{D})} = \frac{\frac{\mathbb{P}(\mathcal{D}|M_1)\mathbb{P}(M_1)}{\mathbb{P}(\mathcal{D})}}{\frac{\mathbb{P}(\mathcal{D}|M_2)\mathbb{P}(M_2)}{\mathbb{P}(\mathcal{D})}} = \underbrace{\frac{\mathbb{P}(M_1)}{\mathbb{P}(M_1)}}_{\text{Prior odds}} \underbrace{\frac{\mathbb{P}(\mathcal{D} \mid M_1)}{\mathbb{P}(\mathcal{D} \mid M_2)}}_{\text{Prior odds}} \underbrace{\frac{\mathbb{P}(\mathcal{D} \mid M_1)}{\mathbb{P}(\mathcal{D} \mid M_2)}}_{\text{Prior odds}}$$

- ullet $\mathbb{P}(\mathcal{D} \mid M_k)$: How well the data is predicted by the model M_k
- ullet With the uniform model prior, the prior odds =1
- Computation of Bayes factor requires the complex integration (***)
 in the previous slide. In this case, we rely on some approximations
 such as MCMC (Markov Chain Monte Carlo).



Summary



THANKS FOR YOUR ATTENTION



Discussions



