

Nonparametric-efficient Causal Mediation Analysis for Stochastic Interventions Nima Hejazi, Mark van der Laan, and Iván Díaz

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OVERVIEW & MOTIVATIONS

- Using stochastic interventions, we present a decomposition of the *population intervention effect* into <u>direct</u> and <u>indirect</u> effects.
 - Define causal contrasts of effects of continuous and categorical exposures
 - **–** ...
- We propose estimators of these direct and indirect effects:
 - Classical parametric: substitution and re-weighted (IPW) estimators
 - Nonparamtric-efficient: one-step and TMLE using data adaptive regression
- Our efficient estimators are asymptotically linear under a condition requiring $n^{1/4}$ -consistency of certain regression functions.

SOFTWARE IMPLEMENTATION

- The medshift R package [3] implements these estimators and leverages state-of-theart machine learning in the procedure.
 - Construction of all estimators via the eponymous medshift () function.
 - Uses the s13 R package to provide machine learning facilities.
- Construction of TML estimators using tools from the tlverse software ecosystem.
- ...

CONSTRUCTION OF NONPARAMETRIC-EFFICIENT ESTIMATORS

- To avoid entropy conditions on initial estimators, we rely on cross-fitting [6, 1]. Let $\hat{\eta}_j$ be the estimator of $\eta = (g, m, e, \phi)$ and j(i) the index of the validation set containing observation i.
- A one-step estimator [4] may be constructed by augmenting the substitution estimator with the efficient influence function:

$$\hat{\theta}_{OS}(\delta) = \frac{1}{n} \sum_{i=1}^{n} D_{\hat{\eta}_{j(i)}, \delta}(O_i) = \frac{1}{n} \sum_{i=1}^{n} \left\{ D_{\hat{\eta}_{j(i)}, \delta}^{Y}(O_i) + D_{\hat{\eta}_{j(i)}, \delta}^{A}(O_i) + D_{\hat{\eta}_{j(i)}, \delta}^{Z, W}(O_i) \right\}.$$

- ..
- **–** ..
- A targeted minimum loss-based estimator may be constructed by using the efficient influence function as an estimating equation, updating estimates of nuisance components:

$$\hat{\theta}_{\text{TMLE}}(\delta) = \dots,$$

where

- Unlike the one-step estimator, the TMLE is a substitution estimator.
- Use unversal least favorable submodels for one-step estimation [5].

STOCHASTIC POPULATION INTERVENTION (IN)DIRECT EFFECTS

- Consider $O = (W, A, Z, Y) \sim P_0 \in \mathcal{M}$, where W is a set of baseline covariates, A an intervention, Z a mediator between A and outcome, and Y the outcome, with no assumptions on model \mathcal{M} .
- We may decompose the PIE in terms of a population intervention direct effect (PIDE) and a population intervention indirect effect (PIIE):

$$\psi(\delta) = \underbrace{\mathbb{E}\{Y(g,q) - Y(g_{\delta},q)\}}_{\text{PIDE}} + \underbrace{\mathbb{E}\{Y(g_{\delta},q) - Y(g_{\delta},q_{\delta})\}}_{\text{PIDE}}.$$

• We show the causal parameter $\mathbb{E}\{Y(g_{\delta},q)\}$ is identified by the observed data parameter [2]:

$$\theta(\delta) = \int m(a, z, w) g_{\delta}(a \mid w) p(z, w) d\nu(a, z, w)$$

• Letting $\eta = (g, m, e, \phi)$, the efficient influence function for $\theta(\delta)$ in the nonparametric model M is $D_{n,\delta}^Y(o) + D_{n,\delta}^A(o) + D_{n,\delta}^{Z,W}(o) - \theta(\delta)$, where

$$D_{\eta,\delta}^{Z,W}(o) = \int m(z, a, w) g_{\delta}(a \mid w) d\kappa(a),$$

$$D_{\eta,\delta}^{A}(o) = \frac{g_{\delta}(a \mid w)}{g(a \mid w)} \left\{ \phi(a, w) - \int \phi(a, w) g_{\delta}(a \mid w) d\kappa(a) \right\}$$

$$D_{\eta,\delta}^{Y}(o) = \frac{g_{\delta}(a \mid w)}{e(a \mid z, w)} \{ y - m(z, a, w) \},$$

where
$$\phi_0(a, w) = \mathbb{E}\left\{\frac{g(A|W)}{e(A|Z,W)}m(Z, A, W) \mid A = a, W = w\right\}.$$

RESULTS & DISCUSSION

- All estimators approx. unbiased in large samples; however, inefficient TMLE with HAL has bias not converging at $n^{-\frac{1}{2}}$.
- Fitting Π with HAL or GLM, efficient TMLE has lower variance than the inefficient.

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BUT WAIT, THERE'S MORE!

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- https://arxiv.org/abs/1901.02776
- Check out Iván's talk tomorrow morning!