## SOME DISCRETE DISTRIBUTIONS

| Name               | Parameters   | Genesis / Usage  | Notation                          | $p(x) = \mathbb{P}[X = x]$ (and non-zero range)                               | $\mathbb{E}[X]$           | Var(X)                                       | Comments   |
|--------------------|--|--|-----------------------------------|---|---------------------------|--|--|
| Uniform (discrete) | $k \in \mathbb{N}$   | Set of $k$ equally likely outcomes   | $Unif(1,\ldots,k)$ (not standard) | p(x) = 1/k $x = 1,, k$  | $\frac{k+1}{2}$           | $\frac{k^2-1}{12}$                           | Fair dice roll $(k=6)$   |
| Bernoulli trial    | $\theta \in [0,1]$   | Experiment with two outcomes (typically, success $= 1$ , fail $= 0$ )  | $Bernoulli(\theta)$               | $p(x) = \theta^x (1 - \theta)^{1-x}$ $x = 0, 1$                               | θ                         | $\theta(1-\theta)$                           | Coin toss  |
| Binomial           | $n \in \mathbb{N},  \theta \in [0, 1]$                             | Number of successes in $n$ i.i.d. Bernoulli trials   | $Bi(n, \theta)$<br>B(n, p)        | $p(x) = \binom{n}{x} \theta^{x} (1 - \theta)^{n-x}$<br>x = 0, 1, 2,, n        | $n\theta$                 | $n\theta(1-\theta)$                          | Sampling with replacement $Bi(1, \theta) \equiv Bernoulli(\theta)$                                       |
| Geometric          | $\theta \in (0,1]$   | Number of successful i.i.d. Bernoulli trials before first failure  | $Geom(\theta)$<br>$Geo(\theta)$   | $p(x) = \theta^{x}(1 - \theta)$<br>$x = 0, 1, 2, \dots$                       | $\frac{\theta}{1-\theta}$ | $\frac{\theta^2}{(1-\theta)^2}$              | Alternative formulations might swap $\theta$ and $1-\theta$ , or use $X'=X+1$ to include the final trial |
| Negative Binomial  | $k \in \mathbb{N}, \ \theta \in (0,1]$                             |  |                                   | $p(x) = {\binom{x-1}{k-1}} \theta^k (1-\theta)^{x-k}  x = k, k+1, k+2, \dots$ | $\frac{k}{\theta}$        | $\frac{k(1-\theta)}{\theta^2}$               | Several alternative formulations exist.  |
| Hypergeometric     | $N \in \mathbb{N}$ $k \in \{0, \dots, N\}$ $n \in \{0, \dots, n\}$ | Number of special objects in a random sample of $n$ objects, from a population of $N$ objects with $k$ special objects | HypGeom(N,k,n) (not standard)     | $p(x) = {k \choose x} {N-k \choose n-x} / {N \choose n}$ $x = 0,, n$          | $\frac{nk}{N}$            | $n\frac{N-n}{N-1}\frac{k}{N}(1-\frac{k}{N})$ |  |
| Poisson            | $\lambda \in (0, \infty)$  | Counting events occurring 'at random' within space or time   | $Poi(\lambda)$                    | $p(x) = \frac{e^{-\lambda} \lambda^x}{x!}$ $x = 0, 1, 2, \dots$               | λ                         | λ  |  |

## SOME CONTINUOUS DISTRIBUTIONS

| Name                 | Parameters  | Genesis / Usage  | Notation   | $f(x) = \mathbf{p.d.f.}$ (and non-zero range)   | $\mathbb{E}[X]$   | Var(X)   | Comments   |
|----------------------|---|--|--|---|---|--|--|
| Uniform (continuous) | $\alpha, \beta \in \mathbb{R} \text{ with } \alpha < \beta$ | The uniform distribution for a continuous interval                             | $Unif(\alpha,\beta) \\ U(a,b)$                                       | $f(x) = \frac{1}{\beta - \alpha}$ $x \in (\alpha, \beta)$   | $\frac{\alpha+\beta}{2}$                                  | $\frac{(\beta - \alpha)^2}{12}$  | Also written as $U[\alpha, \beta]$ and similarly for open and half-open intervals.   |
| Normal               | $\mu \in \mathbb{R}, \ \sigma \in (0, \infty)$              | Empirically and theoretically (via CLT, etc.) a good model in many situations. | $N(\mu, \sigma^2)$   | $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$ $x \in \mathbb{R}$                       | μ   | $\sigma^2$   | $N(0,1) \equiv \text{standard normal.}$<br>$X \sim N(\mu, \sigma^2) \Rightarrow$<br>$aX + b \sim N(a\mu + b, a^2\sigma^2)$<br>Hence $Z = \frac{X - \mu}{\sigma} \sim N(0,1)$ |
| Exponential          | $\lambda \in (0, \infty)$                                   | Inter-arrival times of random events   | $Exp(\lambda)$   | $ f(x) = \lambda e^{-\lambda x} $ $ x > 0 $   | $\frac{1}{\lambda}$                                       | $\frac{1}{\lambda^2}$  | Alternative parametrization: $\theta = \frac{1}{\lambda}$  |
| Gamma                | $\alpha, \beta \in (0, \infty)$                             | Lifetimes of ageing items, multi-inter-arrival times                           | $Ga(\alpha, \beta)$<br>$\Gamma(\alpha, \beta)$                       | $f(x) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x}$ $x > 0$  | $\frac{\alpha}{\beta}$                                    | $\frac{\alpha}{\beta^2}$   | Alternative parametrization: $\theta = 1/\beta$ , $Ga(1,\lambda) \equiv Exp(\lambda)$ , $Ga(n/2,1/2) \equiv \chi_n^2$  |
| Log-normal           | $\mu \in \mathbb{R}, \sigma \in (0, \infty)$                | Quantities related to exponential growth                                       | $\begin{array}{c} logN(\mu,\sigma^2) \\ (not  standard) \end{array}$ | $f(x) = \frac{1}{x\sigma\sqrt{2\pi}} \exp\left(-\frac{(\log x - \mu)^2}{\sqrt{2}\sigma}\right)$ $x > 0$                       | $e^{\mu + \frac{1}{2}\sigma^2}$                           | $(e^{\sigma^2} - 1)e^{2\mu + \sigma^2}$  | If $X \sim log N(\mu, \sigma^2)$ then $log X \sim N(\mu, \sigma^2)$  |
| Chi-squared          | $n \in \mathbb{N}$  | Squared (normally distributed) errors, statistical tests                       | $\chi_n^2$   | $f(x) = \frac{1}{2^{n/2}\Gamma(n/2)}x^{n/2-1}e^{-x/2}$ $x > 0$  | n   | 2n   | $ X_n^2 \equiv Ga(n/2, 1/2) $ $X_i \sim N(0, 1) \text{ i.i.d. } \Rightarrow \sum_{i=1}^n X_i^2 \sim \chi_n^2 $   |
| Beta                 | $\alpha, \beta \in (0, \infty)$                             | Quantities constrained to be within intervals                                  | $Be(\alpha,\beta) \\ Beta(\alpha,\beta)$                             | $f(x) = \frac{x^{\alpha - 1}(1 - x)^{\beta - 1}}{B(\alpha, \beta)}$ $x \in [0, 1]$  | $\frac{\alpha}{\alpha + \beta}$                           | $\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$                                   | $Be(1,1) \equiv Unif(0,1)$   |
| Cauchy               | $a,b \in \mathbb{R}$  | Heavy tailed, pathological examples  | Cauchy(a,b)  | $f(x) = \frac{1}{\pi b} \frac{b^2}{(x-a)^2 + b^2}$ $x \in \mathbb{R}$   | undefined   | undefined  | Cauchy(0,1) is often called 'the' Cauchy distribution  |
| Pareto               | $\alpha, \beta \in (0, \infty)$                             | Heavy tailed quantities  | $Pareto(\alpha, \beta)$  | $f(x) = \frac{\alpha \beta^{\alpha}}{x^{\alpha+1}}$ $x > \beta$   | $\frac{\alpha\beta}{\alpha+1}$ if $\alpha > 1$            | $\frac{\alpha^2 \beta}{(\alpha - 1)^2 (\alpha - 2)}$ if $\alpha > 2$                     | If $X \sim Pareto(\alpha, \beta)$ then $\log \frac{X}{\beta} \sim Exp(\alpha)$   |
| Weibull              | $\lambda, k \in (0, \infty)$                                | Lifetimes, extreme values, particle sizes                                      | $Weibull(\lambda,k)$   | $f(x) = \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{k-1} e^{-(x/\lambda)^k}$ $x > 0$                                    | $\lambda\Gamma(1+1/k)$                                    | $\lambda^{2} \left[ \Gamma(1+2/k) + \Gamma(1+1/k)^{2} \right]$                           | If $X \sim Weibull(\lambda, k)$ then $(X/\lambda)^k \sim Exp(1)$   |
| Student t            | $n \in \mathbb{N}$  | Statistical tests  | $t_n$  | $f(x) = \frac{\Gamma(\frac{n+1}{2})}{\sqrt{n\pi}\Gamma(\frac{n}{2})} (1 + \frac{x^2}{n})^{-\frac{n+1}{2}}$ $x \in \mathbb{R}$ | 0 if $n > 1$  | $\frac{n}{n-2}$ if $n > 2$   | $t_1 \equiv Cauchy(0,1)$<br>Can take $n \in (0,\infty)$  |
| F                    | $\nu, \delta \in (0, \infty)$                               | Statistical tests  | $F_{ u,\delta}$  | $f(x) = \frac{\nu^{\nu/2} \delta^{\delta/2} x^{\nu/2 - 1}}{B(\nu/2, \delta/2)(\nu x + \delta)^{(\nu + \delta)/2}}$ $x > 0$    | $\frac{\frac{\delta}{\delta - 2}}{\text{if } \delta > 2}$ | $\frac{\frac{2\delta^2(\nu+\delta-2)}{\nu(\delta-2)^2(\delta-4)}}{\text{if }\delta > 4}$ | If $X \sim \chi^2_{\nu}$ and $Y \sim \chi^2_{\delta}$ are independent then $\frac{X/\nu}{Y/\delta} \sim F_{\nu,\delta}$ .  |
|                      |   |  |  |   |   |  | If $T \sim t_{\nu}$ then $T^2 \sim F_{1,\nu}$ .<br>If $Z \sim Be(\alpha, \beta)$ then $\frac{\beta Z}{\alpha(1-Z)} \sim F_{2\alpha,2\beta}$ .                                |