## MAS350: Assignment 3

Solutions and discussion are written in blue. A sample mark scheme, with a total of 25 marks, is given in red, with each mark placed after the statement/deduction for which the mark would be given. As usual, mathematically correct solutions that follow a different method would be marked analogously.

1. Let  $f_n, f: [0,1] \to \mathbb{R}$ . In each of the following cases, explain whether the Monotone and/or Dominated Convergence Theorems can be used to prove that  $\int_0^1 f_n(x) dx \to \int_0^1 f(x) dx$ .

(a) 
$$f_n(x) = \cos(\frac{x}{n}) + \sin(\frac{x}{n})$$
 and  $f(x) = 1$ .

(b) 
$$f_n(x) = \mathbb{1}_{[1,1]}(x) x^{-1}$$
 and  $f(x) = \mathbb{1}_{(0,1]}x^{-1}$ .

(c) 
$$f_n(x) = \mathbb{1}_{[0,\frac{1}{2}]}(x) n$$
 and  $f(x) = 0$ .

Solution.

- (a) DCT only (the MCT can't be used here because  $f_n \leq f_{n+1}$  doesn't hold). [2]
- (b) MCT only (the DCT can't be used here because  $\int_0^1 f(x) dx = \infty$ ). [2]
- (c) Neither, in this case  $\int_0^1 f_n(x) dx = 1$  and  $\int_0^1 f(x) dx = 0$ . [2]

2. Let  $(S, \Sigma, m)$  be a measure space. Let  $f: S \to [0, \infty)$  be measurable and let c > 0. Consider the following two facts, which were stated (and proved) within the lecture notes:

(a) 
$$\left| \int_{S} f \, dm \right| \leq \int_{S} |f| \, dm$$
,

(b) 
$$m(\{x \in S : f(x) \ge c\}) \le \frac{1}{c} \int_S f \, dm.$$

You do *not* need to prove these facts here.

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space and let  $X : \Omega \to [0, \infty)$  be a random variable. Let  $\mathbb{E}$  denote expectation with respect to  $\mathbb{P}$ . Use this notation to write down probabilistic versions of statements (a) and (b).

Solution.

(a) 
$$|\mathbb{E}[X]| \le \mathbb{E}[|X|]$$
. [2]

(b) 
$$\mathbb{P}[X \ge c] \le \frac{1}{c} \mathbb{E}[X]$$
. [2]

3. Consider the probability space ([0, 1],  $\mathcal{B}([0, 1])$ ,  $\lambda$ ) where  $\lambda$  denotes the restriction of Lebesgue measure to the Borel  $\sigma$ -field  $\mathcal{B}([0, 1])$  on [0, 1].

1

Let 
$$X_n(\omega) = \begin{cases} 1 & \text{if } \omega = 0\\ \omega n^{3/2} & \text{if } \omega \in (0, \frac{1}{n}]\\ 0 & \text{if } \omega \in (\frac{1}{n}, 1]. \end{cases}$$

Determine in which modes of convergence we have  $X_n \to 0$ .

Solution.

We first check almost sure convergence. Fix some  $\omega \in [0,1]$ . If  $\omega \in (0,1]$  then for all large enough n we have  $\frac{1}{n} < \omega$ . For such n we have  $X_n(\omega) = 0$ , [1] which means  $X_n(\omega) \to 0$ . We thus obtain that  $\{X_n(\omega) \to 0\} \subseteq (0,1]$  [1] so

$$\lambda(\{X_n(\omega) \to 0\}) \ge \lambda((0,1]) = 1,$$

which means that  $X_n \stackrel{a.s.}{\to} 0$ . [1] It follows that  $X_n \stackrel{\mathbb{P}}{\to} 0$  [1] and also that  $X_n \stackrel{d}{\to} 0$ . [1] Lastly, the expectation of  $|X_n|^p = X_n^p$  is given by

$$\mathbb{E}[X_n^p] = \int_0^1 X_n(\omega)^p \, d\lambda(\omega)$$

$$= \int_0^{\frac{1}{n}} \omega^p n^{3p/2} \, d\lambda(\omega)$$

$$= n^{3p/2} \left[ \frac{\omega^{p+1}}{p+1} \right]_0^{\frac{1}{n}}$$

$$= n^{3p/2} \frac{(1/n)^{p+1}}{p+2}$$

$$= n^{p/2-1}.$$

[1] Here we use that  $\{0\}$  is a  $\lambda$ -null subset of [0,1] (so values of  $X_n$  here have no effect on the integral) [1] and that  $X_n(\omega) = 0$  when  $\omega > \frac{1}{n}$ . [1]

Noting that  $n^{p/2-1} \to 0$  if and only if p < 2, we have that  $X_n \stackrel{L^p}{\to} 0$  if and only if p < 2. [1] [I would accept "p = 1 works but  $p = 2, 3, 4 \dots$  does not"]

- 4. (a) Let  $(U_n)_{n\in\mathbb{N}}$  be a sequence of independent, identically distributed uniform random variables on (0,1). Prove that,  $\mathbb{P}[U_n < 1/n \text{ i.o.}] = 1$  and  $\mathbb{P}[U_n < 1/n^2 \text{ i.o.}] = 0$ .
  - (b) Let  $(X_n)_{n\in\mathbb{N}}$  be the sequence of results obtained from infinitely many rolls of a fair six sided dice. Prove that the (consecutive) pattern 123456 will occur infinitely often.

Solution.

- (a) We have  $\mathbb{P}[U_n \leq a] = a$ . For any (deterministic) sequence  $(x_n)$  the events  $\{U_n < x_n\}$  are independent, because the  $U_n$  are independent. [1]

  Noting that  $\sum 1/n = \infty$  and  $\sum 1/n^2 < \infty$ , we have  $\sum_n \mathbb{P}[U_n < 1/n] = \infty$  and  $\sum_n \mathbb{P}[U_n < 1/n^2] < \infty$ . [1]

  By the second Borel-Cantelli lemma  $\mathbb{P}[U_n < 1/n \text{ i.o.}] = 1$  and by the first Borel-Cantelli lemma  $\mathbb{P}[U_n < 1/n^2 \text{ i.o.}] = 0$ . [1]
- (b) Let  $E_n = \{X_n + i = i \text{ for } i = 1, 2, 3, 4, 5, 6\}$ . We have  $\mathbb{P}[E_n] = (1/6)^6 > 0$ . Note that  $E_n$  and  $E_{n+6}$  are independent (but  $E_n$  and  $E_{n+1}$  are not!). [1] We have  $\sum_{n=1}^{\infty} \mathbb{P}[E_{6n}] = \sum_{n=1}^{\infty} (1/6)^n = \infty$ , [1] hence by the second Borel-Cantelli lemma we have  $\mathbb{P}[E_{6n} \text{ i.o.}] = 1$ . [1] Noting that  $\{E_{6n} \text{ i.o.}\} \subseteq \{E_n \text{ i.o.}\}$ , we have  $\mathbb{P}[E_n \text{ i.o.}] = 1$ .