Probability and Random Process (SWE3026)

Joint Distributions

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H. Pishro-Nik, "Introduction to probability, statistics, and random processes", available at https://www.probabilitycourse.com, Kappa Research LLC, 2014.

The covariance between X and Y is defined as

$$Cov(X,Y) = E[(X - EX)(Y - EY)] = E[XY] - (EX)(EY).$$

Proof:

$$E[(X - EX)(Y - EY)] = E[XY - X(EY) - (EX)Y + (EX)(EY)]$$

= $E[XY] - (EX)(EY) - (EX)(EY) + (EX)(EY)$
= $E[XY] - (EX)(EY)$.

Example. Suppose $X \sim Uniform(1,2)$ and given X=x , Y is

$$Y|X=x \sim Exponential(\lambda=x).$$

Find Cov(X, Y).

$$f_X(x) = egin{cases} 1 & 1 \leq x \leq 2 \ 0 & ext{else} \end{cases}, \quad f_{Y|X}(y|x) = \lambda e^{-\lambda y} u(y).$$

Lemma. The covariance has the following properties:

1)
$$Cov(X, X) = E[XX] - EXEX = E[X^2] - (EX)^2 = Var(X)$$
.

2) X&Y independent:

$$Cov(X, X) = E[XY] - EXEY = E[X]E[Y] - EXEY = 0.$$

- 3) Cov(X, Y) = Cov(Y, X)
- 4) $\operatorname{Cov}(aX,Y) = a\operatorname{Cov}(X,Y)$ $a \in \mathbb{R}$

- 5) $\operatorname{Cov}(X+c,Y) = \operatorname{Cov}(X,Y)$
- 6) $\operatorname{Cov}(X + Y, Z) = \operatorname{Cov}(X, Z) + \operatorname{Cov}(Y, Z)$
- 7) $\operatorname{Cov}(X+Y,Z+W) = \operatorname{Cov}(X,Z) + \operatorname{Cov}(X,W) + \operatorname{Cov}(Y,Z) + \operatorname{Cov}(Y,W)$

$$Cov(2X + Y, 3Z + W) = 6Cov(X, Z) + 2Cov(X, W) +$$
$$3Cov(Y, Z) + Cov(Y, W)$$

More generally

$$\operatorname{Cov}\left(\sum_{i=1}^m a_i X_i, \sum_{j=1}^n b_j Y_j\right) = \sum_{i=1}^m \sum_{j=1}^n a_i b_j \operatorname{Cov}(X_i, Y_j).$$

Variance of a sum:

If
$$Z = X + Y$$
, then

$$\begin{aligned} \operatorname{Var}(Z) &= \operatorname{Cov}(Z, Z) \\ &= \operatorname{Cov}(X + Y, X + Y) \\ &= \operatorname{Cov}(X, X) + \operatorname{Cov}(X, Y) + \operatorname{Cov}(Y, X) + \operatorname{Cov}(Y, Y) \\ &= \operatorname{Var}(X) + \operatorname{Var}(Y) + 2\operatorname{Cov}(X, Y). \end{aligned}$$

More generally, for $a,b\in\mathbb{R}$, we conclude:

$$Var(aX + bY) = a^{2}Var(X) + b^{2}Var(Y) + 2abCov(X, Y).$$

Correlation Coefficient:

$$\rho_{XY} = \rho(X, Y) = \frac{\operatorname{Cov}(X, Y)}{\sqrt{\operatorname{Var}(X) \operatorname{Var}(Y)}} = \frac{\operatorname{Cov}(X, Y)}{\sigma_X \sigma_Y}.$$

Properties of the correlation coefficient:

1)
$$-1 \le \rho(X, Y) \le 1$$
;

2)
$$\rho(aX + b, cY + d) = \rho(X, Y)$$
 for $a, c > 0$;

3)
$$ho(X,Y)=1$$
 if $Y=aX+b$ $a>0;$ $ho(X,Y)=-1$ if $Y=aX+b$ $a<0.$

Definition. Consider two random variables X and Y:

- 1) If ho(X,Y)=0 , we say that X and Y are uncorrelated.
- 2) If ho(X,Y)>0 , we say that X and Y are positively correlated.
- 3) If ho(X,Y) < 0 , we say that X and Y are negatively correlated.

Definition. Two random variables X and Y are said to be bivariate normal, or jointly normal, if aX+bY has a normal distribution for all $a,b\in\mathbb{R}$.

$$b=0, a=1 \longrightarrow X: Normal$$

$$a=0,\ b=1\longrightarrow Y: Normal$$

Definition. Two random variables X and Y are said to have a bivariate normal distribution with parameters $\mu_X, \sigma_X^2, \mu_Y, \sigma_Y^2$ and ρ , if their joint PDF is given by

$$\begin{split} f_{XY}(x,y) &= \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \cdot \\ &\exp\left\{-\frac{1}{2(1-\rho^2)} \left[\left(\frac{x-\mu_X}{\sigma_X}\right)^2 + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2 - 2\rho \frac{(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y} \right] \right\} \end{split}$$

Where $\mu_X, \mu_Y \in \mathbb{R}, \; \sigma_X, \sigma_Y > 0 \;$ and $\rho \in (-1,1)$ are all constants.

If ho=0 (X&Y are uncorrelated):

$$\begin{split} f_{XY}(x,y) &= c \cdot \exp\left\{-\frac{1}{2} \bigg(\frac{x-\mu_X}{\sigma_X}\bigg)^2 - \frac{1}{2} \bigg(\frac{y-\mu_Y}{\sigma_Y}\bigg)^2\right\} \\ &= c' \ \exp\left\{-\frac{1}{2} \bigg(\frac{x-\mu_X}{\sigma_X}\bigg)^2\right\} \cdot d \ \exp\left\{-\frac{1}{2} \bigg(\frac{y-\mu_Y}{\sigma_Y}\bigg)^2\right\} \end{split}$$
 Function of X

Theorem. If \boldsymbol{X} and \boldsymbol{Y} are bivariate normal and uncorrelated, then they are independent.

$$f_{XY}(x,y) = f_X(x)f_Y(y).$$

Theorem. Let X and Y be two bivariate normal random variables. Then, there exist independent standard normal random variables Z_1 and Z_2 such that

$$\begin{cases} X = \sigma_X Z_1 + \mu_X \\ Y = \sigma_Y \left(\rho Z_1 + \sqrt{1 - \rho^2} Z_2 \right) + \mu_Y \end{cases}$$

Theorem. Suppose X and Y are jointly normal random variables with parameters $\mu_X, \sigma_X^2, \mu_Y, \sigma_Y^2$ and ρ . Then, given X=x,Y is normally distributed with

$$E[Y|X=x] = \mu_Y +
ho\sigma_Y rac{x-\mu_X}{\sigma_X},$$
 $ext{Var}(Y|X=x) = (1-
ho^2)\sigma_Y^2.$

Theorem. If X and Y are bivariate normal & uncorrelated, then they are independent.

$$E[Y|X=x] = \mu_Y +
ho\sigma_Y rac{x-\mu_X}{\sigma_X},$$
 $ext{Var}(Y|X=x) = (1-
ho^2)\sigma_Y^2.$