# npark62 - ISYE6502 Homework 2 - 2/16/2020

(All discussions are included as part of code as #comments)

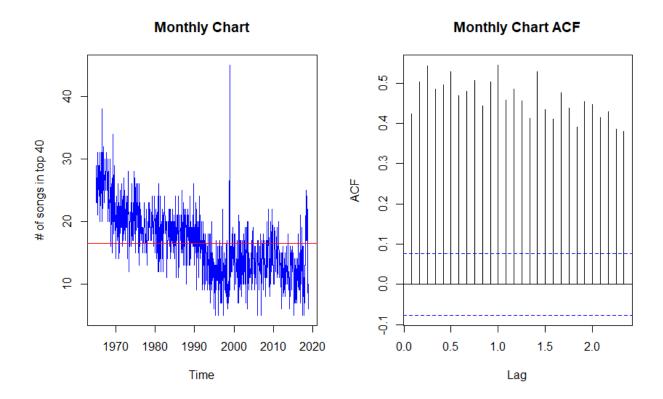
```
#npark62 ISYE6402 HW2
#Load chart_monthly.csv

library(TSA)
library(mgcv)

fname <- file.choose("C://Users//nicholas.park//Downloads//chart_monthly")
data <- read.csv(fname)
data <- data[,2]
chart = ts(data,start=c(1965,1),freq=12)
chart.dif = diff(chart)

#1a
par(mfrow=c(1,2))

ts.plot(chart,main='Monthly Chart',ylab='# of songs in top 40', col='blue')
abline(a=mean(chart),b=0,col='red')
acf(chart,main='Monthly Chart ACF')</pre>
```



<sup>#</sup> Unclear/inconsistent downward trend. Needs more quantitative testing to understand trend.

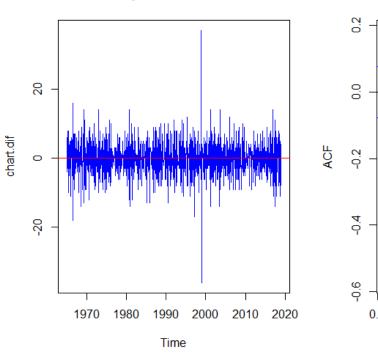
<sup># 1.</sup> Non-constant mean.

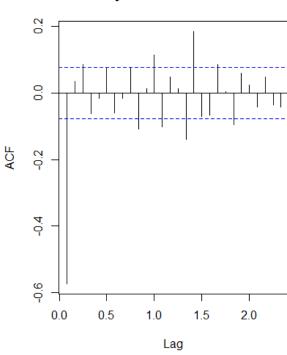
```
# 2. Seemingly constant (& finite) variance except for a suddenly large
variance in 2000 due to the spike in the data.
# 3. Significant autocorrelation is present.
# Thus, not stationary.

par(mfrow=c(1,2))
ts.plot(chart.dif,main='Monthly Chart Differenced', col='blue')
abline(a=mean(chart.dif),b=0,col='red')
acf(chart.dif,main='Monthly Chart Differenced ACF')
pacf(chart.dif,main='Monthly Chart Differenced PACF')
```

# **Monthly Chart Differenced**

# Monthly Chart Differenced ACF





# Monthly Chart Differenced PACF

```
0.7
       0.0
      Ö.
Partial ACF
       Ŋ
       q
       ന
       ö
       о
4.
       LO.
       o.
           0.0
                       0.5
                                   1.0
                                               1.5
                                                           2.0
                                        Lag
```

```
# Analysis:
# 1. Constant mean.
# 2. Heteroscedasticity is very minor. Seemingly constant (& finite) variance except for near year 2000.
# 3. ACF shows autocorrelation at lag 0.1.
# 4. PACF shows significant partial autocorrelation up to lag 1.
# Due to the ACF shown, the different time series is not so stationary and doesn't quite resemble white noise.
# Additional modeling is likely necessary for further improvement, as we will prove in the next question.
```

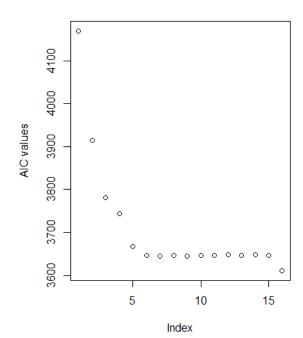
#1b

```
n = length(chart)
norder = 4 # since max order is 3.
p = c(1:norder)-1; q = c(1:norder)-1
aic = matrix(0,norder,norder)
for(i in 1:norder){
    for(j in 1:norder){
        modij = arima(chart,order = c(p[i],1,q[j]), method='ML')
            aic[i,j] = modij$aic-2*(p[i]+q[j]+1)+2*(p[i]+q[j]+1)*n/(n-p[i]-q[j]-2)
    }
}
aic2 = matrix(0,norder,norder)
for(i in 1:norder){
    for(j in 1:norder){
        modij = arima(chart,order = c(p[i],0,q[j]), method='ML')
        aic2[i,j] = modij$aic-2*(p[i]+q[j]+1)+2*(p[i]+q[j]+1)*n/(n-p[i]-q[j]-2)
```

```
min(aic) #3610.863
min(aic2) #3634.947

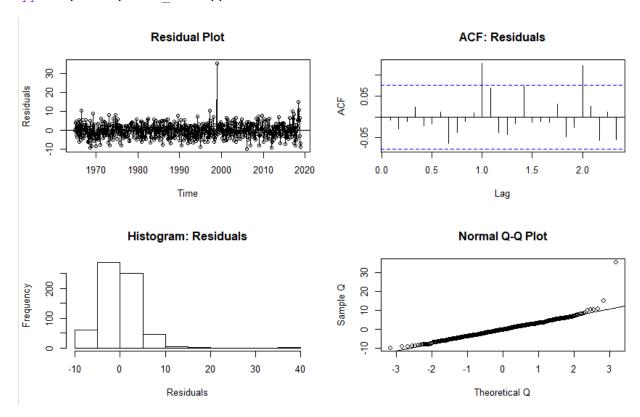
# since the minimum occurs when the differencing order d = 1 as opposed to 0.
# We move on to finding the p & q values of ARIMA(p, 1, q)

aicv = as.vector(aic)
plot(aicv,ylab="AIC values")
which(aic==min(aic)) # 16 which is (i,j)=(4,4) where 3610.863 occurs.
# Therefore ARIMA(3,1,3) is suitable due to lowest AIC as shown above.
```



```
final model = arima(chart, order = c(3,1,3), method = "ML")
# Call:
   arima(x = chart, order = c(3, 1, 3), method = "ML")
#
# Coefficients:
            ar2
                     ar3
                            ma1
                                     ma2
# -1.3082 -1.1759 -0.1525 0.2440 -0.0409 -0.9053
# s.e. 0.0420
                0.0486 0.0420 0.0194
                                         0.0212
                                                   0.0173
\# sigma^2 estimated as 15.05: log likelihood = -1799.34, aic = 3610.69
par(mfrow=c(2,2))
plot(resid(final model), ylab='Residuals',type='o',main="Residual Plot")
abline(h=0)
acf(resid(final model), main="ACF: Residuals")
```

```
hist(resid(final_model),xlab='Residuals',main='Histogram: Residuals')
qqnorm(resid(final_model),ylab="Sample Q",xlab="Theoretical Q")
qqline(resid(final_model))
```



- # Residual plot does not seem to have a pattern.
- $\sharp$  Variance is constant. ACF is not too significant beyond the confidence bands for the most part.
- $\mbox{\# QQ}$  Norm seems to have a tail on both left and right, but for the most part is symmetric.

```
Box.test(final_model$resid, lag = (3+3+1), type = "Box-Pierce", fitdf = (3+3))
Box.test(final_model$resid, lag = (3+3+1), type = "Ljung-Box", fitdf = (3+3))
# P-values are 0.2027 and 0.2007, which are high enough to accept the null hypthesis that the residuals consist of uncorrelated variables.
```

### # Question 1c

```
initial_model = arima(chart, order = c(2,1,4),method = "ML")

# Call:
# arima(x = chart, order = c(2, 1, 4), method = "ML")

# Coefficients:
# ar1 ar2 ma1 ma2 ma3 ma4
# -1.1589 -0.9993 0.0912 -0.0963 -0.8896 0.1598
```

```
0.0011 0.0397
# s.e.
         0.0017
                                    0.0186
                                            0.0187 0.0387
\# sigma^2 estimated as 15.05: log likelihood = -1799.39, aic = 3612.78
# ARIMA n full form: y t where [x t = y t - y (t-1)] &
\# \times t + 1.1589 \times (t-1) + 0.9993 \times (t-2) = 0.0912 \times (t-1) - 0.0963 \times (t-2) -
0.8896*w (t-3) + 0.1598*w (t-4)
time.pts = c(1:length(chart))
n = length(chart)
nfit = n-6
model = arima(chart[1:nfit], order = c(2,1,4), method = "ML")
outpred = predict(model, n.ahead=6)
ubound = outpred$pred+1.96*outpred$se
lbound = outpred$pred-1.96*outpred$se
ymin = min(lbound)
ymax = max(ubound)
plot(time.pts[(n-25):n],chart[(n-25):n],type="1", ylim=c(ymin,ymax),
xlab="Month #", ylab="# of songs in top 40")
points(time.pts[(nfit+1):n],outpred$pred, col="red")
lines(time.pts[(nfit+1):n],ubound,lty=3,lwd= 2, col="blue")
lines(time.pts[(nfit+1):n],lbound,lty=3,lwd= 2, col="blue")
songs in top 40
    8
    ß
                                                00
    9
₽
    LO.
           625
                   630
                            635
                                    640
                                            645
                           Month #
# > outpred$se
# Time Series:
    Start = 643
# End = 648
# Frequency = 1
# [1] 3.826478 3.844869 3.855123 3.870862 3.885505 3.895783
for (i in 1:6){
  cat("The ", i, "th prediction's confidence interval is (" , lbound[i], ","
, ubound[i], ") \n")
}
\# The 1 th prediction's confidence interval is ( 3.896225 , 18.89602 )
\# The 2 th prediction's confidence interval is ( 6.301965 , 21.37385 )
\# The 3 th prediction's confidence interval is ( 6.926165 , 22.03825 )
# The 4 th prediction's confidence interval is ( 4.752241 , 19.92602 )
# The 5 th prediction's confidence interval is (6.563145, 21.79432)
# The 6 th prediction's confidence interval is (6.552773, 21.82424)
```

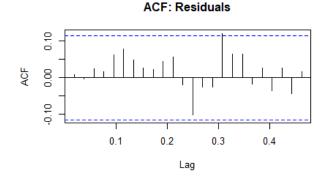
```
## Compute Accuracy Measures
obsprice = chart[(nfit+1):n]
predprice = outpred$pred
### Mean Squared Prediction Error (MSPE)
mean((predprice-obsprice)^2)
# 51.93373
### Mean Absolute Prediction Error (MAE)
mean(abs(predprice-obsprice))
# 6.165023
### Mean Absolute Percentage Error (MAPE)
mean (abs (predprice-obsprice) / obsprice)
# 0.4927662
### Precision Measure (PM)
sum((predprice-obsprice)^2)/sum((obsprice-mean(obsprice))^2)
# 1.217197
# The 4 accuracy measures: MSPE & MAE & MAPE & PM were not that low, so we
can say that the prediction was not the best using this model.
```

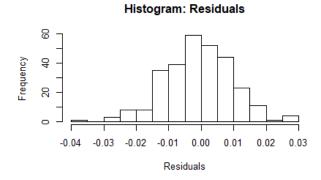
```
# Question 2a
#Libraries and Data
library (TSA)
library (mgcv)
#USD to EU
fname <- file.choose("C://Users//nicholas.park//Downloads//USD to EU")</pre>
data1 <- read.csv(fname)</pre>
data1 <- data1[,2]</pre>
EU = ts(data1, start=c(2014), freq=52)
#USD to GBP
fname2 <- file.choose("C://Users//nicholas.park//Downloads//USD to GBP")</pre>
data2 <- read.csv(fname2)</pre>
data2 <- data2[,2]</pre>
GBP = ts(data2, start=c(2014), freq=52)
# Ouestion 2a
test modelA <- function(p,d,q){
  mod = arima(EU, order=c(p,d,q), method="ML")
  current.aic = AIC(mod)
 df = data.frame(p,d,q,current.aic)
  names(df) <- c("p","d","q","AIC")</pre>
  print(paste(p,d,q,current.aic,sep=" "))
  return (df)
}
orders = data.frame(Inf, Inf, Inf, Inf)
names(orders) <- c("p", "d", "q", "AIC")</pre>
for (p in 0:3) {
  for (d in 0:2){
    for (q in 0:3) {
      possibleError <- tryCatch(</pre>
        orders<-rbind(orders, test modelA(p,d,q)),
        error=function(e) e
      if(inherits(possibleError, "error")) next
    }
  }
# [1] "0 0 0 -569.420853736522"
# [1] "0 0 1 -917.969448175781"
# [1] "0 0 2 -1172.65879202353"
# [1] "0 0 3 -1355.38383000796"
# [1] "0 1 0 -1784.85289049491"
```

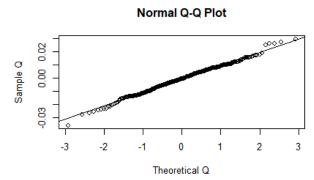
```
# [1] "0 1 1 -1788.70792245714"
# [1] "0 1 2 -1786.81923283967"
# [1] "0 1 3 -1786.27635631268"
# [1] "0 2 0 -1622.22817773909"
# [1] "0 2 1 -1774.06048056079"
# [1] "0 2 2 -1776.50757790129"
# [1] "0 2 3 -1774.92343233889"
# [1] "1 0 0 -1782.87765210213"
# [1] "1 0 1 -1786.92093245319"
# [1] "1 0 3 -1787.57530512863"
# [1] "1 1 0 -1788.22872828377"
# [1] "1 1 1 -1786.76302109289"
# [1] "1 1 2 -1784.94794213084"
# [1] "1 1 3 -1784.58680033341"
# [1] "1 2 0 -1669.12059630621"
# [1] "1 2 1 -1775.9706223139"
# [1] "1 2 2 -1774.67378352835"
# [1] "1 2 3 -1773.2960304864"
# [1] "2 0 0 -1786.46590069187"
# [1] "2 0 1 -1784.96352201317"
# [1] "2 0 3 -1778.6280106619"
# [1] "2 1 0 -1787.24390722201"
# [1] "2 1 1 -1785.32555803117"
# [1] "2 1 2 -1798.00812015844"
# [1] "2 1 3 -1796.10009662035"
# [1] "2 2 0 -1694.18164066097"
# [1] "2 2 1 -1775.66283339909"
# [1] "2 2 2 -1773.98221697775"
# [1] "2 2 3 -1786.40920002392"
# [1] "3 0 0 -1572.09389517767"
# [1]
     "3 0 1 -1782.57505604075"
# [1] "3 0 2 -1787.57246385201"
# [1] "3 0 3 -1783.9558167744"
# [1] "3 1 0 -1785.65739135954"
# [1] "3 1 1 -1784.37651976233"
# [1] "3 1 2 -1796.0987647007"
# [1] "3 1 3 -1796.76079749679"
# [1] "3 2 0 -1727.84285678383"
# [1] "3 2 1 -1774.92023437415"
# [1] "3 2 2 -1773.5370075969"
# [1] "3 2 3 -1784.41245565212"
orders <- orders[order(-orders$AIC),]
tail (orders)
#pdq
              AIC
# 17 1 1 0 -1788.229
# 7 0 1 1 -1788.708
# 42 3 1 2 -1796.099
# 31 2 1 3 -1796.100
# 43 3 1 3 -1796.761
# 30 2 1 2 -1798.008
# We can observe that the minimum AIC occurs in (p,d,q) = (2,1,2)
EU final model = arima(EU, order = c(2,1,2), method = "ML")
```

```
# Call:
    arima(x = EU, order = c(2, 1, 2), method = "ML")
#
#
 Coefficients:
             ar2
                     ma1
 -0.4504 -0.8491 0.5958 0.9286
        0.0798
                  0.0582 0.0614 0.0573
\# sigma^2 estimated as 0.0001072: log likelihood = 904, aic = -1800.01
par(mfrow=c(2,2))
plot(resid(EU final model), ylab='Residuals',type='o',main="Residual Plot")
abline(h=0)
acf(resid(EU final model), main="ACF: Residuals")
hist(resid(EU final_model), xlab='Residuals', main='Histogram: Residuals')
qqnorm(resid(EU final model), ylab="Sample Q", xlab="Theoretical Q")
qqline(resid(EU final model))
```

# Residual Plot 2014 2015 2016 2017 2018 2019 Time





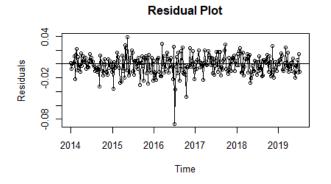


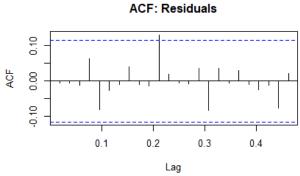
- # Residual plot does not seem to have a pattern. Heteroscedasticity is minor.
- # Variance is somewhat constant. ACF is not significant beyond the confidence bands.
- # Residual plot appears normal from histogram.
- # QQ Norm seems to indicate a symmetric normal distribution.
- # Same procedure with GBP.

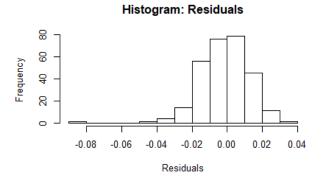
test\_modelB <- function(p,d,q){</pre>

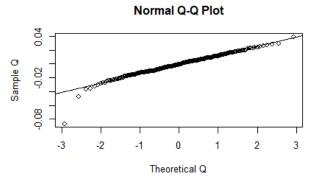
```
mod = arima(GBP, order=c(p,d,q), method="ML")
  current.aic = AIC(mod)
  df = data.frame(p,d,q,current.aic)
  names(df) <- c("p","d","q","AIC")</pre>
  print(paste(p,d,q,current.aic,sep=" "))
  return (df)
orders = data.frame(Inf, Inf, Inf, Inf)
names(orders) <- c("p", "d", "q", "AIC")</pre>
for (p in 0:3) {
  for (d in 0:2){
    for (q in 0:3) {
      possibleError <- tryCatch(</pre>
        orders<-rbind(orders,test_modelB(p,d,q)),</pre>
        error=function(e) e
      )
      if(inherits(possibleError, "error")) next
    }
  }
}
# [1] "0 0 0 -294.663964544714"
# [1] "0 0 1 -658.997635196966"
# [1] "0 0 2 -939.93637211017"
# [1] "0 0 3 -1111.94375778033"
# [1] "0 1 0 -1619.85724246706"
# [1] "0 1 1 -1631.80095196078"
# [1] "0 1 2 -1631.01849308131"
# [1] "0 1 3 -1630.53287452623"
# [1] "0 2 0 -1476.94141415387"
# [1] "0 2 1 -1608.08134363572"
# [1] "0 2 2 -1619.63606508976"
# [1] "0 2 3 -1618.99324218288"
# [1] "1 0 0 -1616.91059168528"
# [1] "1 0 1 -1629.12875099032"
# [1] "1 0 2 -1628.21555964443"
# [1] "1 0 3 -1625.63262130587"
# [1] "1 1 0 -1628.84298539492"
# [1] "1 1 1 -1630.53419482776"
# [1] "1 1 2 -1629.50656416397"
# [1] "1 1 3 -1630.3381509502"
# [1] "1 2 0 -1506.16260565096"
# [1] "1 2 1 -1616.6637017359"
# [1] "1 2 2 -1618.4320937032"
# [1] "1 2 3 -1617.6129365503"
# [1] "2 0 0 -1626.18347599749"
# [1] "2 0 1 -1627.79422165298"
# [1] "2 0 2 -1630.63340960347"
# [1] "2 0 3 -1627.18007540341"
# [1] "2 1 0 -1632.05834292573"
# [1] "2 1 1 -1630.10642864094"
# [1] "2 1 2 -1629.9632069891"
# [1] "2 1 3 -1628.65142261479"
# [1] "2 2 0 -1542.94155479698"
```

```
# [1] "2 2 1 -1620.15552253138"
# [1] "2 2 2 -1618.18511908651"
# [1] "2 2 3 -1617.95135948199"
# [1] "3 0 0 -1512.31858802444"
# [1] "3 0 1 -1627.07251843573"
# [1] "3 0 3 -1605.79059387389"
# [1] "3 1 0 -1630.14796794164"
# [1] "3 1 1 -1628.22885253255"
# [1] "3 1 2 -1629.68453359146"
# [1] "3 1 3 -1628.14650991635"
# [1] "3 2 0 -1565.58861769817"
# [1] "3 2 1 -1618.20782422363"
# [1] "3 2 2 -1616.78165254184"
# [1] "3 2 3 -1616.67399964506"
orders <- orders[order(-orders$AIC),]</pre>
tail(orders)
#pdq
             AIC
# 9 0 1 3 -1630.533
# 19 1 1 1 -1630.534
# 28 2 0 2 -1630.633
# 8 0 1 2 -1631.018
# 7 0 1 1 -1631.801
# 30 2 1 0 -1632.058
# We can observe that the minimum AIC occurs in (p,d,q) = (2,1,0)
GBP final model = arima(GBP, order = c(2,1,0), method = "ML")
par(mfrow=c(2,2))
plot(resid(GBP final model), ylab='Residuals',type='o',main="Residual Plot")
abline (h=0)
acf(resid(GBP final model), main="ACF: Residuals")
hist(resid(GBP final model), xlab='Residuals', main='Histogram: Residuals')
qqnorm(resid(GBP final model),ylab="Sample Q",xlab="Theoretical Q")
qqline(resid(GBP final model))
```







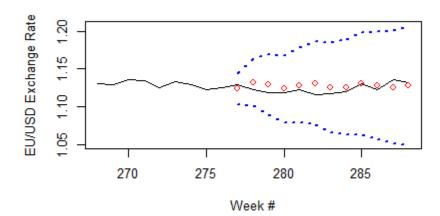


- # Residual plot does not seem to have a pattern.
- # Variance is mostly constant. ACF is not significant beyond the confidence bands.
- # Residual plot appears approximately normal from histogram.
- # QQ Norm seems to indicate a symmetric normal distribution, except for a slight left tail.

### # Question 2b

```
EU final model
# Call:
    arima(x = EU, order = c(2, 1, 2), method = "ML")
 Coefficients:
    ar1
            ar2
                    ma1
                            ma2
 -0.4504 -0.8491 0.5958 0.9286
                 0.0582 0.0614 0.0573
       0.0798
\# sigma^2 estimated as 0.0001072: log likelihood = 904, aic = -1800.01
GBP final model
# Call:
    arima(x = GBP, order = c(2, 1, 0), method = "ML")
# Coefficients:
```

```
ar1
            ar2
# 0.2199 -0.1340
# s.e. 0.0585
                0.0584
\# sigma^2 estimated as 0.0001944: log likelihood = 819.03, aic = -1634.06
eu.time.pts = c(1:length(EU))
n = length(EU)
nfit = n-12
EU model = arima(EU[1:nfit], order = c(2,1,2), method = "ML")
EU outpred = predict(EU model, n.ahead=12)
eu ubound = EU outpred$pred+1.96*EU outpred$se
eu lbound = EU outpred$pred-1.96*EU outpred$se
ymin = min(eu lbound)
ymax = max(eu ubound)
plot(eu.time.\overline{pts}[(n-20):n], EU[(n-20):n], type="1", ylim=c(ymin,ymax),
xlab="Week #", ylab= "EU/USD Exchange Rate")
points(eu.time.pts[(nfit+1):n],EU outpred$pred,col="red")
lines(eu.time.pts[(nfit+1):n],eu ubound,lty=3,lwd= 2, col="blue")
lines (eu.time.pts[(nfit+1):n],eu lbound,lty=3,lwd= 2, col="blue")
```



```
for (i in 1:12) {
  cat("The ", i, "th prediction's confidence interval is (" , eu lbound[i],
"," , eu ubound[i], ") \n")
}
\# The 1 th prediction's confidence interval is (1.103374, 1.144223)
\# The 2 th prediction's confidence interval is ( 1.100951 , 1.163255 )
\# The 3 th prediction's confidence interval is ( 1.08962 , 1.168672 )
      4 th prediction's confidence interval is ( 1.078527 , 1.168353 )
# The
# The
      5 th prediction's confidence interval is ( 1.07854 , 1.178575 )
\# The 6 th prediction's confidence interval is ( 1.075484 , 1.186656 )
\# The 7 th prediction's confidence interval is ( 1.065552 , 1.185596 )
# The 8 th prediction's confidence interval is ( 1.062199 , 1.189697 )
# The 9 th prediction's confidence interval is (1.062548, 1.198346)
\# The 10 th prediction's confidence interval is ( 1.056225 , 1.199927 )
# The 11 th prediction's confidence interval is ( 1.050241 , 1.20043 )
# The 12 th prediction's confidence interval is ( 1.050213 , 1.20699 )
```

```
## Compute Accuracy Measures
EU obsprice = EU[(nfit+1):n]
EU predprice = EU outpred$pred
### Mean Squared Prediction Error (MSPE)
mean((EU predprice-EU obsprice)^2)
# 6.395014e-05
### Mean Absolute Prediction Error (MAE)
mean (abs (EU predprice-EU obsprice))
# 0.007161764
### Mean Absolute Percentage Error (MAPE)
mean(abs(EU predprice-EU obsprice)/EU obsprice)
# 0.006380015
### Precision Measure (PM)
sum((EU predprice-EU obsprice)^2)/sum((EU obsprice-mean(EU obsprice))^2)
# 1.616\overline{3}09
# The 4 accuracy measures: MSPE & MAE & MAPE & PM were very low. Prediction
was quite accurate.
# Same procedure for GBP
gbp.time.pts = c(1:length(GBP))
n = length (GBP)
nfit = n-12
GBP model = arima(GBP[1:nfit], order = c(2,1,0), method = "ML")
GBP outpred = predict (GBP model, n.ahead=12)
gbp ubound = GBP outpred$pred+1.96*GBP outpred$se
gbp lbound = GBP outpred$pred-1.96*GBP outpred$se
ymin = min(gbp lbound)
ymax = max(gbp ubound)
plot(gbp.time.\overline{pts}[(n-20):n],GBP[(n-20):n],type="l", ylim=c(ymin,ymax),
xlab="Week #", ylab= "GBP/USD Exchange Rate")
points(gbp.time.pts[(nfit+1):n],GBP outpred$pred,col="red")
lines(gbp.time.pts[(nfit+1):n],gbp ubound,lty=3,lwd= 2, col="blue")
lines(gbp.time.pts[(nfit+1):n],gbp lbound,lty=3,lwd= 2, col="blue")
  3BP/USD Exchange Rate
      8
      8
      8
```

270

275

280

Week#

285

```
for (i in 1:12) {
 cat("The ", i, "th prediction's confidence interval is (" , gbp lbound[i],
"," , gbp ubound[i], ") \n")
# The 1 th prediction's confidence interval is ( 1.278195 , 1.333416 )
# The 2 th prediction's confidence interval is (1.262772, 1.350049)
# The 3 th prediction's confidence interval is ( 1.252742 , 1.360283 )
# The 4 th prediction's confidence interval is ( 1.244859 , 1.368047 )
\# The 5 th prediction's confidence interval is (1.237889, 1.374962)
\# The 6 th prediction's confidence interval is (1.231512, 1.381343)
\# The 7 th prediction's confidence interval is ( 1.225625 , 1.387239 )
\# The 8 th prediction's confidence interval is (1.220141, 1.392724)
# The 9 th prediction's confidence interval is (1.214988, 1.397876)
\# The 10 th prediction's confidence interval is ( 1.21011 , 1.402754 )
\# The 11 th prediction's confidence interval is ( 1.205467 , 1.407396 )
# The 12 th prediction's confidence interval is ( 1.201029 , 1.411835 )
## Compute Accuracy Measures
GBP obsprice = GBP[(nfit+1):n]
GBP predprice = GBP outpred$pred
### Mean Squared Prediction Error (MSPE)
mean ((GBP predprice-GBP obsprice)^2)
# 0.0009270949
### Mean Absolute Prediction Error (MAE)
mean(abs(GBP predprice-GBP obsprice))
# 0.02527872
### Mean Absolute Percentage Error (MAPE)
mean(abs(GBP predprice-GBP obsprice)/GBP obsprice)
# 0.01990941
### Precision Measure (PM)
sum((GBP predprice-GBP obsprice)^2)/sum((GBP obsprice-mean(GBP obsprice))^2)
# 3.133436
# The 4 accuracy measures: MSPE & MAE & MAPE & PM are quite low, but not as
low as prediction done by EU model.
# Prediction was still quite accurate but less accurate toward later points.
```

- # We can see that ARIMA modeling sometimes performs well with AIC model selection criteria, sometimes not so much. Overall the effectiveness of it is better than just differencing or using AR or MA models.
- # For ARIMA to work well, there must be a time series that contains a significant number of datapoints to forecast values using its data's inertia. It would be appropriate to use for contexts with a lot of datapoints such as a 20+ year old stock's price. I would recommend against it for newer stock's with less datapoints.
- # It appears models with the most amount of coefficients do not necessarily perform better (i.e., models with p & Q values close to the max value). Perhaps it has to do with the possibility of overfitting the data that can turn out to predict rather poorly on test data. The EU model which uses ARIMA(2,1,2) performs slightly better than GBP model which uses ARIMA(2,1,0), by taking into account MA coefficients which happened due to model selection via low resulting AIC. This makes me think the context and structure of the data is also important in determining the criteria for model selection + interpretation. Currently, the model selection simply relies on searching for low values of AIC on different combinations of (p,d,q), but there are better models that can better select the model to increase performance even more. I could've further taken into consideration seasonal ARIMA instead of normal ARIMA. I also want to note that the residual analysis after ARIMA was not always a perfect white noise, due to non-ideal aspects of real-world data.