Supplementary Methods

Attention check

Attention checks were added to phase one (i.e., the ratings phase) of Pilot full and the Study 1 full and Study 1 ratings conditions, to compensate for the unsupervised nature of online data collection. Every attention check showed a cross, a 'next 'button, and the text "press 'next 'when the cross disappears". The cross disappeared at a random time interval between one and five seconds. The 'next 'button was active the whole time. If participants were paying attention, they would not press the 'next 'button as soon as it appeared but would instead read the text and respond only after the cross had disappeared. Thus, if participants' response time exceeded the cross-display time, they passed the attention check. Nevertheless, we found high correlations between phase 1 and phase 2 ratings (i.e., > .8) across our studies (see average correlations between phase1 ratings reported in Methods to pilot studies and Study 1) and so we elected to not remove participants based on attention check data in Pilot full, Study 1 full and Study 1 ratings and we discontinued the use of attention check trials in Studies 2 and 3.

Detail on backwards induction

The action value associated with continued sampling (and, therefore, effectively the decision threshold for each sequence position) was computed based on a backwards induction algorithm. Here, we give an intuitive description of this algorithm, using the simplest incentivisation scheme, where the reward value of each option equals its option value and there is zero cost to sample. Backwards induction begins by considering the final option. When the agent encounters the final option, the decision threshold T_1 must be zero, as the paradigm forces choice of this option and there are no future option values to consider. Before that, when the agent encounters the second to last option, the computation of the threshold T_2 for this option remains relatively simple: the agent must compare the value of the second to last option against the expected value of the final option (i.e., the mean of generating distribution of option values). For the third to last option, its decision threshold T_3 depends on the expected value of continuing to the second to last option, with the opportunity to also continue to the final option. Computation of T_3 therefore involves integrating the density function of the generating distribution of option values to compute the expected value of options above the decision threshold T_2 and then adding the expected reward for the final position. This process can continue backwards to obtain the expected value of getting any option in the future that is better than the current one, wherever in the sequence the current one is.

Supplementary Results

Parameter recovery

To ascertain the ability of our models to derive the correct parameters from individual participant data, we performed parameter recovery analyses, in which we simulated model choices (take option or sample again) in response to randomly generated option values. We simulated models with 12 possible levels of each parameter, within which there were 25 simulated participants with five sequences per participant and 12 options per sequence. To parallel the structure of our empirical paradigms, we created a generating distribution (separately for each simulated participant) of 90 option values, randomly produced from a Gaussian distribution with mean 40 and standard deviation of 20 and within the range of 1 to 100 (recall that we normalised all our prices to this same range when fitting models to human participants). Then, we populated the sequences of input option values for the optimal stopping task from this participant-specific generating distribution. We configured our models with ranges of the key theoretical parameters (Supplementary Figure 4, x axis) that produced sampling rates between roughly two and ten samples to decision (Supplementary Figure 5). The aforementioned randomly generated option values were then presented to every configured model to extract simulated sampling rates associated with each configured parameter value. We then fitted the models to these simulated take option / sample again decisions in the same way as we fitted human participants to obtain parameter estimates of the configured parameters. For the three models we cover in the main text, Cost to Sample, Cut Off and Biased Prior models, configured and estimated parameters tended to correlate (Supplementary Figure 4 and lower panel of Supplementary Figure 5). Also, the sampling rates simulated using configured parameters highly correlated with sampling rates simulated using the estimated parameters (middle panel of Supplementary Figure 5 and Supplementary Figure 6).

We formulated two more theoretical models, but these did not perform so well during parameter recovery and so were excluded from model fitting and comparison. In the *Biased Values*

model, we considered the possibility that, although participants may use the optimal solution to solve the task, they might instead be biased to misperceive the magnitudes of the option values that are input into this optimal solution. This might especially be the case if participants perceive only the very most valued options as worthy of consideration at all, as might be the case in "high threshold" models of optimal stopping in mate choice¹. We incorporated a threshold into the Biased Values model: Option values above this option value threshold are transformed to 100 (the maximal option value) and option values below this threshold are transformed to 1 (the minimal option value). The transformed option values, once rendered attractive only when above threshold, are then submitted to the Ideal Observer. We picked the centre of the input value range (i.e., 50) as the starting value for the threshold parameter when fitting to participants' choices and bounded the parameter fitting to be within the option value range 1 to 100.

The *Biased Rewards model* is based on similar logic as the Biased Values model. However, instead of assuming participants place a threshold on the option values being input to the model, we instead assumed such a threshold on the reward function *R* (See Equation 3 in Methods). Recall that this function assigns reward payoffs the relative ranks of choices. As with the Biased Values model, we adopted a threshold, which was the free parameter, above which choices received a maximal reward of 100 and below which choices received a reward of 1. As in the Biased Values model, the starting value of this threshold parameter was initialised at 50 and fitting was bounded between 1 and 100.

Model Recovery

To confirm the validity of our model fitting procedures, we employed out-of-sample testing of model performance. We simulated 1000 participants' data using the Cost to Sample, Cut Off and Biased Prior models. Each simulated participants' models processed five sequences of options created in the same way as for the parameter recovery. Parameters (the first theoretical parameter)

and beta (the Softmax inverse temperature) were sampled from uniform distributions using the same ranges we used to simulate model behaviour in the parameter recovery. Then, the same three models were fit to all simulated participants' models. The results² (Supplementary Figure 7) show successful overall recovery, with the simulated models largely better (smaller) BIC than other fitted models, despite a small degree of confusability between Cost to Sample and Biased Prior models.

Supplementary Discussion

Divergent model predictions

As mentioned in the Discussion in the main text, all the models we tested excelled at predicting participants' mean sampling rates. Our supplementary analyses of the large sample in Study 2 further shows that individual differences in participant sampling rates were highly correlated with sampling rates predicted by all three models (Supplementary Figure 12). And, participants' sequence specific thresholds were approximated by all three models (Supplementary Figure 13). The models did, however, diverge in their predictions of participants' mean rank of chosen options (Supplementary Figures 8-11, 14). Here, the Cut Off heuristic was unable to obtain similar levels as participants, while Biased Prior and Cost to Sample models could. Given the high predictivity of all our models, it can be difficult to discern exactly what choices the Biased Prior is superior at predicting, compared to other models. Our recommendation is that these models be compared on paradigms specifically designed to test this hypothesis. For example, manipulations of participants' expectations about upcoming option values (i.e., their prior) should produce the types of systematically different decisions that would be predictable from a Biased Prior model.

Supplementary References

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Supplementary Tables

Supplementary Table 1Details of pairwise tests comparing human participant and Ideal Observer sampling

Ideal Observer	n	Bayes r = .5	Factor 10 r = .71	(BF10) r = 1	t test	Cohen's		dence rval	Lilliefors	
0.0001701		5	, ./ ±		Pilot baseline	•	11100			
Objective values	47	2.98E18	4.04E18	5.30E18	t(46) = -17.24, p < .001	-2.52	-3.89	-3.08	D = .12, p = .10	
				·	Pilot full				1	
Objective values	50	.41	.31	.23	t(49) = 1.23, p = .45	.17	-0.26	0.84	D = .10, p = .55	
Subjective values	50	12.86	11.37	9.31	t(48) = 3.15, p = .01	.45	0.18	1.13	D = .09, p = .86	
	Study 1 baseline									
Objective values	50	7.00E15	9.31E15	1.18E16	t(49) = -14.19, p < .001	-2.01	-3.33	-2.5	D = .06, p = .50	
					Study 1 squares	;				
Objective values	50	1.9E13	2.49E13	3.04E13	t(49) = -12.10, p < .001	-1.71	-3.29	-2.35	D = .07, p = .50	
					Study 1 timing					
Objective values	50	3.7E13	4.86E13	5.97E13	t(49) = -12.33, p < .001	-1.74	-3.32	-2.39	D = .09, p = .36	
					Study 1 payoff					
Objective values	51	6.41E15	8.51E15	1.07E16	t(50) = -14.01, p < .001	-1.96	-3.98	-2.99	D = .08, p = .50	
					Study 1 ratings					
Objective values	51	4.5E13	5.81E13	7.11E13	t(50) = -12.28 p < .001	-1.72	-3.75	-2.56	D = .16, p < .001	
Subjective values	51	3.9E8	4.73E8	5.25E8	t(50) = -8.60, p < .001	-1.2	-2.8	-1.62	D = .08, p = 1	
					Study 1 full					
Objective values	50	326.18	320.01	284.2	t(49) = -4.35, p < .001	61	-1.84	-0.56	D = .09, p = .72	
Subjective values	50	.48	.37	.27	t(49) = 1.37, p = .36	.19	-0.33	1.27	D = .07, p = 1	
			Ī	T	Study 2	1		Ī		
values	151	149.5	122.97	95.03	t(150) = -3.93, p < .001	32	-0.32	-0.99	D = .04, p = 1	
Subjective values	151	.15	.11	.08	t(150) = .62, p = 1	.05	0.05	-0.26	D = .06, p = .51	
			1	S	tudy 3 10 option	ns		1	.	
Objective values	75	1.81	1.43	1.08	t(74) = 2.28, p = .05	.26	0	0.98	D = .08, p = .42	
Subjective values	75	6.99	5.78	4.49	t(74) = 2.89, p = .01	.33	0.16	1.36	D = .06, p = 1	
				S	tudy 3 14 option	ns				
Objective values	65	5056.4	5042.47	4529.57	t(64) = -5.10, p < .001	63	-2.3	-0.87	D = .05, p = 1	
Subjective values	65	38.66	34.25	27.99	t(64) = -3.55, p < .001	44	-1.57	-0.34	D = .08, p = .78	

Note: All p and confidence interval alpha values are Bonferroni corrected for numbers of tests per study. *r* values columns represent the Cauchy scale factor of the prior on the effect size.

Supplementary Table 2

Details of tests between pairs of mean BIC values in pilot studies

Model 1	Model 2	n	Baye	Cohen's d					
iviodel 1	iviodei z	TI .	r = .5	r = .71	r = 1	Conen's a			
		Pilot ba	seline						
Biased Prior obj	Cut Off obj	47	3.49E6	4.10E4	4.37E6	-1.66			
Biased Prior obj	Cost to Sample obj	47	0.26	0.19	0.14	32			
Cut Off obj	Cost to Sample obj	47	4.91E8	6.07E8	6.88E8	1.42			
Pilot full									
Biased Prior obj	Biased Prior subj	50	0.21	0.16	0.11	0.02			
Biased Prior obj	Cut Off obj	50	0.48	0.36	0.27	-0.19			
Biased Prior obj	Cut Off subj	50	1.86	1.52	1.17	-0.32			
Biased Prior obj	Cost to Sample obj	50	2.4	1.99	1.55	-0.34			
Biased Prior obj	Cost to Sample subj	50	0.22	0.16	0.11	0.03			
Biased Prior subj	Cut Off obj	50	0.34	0.25	0.19	-0.15			
Biased Prior subj	Cut Off subj	50	0.5	0.38	0.29	-0.2			
Biased Prior subj	Cost to Sample obj	50	0.24	0.18	0.13	-0.08			
Biased Prior subj	Cost to Sample subj	50	0.21	0.15	0.11	-0.01			
Cut Off obj	Cut Off subj	50	0.52	0.4	0.29	-0.2			
Cut Off obj	Cost to Sample obj	50	0.24	0.18	0.13	0.08			
Cut Off obj	Cost to Sample subj	50	0.51	0.39	0.29	0.2			
Cut Off subj	Cost to Sample obj	50	0.57	0.44	0.33	0.21			
Cut Off subj	Cost to Sample subj	50	2.69	2.23	1.75	0.34			
Cost to Sample obj	Cost to Sample subj	50	0.93	0.73	0.55	0.26			

Note. All *p* values Bonferroni corrected for the number of tested pairs per study. *r* values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. Compare to middle row of Figure 3. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that make choices about objective values.

Supplementary Table 3.

Details of tests between pairs of mean BIC values in Study 1 models with no first phase

Model 1	Model 2	n	Baye	Cohen's d					
iviodei 1	Model 2	n	r = .5	r = .71	r = 1	Conen's a			
Study 1 baseline									
Biased Prior obj	Cut Off obj	50	5.51E9	6.86E9	7.86E9	-1.34			
Biased Prior obj	Cost to Sample obj	50	4.5	3.81	3.03	-0.38			
Cut Off obj	Cost to Sample obj	50	1.18E6	1.35E6	1.40E6	0.98			
	Study 1 squares								
Biased Prior obj	Cut Off obj	50	3.35E10	4.2210	4.91E10	-1.42			
Biased Prior obj	Cost to Sample obj	50	2085.22	2142.51	1981.89	-0.7			
Cut Off obj	Cost to Sample obj	50	5.22E4	5.72E4	5.62E4	0.84			
		Study 1	timing						
Biased Prior obj	Cut Off obj	50	9.48E9	1.19E10	1.37E10	-1.36			
Biased Prior obj	Cost to Sample obj	50	10.32	9.05	7.36	-0.43			
Cut Off obj	Cost to Sample obj	50	1.90E4	2.05E4	1.98E4	0.8			
Study 1 payoff									
Biased Prior obj	Cut Off obj	51	2.48E14	3.26E14	4.03E14	-1.8			
Biased Prior obj	Cost to Sample obj	51	6.99	6.02	4.83	-0.4			
Cut Off obj	Cost to Sample obj	51	1.36E9	1.67E9	1.88E9	1.26			

Note: All *p* values Bonferroni corrected for the number of tested pairs per study. *r* values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. Compare to middle row of Figure 4. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that make choices about objective values.

Supplementary Table 4

Details of tests between pairs of mean BIC values in Study 1 models with a first phase

	Madal 2		Baye			
Model 1	Model 2	n	r = .5	r = .71	r = 1	Cohen's d
	S	tudy 1 ra	atings	T		T
Biased Prior obj	Biased Prior subj	51	2593.14	2667.58	2469.76	-0.7
Biased Prior obj	Cut Off obj	51	1.13E11	1.43E11	1.68E11	-1.45
Biased Prior obj	Cut Off subj	51	7.54E11	9.65E11	1.15E12	-1.53
Biased Prior obj	Cost to Sample obj	51	0.97	0.77	0.58	-0.26
Biased Prior obj	Cost to Sample subj	51	5.85E5	6.63E5	6.76E5	-0.93
Biased Prior subj	Cut Off obj	51	37.68	34.52	29.01	-0.5
Biased Prior subj	Cut Off subj	51	4.16E7	4.95E7	5.35E7	-1.11
Biased Prior subj	Cost to Sample obj	51	43.92	40.45	34.11	0.51
Biased Prior subj	Cost to Sample subj	51	2.45	2.02	1.57	-0.34
Cut Off obj	Cut Off subj	51	29.43	26.74	22.33	-0.49
Cut Off obj	Cost to Sample obj	51	3.11E8	3.78E8	4.18E8	1.19
Cut Off obj	Cost to Sample subj	51	2.35	1.94	1.51	0.33
Cut Off subj	Cost to Sample obj	51	3.02E9	3.73E9	4.23E9	1.29
Cut Off subj	Cost to Sample subj	51	1.19E6	1.36E6	1.40E6	0.96
Cost to Sample obj	Cost to Sample subj	51	3554.51	3681.92	3430.71	-0.72
		Study 1	full			
Biased Prior obj	Biased Prior subj	50	0.27	0.2	0.15	0.11
Biased Prior obj	Cut Off obj	50	0.83	0.65	0.49	-0.25
Biased Prior obj	Cut Off subj	50	3.07	2.56	2.01	-0.35
Biased Prior obj	Cost to Sample obj	50	4.90E6	5.73E6	6.05E6	-1.04
Biased Prior obj	Cost to Sample subj	50	0.32	0.24	0.17	-0.13
Biased Prior subj	Cut Off obj	50	1	0.79	0.6	-0.27
Biased Prior subj	Cut Off subj	50	3.57	3	2.37	-0.36
Biased Prior subj	Cost to Sample obj	50	5.32	4.55	3.63	-0.39
Biased Prior subj	Cost to Sample subj	50	0.61	0.47	0.35	-0.22
Cut Off obj	Cut Off subj	50	1.24	1	0.76	-0.29
Cut Off obj	Cost to Sample obj	50	0.31	0.23	0.17	-0.13
Cut Off obj	Cost to Sample subj	50	0.34	0.25	0.19	0.15
Cut Off subj	Cost to Sample obj	50	0.21	0.15	0.11	0.01
Cut Off subj	Cost to Sample subj	50	1.13	0.9	0.69	0.28
Cost to Sample obj	Cost to Sample subj	50	5.07	4.32	3.45	0.39

Note. All *p* values Bonferroni corrected for the number of tested pairs per study. *r* values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. Compare to middle row of Figure 5. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that make choices about objective values.

Supplementary Table 5.Details of pairwise comparisons between participant sampling rates in Studies 1 and 3.

Condition 1	n	Condition 2	n	Bayes I	actor 10 (BF10)	t test	Cohen's	Confid		Brown-
1		2		r = .5	r = .71	r = 1		u	inte	IVdI	Forsythe
						Study	1				
Payoff	51	Timing	50	.28	.21	.15	t(99) = .07, p = 1	.01	-0.77	0.81	F(1,99) = .36, p = 1
Payoff	51	Squares	50	.28	.21	.15	t(99) = - .07, p = 1	01	-0.8	0.76	F(1,99) = .60, p = 1
Payoff	51	Baseline	50	.29	.22	.16	t(99) = .33, p = 1	.07	-0.64	0.79	F(1,99) = 4.25, p = .63
Payoff	51	Full	50	.68	.55	.42	t(99) = - 1.47, p = 1	29	-1.49	0.51	F(1,99) = 6.16, p = .22
Payoff	51	Ratings	51	.4	.31	.23	t(100) = .92, p = 1	.18	-0.56	1.06	F(1,100) = .02, p = 1
Timing	50	Squares	50	.29	.21	.16	t(98) = - .14, p = 1	03	-0.81	0.74	F(1,98) = .02, p = 1
Timing	50	Baseline	50	.29	.22	.16	t(98) = .25, p = 1	.05	-0.65	0.77	F(1,98) = 1.65, p = 1
Timing	50	Full	50	.73	.59	.46	t(98) = - 1.53, p = 1	31	-1.51	0.49	F(1,98) = 8.18, p = .08
Timing	50	Ratings	51	.38	.29	.22	t(99) = .85 p = 1	.17	-0.58	1.04	F(1,99) = .18, p = 1
Squares	50	Baseline	50	.3	.23	.17	t(98) = .42 p = 1	.08	-0.59	0.79	F(1,98) = 1.33, p = 1
Squares	50	Full	50	.65	.52	.4	t(98) = - 1.43, p = 1	29	-1.46	0.52	F(1,98) = 9.24, p = .05
Squares	50	Ratings	51	.43	.33	.25	t(99) = 1.01, p = 1	.2	-0.53	1.06	F(1,99) = .35, p = 1
Baseline	50	Full	50	1.09	.91	.71	t(98) = - 1.82, p = 1	36	-1.51	0.37	F(1,98) = 17.35, p = .00
Baseline	50	Ratings	51	.34	.26	.19	t(99) = .70, p = 1	.14	-0.56	0.9	F(1,99) = 2.99, p = 1
Full	50	Ratings	51	2	1.73	1.4	t(99) = 2.19, p = .46	.44	-0.28	1.75	F(1,99) = 6.05, p = .24
	•					Study	•			•	-
14 options	65	10 options	75	1.55	1.29	1.01	t(138) = 2.08, p = .04	.35	.03	1.33	F(1,138) = 3.55, p = .06

Note. All p values and alpha values used in confidence intervals are Bonferroni corrected for 15 tested pairs. r values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. Compare to Figure 2.

Supplementary Table 6.

Lilliefors tests of deviations from normality of sampling rates in Study 1 conditions

Condition	n	D	р
Baseline	50	.09	1
Squares	50	.08	1
Timing	50	.12	.42
Payoff	51	.1	1
Ratings	51	.11	.72
Full	50	.09	1

Note. All p values are Bonferroni corrected for tests of six conditions.

Supplementary Table 7.Details of tests between pairs of mean BIC values in Study 2

Model 1	Model 2	n	Baye	Cohen's d			
Wiodel 1	iviouei z	n	r = .5	r = .71	r = 1	concil 3 d	
Biased Prior obj	Biased Prior subj	151	.15	.11	.08	05	
Biased Prior obj	Cut Off obj	151	9.74E16	1.09E17	1.10E17	87	
Biased Prior obj	Cut Off subj	151	2.58E19	2.97E19	3.07E19	94	
Biased Prior obj	Cost to Sample obj	151	96269.85	86925.33	71670.29	45	
Biased Prior obj	Cost to Sample subj	151	2.35	1.79	1.32	2	
Biased Prior subj	Cut Off obj	151	9.57E14	1.05E15	1.03E15	8	
Biased Prior subj	Cut Off subj	151	2.18E19	2.51E19	2.59E19	94	
Biased Prior subj	Cost to Sample obj	151	4.16	3.21	2.38	22	
Biased Prior subj	Cost to Sample subj	151	7.93	6.19	4.63	24	
Cut Off obj	Cut Off subj	151	4.75	3.67	2.73	23	
Cut Off obj	Cost to Sample obj	151	5.08E10	5.22E10	4.79E10	.66	
Cut Off obj	Cost to Sample subj	151	1.19E11	1.23E11	1.14E11	.68	
Cut Off subj	Cost to Sample obj	151	6.78E13	7.32E13	7.05E13	.77	
Cut Off subj	Cost to Sample subj	151	7.29E14	7.98E14	7.79E14	.8	
Cost to Sample obj	Cost to Sample subj	151	.21	.15	.11	.08	

Note. All *p* values Bonferroni corrected for the number of tested pairs per study. *r* values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. Compare to middle row of Figure 6. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that make choices about objective values.

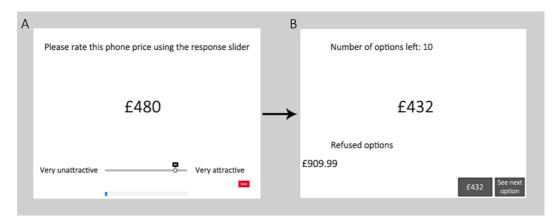
Supplementary Table 8.

Details of tests between pairs of mean BIC values in Study 3

Model 1	Model 2	n		3F10)	Cohen's d	
Wodel 1			r = .5	r = .71	r = 1	concil 5 d
		i –	options	I		ı
Biased Prior obj	Biased Prior subj	75	.31	.23	.17	13
+Biased Prior obj	Cut Off obj	75	6.05E6	6.55E6	6.35E6	79
Biased Prior obj	Cut Off subj	75	8.4E9	9.8E9	1E10	-1
Biased Prior obj	Cost to Sample obj	75	1.89	1.49	1.13	27
Biased Prior obj	Cost to Sample subj	75	.26	.19	.14	1
Biased Prior subj	Cut Off obj	75	695.77	649.11	551.07	52
Biased Prior subj	Cut Off subj	75	5.9E7	6.6E7	6.6E7	86
Biased Prior subj	Cost to Sample obj	75	.19	.14	.1	.04
Biased Prior subj	Cost to Sample subj	75	.21	.15	.11	.07
Cut Off obj	Cut Off subj	75	3.3	2.66	2.04	3
Cut Off obj	Cost to Sample obj	75	8.44E4	8.61E4	7.89E4	.67
Cut Off obj	Cost to Sample subj	75	2.7E4	2.72E4	2.45E4	.63
Cut Off subj	Cost to Sample obj	75	1.2E9	1.4E9	1.4E9	.94
Cut Off subj	Cost to Sample subj	75	4.5E11	5.4E11	5.9E11	1.11
Cost to Sample obj	Cost to Sample subj	75	.18	.13	.09	.01
	Stud	y 3 14	options	•		•
Biased Prior obj	Biased Prior subj	65	2.3	1.85	1.42	3
Biased Prior obj	Cut Off obj	65	3.8E10	4.5E10	5E10	-1.15
Biased Prior obj	Cut Off subj	65	2.1E10	2.5E10	2.7E10	-1.13
Biased Prior obj	Cost to Sample obj	65	17.62	15.24	12.24	4
Biased Prior obj	Cost to Sample subj	65	.83	.65	.48	23
Biased Prior subj	Cut Off obj	65	6436.71	6449.61	5817.81	64
Biased Prior subj	Cut Off subj	65	4.64E6	5.15E6	5.12E6	86
Biased Prior subj	Cost to Sample obj	65	.22	.16	.12	.08
Biased Prior subj	Cost to Sample subj	65	.29	.22	.16	.12
Cut Off obj	Cut Off subj	65	.62	.47	.35	2
Cut Off obj	Cost to Sample obj	65	2.4E7	2.7E7	2.7E7	.92
Cut Off obj	Cost to Sample subj	65	8.52E4	8.94E4	8.41E4	.73
Cut Off subj	Cost to Sample obj	65	9.1E7	1E8	1.1E8	.96
Cut Off subj	Cost to Sample subj	65	1.6E8	1.8E8	1.9E8	.98
Cost to Sample obj	Cost to Sample subj	65	.19	.14	.1	0

Note. All p values Bonferroni corrected for the number of tested pairs per study. r values in the Bayes Factor columns represent the scale factor of the Cauchy prior on the effect size. obj = Compare to middle row of Figure 7. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that make choices about objective values.

Supplementary Figure 1. Pilot full and Study 1 full paradigms



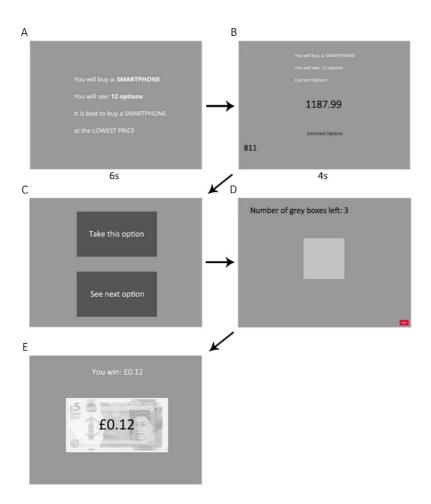
In an initial phase participants rate potential options for their subjective value (A). Then, in (B), participants view option screens as part of the optimal stopping decision task.

Supplementary Figure 2. Study 1 payoff condition



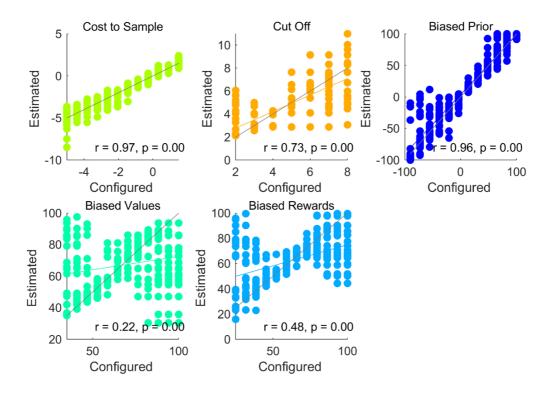
Participants receive instructions (A) before each new sequence of options screens (B), for which they choose to take each option or sample another (C). In the payoff condition, taking an option leads to feedback about the reward value of the choice, in the form of a number of stars that have been won.

Supplementary Figure 3. Study 1 squares condition



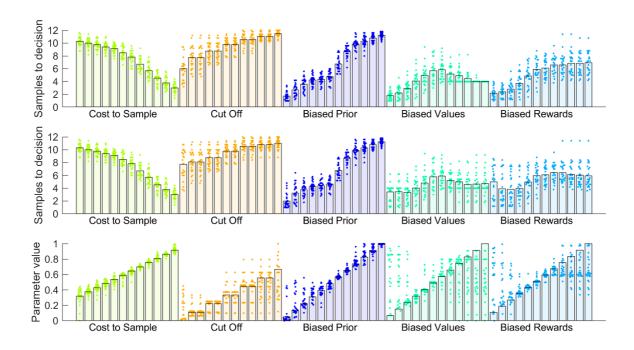
After receiving instructions (A), participants view option screens (B), followed by choices (C) to take the option or sample another. In the squares condition, when a choice is made, participants then must page through grey squares (D), which replace the remaining option screens, before they can continue to receive feedback on their choice in that sequence (E).

Supplementary Figure 4. Recovered parameters



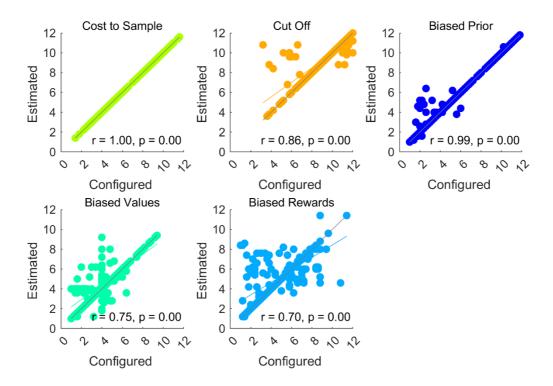
For parameter recovery analysis, we examined parameter values (vertical axes) estimated from fitting models to decisions simulated using configured parameter values (horizontal axes) for simulated participants, each shown as an individual scatter point. The grey diagonal indicates when configured and estimated parameters would be exactly equal. The coloured line indicates the regression line relating configured and estimated parameter values. Each regression is based on n = 300 observations (12 configured parameter levels \times 25 simulated participants). Biased Rewards and Biased Values models showed too poor parameter recovery to be entered into formal model comparison.

Supplementary Figure 5. Parameter recovery details



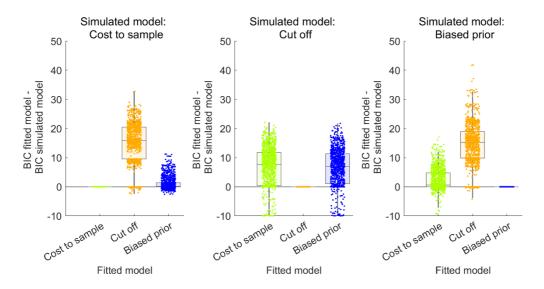
Top panel: We plot sampling rates for the 25 individual simulated participants (points) and their mean values (bars) for each of the 12 configured parameter level in the parameter recovery analysis. Systematically varying configured parameter values successfully increases or decreases simulated sampling rates for Cost to Sample and Biased Prior models and the Cut Off heuristic. Middle panel: Models were fitted to the data in the top panel and parameters estimated. We plot the sampling rates simulated using each estimated parameter (points) and their mean sampling rates (bars). Lower panel: The estimated parameters (points) are plotted relative to their target configured parameter values (bars). Each model's parameter values are normalised to a 0 to 1 range to facilitate plotting on one scale.

Supplementary Figure 6. Parameter recovery sampling rates



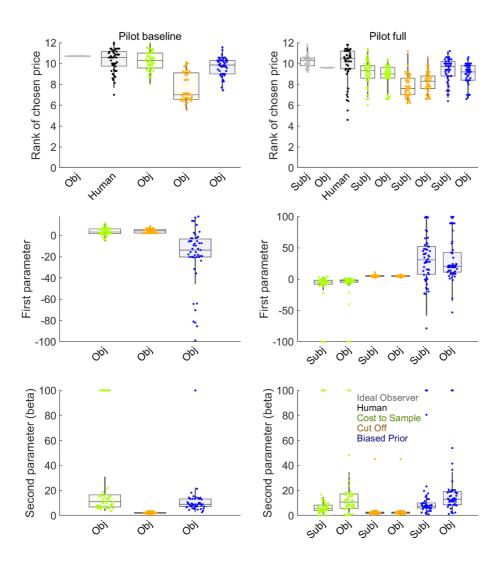
Sampling rates simulated using configured parameters (horizontal axis) are plotted against sampling rates computed from estimated parameters. The grey diagonal indicates when sampling rates based on configured and estimated parameters would be exactly equal. The coloured line indicates the regression line relating sampling rates based on configured and estimated parameter values. Each regression is based on n = 300 observations (12 configured parameter levels \times 25 simulated participants).

Supplementary Figure 7. Model recovery



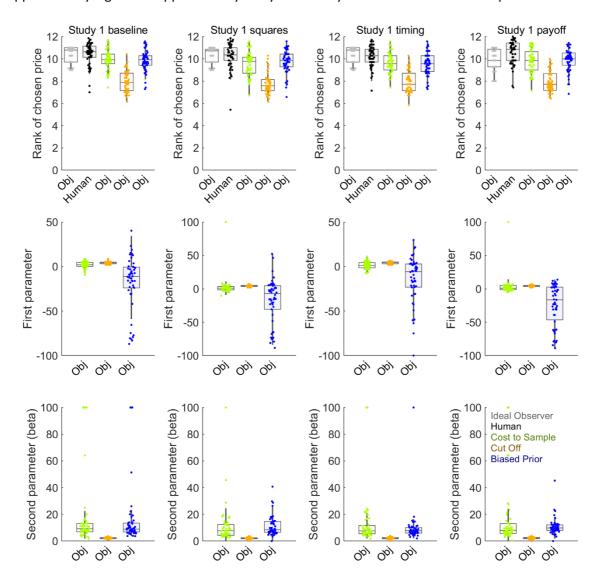
Choice data were simulated using the three models. Shown are Bayesian information criterion (BIC) values for each fitted model with the BIC value for the fitted model corresponding to the simulated model subtracted. The results show fits were largely better when the fitted and the simulated model were the same, suggesting successful model recovery.

Supplementary Figure 8. Supplementary analysis of pilot studies



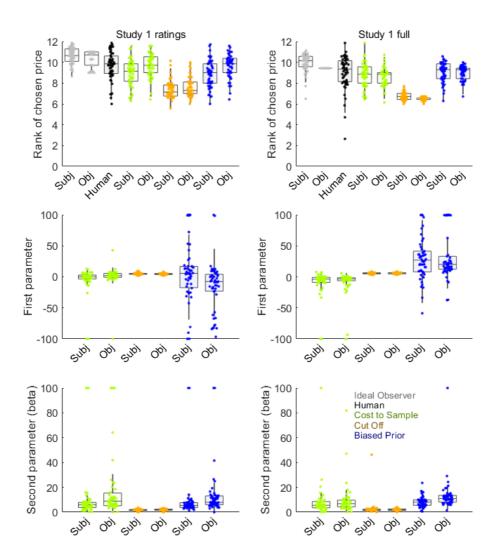
Model comparison for Pilot baseline (n = 47 participants) is in the left column and Pilot full (n = 50) is in the right column. The top row shows the ranks of chosen prices. The Cut Off model was not able to approximate participants' achieved ranks. The second row plots the "first" or key theoretical parameter values, estimated for each fitted model. For pilot baseline, where there was clear undersampling (Figure 3), the Biased Prior parameters tend negatively, suggesting a pessimistic outlook. The third row shows the "second," or inverse temperature parameter beta, estimated for each fitted model. Boxplots reflect first, second (median) and third quartiles, while whiskers reflect 1.5 interquartile range. Points reflect individual participant mean values. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.





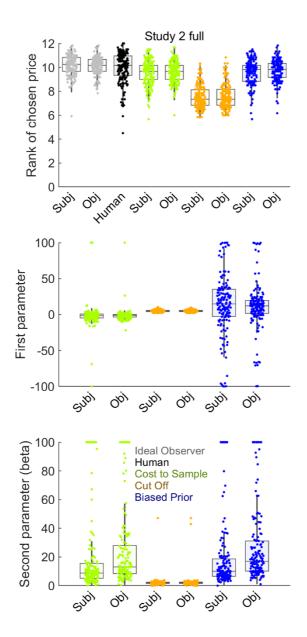
From left to right, columns show results from baseline (n = 50 participants), squares (n = 50), timing (n = 50) and payoff (n = 51) conditions. The top row shows ranks of chosen prices. The Cut Off model could never approximate participants' achieved ranks in any condition. The second row plots the "first" or theoretical parameter values, estimated for each fitted model. Every condition's Biased Prior model shows a negative (pessimistic) parameter, likely associated with undersampling in these conditions (Figure 4). The third row shows the "second," or inverse temperature parameter beta, estimated for each fitted model. Boxplots reflect first, second (median) and third quartiles, while whiskers reflect 1.5 interquartile range. Points reflect individual participant data. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.

Supplementary Figure 10. Supplementary analysis of Study 1 conditions with a first phase



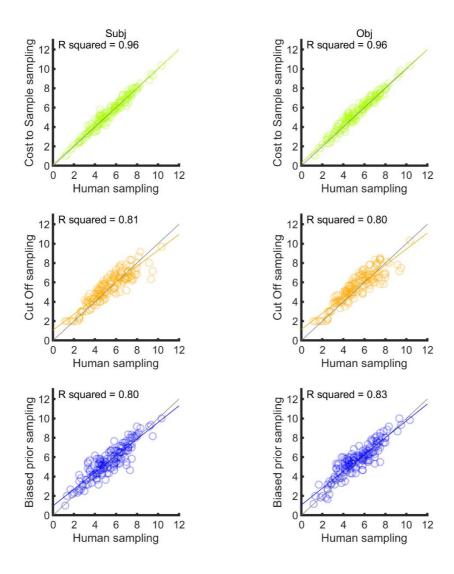
Results from the ratings condition (n = 51 participants) are shown in the left column and from the full condition (n = 50) on the right. The top row shows the ranks of chosen prices. The Cut Off model was not able to approximate participants' achieved ranks. The second row plots the "first" or theoretical parameter values, estimated for each fitted model. The condition showing the clearest undersampling, ratings (Figure 5) shows more negative (pessimistic) Biased prior parameter values. The third row shows the "second," or inverse temperature parameter beta, estimated for each fitted model. Boxplots reflect first, second (median) and third quartiles, while whiskers reflect 1.5 interquartile range. Points reflect individual participant data. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.

Supplementary Figure 11. Supplementary analysis of Study 2



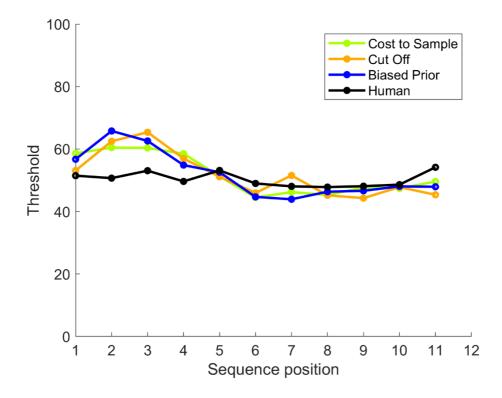
The top row shows ranks of chosen prices for n=151 participants. The top row shows the ranks of chosen prices. The Cut Off model was not able to approximate participants' achieved ranks. The second row plots the "first" or theoretical parameter values, estimated for each fitted model. The third row shows the "second," or inverse temperature parameter beta, estimated for each fitted model. Boxplots reflect first, second (median) and third quartiles, while whiskers reflect 1.5 interquartile range. Points reflect individual participant data. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.

Supplementary Figure 12. Model predictions of sampling rates in Study 2



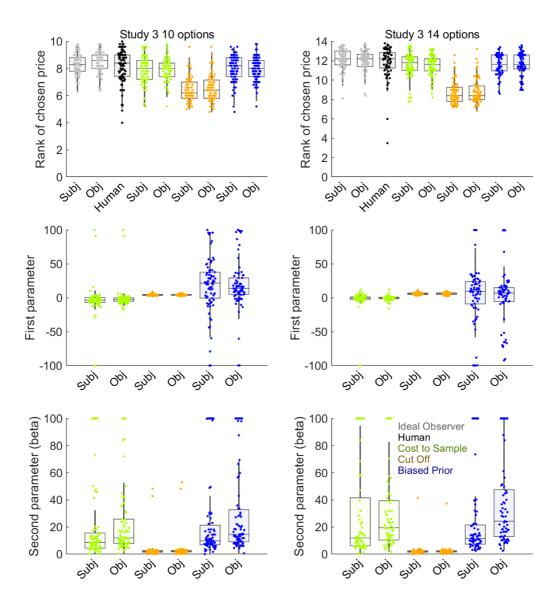
Relationships between individual differences in fitted model sampling and participant sampling for models (rows) operating on subjective values (left column) or objective values (right column). All models can predict human sampling from the n=151 participants reasonably well. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.

Supplementary Figure 13. Choice thresholds in Study 2



Decision thresholds (option values above which search is discontinued and the current option chosen) estimated for n=151 participants and corresponding fitted models and averaged for each sequence position. Outside of the first few sequence positions, all models reasonably approximate participants' linearly-changing thresholds. In this example, models processed objective values of prices. Thresholds are estimated by fitting a model resembling the independent threshold model³, which estimates each sequence position's threshold as a free parameter. Fitted data were either participants' choices or choices simulated by our models. Fitting was performed using fminearch in MATLAB. Though we recognise that choice thresholds are already explicitly computed within all our models and could be directly extracted from them, this threshold estimation procedure facilitated direct comparison with participants, by using the same procedure to estimate thresholds for participants and for models.

Supplementary Figure 14. Supplementary analysis of Study 3



The left columns shows the 10 option condition (n = 75 participants) and the right column shows the 14 option condition (n = 65). The top row shows ranks of chosen prices, where the Cut Off model underperforms, compared to participants. The second row plots the "first" or theoretical parameter values, estimated for each fitted model. The third row shows the "second," or inverse temperature parameter beta, estimated for each fitted model. Boxplots reflect first, second (median) and third quartiles, while whiskers reflect 1.5 interquartile range. Points reflect individual participant data. Abbreviations: Subj = Models that make choices about subjective values; Obj = Models that makes choices about objective values.