

# **1 Pattern-based downscaling of snowpack variability in 2 the western United States**

**3 Nicolas Gauthier · Kevin J. Anchukaitis ·  
4 Bethany Coulthard**

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**6 Received: date / Accepted: date**

**7 Abstract** The decline in snowpack across the western United States is one  
8 of the most pressing threats posed by climate change to regional economies  
9 and livelihoods. Earth system models are important tools for exploring past  
10 and future snowpack variability, yet their coarse spatial resolutions distort lo-  
11 cal topography and bias spatial patterns of accumulation and ablation. Here,  
12 we explore pattern-based statistical downscaling for spatially-continuous in-  
13 terannual snowpack estimates. We find that a few leading patterns capture  
14 the majority of snowpack variability across the western US in observations,  
15 reanalyses, and free-running simulations. Pattern-based downscaling methods  
16 yield accurate, high resolution maps that correct mean and variance biases  
17 in domain-wide simulated snowpack. Methods that use large-scale patterns  
18 as both predictors and predictands perform better than those that do not  
19 and all are superior to an interpolation-based “delta change” approach. These  
20 findings suggest that pattern-based methods are appropriate for downscaling  
21 interannual snowpack variability and that using physically meaningful large-  
22 scale patterns is more important than the details of any particular downscaling  
23 method.

**24 Keywords** snow water equivalent · empirical orthogonal functions · canonical  
25 correlation analysis · teleconnections · water resources

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This research was supported by the National Science Foundation Paleo Perspectives on Climate Change (P2C2) grant AGS-1803995.

N. Gauthier  
School of Geography, Development & Environment and Laboratory of Tree-Ring Research,  
University of Arizona, Tucson, AZ E-mail: ngauthier@arizona.edu

K. J. Anchukaitis  
School of Geography, Development & Environment and Laboratory of Tree-Ring Research,  
University of Arizona, Tucson, AZ

B. Coulthard  
Department of Geoscience, University of Nevada Las Vegas, Las Vegas, NV

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**26 1 Introduction**

27 The decline in snowpack across the western United States is one of the most  
28 pressing threats posed by climate change to regional economies and liveli-  
29 hoods (Mankin and Diffenbaugh, 2015; Mote et al, 2018; Xiao et al, 2018;  
30 Huning and AghaKouchak, 2020). Spring snowmelt is critical for regional wa-  
31 ter managers—more than half of annual runoff in the western US derives from  
32 snowpack (Li et al, 2017). Snow plays a central role in local and regional  
33 climates and ecosystems, from its cooling effect on temperatures to its mod-  
34 ulation of the timing and intensity of streamflow and soil moisture anomalies  
35 (Walsh et al, 1982; Marks and Dozier, 1992; Bales et al, 2006; Maurer and  
36 Bowling, 2014; Li et al, 2017). The observed decline in snowpack is the result  
37 of several interacting factors including shifts in the timing and intensity of sea-  
38 sonal precipitation and temperature patterns, each of which are exacerbated  
39 by warming temperature trends and the attendant changes in accumulation  
40 and ablation (Pierce et al, 2008; Kapnick and Hall, 2012; Pederson et al, 2013;  
41 Klos et al, 2014; Xiao et al, 2018). These snowpack deficits are of a magnitude  
42 and extent unprecedented in the observational period (McCabe and Wolock,  
43 2009; Mote et al, 2018; Schoenemann et al, 2020) and are expected to worsen  
44 in the future (Fyfe et al, 2017; Marshall et al, 2019; Siler et al, 2019).

45 Yet it remains difficult to observe snowpack uniformly across large spatial  
46 domains. Spatially-continuous high-resolution maps of snowpack are therefore  
47 a challenge to produce, particularly in areas with complex terrain (Erickson  
48 et al, 2005; Meromy et al, 2013). Different sensor types and measurement  
49 strategies focus on distinct—if related—facets of the system, such as snow  
50 water equivalent (SWE), snow-covered area, and snow depth. Each has unique  
51 uncertainties, coverage, and observational spans, making them a challenge to  
52 integrate (Dozier et al, 2016; Dong, 2018). In most locations the observational  
53 record only extends for a few decades into the past (e.g. Serreze et al, 1999),  
54 making it difficult to place observed variability in a long-term context.

55 An array of modeling approaches provides ways to estimate gaps in the  
56 observational record and produce continuous spatiotemporal data. From stan-  
57 dalone hydrological bucket models to the complex land-surface components of  
58 Earth system models, snowpack simulations attempt to capture the interact-  
59 ing drivers of snowpack variability across spatial and temporal scales. These  
60 models allow for assessments of the mechanistic uncertainty of these drivers  
61 and uncertainty in their observations (Clark et al, 2011). Even simple models  
62 provide useful information for constraining noisy observations (Broxton et al,  
63 2016a). Although the skill of current-generation snow models is high overall,  
64 issues remain in the representation of processes like ablation at near-freezing  
65 temperatures (Rutter et al, 2009; Broxton et al, 2016b; Krinner et al, 2018).  
66 Regional and global snow models must run on daily to sub-daily time scales, so  
67 a reduction in spatial resolution may be required to minimize computational  
68 costs. This tradeoff makes accurate spatial modeling of snowpack difficult, even  
69 when the underlying process models are physically appropriate.

70 Snow accumulation and ablation is sensitive to local topography, particularly  
71 in the mountainous regions that receive the most snowfall (Anderson  
72 et al, 2014; Tennant et al, 2017; Jennings and Molotch, 2019). The resolution  
73 of most simulations smooth this topography, eliminating mountain peaks and  
74 introducing temperature biases that prevent snow from accumulating where  
75 it otherwise would (Rhoades et al, 2018). The tendency for snow models to  
76 underpredict accumulated SWE has been well documented. Xu et al (2019)  
77 showed that increasing model resolution from 0.44° to 0.11° increases the accu-  
78 racy of simulated SWE by 35%. Such low-snow biases in regional and global  
79 snow simulations preclude their use by local water managers without correc-  
80 tions to this fundamental scale mismatch. Some form of downscaling is required  
81 to estimate fine-resolution snowpack maps from coarser-resolution simulation  
82 outputs (McGinnis, 1997; Pons et al, 2010; Tryhorn and Degaetano, 2013).  
83 However, this is increasingly accomplished via an additional high-resolution  
84 regional climate model or by forcing a hydrological model with atmospheric  
85 forcing data downscaled by constructed analogue methods, both of which re-  
86 quire data on hourly to daily time scales, making them computationally in-  
87 feasible for assessing variability on time horizons greater than a few decades  
88 (Rhoades et al, 2018; Chegwidden et al, 2019; Fiddes et al, 2019; Ikeda et al,  
89 2021).

90 Non-local “pattern-based” statistical downscaling methods are an effective  
91 alternative to quickly generate fine-scale, long-term ensembles from existing  
92 coarse-resolution climate model simulations. Pattern-based methods decom-  
93 pose observed and simulated climate fields into a limited number of spa-  
94 tiotemporal patterns or “modes of variability,” finding statistical relationships  
95 that translate one set of modes into the other (Bretherton et al, 1992; Tip-  
96 ppett et al, 2008; Simon et al, 2013; Maraun and Widmann, 2018). Because  
97 they find associations between internally-consistent predictor and predictand  
98 fields, pattern-based statistical methods share some benefits with more compu-  
99 tionally expensive dynamic downscaling methods that preserve the physical  
100 consistency of the simulated climate fields. These methods are “non-local” in  
101 that they focus on associations between large-scale patterns, rather than local  
102 associations between an observed location and the overlapping simulation grid  
103 cell. The simulation grid cell that best captures the observed variability at  
104 a given location is often *not* the corresponding local grid cell (van den Dool  
105 et al, 2000; Maraun and Widmann, 2015; Nicholson et al, 2019). While local  
106 mean conditions reflect local terrain, year-to-year departures from the mean  
107 often reflect teleconnections to remote, large-scale atmosphere-ocean variabil-  
108 ity (van den Dool et al, 2000; Hewitt et al, 2018). Anchoring the downscaling  
109 process in these large-scale physical mechanisms leads to a higher signal to  
110 noise ratio (Benestad et al, 2015), ensuring the estimated statistical relation-  
111 ships are internally consistent and likely to remain stable over time.

112 Here, we explore pattern-based statistical methods for downscaling inter-  
113 annual variability in March mean SWE across the western United States.  
114 We find that a few leading modes—present in observations, simulations, and  
115 reanalyses—capture the majority of snowpack variability in this domain. We

116 compare several related regression methods for finding associations between  
117 observed and simulated patterns and show that even simple linear models  
118 perform well under cross validation. These methods yield accurate high reso-  
119 lution maps that correct mean and variance biases in domain-wide simulated  
120 SWE. Methods that use large-scale patterns as both predictors and predic-  
121 tands perform better than those that use those patterns on only one side of  
122 the regression equation, and all pattern-based methods are superior to a local  
123 “delta change” approach. These findings suggest that pattern-based methods  
124 are indeed appropriate for downscaling interannual snowpack variability, and  
125 that employing physically-meaningful large-scale patterns is more important  
126 for accuracy than the details of any particular downscaling method. Our find-  
127 ings here demonstrate the utility of applying these approaches where more  
128 computational- or data-intensive methods are impractical, including paleocli-  
129 mate modeling and data assimilation.

## 130 2 Data

### 131 2.1 Observations

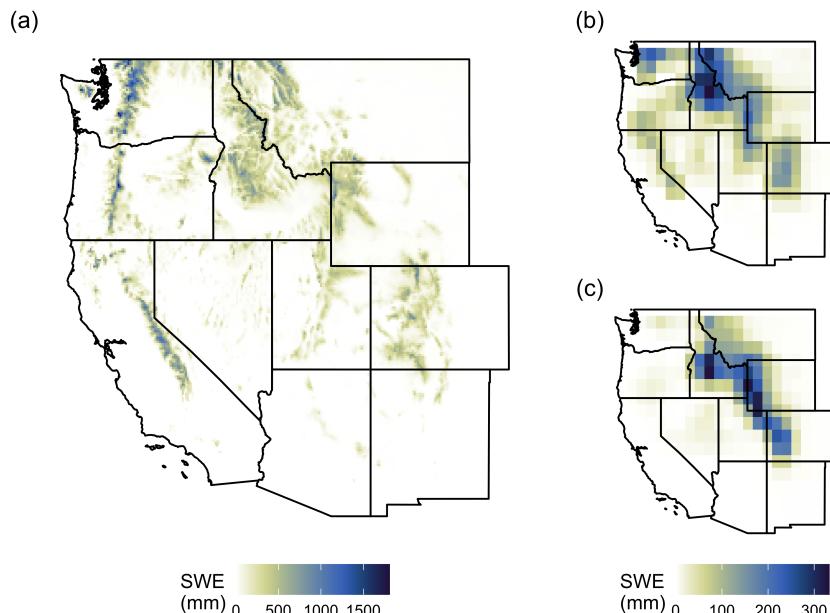
132 We focus on a domain between 125°W–102°W and 31°N–49°N, covering the  
133 western US states of Arizona, California, Colorado, Idaho, Montana, Nevada,  
134 New Mexico, Oregon, Utah, Washington, and Wyoming. Observed March SWE  
135 was calculated from the University of Arizona (UA) Daily 4km SWE data  
136 product, a gridded record of daily SWE and snow depth for water years 1982-  
137 2017 at 4km resolution across the conterminous US (Broxton et al, 2019).  
138 March mean SWE has been shown to approximate the more commonly used  
139 April 1st SWE measure, but is less sensitive to sampling variability than a sin-  
140 gle daily value (Mankin and Diffenbaugh, 2015; Ye, 2019). The UA SWE data  
141 were based on a simple ablation and accumulation model driven by gridded  
142 daily PRISM temperature and precipitation fields (Daly et al, 2008), rescaled  
143 by relative anomalies from thousands of *in situ* observations from the SNO-  
144 TEL and COOP networks (Broxton et al, 2016a; Zeng et al, 2018). We also  
145 acquired the raw PRISM temperature and precipitation fields to assess local  
146 relationships between SWE accumulation and seasonal hydroclimate variabil-  
147 ity.

### 148 2.2 Reanalyses and Simulations

149 Modeled SWE for the downscaling experiments was derived from the CERA  
150 20th century (CERA-20C) reanalysis product (variable name *SD*) (Laloyaux  
151 et al, 2018). CERA-20C is a long-term reanalysis product that uses the Euro-  
152 pean Centre for Medium-Range Weather Forecast (ECMWF) system spanning  
153 1901–2010 at six-hourly temporal resolution and ~1°spatial resolution. It as-  
154 simulates sea level pressure pressure and ocean temperature observations from

across this period in order to avoid temporal inconsistencies from the later introduction of, for example, satellite observations. We also acquired monthly sea surface temperatures and 500mb geopotential heights from the same reanalysis to assess large-scale atmosphere-ocean teleconnections. We used the means of the 10-member ensemble for all analyses as the individual ensemble members showed few major differences during the period of interest.

As a preliminary evaluation of whether these methods could be applied to free-running paleoclimate model simulations, we also analyzed outputs from the CCSM4 Last Millennium simulation (Landrum et al, 2013) and the CESM Last Millennium Ensemble (Otto-Briesner et al, 2016), their associated 20th century extensions (variable name *H2OSNO*, CMIP5 standard name *SNW*), and version 3 of the NOAA-CIRES-DOE 20th Century Reanalysis (20CRv3) (variable name *WEASD*) (Slivinski et al, 2020) in order to assess modes of snowpack variability in free-running Earth system models of different native resolutions ( $\sim 1^\circ$  and  $\sim 2^\circ$ ) and reanalysis data from different modeling groups, respectively. Herein, we collectively refer to both reanalyses and free-running climate models as “simulations” for simplicity.



**Fig. 1** Mean March snow water equivalent (SWE) in mm from A) UA 4km daily SWE observations (Broxton et al, 2019), B) CERA-20C reanalysis (Laloyaux et al, 2018), C) CCSM4 Last Millennium simulation (Landrum et al, 2013). Note the scale of the observations differs from the simulations by nearly an order of magnitude due to differences in model resolution.

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172 2.3 Preprocessing

173 Both observed and simulated data were truncated to the overlapping period  
174 of 1982-2010 and aggregated from daily to monthly timescales by calculating  
175 the average March SWE value for each grid cell and year (Figure 1). We used  
176 bilinear interpolation to resample each of the large-scale simulation outputs to  
177 a common 1°grid. We also resampled the 4km snow observations to an 8km  
178 grid to decrease computational costs without degrading the high-resolution  
179 spatial signal. Grid cells that experienced no SWE accumulation throughout  
180 the observational period were masked from successive analyses.

181 3 Methods

182 3.1 Estimating modes of snowpack variability

183 We isolated key modes of observed snowpack variability using principal com-  
184 ponents analysis (PCA). The observed and simulated data were area weighted  
185 to prevent undue influence from grid cells at higher latitudes by multiplying  
186 the observations at each grid cell by the square root of the cosine of the  
187 cell's latitude in radians (Livezey and Smith, 1999). We calculated interannual  
188 SWE anomalies by mean-centering the data before analysis. We do not use  
189 standardized or detrended anomalies in order to preserve spatial patterns of  
190 variance across the field (Zeng et al, 2018).

191 The PCA results in a set of orthogonal principal component time series  
192 or “amplitudes,” eigenvalues representing the variation accounted for by each  
193 amplitude time series, and eigenvectors or “empirical orthogonal functions”  
194 (EOFs) mapping the amplitude time series back onto the original spatial grid.  
195 We standardized the PC amplitudes to unit variance and reweighted the eigen-  
196 vectors by the square root of their corresponding eigenvalues to give higher  
197 weight to the leading spatial modes (Hannachi et al, 2007). Thus, the origi-  
198 nal dataset could be reconstructed by multiplying each amplitude time series  
199 by its corresponding EOF spatial pattern, summing the results to get SWE  
200 anomalies, and adding in the sample mean of the grid cell. Using only a subset  
201 of these spatiotemporal patterns to reconstruct the original SWE field effec-  
202 tively removes “noise” associated with the higher order modes, limiting the  
203 data to a subspace representing only the most important axes of variation.  
204 The truncation level  $k$  for each field was selected by cross validation (see sec-  
205 tion 3.3).

206 We used several techniques to examine the leading spatiotemporal modes.  
207 We visualized the EOF modes by calculating the Pearson correlation coeffi-  
208 cient between each PC amplitude time series and each grid cell's original time  
209 series. We explored potential atmosphere-ocean teleconnections by calculat-  
210 ing the correlation between each PC amplitude and average October-March  
211 global sea surface temperatures (SSTs) and 500mb geopotential heights from  
212 the CERA-20C reanalysis (Laloyaux et al, 2018) and regional temperature

and precipitation observations from PRISM (Daly et al, 2008), assessing statistically significant correlations using the false discovery rate (Wilks, 2006, 2016). We also applied a varimax rotation to the leading PCs to examine regional response patterns (Richman, 1986), although unrotated PCs were used for downscaling due to their favorable statistical properties and similarity to the rotated PCs.

Although we attempted to find physically-meaningful patterns where they were present, we did not consider the lack of physical interpretation to be a criterion for excluding a particular mode from the downscaling model. We ensured only that the retained modes *collectively* reflected large-scale atmosphere-ocean variability. In other words, the choice of truncation level  $k$  and the combined set of coupled patterns were more important to our downscaling process than the physical interpretation of any particular mode.

### 3.2 Pattern-based downscaling

Pattern-based downscaling models use some combination of observed and simulated PC time series to predict one climate field from another. There are multiple statistical methods capable of doing so, many of which are variants on multiple linear regression (Bretherton et al, 1992; Tippett et al, 2008). They generally differ in whether they maximize explained variance in the observations as opposed to the shared variance between observations and simulations, and whether they use PCs as predictors, predictands, or both (Table 1). We compared four downscaling methods that spanned this methodological spectrum along with an additional “local” null model.

**Table 1** Pattern-based downscaling methods: canonical correlation analysis (CCA), principal components regression (PCR), principal components regression via generalized additive models (PCR-GAM), and empirical orthogonal teleconnections (EOT). Either the predictors ( $x$ ), predictands ( $y$ ), or both are subjected to PCA prefiltering prior to downscaling. Asymmetric models seek to explain variance of the predictands while symmetric models explain the shared correlation. Cross-validated performance metrics for the best-performing model of each class are the space-time root mean square error and the Pearson correlation between observed and simulated domain-wide SWE. The additive delta change approach using bilinearly interpolated anomalies is also including here as a local baseline for the nonlocal downscaling approaches.

Method	PCA Prefiltering	Symmetric	RMSE	Correlation
CCA	$x, y$	yes	41.4	0.940
PCR	$x, y$	no	43.1	0.949
PCR-GAM	$x, y$	no	42.7	0.932
EOT	$y$	no	48.5	0.918
DELTA	<i>none</i>	no	53.2	0.912

Canonical correlation analysis (CCA) is one of the most common approaches to coupled pattern analysis (Maraun and Widmann, 2018). It yields

a set of patterns that maximizes the shared correlation between the predictor and predictand fields (Tippett et al., 2008). We applied CCA to the leading predictor and predictand modes of variability to regularize the model and make it computationally tractable (Barnett and Preisendorfer, 1987; Bretherton et al., 1992; Benestad, 2001; Tippett et al., 2008). Downscaling models are prone to overfitting on shorter calibration windows, so this PCA prefiltering step increases the signal-to-noise ratio to ensure the resulting patterns are statistically robust.

Principal components regression (PCR) is a similar method that uses the PC time series in independent multiple linear regressions. Traditional PCR fits a different model to the predictor PCs for each predictand grid cell, although here we take the more efficient approach of using predictand PCs directly (Benestad et al., 2015). Because the PC time series are mutually uncorrelated each predictand PC can be modeled independently and there is no concern of multicollinearity. PCR is asymmetric in that it only explains the variance of the predictands, contrary to CCA, although both methods are linear and are equivalent under certain conditions (Tippett et al., 2008). We also tested a nonlinear variant of PCR which replaces the linear models with penalized piecewise polynomials estimated in a generalized additive model (PCR-GAM).

Empirical orthogonal teleconnections (EOT) finds a set of grid cells that explain the most variance in the observation domain by fitting a linear model between all pairs of predictor and predictand grid cells (van den Dool et al., 2000; Appelhans et al., 2015). The simulation grid cell that predicts the most variance in all of the predictand grid cells is selected as the first pattern. Then the algorithm is run again on the residuals from the regressions on the first pattern, and the process is repeated until a set number of patterns is reached. EOT yields more localized spatial patterns, similar to rotated EOFs, than methods that use predictor and predictand PCs directly. Although PCA can be used to denoise both fields prior to the analysis, EOT focuses on the grid-cell level time series and is not constrained to fit the large scale patterns used by CCA and PCR.

We compared these non-local pattern-based techniques to a simple null model using interpolated simulation anomalies. This “delta change” model involved calculating the yearly simulated SWE anomalies relative to the mean fields, bilinearly interpolating the low-resolution simulated anomalies to the high resolution of the observations, and combining the interpolated anomalies with the high resolution observed mean. We tested delta change models with both additive and multiplicative anomalies by either subtracting the simulated mean and adding the observed mean or by dividing by the simulated mean and multiplying by the observed mean, respectively. While conceptually similar to the pattern-based methods, the delta change approach uses only local information and cannot correct any spatial biases caused by the smoothed topography of the simulation. We used these model to assess the added value of the non-local downscaling approaches relative to these common local methods.

All downscaling methods were implemented in R version 4.0.3 (R Core Team, 2020) using the packages `stars`, `tidyverse`, `mgcv`, `remote`, and `MuMIN`

(Wood, 2006; Appelhans et al, 2015; Wickham et al, 2019; Bartoní, 2020; Pebesma, 2021). Code for reproducing the main analysis and figures is available at <https://github.com/nick-gauthier/tidyEOF>.

### 287 3.3 Cross Validation

288 Each of the pattern-based downscaling methods required the number of coupled  
289 patterns to be defined by a hyperparameter  $k$ . The methods that used  
290 PCA prefiltering also required selection of a truncation level for the predictor  
291 and predictand PCs. We used a five-fold cross validation routine to tune the  
292 hyperparameters of each model, fitting and predicting from models with all  
293 possible combinations of up to ten predictor patterns  $k_x$ , predictand patterns  
294  $k_y$ , and coupled patterns  $k_{xy}$ , with the constraint that  $k_{xy} \leq \min(k_x, k_y)$ .

295 We divided the 29-year calibration period into five contiguous folds, four of  
296 which contained six years and one of which contained five. We held out one fold  
297 at a time, fitting each model and parameter combination on the remaining folds  
298 and using them to predict the held out fold. The entire modeling workflow—  
299 anomaly calculation, PCA truncation, and model fitting—was repeated for  
300 each training and testing fold independently to prevent leaking information  
301 among the folds (Van Den Dool, 1987; Livezey and Smith, 1999; Smerdon  
302 et al, 2010). We repeated this process until each fold had been used four  
303 times for training and once for testing, after which we combined the test folds  
304 into a single 29 year sequence from which we calculated the prediction error  
305 against the observed sequence. It is often preferable to use a nested cross  
306 validation routine when doing model selection and performance assessment  
307 simultaneously, but we did not do so in this case because our sample size was  
308 limited and the different models were of broadly the same type with a low  
309 number of similar hyperparameters (Wainer and Cawley, 2018).

310 We used two metrics to assess the skill of each model and parameter combi-  
311 nation. First we examined the correlation between the observed and predicted  
312 domain-wide total SWE time series. We calculated total domain SWE by mul-  
313 tiplying each SWE value by the area of its grid cell and summing the result.  
314 We then assessed the local spatial skill of the downscaled product by calculat-  
315 ing the total space-time root mean square error (RMSE) between all observed  
316 and predicted grid cells. We selected the models and parameter combinations  
317 that maximized domain-wide correlation and minimized RMSE under cross  
318 validation, and refit the best performing model to the entire data series. We  
319 compared the predictions from this final model to the raw CERA-20C re-  
320 analysis to assess the added value of downscaling for correcting mean and  
321 variance biases in domain-wide SWE. We demonstrated the spatial skill of  
322 the model by comparing the spatial anomalies of observed, reanalysis, and  
323 downscaled fields during a known extreme year. To test its sensitivity to re-  
324 cent warming trends, we refit the best model holding out the years with the  
325 top 20% warmest October-March average temperatures in the PRISM obser-  
326 vations. We also compared these CERA-20C based reconstructions to models

327 using the NOAA-CIRES-DOE 20CRv3 reanalysis (Slivinski et al, 2020) as an  
328 alternative predictor to assess the sensitivity of the outputs to the specific  
329 reanalysis methodology.

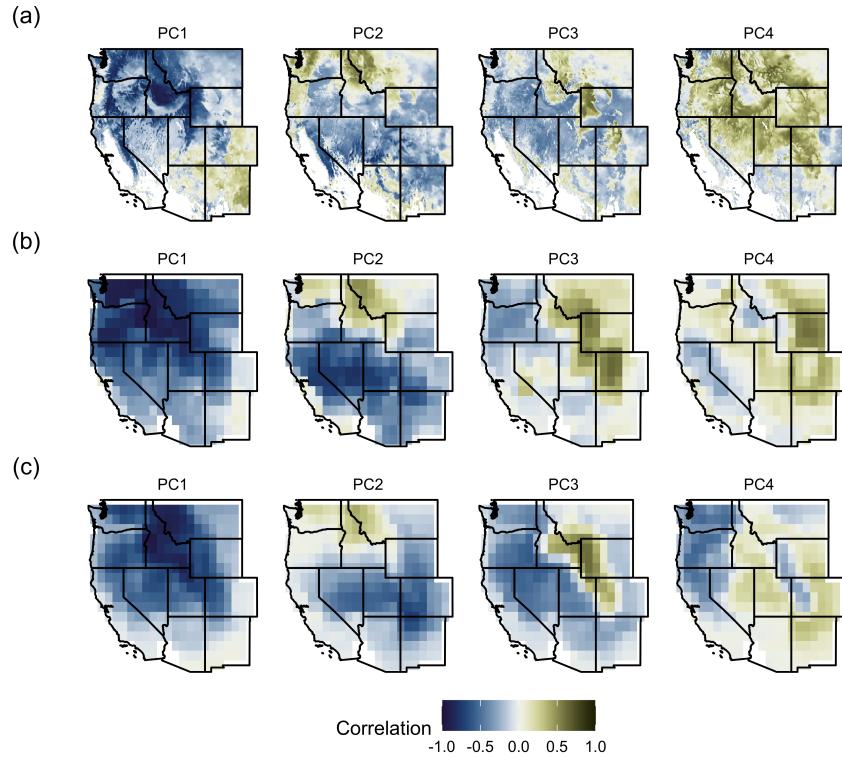
330 A downscaling model trained on reanalysis data must also be able to make  
331 predictions from unseen, free-running simulations to make skillful climate-  
332 change impact assessments beyond the observational period (Maraun and  
333 Widmann, 2018). As a proof-of-concept of the generalizability of the final  
334 model and EOF patterns, we used it to downscale additional 300-year sim-  
335 ulated snowpack sequences by projecting data from the CCSM4 and CESM  
336 Last Millennium simulations onto the reanalysis PC patterns. As these free-  
337 running simulations were not constrained to match the year-to-year evolution  
338 of the observations as were the reanalyses, the added value of downscaling was  
339 assessed through improvements in the mean and variance biases on a 50-year  
340 distributional basis.

## 341 4 Results

342 *A limited set of climate modes explain the majority of observed and simulated  
343 March SWE variance.* Four spatiotemporal patterns explain 76% of the ob-  
344 served variance in March snowpack over the western United States (Figure 2a).  
345 The leading ten patterns explain nearly 90% of the observed variance. These  
346 patterns represent recurring modes of spatiotemporal variability and are an  
347 efficient means of capturing the high dimensional spatiotemporal snowpack  
348 field in a limited subspace of patterns.

349 Similar patterns are found in coarse-resolution simulations. The leading 10  
350 PCs of the 110 year CERA-20C reanalysis explain 96% of the variance in sim-  
351 ultated snowpack, and the leading four explain 89% of the variance (Figure 2b).  
352 These reanalysis PCs are associated with the same broad spatial patterns as  
353 the observed PCs, but the longer sample windows allows for greater separation  
354 between the leading modes than with the 36 year observational record.

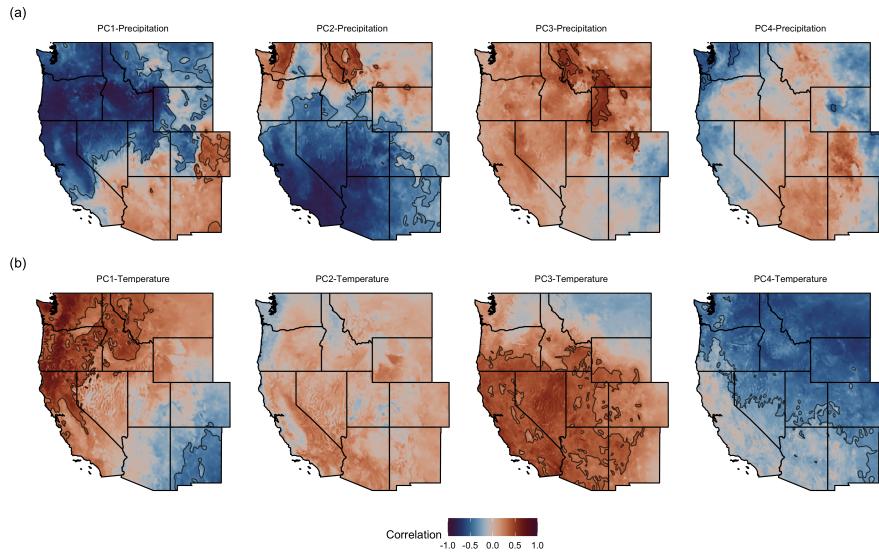
355 *Large-scale snow patterns reflect orography and atmosphere-ocean variability.*  
356 The spatial EOF patterns associated with the leading snowpack PCs exhibit  
357 clear relationships to regional precipitation and temperature (Figure 3) as well  
358 as global pressure systems and sea surface temperatures (Figure 4). EOF/PC1  
359 is a domain-wide signal with high loadings in the Rocky Mountains, Sierra  
360 Nevada, and Cascade ranges. It is associated with simultaneous cold and wet  
361 conditions (or vice versa) over the domain and anomalous pressure systems  
362 over northwestern North America. EOF/PC2 exhibits a north-south dipole  
363 pattern with opposite-sign loadings in the Cascades and northern Rockies  
364 and the Sierra Nevada and southern Rockies, respectively. Unlike EOF1, this  
365 pattern is associated the precipitation, not temperature, anomalies over the  
366 domain and a far more zonal geopotential height anomaly over North America.  
367 EOF3 is localized to the Rocky Mountains and is associated with domain-  
368 wide temperature anomalies and geopotential and SST dipoles over the north



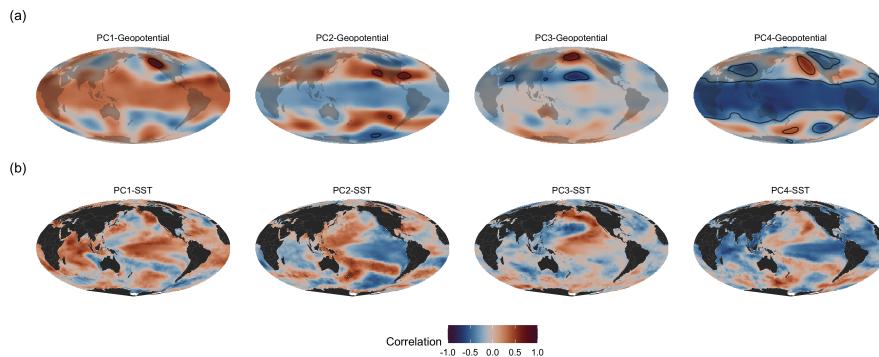
**Fig. 2** The leading four EOF spatial patterns expressed as the Pearson correlation coefficient between each PC time series and March snow water equivalent in (A) UA SWE observations (1982-2017) (Broxton et al., 2019), (B) CERA-20C reanalysis (1901-2010) (Laloyaux et al., 2018), and (C) the CCSM4 Last-Millennium simulation and historical extension (850-2005) (Landrum et al., 2013). These patterns represent between 76% and 90% of the variance in their respective spatiotemporal fields.

369 Pacific. EOF4 is a domain-wide mode associated with temperature anomalies  
 370 and SST and geopotential height anomalies off the Pacific coast and in the  
 371 tropics. Although the SST correlations exhibit spatial structure resembling  
 372 ENSO and other modes of Pacific SST variability (Figure 4b), none of these  
 373 are significant during the 1982-2010 period (although the horseshoe-shaped  
 374 PC3-SST and coastal PC4-SST patterns are significant in SST observations  
 375 that extend to 2017 (Huang et al., 2017a)).

376 Higher order PCs/EOFs beyond the leading four also show spatially co-  
 377 herent variability. While these PC/EOF pairs may resemble physical climate  
 378 patterns, they are not interpreted here as the orthogonality constraints may  
 379 lead to mixed or otherwise poorly resolved patterns spread across multiple  
 380 PCs. Given the small sample size, it can be difficult to distinguish such “de-  
 381 generate multiplets” from proper modes (North et al., 1982). While the first two  
 382 observed PCs are distinct modes of variability, PCs 3-4 and 5-10 are degen-  
 383 erate multiplets that cannot be readily distinguished from one another given

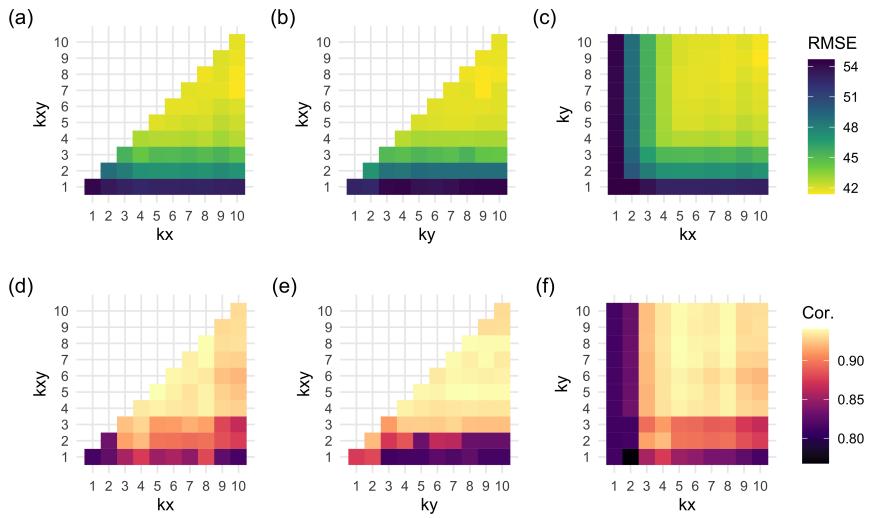


**Fig. 3** Pearson correlation coefficients between the leading four observed PC time series and October–March (A) total precipitation and (B) average temperature from PRISM (Daly et al. 2008) over the 1982–2017 period. Contour lines indicate regions of statistically significant correlation with a false discovery rate below 0.1.



**Fig. 4** Pearson correlation coefficients between the leading four observed PC time series and October–March (A) 500mb geopotential height and (B) sea surface temperature from CERA-20C (Laloyaux et al. 2018) over the 1982–2010 period. Contour lines indicate regions of statistically significant correlation with a false discovery rate below 0.1.

384 the limited 36-year observational period (1982–2017). Likewise, the first four  
 385 reanalysis PCs represent distinct modes while PCs 5–7 and 8–10 are degenerate  
 386 multiplets. A varimax rotation of the leading ten PCs alleviates some of these  
 387 concerns, yielding more discrete zones reflecting topographic interception of  
 388 different directions of atmospheric flow. Regardless, that these patterns are  
 389 present in reanalysis and simulation data from much longer time spans (1901–

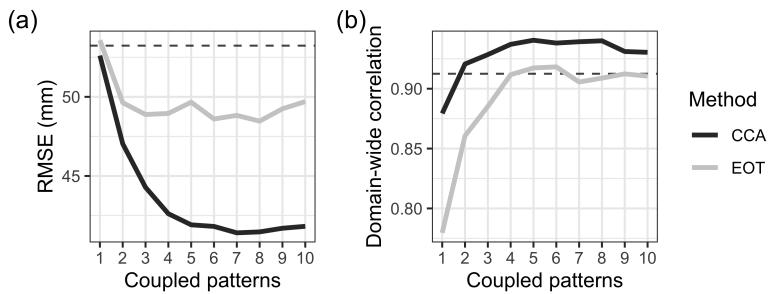


**Fig. 5** Cross validation results for the three CCA parameters after Smerdon et al (2010): the number of predictor PCs  $k_x$ , the number of predictand PCs  $k_y$ , and the number of coupled patterns  $k_{xy}$ . (A)–(C) cross validation results for space-time root mean square error in millimeters (lower is better). (D)–(E) correlation between observed and downscaled total domain SWE (higher is better).

390 2010 and 840-2005, respectively) suggests the observed patterns are robust in  
 391 time and can be used as anchoring points for a non-local downscaling approach.

392 *Downscaling with coupled patterns has higher cross-validated skill than similar  
 393 local and non-local methods.* CCA is the best-performing downscaling model  
 394 under cross validation, with the lowest space-time root mean square error and  
 395 effectively tied for the highest correlation with total western US SWE (Table  
 396 1). The most important parameter for model skill is the number of coupled  
 397 patterns  $k_{xy}$ , while the precise number of prefiltering patterns  $k_x$  and  $k_y$  is less  
 398 important as long as they are greater than or equal to the optimal number of  
 399 coupled patterns (Figure 5). A CCA model with five coupled patterns maxi-  
 400 mizes the domain-wide correlation, but even one coupled pattern yields a high  
 401 correlation coefficient. Likewise, a model with seven coupled patterns is the  
 402 most accurate in reconstructing the entire spatiotemporal field (lowest cross  
 403 validated RMSE), but a five-pattern model also performs reasonably well.

404 All models have comparable skill to CCA for domain-wide SWE correla-  
 405 tions, yielding a cross validated correlation of around 0.9, but there is greater  
 406 spread for space-time RMSE. Both PCR models perform similarly to CCA  
 407 for domain-wide SWE correlation, but the spatial skill is degraded due to the  
 408 asymmetrical relationships between the predictors and predictands. PCR and  
 409 PCR-GAM models produced largely similar reconstructions, yet the nonlinear



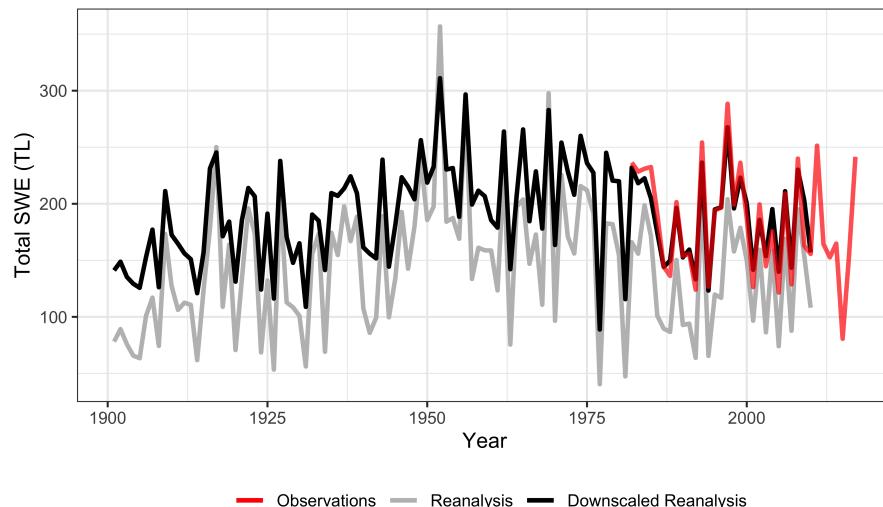
**Fig. 6** Comparison of CCA and EOT downscaling under five-fold cross validation. (A) Space-time root mean square error, in millimeters of SWE, for increasing number of coupled patterns. Lower RMSE corresponds to more accurate reconstructions. (B) Correlation between observed and reconstructed total SWE over western North America. The dashed horizontal line indicates the cross validated skill of the additive delta change model, a “local” interpolation-based downscaling approach. The curves for the PCR and PCR-GAM models (not shown) resemble those of the CCA model.

410 PCR-GAM consistently performs slightly worse than the linear PCR method  
411 due to its potential to overfit.

412 EOT yielded spatial patterns similar to the coupled-pattern methods but  
413 with notably more instability under cross validation than the pattern methods  
414 because the base grid cell tended to vary between folds (Figure 6). All meth-  
415 ods are better than the delta change approach with additive anomalies, which  
416 performed similar to the pattern-based methods with only one or two pat-  
417 terns. The multiplicative delta change approach was by far the least effective,  
418 as the use of multiplicative anomalies introduced artifacts in years with un-  
419 usually high SWE over areas with SWE averages close to zero. These artifacts  
420 significantly degraded the overall temporal and spatial skill, and were partic-  
421 ularly severe under cross validation. These results support the interpretation  
422 that anchoring downscaling relationships in spatial patterns, rather than grid-  
423 cell level relationships, increases the robustness of the resulting downscaled  
424 predictions.

425 *Downscaling reduces spatial and temporal biases in simulated snowpack.* Down-  
426 scaling the CERA-20C reanalysis with any of the above pattern-based methods  
427 considerably reduces spatial and temporal biases in the raw reanalysis. With-  
428 out downscaling, the CERA-20C reanalysis tends to underpredict domain-wide  
429 total SWE averages and overpredict their variance. CCA downscaling with five  
430 coupled patterns reduces this mean and variance bias relative to observations  
431 (Figure 7). By construction, pattern-based downscaling also improves the spa-  
432 tial structure of simulated SWE anomalies and removes spatial biases caused  
433 by the coarse resolution of the simulated topography in a way that interpolat-  
434 ing the simulated anomalies does not (Figure 8).

435 The spatial skill of the best-performing CCA model does not appear to  
436 be sensitive to recent warming trends. Using a model fit on the 80% coolest

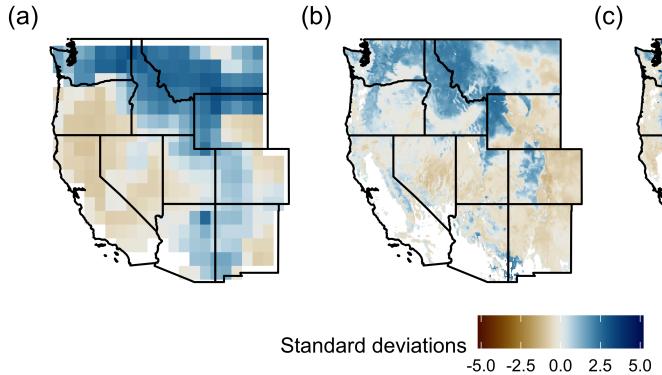


**Fig. 7** Total Western US March SWE in teraliters ( $km^3$ ) from the CERA-20C reanalysis with five-pattern CCA downscaling (black) and without (gray), compared to recent observations (red). Downscaling adds value to the raw reanalysis by increasing the mean and decreasing the variance relative to observations.

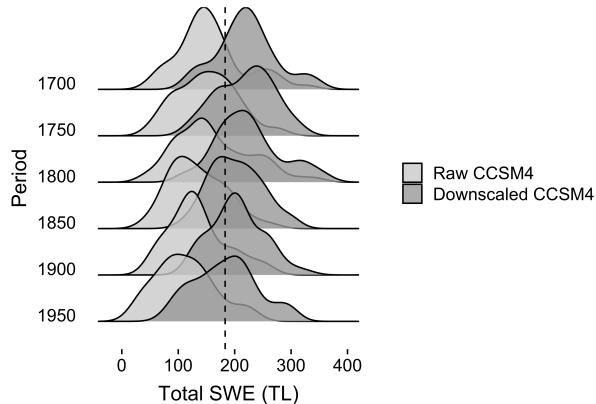
437 years to predict the 20% warmest years in the calibration period (1992, 1999,  
 438 2000, 20003, 2004, 2005) yields a space-time RMSE of 40.9mm, with virtually  
 439 no spatial bias between the performance of this “cool” model and the full  
 440 one. However, both models do tend to underestimate the total domain SWE  
 441 deficits in the driest years, suggesting that while the pattern-based methods  
 442 can represent recent warming trends in space, they may still be inheriting  
 443 small temporal biases from the underlying reanalysis.

444 Reconstructions driven instead by the NOAA-CIRES-DOE 20th century  
 445 reanalysis are consistent with those downscaled from CERA-20C. The raw  
 446 CERA-20C and 20CRv3 SWE fields have a domain-wide SWE correlation of  
 447 0.90 and a space-time RMSE of 77mm, while the downscaled fields have a cor-  
 448 relation of 0.88 and RMSE of 31mm, indicating that downscaling substantially  
 449 improves the spatial coherence of the reanalysis data while leaving temporal  
 450 coherence largely the same. Notably, the RMSE among the two downscaled  
 451 reanalysis fields is well below that of the best performing downscaling model  
 452 under cross validation, suggesting that uncertainty due to changing calibra-  
 453 tion windows is greater than that from the selection of the particular predictor  
 454 dataset.

455 A CCA model fit to the reanalysis data also reduces biases in the free-  
 456 running CCSM4 Last Millennium simulation. Downscaling CCSM4 outputs  
 457 by simply projecting them onto the patterns estimated from CERA-20C cor-  
 458 rects mean and variance biases in total domain-wide SWE relative to the raw  
 459 simulation (Figure 9). Domain-wide SWE downscaled from CCSM4 exhibits



**Fig. 8** Standardized SWE anomalies for the 1997 El Niño in (A) CERA-20C reanalysis, (B) downscaled CERA-20C reanalysis using CCA with seven coupled patterns, and (C) observations. (B) and (C) are both scaled by the observed SWE standard deviation to allow comparison. Note the lower standardized anomalies in the downscaled reanalysis relative to the observations, due to the residual variance unexplained by the leading patterns.



**Fig. 9** Total Western US March SWE in teraliters ( $km^3$ ) from the CCSM4 Last Millennium simulation (Landrum et al., 2013) with CCA downscaling (dark gray) and without (light gray), compared to the 1982–2017 observed mean (dashed line). Unlike CERA-20C, CCSM4 is not constrained to be synchronous with observations and is instead assessed on a 50-year distributional basis. The same model fit from Figure 7 is used here, with the CCSM4 data simply projected onto the reanalysis PC space to enable downscaling. This approach was less successful when applied to CESM-LME outputs (not shown), as its  $\sim 2^\circ$  native resolution was too coarse to meaningfully project onto the  $1^\circ$  reanalysis patterns.

460 the same broad temporal correlations to simulated temperature and precipitation  
 461 trends internal to the raw CCSM4 simulation, indicating that downscaling  
 462 does not break the physical consistency of the water balance from the free-  
 463 running simulation.

That simply projecting the CCSM4 data onto the CERA-20C patterns, without additional transformations, results in reasonable estimates at all is informative. For a statistical model fit on large-scale patterns from one simulation to meaningfully generalize to those from a different simulation is not guaranteed. Indeed, this is not the case for the coarser 2°CESM-LME simulation. Although the spatial patterns from CESM-LME are visually similar to those in Figure 2, they are too different at the grid cell level to be used directly for downscaling. This constraint holds regardless of whether the CESM-LME data are first resampled to the 1°resolution of CERA-20C and CCSM4 or when CERA-20C is resampled to the lower CESM resolution. This indicates that the problem is not due grid-size *per se*, but rather the impact of the simulation’s native resolution on the underlying dynamics. That the downscaling model can generalize to both a distinct reanalysis dataset (20CRv3) and free-running climate model (CCSM4) at the same native resolution as the CERA-20C data used to fit the model, but not to the coarser CESM data, suggests the model generalizes well only to simulations run with a similar native resolution to the training data.

## 5 Discussion

A small number of climate modes explain the majority of observed and simulated interannual variance in snowpack across the western United States. Five to seven of these coupled modes are sufficient to downscale accurate high-resolution maps of regional snow water equivalent from coarse-resolution climate simulations. Even an extremely simple model with only one mode is able to reproduce the time evolution of the total volume of water stored in snow across the whole domain, although this is unlikely to be sufficient for full field spatiotemporal analyses. In spite of known biases in simulated SWE arising from issues of scale and process uncertainty, these findings suggest modern numerical simulations capture enough of the large scale atmosphere-ocean dynamics that drive interannual snowpack variability to be appropriate predictors for high resolution downscaling products.

Given judicious choice of physically meaningful patterns as predictors and predictands, even a simple linear downscaling method yields skillful hindcasts of observed SWE variability. This approach relies on the ability of climate and weather models to accurately simulate large-scale atmosphere-ocean variability. Rather than deriving complex transfer functions between a variety of local variables—a process that often breaks the physical consistency of climate model outputs—this approach uses the internal physical consistency of those simulations to its advantage by finding a simple mapping between simulated and observed patterns. Anchoring statistical downscaling methods in a mechanistic understanding of the climate system, instead of using downscaling as a replacement for that understanding, is of paramount importance to any downscaling project.

The leading two principal modes of variability highlighted in this study—a coherent domain-wide signal and a north/south dipole—have been identified previously in observational data of snow and several variables (Redmond and Koch, 1991; Cayan, 1996; McCabe and Dettinger, 2002; Jin et al, 2006; McCabe et al, 2013; Pederson et al, 2013; Malevich and Woodhouse, 2017). The first mode represents a domain-wide temperature anomaly associated with PNA-type atmospheric circulation. The second represents the influence of tropical Pacific SST variability (ENSO, PDO) deflecting storm tracks north or south and causing coincident temperature and precipitation anomalies in each region. This pair of influences is robust over time and appears in long-term tree-ring reconstructions from similar domains (Woodhouse, 2003; Pederson et al, 2011; Coulthard, 2015; Barandiaran et al, 2017).

There is less certainty as to the drivers of the successive modes of variability. Possible influences include cold vs. warm El Niño years, atmospheric rivers, temperature anomalies due to the Northern Annular Mode and North Atlantic Oscillation, or overlapping multidecadal modes of Pacific SST variability (QDO, PDO, IPO) (Ghatak et al, 2010; Seager et al, 2010; Barrett et al, 2015; Barandiaran et al, 2017; Goldenson et al, 2018). Complicating matters further is that the same large scale pattern can influence snowpack through multiple physical pathways and different teleconnections can act through the same pathway (Mote, 2003; Ge et al, 2009; Ghatak et al, 2010). For example, ENSO variability influences both temperature and precipitation, and by extension snow accumulation and ablation, simultaneously. Likewise, Pacific SST variability can influence storm tracks across multiple spatial and temporal scales.

Ultimately, these large-scale patterns represent the outcome of nonlinear, interacting processes that may not necessarily be well represented by linear statistical methods like PCA and CCA. What may appear to be distinct climatic modes in a PCA may instead reflect the method's linearity assumptions and orthogonality constraints. While these methods are nevertheless useful for downscaling because they isolate the parsimonious subspace of variability most influenced by these large-scale dynamics, interpretations of the individual modes must always be treated with caution. An alternative approach would be to use nonlinear feature extraction methods such as independent components analysis, self-organizing maps, or variational autoencoders to generate statistically independent patterns with increased interpretability and out-of-sample predictability (Reusch et al, 2005; Fassnacht and Derry, 2010; Henderson et al, 2017; Baño-Medina et al, 2020; He and Eastman, 2020). However, the risk of overfitting nonlinear methods remains high given the short observational window, and standard linear methods are already highly skillful.

Regardless of whether this large-scale variability is captured by linear or nonlinear methods, a degree of unexplained local variability will remain. About 20% of the local SWE variance observed at the grid cell level is left unexplained by the large-scale patterns. By definition, methods that use a restricted number of patterns on the left hand side of the regression equation will explain only a subset of the observed variance. Ideally, a downscaled SWE product would

552 preserve this full range of variability and give some insight into the uncertainty  
553 in the downscaled estimates (Hewitt et al, 2018). An intuitive approach would  
554 be to add the residual variance back to each grid cell as uncorrelated white  
555 noise. However, we find here that the residual fraction is non-normal, spatially  
556 autocorrelated, and varies in magnitude across the study domain. While an  
557 analytical solution to the CCA noise fraction exists (Wilks, 2014), a more  
558 pragmatic approach may be to fit Gaussian process or copula models to the  
559 cross-validated errors directly. Regardless of the precise method, this residual  
560 internal variability should be modeled in order to yield downscaled data ap-  
561 propriate for localized climate-change impact assessments (Towler et al, 2017).

562 To be truly useful to researchers, stakeholders, and policy makers in the  
563 western US, downscaled snowpack products should take advantage of the wide  
564 range of long-term paleoclimate simulations to generate long-term ensembles  
565 of high-resolution snowpack variability. Such products would provide a crucial  
566 baseline for assessing present and future climate changes. Downscaled SWE es-  
567 timates can also serve as spatially-explicit priors for data-assimilation (Huang  
568 et al, 2017b; Devers et al, 2019; Fiddes et al, 2019; Girotto et al, 2020), com-  
569 bining high-resolution snowpack fields with snow-sensitive tree-ring proxies  
570 (Coulthard et al, 2021) to generate integrated paleoclimate reconstructions  
571 (Hakim et al, 2016). We applied our reanalysis-based downscaling approach to  
572 a free-running CCSM4 simulation to test the generality of the leading SWE  
573 patterns, suggesting that pattern-based downscaling of long-term paleoclimate  
574 simulations is indeed possible. Operational downscaling for long-term climate-  
575 change impact assessments will require further steps to ensure the robustness  
576 of the coupled patterns, such as Common EOF analysis on combined reanaly-  
577 sis and GCM fields (Benestad, 2001) and perfect model experiments (Maraun  
578 and Widmann, 2018) to determine whether a long-term climate change sig-  
579 nal can be captured by changes in the relative expression of existing spatial  
580 patterns. Nevertheless, our results indicate that leading modes of snowpack  
581 variability have been sufficiently stable for at least the past few centuries,  
582 and that pattern-based downscaling provides clear added value for assessing  
583 changing snowpack over the long-term.

## 584 **Declarations**

### 585 **Funding**

586 This research was supported by the National Science Foundation Paleo Per-  
587 spectives on Climate Change (P2C2) grant AGS-1803995.

### 588 **Conflicts of interest/Competing interests**

589 The authors have no relevant financial or non-financial interests to disclose.

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590 Availability of data and material

591 The daily 4km UA-SWE product can be accessed at <https://nsidc.org/data/nsidc-0719> and the CERA-20C reanalysis at <https://www.ecmwf.int/en/forecasts/datasets/reanalysis-datasets/cera-20c>. Although the entire analysis can be made using publicly available data, intermediate data required for the main analysis are included with the R research compendium linked below. The downscaled CERA-20C SWE estimates for 1901-2010 generated in this study will be made available at doi:10.5281/zenodo.5110319 upon acceptance.

598 Code availability

599 Code for reproducing the main analysis and figures is available at <https://github.com/nick-gauthier/tidyEOF> and will be permanently archived on Zenodo upon acceptance at doi:10.5281/zenodo.5110395.

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