Self-Exciting Spatio-Temporal Models for Count Data

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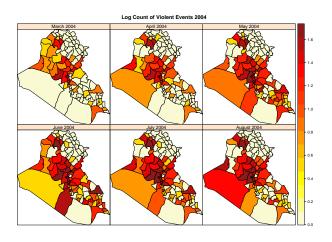
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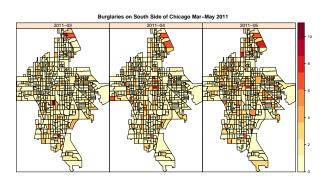
Motivation: The Evolution of Violence in Space and Time

At present, the most under-researched area of spatial criminology is that of spatio-temporal crime patterns... the temporal component of the underlying crime distributions has languished as a largely ignored area of study - Crime mapping: Spatial and Temporal Challenges, Ratcliffe (2010)

The Spread of Violence in Iraq 2004



Burglaries South Side of Chicago



Goals

- General statistical model for diffusion of violence in space-time
 - Accurately reflects beliefs on how violence/crime evolves
 - Extends traditional statistical models for count data
 - Stationary with extremely flexible second order properties
 - Inference via traditional MCMC techniques

Overview

- Mathematical Model for Diffusion of Crime and Related Statistical Models
 - Issues with INGARCH (1,1) Model
- SPINGARCH Model
- SPINGARCH Stationarity and Model Properties
- Inference
- Simulation
- Burglaries in South Side of Chicago

A Model of Criminal Behavior (Short et al. 2008)

- $Z(s_i, t)$ number of observed burglaries from $(t \Delta t, t)$
- ullet $s_i \in \{s_1, \cdots, s_n\}$ fixed regions in \mathbb{R}^2
- ullet $t \in \{1, \cdots, T\}$ discrete time
- Define $A(s_i, t) \equiv A(s_i, 0) + B(s_i, t)$ as attractiveness

$$B(s_i,t) = (1 - \chi \Delta t) B(s_i, t - \Delta t) + \eta Z(s_i, t - \Delta t)$$
 (1)

- Probability of occurrence at each time interval, $(t, t + \Delta t)$ is Poisson with rate, $A(s_i, t)$
- Three factors impact change in crime rate, base attractiveness , decay $\chi,$ and repeat victimization, η



Relationship to INGARCH Model

Integer Auto-Regressive Conditionally Heteroskedastic, INGARCH (1,1), or Poisson Auto-Regression Model

$$Z(s_i, t) \sim \text{Pois}(\lambda(s_i, t))$$
 (2)

$$\lambda(s_i, t) = d + a\lambda(s_i, t - 1) + bZ(s_i, t - 1)$$
(3)

- Unlike GARCH, not solely a variance property
- Short model is similar to INGARCH(1,1) with $A(s_i,0) = \sum_{k=0}^{t} a^k d$, $a = (1 \chi \Delta t)$, and $b = \eta$

Relationship to Self-Exciting Models

Point process introduced by Alan Hawkes with intensity

$$\lambda(t) = \nu(t) + \int_0^t g(t - u) N(ds) \tag{4}$$

Commonly discretized as

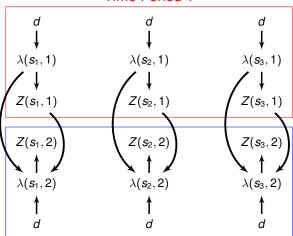
$$Z(s_i,t) \sim \text{Pois}(\lambda(s_i,t))$$
 (5)

$$\lambda(\mathbf{s}_i, t) = \nu + \sum_{j < t} \eta^{t-j} Z(\mathbf{s}_i, t - j)$$
 (6)

Equivalent to stationary INGARCH(1,1)

Structural Diagram - INGARCH(1,1)

Time Period 1



Time Period 2

Short (2008) Extension - Spatial Spread

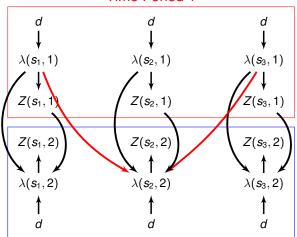
- Motivated by Reaction-Diffusion PDE
- Recall $A(s_i, t) \equiv A(s_i, 0) + B(s_i, t)$ as attractiveness, let N_{s_i} be spatial adjacency matrix

$$B(s_i, t) = \kappa B(s_i, t - \Delta t) + \frac{\psi}{|N_{s_i}|} \sum_{s_j \in N_{s_i}} \left[B(s_j, t - \Delta t) - B(s_i, t - \Delta t) \right] + \eta Z(s_i, t - \Delta t)$$
(7)

- Four factors impact change in crime rate, base attractiveness, decay χ , and repeat victimization, η , and spatial spread ψ
- Resulting model is MINGARCH (1,1)

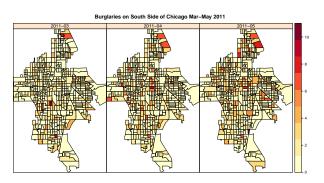
Short (2008) Extension - Spatial Spread

Time Period 1



Time Period 2

Applied to Residential Burglaries in Chicago



552 Spatial Locations, 72 Months, residential burglaries

Applied to Residential Burglaries in Chicago

$$B(s_i, t) = \kappa B(s_i, t - \Delta t) + \frac{\psi}{|N_{s_i}|} \sum_{s_j \in N_{s_i}} \left[B(s_j, t - \Delta t) - B(s_i, t - \Delta t) \right] + \eta Z(s_i, t - \Delta t)$$
(8)

- Further structure $A(s_i, 0)$ to account for socio-economic factors
- MLEs are $\hat{\psi} = .02$, $\hat{\eta} = .155 \ \hat{\kappa} = .773$
- Simulate data from asymptotic distribution unable to replicate lag-one autocorrelation, spatial correlation, or variance to mean ratio of original data

Properties of INGARCH Model

$$Z(s_i,t) \sim ext{Pois } (\lambda(s_i,t)) \ \lambda(s_i,t) = d + a\lambda(s_i,t-1) + bZ(s_i,t-1)$$

Stationarity yields:

$$E[Z(s_i,t)] = \frac{d}{1-(a+b)} \tag{9}$$

$$Var[Z(s_i,t)] = \frac{1 - (a+b)^2 + b^2}{1 - (a+b)^2} E[Z(s_i,t)]$$
 (10)

$$Cov[Z(s_i,t),Z(s_i,t-h)] = \frac{b(1-a(a+b))(a+b)^h}{1-(a+b)^2}E[Z(s_i,t)]$$
 (11)

Var-Mean Ratio[
$$Z(s_i, t)$$
] = 1 + $\frac{b^2}{1 - (a + b)^2}$ (12)

Issues

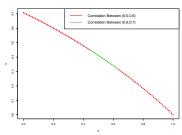
Allows for Overdispersion... But at a cost!

$$Cor[Z(s_i, t), Z(s_i, t - 1)] = \frac{b(a+b)(a^2 + ab - 1)}{a^2 + 2ab - 1}$$
(13)

Var-Mean Ratio[
$$Z(s_i, t)$$
] = 1 + $\frac{b^2}{1 - (a + b)^2}$ (14)

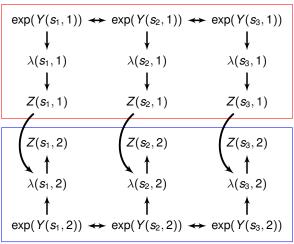
(15)

For fixed Var-Mean Ratio at 2 $\implies b = 1/2(-a + \sqrt{2 - a^2})$.



$Y(s_i, t)$ - Spatially Correlated Latent Gaussian

Time Period 1



Spatially Correlated Self-Exciting Model (Clark & Dixon, 2018)

 Theory: Exist common cause between geographically similar locations, regions that experience uptick in violence likely to have short term self-excitation

Spatially Correlated Self-Exciting Model (Clark & Dixon, 2018)

- Theory: Exist common cause between geographically similar locations, regions that experience uptick in violence likely to have short term self-excitation
- Mixture of two processes that influence expectation: LGCP and Hawkes process
- Hawkes process letting $g(t j) = \eta$ if (t j) = 1, 0 otherwise

$$Z(\mathbf{s_i}, t) | \lambda(\mathbf{s_i}, t) \sim \text{Pois} (\lambda(\mathbf{s_i}, t))$$

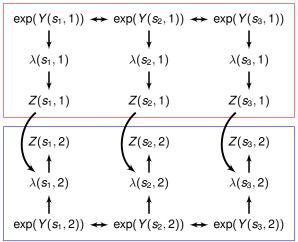
$$\lambda(\mathbf{s_i}, t) = \exp(Y(\mathbf{s_i}, t)) + \eta Z(\mathbf{s_i}, t - 1)$$

$$Y(\mathbf{s_i}, t) = \theta_1 \sum_{\mathbf{s_j} \in N(\mathbf{s_i})} Y(\mathbf{s_j}, t) + \epsilon(\mathbf{s_i}, t)$$

$$\epsilon(\mathbf{s_i}, t) \sim Gau(0, \sigma^2)$$
(16)

$Y(s_i, t)$ - Spatially Correlated Latent Gaussian

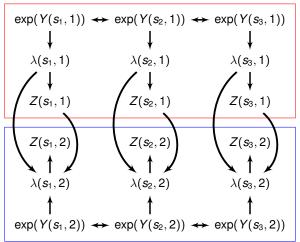
Time Period 1



SPINGARCH(1,1) Model

Spatially Correlated INGARCH(1,1) Model

Time Period 1



SPINGARCH Model

Theory: Exist common cause between geographically similar locations, regions that experience uptick in violence likely to have short term self-excitation, **absence of violence or exogeneous effects reduces tension**

SPINGARCH Model

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• $Z(s_i,t)|Y(s_i,t),\mathcal{H}_{Z(s_i)}\sim \text{Pois}\ (\lambda(s_i,t))$ where $\mathcal{H}_{Z(s_i)}$ is history of process at location s_i

$$\lambda(\mathbf{s}_i, t) = \exp\left[Y(\mathbf{s}_i, t)\right] + \eta Z(\mathbf{s}_i, t - 1) + \kappa \lambda(\mathbf{s}_i, t - 1)$$
(17)

• Define: $N_i = \{s_j : s_j \text{ is a spatial neighbor of } s_i\}$

$$Y(s_i, t)|Y(N_i) \sim N(\mu(s_i, t), \sigma_{sp}^2)$$

$$\mu(s_i, t) = \alpha(s_i) + \zeta \sum_{s_i \in N_i} \{Y(s_j, t) - \alpha(s_j)\}.$$
(18)

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(18)

• $\eta=0, \kappa=0$ Poisson - CAR, $\sigma_{sp}^2 \to 0$, INGARCH(1,1)/Short model

SPINGARCH Model as Stochastic Difference Equation

$$\frac{\lambda(s_i,t)-\lambda(s_i,t-1)}{\Delta t}=d-\chi\lambda(s_i,t-1)+\eta Z(s_i,t-1)$$
(19)

- Change in violence due to exogeneous d, natural decay, χ , and excitement, η
- Assume each time period, exogeneous impact is stochastic and spatially correlated yields SPINGARCH

$$\frac{\lambda(s_i,t) - \lambda(s_i,t-1)}{\Delta t} = \exp(Y(s_i,t)) - \chi\lambda(s_i,t-1) + \eta Z(s_i,t-1)$$
(20)

 Change in intensity due to three factors, CAR, natural decay, and excitement

SPINGARCH Model - Parameter Space

• $Z(s_i, t)|Y(s_i, t), \mathcal{H}_{Z(s_i)} \sim \text{Pois}(\lambda(s_i, t))$ where $\mathcal{H}_{Z(s_i)}$ is history of process at location s_i

$$\lambda(s_i, t) = \exp[Y(s_i, t)] + \eta Z(s_i, t - 1) + \kappa \lambda(s_i, t - 1)$$

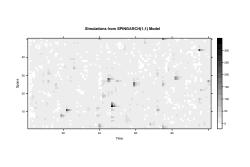
$$Y(s_i, t) | Y(N_i) \sim N(\mu(s_i, t), \sigma_{sp}^2)$$

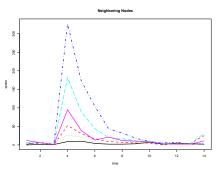
$$\mu(s_i, t) = \alpha(s_i) + \zeta \sum_{s_i \in N_i} \{Y(s_i, t) - \alpha(s_i)\}.$$

- $\zeta \in (\psi_{(1)}^{-1}, \psi_{(n)}^{-1})$ where $\psi_{(i)}$ is the *i*th largest eigenvalue of adjacency matrix
- For stationarity, $\eta > 0$, $\kappa > 0$, $\eta + \kappa < 1$

Data Realizations

50 Spatial Observations on \mathbb{R}^1 , 100 Temporal Observations





$$\lambda(s_{i},t) = \exp[Y(s_{i},t)] + 0.1 Z(s_{i},t-1) + 0.4 \lambda(s_{i},t-1)$$

$$Y(s_{i},t)|Y(N_{i}) \sim \text{Gau}(\mu(s_{i},t),0.5)$$

$$\mu(s_{i},t) = 0 + 0.49 \sum_{s_{j} \in N_{i}} \{Y(s_{j},t)\}.$$
(21)

SPINGARCH Model as Markov Chain

• Let $\lambda_t = (\lambda(s_1, t), \lambda(s_2, t), \cdots, \lambda(s_{n_d}, t))^T$

$$Z(s_{i}, t)|\lambda(s_{i}, t) \sim \mathsf{Pois}(\lambda(s_{i}, t))$$

$$E[Z(s_{i}, t)] = \lambda(s_{i}, t)$$

$$\lambda_{t} = \exp(Y_{t}) + \eta Z_{t-1} + \kappa \lambda_{t-1}$$

$$Y_{t} \sim \mathsf{Gau}(\alpha_{t}, (I_{n_{d}, n_{d}} - \mathbf{C})^{-1}\mathbf{M})$$
(22)

• Markov chain for λ_t on State space, $(\mathbb{R}^+)^{n_d}$

Impact of Initial Condtions and Recursion

By recursion

$$\begin{aligned} &[\lambda(s_{i},t)|\lambda(s_{i},0) = B] = \exp(Y(s_{i},t)) + \kappa\lambda(s_{i},t-1) + \eta Z(s_{i},t-1) \\ &= \exp(Y(s_{i},t)) + \kappa\left[\exp(Y(s_{i},t-1)) + \kappa\lambda(s_{i},t-2) + \eta Z(s_{i},t-2)\right] + \eta Z(s_{i},t-1) \\ &\cdots \\ &= \sum_{i=1}^{t-1} \kappa^{k} \exp(Y(s_{i},t-k)) + \sum_{i=1}^{t-1} \kappa^{k} \pi^{2} Z(s_{i},t-k-1) + \kappa^{t} B \end{aligned}$$

Geometric Ergodicity with Finite Moments

Under the parameter space restriction, $\eta, \kappa \geq 0$ and $\eta + \kappa < 1$, the SPINGARCH (1,1) is geometrically ergodic and admits a unique stationary distributions that has finite first two moments.

Sketch of Proof.

Meyn and Tweedie (15.0.1) need to show aperiodic, ϕ -irreducible and \exists test function V(.) such that $E[V(\lambda_{t+1})|\lambda_t = \lambda_*] \leq \psi V(\lambda_*) + L I(\lambda_* \in C)$ holds where $\psi \in (0,1)$,

 $E[V(\lambda_{t+1})|\lambda_t = \lambda_*] \le \psi V(\lambda_*) + L I(\lambda_* \in C)$ holds where $\psi \in (0,1)$, $L \in (0,\infty)$ and I(.) is the indicator function and C is a petite set.

Basic Idea: With positive probability, \exists a realization $Z(s_i,1)=Z(s_i,2)=\cdots=Z(s_i,t-1)=0$. Along that chain, $P(\lambda(s_i,t))\in A=P(\exp(Y(s_i,t))\in A-\kappa^t B)$. If $\kappa^T B>\sup A$ run chain longer.

Geometric Ergodicity with Finite Moments

Under the parameter space restriction, $\eta, \kappa \geq 0$ and $\eta + \kappa < 1$, the SPINGARCH (1,1) is geometrically ergodic and admits a unique stationary distributions that has finite first two moments.

Sketch of Proof Cont.

Test function $V(\lambda) = 1 + \lambda^2$ works for $E[V(\lambda_{t+1})|\lambda_t = \lambda_*] \le \psi V(\lambda_*) + L I(\lambda_* \in C)$.

 \implies Unique stationary distribution, goes to geometrically fast. Specific choice of V(.) gives (at least) finite first two moments (can be extended likely as in Fokianos, 2009.)

$$egin{aligned} & Z(oldsymbol{s}_i,t) \sim \operatorname{Pois}\left(\lambda(oldsymbol{s}_i,t)
ight) \ & \lambda(oldsymbol{s}_i,t) = \exp\left[Y(oldsymbol{s}_i,t)
ight] + \eta Z(oldsymbol{s}_i,t-1) + \kappa E\left[Z(oldsymbol{s}_i,t-1)
ight] \ & Y(oldsymbol{s}_i,t) | oldsymbol{Y}(oldsymbol{N}_i) \sim N(\mu(oldsymbol{s}_i,t),\sigma_{\operatorname{sp}}^2) \ & \mu(oldsymbol{s}_i,t) = lpha(oldsymbol{s}_i) + \zeta \sum_{oldsymbol{s}_j \in \mathcal{N}_i} \{Y(oldsymbol{s}_j,t) - lpha(oldsymbol{s}_j)\} \end{aligned}$$

Define
$$\Sigma_{i,j}$$
 as i,j entry of $(I_{n_d,n_d} - \boldsymbol{C})^{-1}\boldsymbol{M}$

$$E[Z(s_i,t)] = \frac{1}{1-\eta-\kappa} \exp(\alpha + \frac{\Sigma_{1,1}}{2})$$
 (23)

$$Var(Z(s_i, t)) = \frac{1}{1 - (\kappa + \eta)^2} Var(\exp(Y(s_i, t))) + \frac{1 - \kappa^2 - 2\kappa\eta}{1 - (\kappa + \eta)^2} E(Z(s_i, t))$$

Temporal Covariance:

$$Cov (Z(s_i, t), Z(s_i, t - 1) = (\eta + \kappa) Var(Z(s_i, t)) - \kappa E[Z(s_i, t)]$$
 (25)

Temporal Covariance:

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Recall: Var-Mean Ratio at 2 \implies Lag-1 correlation is between 0.5 and $\sqrt{1/2}$ for INGARCH(1,1)

Temporal Covariance:

Cov
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Recall: Var-Mean Ratio at 2 \implies Lag-1 correlation is between 0.5 and $\sqrt{1/2}$ for INGARCH(1,1)

Let $\kappa = 0$ (SPINGARCH(0,1)), Var-Mean Ratio at 2

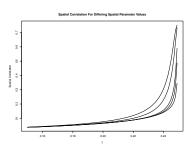
$$\implies 2 = \frac{\text{Var}(\exp(Y(s_i, t)))}{(1 - \eta)^2 E[Z(s_i, t)]} + \frac{1}{1 - \eta^2}$$
 (26)

$$Cor (Z(s_i, t), Z(s_i, t - 1)) = \eta$$
(27)

 $\forall \eta \in (0, \sqrt{1/2}) \quad \exists \alpha, \sigma_{sp}^2 \text{ such that equality holds}$

Spatial Correlation

$$\operatorname{Corr}(Z(s_i,t),Z(s_j,t)) = \frac{\left(\exp(\Sigma_{i,i}+\Sigma_{i,j})-\exp(\Sigma_{i,i})\right)}{\exp(2\Sigma_{i,i})-\exp(\Sigma_{i,i})+\exp(-\alpha+\frac{\Sigma_{i,i}}{2})\frac{1}{1-(\kappa+\eta)}}$$



$$\eta = .3$$
, $\sigma_{sp}^2 = .5$, 4 × 4 to 15 × 15 size lattice



Inference

- Likelihood roots for INGARCH(1,1) easily found, asymptotically Gaussian
- Inclusion of latent process in SPINGARCH(1,1) complicates
- $\bullet \ \theta \equiv (\eta, \alpha, \zeta, \sigma_{sp}^2)$

$$\pi(\theta|\mathbf{Z},\mathbf{Y}) \propto \prod_{t} \pi(\mathbf{Z}_{t}|\lambda_{t})\pi(\lambda_{t}|\lambda_{t-1},\mathbf{Z}_{t-1},\theta,\mathbf{Y}_{t})\pi(\mathbf{Y}_{t}|\theta)\pi(\theta)$$
 (28)

$$\pi(\mathbf{Y}|\mathbf{Z},\theta) \propto \prod_{t} \pi(\mathbf{Z}_{t}|\lambda_{t}) \pi(\lambda_{t}|\lambda_{t-1},\mathbf{Z}_{t-1},\theta,\mathbf{Y}_{t}) \pi(\mathbf{Y}_{t}|\theta).$$
 (29)

Efficient Bayesian Inference

$$\log(\mathbf{Y}|\alpha, \sigma_{sp}, \zeta) \propto \frac{1}{2} \log |\Sigma_f^{-1}(\theta)| - \frac{1}{2} (Y - \alpha)^T \Sigma_f^{-1}(\theta) (Y - \alpha),$$
(30)

- $\bullet \ \Sigma_t^{-1} \equiv \left(I_{n_d \times T, n_d \times T} I_{t,t} \otimes \boldsymbol{C}\right)^{-1} I_{t,t} \otimes \boldsymbol{M}$
- $\log |\Sigma^{-1}(\theta)| = \frac{n_d}{2\log \sigma_{sp}^2} + \log |I_{n_d,n_d} \zeta N|$
- Letting $V \wedge V^T$ be the spectral decomposition of N we have $|I_{n_d,n_d} \zeta N| = |V| |I_{n_d,n_d} \zeta \Lambda| |V^T| = \prod_{j=1}^{n_d} (1 \zeta \chi_j)$ where χ_j are the eigenvalues of the neighborhood matrix

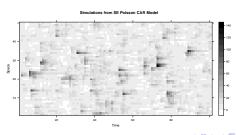
$$\log|\Sigma_f^{-1}(\theta)| = T\log|\Sigma^{-1}(\theta)| \tag{31}$$

$$\propto \frac{n_d \times T}{\log \sigma_{sp}^2} + T \sum_{j=1}^{n_d} (1 - \zeta \chi_j)$$
 (32)

Simulation and Estimation

$$Z(s_{i}, t) \sim \text{Pois}(\lambda(s_{i}, t))$$

 $\lambda(s_{i}, t) = \exp[Y(s_{i}, t)] + 0.66Z(s_{i}, t - 1)$
 $Y(s_{i}, t)|Y(N_{i}) \sim \text{Gau}(\mu(s_{i}, t), 0.5)$
 $\mu(s_{i}, t) = 0 + 0.49 \sum_{s_{j} \in N_{i}} \{Y(s_{j}, t)\}.$ (33)

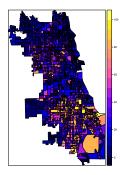


Simulation and Estimation

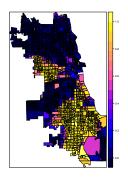
- SPINGARCH(0,1) 95% credible intervals: $\alpha \in (-0.24, 0.1)$, $\sigma^2 \in (0.46, 0.59)$, $\zeta \in (0.486, 0.492)$, and $\eta \in (0.64, 0.66)$
- SPINGARCH(1,0) 95% credible intervals: $\alpha \in (-0.54, -0.2)$, $\sigma^2 \in (0.96, 1.2)$, $\zeta \in (0.47, 0.48)$, and $\kappa \in (0.65, 0.67)$

	SPINGARCH(1,0)	SPINGARCH(0,1)
p ₁ - Moran's I	.05	.46
p_2 - Var to Mean	.99	.65
p₃ - Lag 1 Corr	.45	.7

Burglaries in South Side of Chicago

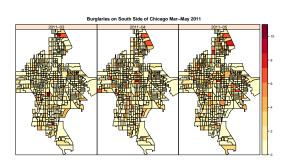


Aggregated Burglaries



Racial Segregation

Burglaries South Side of Chicago



- Crime data from city of Chicago
- 72 months (2010-2015), 552 locations (Census block groups)
- Demographic data from Census bureau

SPINGARCH(1,1) Model

$$\begin{split} & Z(\boldsymbol{s}_i,t) \sim \mathsf{Pois}(\lambda(\boldsymbol{s}_i,t)) \\ & E[Z(\boldsymbol{s}_i,t)] = \lambda(\boldsymbol{s}_i,t) \\ & \lambda_t = \exp(\boldsymbol{Y_t} + \boldsymbol{U}) + \eta \boldsymbol{Z_{t-1}} + \kappa \lambda_{t-1} \\ & \boldsymbol{Y_t} \sim \mathsf{Gau}(\boldsymbol{\alpha}, \sigma_{ind}^2 \boldsymbol{I}_{n_d,n_d}) \\ & \boldsymbol{U} \sim \mathsf{Gau}(\boldsymbol{0}, [\sigma_{sp}^2 (\boldsymbol{N} - \boldsymbol{C})]^{-1}) \end{split}$$

- Removed temporal trend and seasonality
- ullet Single spatial effect, ζ fixed near edge of parameter space
- ullet Additional small scale effect captured in σ_{ind}^2

$$\alpha_{s_i} = \exp\left(\beta_0 + \beta_{pop} \log(\mathsf{Pop}_{s_i}) + \beta_{ym} \mathsf{Young} \, \mathsf{Men}_{s_i} + \beta_{wealth} \mathsf{Wealth}_{s_i} + \beta_{unemp} \mathsf{Unemp}_{s_i}\right)$$
(35)

Impacts of Including Spatial Correlation

Parameter	SPINGARCH(1,1)	INGARCH(1,1)	
β_{0}	(-3.3,-1.0)	(-4.2,-3.4)	
$eta_{ extsf{pop}}$	(0.11,0.34)	(0.33, 0.46)	
$eta_{ extsf{ym}}$	(-0.75, 0.17)	(0.06, 0.09)	
$eta_{ extbf{wealth}}$	(0.05, 0.16)	(-0.04, 0.01)	
etaunemp	(0.006, 0.07)	(0.002,0.03)	
η	(0.04, 0.07)	(0.22, 0.24)	
κ	(0.31,0.39)	(0.44, 0.48)	
σ_{sp}^2	(0.40,0.54)	-	
$\sigma_{ind}^{2^{r}}$	(0.40,0.47)	-	

Model Assessment - Posterior Predictive Checks

	SPINGARCH(1,1)	INGARCH(1,1)
p_1 - Moran's I Statistic	0.43	0
p_2 - Variance to Mean Ratio	0.62	0
p_3 - Lag 1 Auto Correlation	0.67	0.74

- SPINGARCH(1,1) observed maximum (p=.67), number of zeros (p=.49)
- Conclusions on repeat-victimization

Summary

- Addition of spatially correlated effects naturally extends INGARCH model
- Both Poisson-CAR and INGARCH arise from SPINGARCH in limit
- Failure to specify random structure may result in differing conclusions
- Precomputing eigenvalues allows for relatively efficient Bayesian inference (8 hrs for 552 spatial locations, 104 time locations)

Future Work

- Impacts of aggregation
- Laplace approximations greatly speed up SPINGARCH(0,1) Can extend to SPINGARCH(1,1)?
- Reaction Diffusion Self-Exciting Model from (Clark & Dixon, 2018) does not fit in framework (temporally correlated errors)
 - RDSEM captures reaction diffusion process of Short in Latent Process
- Slides and code available at https://github.com/nick3703/Talk
- Questions?