

# Supplementary Materials for *GLS Kernel Regression for Network-Structured Data*

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## I. TABLE OF IDENTITIES

The following is a table of metamathematical identities used in various places within this paper to derive expressions. The equations listed in this section are considered well-known or easily verifiable, and are therefore provided without proof. The majority can be found in standard references such as [1] and [2].

TABLE I

A SERIES OF IDENTITIES INVOLVING MATRICES, KRONECKER PRODUCTS AND VECTORIZATION.

	Expression	Alternative expression
1	$\text{vec}(AXB)$	$(B^\top \otimes A) \text{vec}(X)$
2	$\text{tr}(A^\top B)$	$\text{vec}(A)^\top \text{vec}(B)$
3	$AC \otimes BD$	$(A \otimes B)(C \otimes D)$
4	$(A \otimes B)^{-1}$	$A^{-1} \otimes B^{-1}$
5	$\text{tr}(X^\top AYB)$	$\text{vec}(X)^\top (B^\top \otimes A) \text{vec}(Y)$
6	$\text{vec}(J \circ Y)$	$\text{diag}(\text{vec}(J)) \text{vec}(Y)$
7	$\ln(\det(A \otimes B))$	$\text{rank}(B) \ln  A  + \text{rank}(A) \ln  B $
8	$(B^\top B + A^{-1})^{-1} B^\top$	$A^{-1} B^\top (I + BAB^\top)^{-1}$
9	$\text{diag}^{-1}(A \text{diag}(\mathbf{x}) B)$	$(B^\top \circ A) \mathbf{x}$

TABLE II

THE MATRIX DERIVATIVE OF SEVERAL USEFUL SCALAR FUNCTIONS.

	$f(X)$	$\partial f / \partial X$
1	$\text{tr}((A - BX)^\top (A - BX))$	$-2B^\top A + 2B^\top BX$
2	$\text{tr}(XAX^\top)$	$X(A + A^\top)$
3	$\text{tr}(AXB)$	$A^\top B^\top$
4	$\text{tr}(AXBX^\top C)$	$A^\top C^\top XB^\top + CAXB$
5	$\text{tr}(AX^{-1}B)$	$-(X^{-1}BAX^{-1})^\top$
6	$\ln  \det(X) $	$X^{-\top}$

## II. PROOFS AND DERIVATIONS

**Theorem II.1.** *For a cost function*

$$\xi(F) = \text{tr}((Y - S_N F S_T^\top)^\top (Y - S_N F S_T^\top)) + \gamma \text{tr}(K^{-1} F^\top H^{-2} F)$$

*the minimising value of  $F$  is given by*

$$F^* = H^2 S_N^\top \bar{U} (J \circ (\bar{U}^\top Y \bar{V})) \bar{V}^\top S_T K$$

where

$$S_T^\top K S_T = \bar{V} \bar{\Lambda}_K \bar{V}^\top, \quad \text{and} \quad S_N^\top H^2 S_N = \bar{U} \bar{\Lambda}_H \bar{U}^\top$$

and  $J \in \mathbb{R}^{N' \times T'}$  has the elements given by

$$J_{ij} = \frac{1}{\bar{\lambda}_j^K \bar{\lambda}_i^H + \gamma}$$

*Proof.* The derivative of  $\xi$  with respect to  $F$  is given by

$$\frac{\partial \xi}{\partial F} = 2 S_N^\top (S_N F S_T^\top - Y) S_T^\top + 2 \gamma H^{-2} F K^{-1}$$

Setting this equal to zero and rearranging gives an equation for  $F^*$ .

$$S_N^\top S_N F^* S_T^\top S_T + \gamma H^{-2} F^* K^{-1} = S_N Y S_T^\top$$

Vectorizing both sides and applying identity 1 from table 1 gives:

$$\text{vec}(F^*) = \left( (S_T^\top S_T \otimes S_N^\top S_N + \gamma K^{-1} \otimes H^{-2}) \right)^{-1} (S_T^\top \otimes S_N^\top) \text{vec}(Y)$$

Applying identity 8 from table 1 means this can be rewritten as follows:

$$= (K S_T \otimes H^2 S_N) \times \left( \gamma I_{T'} \otimes I_{N'} + S_T^\top K S_T \otimes S_N^\top H^2 S_N \right)^{-1} \text{vec}(Y)$$

Next, substitute in the eigenvalue decompositions.

$$\begin{aligned} &= (K S_T^\top \otimes H^2 S_N^\top) \times \left( \gamma I_{T'} \otimes I_{N'} + \bar{V} \bar{\Lambda}_K \bar{V}^\top \otimes \bar{U} \bar{\Lambda}_H \bar{U}^\top \right)^{-1} \text{vec}(Y) \\ &= (K S_T^\top \otimes H^2 S_N^\top) \times \left( (\bar{V} \otimes \bar{U}) (\gamma I_{T'} \otimes I_{N'} + \bar{\Lambda}_K \otimes \bar{\Lambda}_H) (\bar{V}^\top \otimes \bar{U}^\top) \right)^{-1} \text{vec}(Y) \\ &= (K S_T^\top \otimes H^2 S_N^\top) (\bar{V} \otimes \bar{U}) \times (\gamma I_{T'} \otimes I_{N'} + \bar{\Lambda}_K \otimes \bar{\Lambda}_H) (\bar{V}^\top \otimes \bar{U}^\top) \text{vec}(Y) \end{aligned}$$

Note that the matrix  $(\gamma I_{T'} \otimes I_{N'} + \bar{\Lambda}_K \otimes \bar{\Lambda}_H)$  is diagonal. This means it can be rewritten as  $\text{diag}(\text{vec}(J))$  where  $J$  is defined above.

$$= (K S_T^\top \bar{V} \otimes H^2 S_N^\top \bar{U}) \text{diag}(\text{vec}(J)) \text{vec}(\bar{U}^\top Y \bar{V})$$

Finally, applying identities 1 and 6 gives the desired expression.  $\square$

**Theorem II.2.** *For a cost function*

$$\xi(F) = \text{tr}((Y - S_N F S_T^\top)^\top \Sigma_N^{-1} (Y - S_N F S_T^\top) \Sigma_T^{-1}) + \gamma \text{tr}(K^{-1} F^\top H^{-2} F)$$

The minimising value for  $\Psi$  can be expressed as

$$F^* = B(J \circ \bar{Y}) C^\top$$

where

$$\Sigma_T = \Psi \Lambda_{\Sigma_T} \Psi^\top, \quad \Sigma_N = \Phi \Lambda_{\Sigma_N} \Phi^\top$$

$$\Lambda_{\Sigma_T}^{-1/2} \Psi^\top S_T K S_T^\top \Psi \Lambda_{\Sigma_T}^{-1/2} = \bar{V} \bar{\Lambda}_K \bar{V}^\top$$

$$\Lambda_{\Sigma_N}^{-1/2} \Phi^\top S_N H^2 S_N^\top \Phi \Lambda_{\Sigma_N}^{-1/2} = \bar{U} \bar{\Lambda}_H \bar{U}^\top$$

and

$$J_{ij} = \frac{1}{\gamma + \bar{\lambda}_j^K \bar{\lambda}_i^H}$$

$$B = H^2 S_N^\top \Phi \Lambda_{\Sigma_N}^{-1/2} \bar{U}$$

$$C = K S_T^\top \Psi \Lambda_{\Sigma_T}^{-1/2} \bar{V}$$

$$\bar{Y} = \bar{U}^\top \Lambda_{\Sigma_N}^{-1/2} \Phi^\top Y \Psi \Lambda_{\Sigma_T}^{-1/2} \bar{V}$$

*Proof.* : Taking the derivative of the cost function and setting the result to zero in a similar fashion to the previous proof results in the following expression for  $F^*$ :  $\square$

$$\text{vec}(F^*) = \left( S_T^\top \Sigma_T^{-1} S_T \otimes S_N^\top \Sigma_N^{-1} S_N + \gamma K^{-1} \otimes H^{-2} \right)^{-1} (S_T^\top \Sigma_T^{-1} \otimes S_N^\top \Sigma_N^{-1}) \text{vec}(Y)$$

Substituting in the eigenvalue decompositions for the covariance matrices  $\Sigma_N$  and  $\Sigma_T$  gives

$$= \left( S_T^\top \Psi \Lambda_{\Sigma_T}^{-1} \Psi^\top S_T \otimes S_N^\top \Phi \Lambda_{\Sigma_N}^{-1} \Phi^\top S_N + \gamma K^{-1} \otimes H^{-2} \right)^{-1} (S_T^\top \Psi \Lambda_{\Sigma_T}^{-1} \Psi^\top \otimes S_N^\top \Phi \Lambda_{\Sigma_N}^{-1} \Phi^\top) \text{vec}(Y)$$

Applying identity 8 from table 1 and substituting the second set of eigenvalue decompositions gives

$$\begin{aligned} &= (K S_T^\top \Psi \Lambda_{\Sigma_T}^{-1/2} \otimes H^2 \Phi \Lambda_{\Sigma_N}^{-1/2}) \times \\ &\quad \left( \gamma I_{T'} \otimes I_{N'} + \bar{V} \bar{\Lambda}_K \bar{V}^\top \otimes \bar{U} \bar{\Lambda}_H \bar{U}^\top \right)^{-1} \times \\ &\quad (\Lambda_{\Sigma_T}^{-1/2} \Psi^\top \otimes \Lambda_{\Sigma_N}^{-1/2} \Phi^\top) \text{vec}(Y) \\ &= (K S_T^\top \Psi \Lambda_{\Sigma_T}^{-1/2} \bar{V} \otimes H^2 \Phi \Lambda_{\Sigma_N}^{-1/2} \bar{U}) \times \\ &\quad \left( \gamma I_{T'} \otimes I_{N'} + \bar{\Lambda}_K \otimes \bar{\Lambda}_H \right)^{-1} \times \\ &\quad (\bar{V}^\top \Lambda_{\Sigma_T}^{-1/2} \Psi^\top \otimes \bar{U}^\top \Lambda_{\Sigma_N}^{-1/2} \Phi^\top) \text{vec}(Y) \\ &= (K S_T^\top \Psi \Lambda_{\Sigma_T}^{-1/2} \bar{V} \otimes H^2 \Phi \Lambda_{\Sigma_N}^{-1/2} \bar{U}) \times \\ &\quad \text{diag}(\text{vec}(J)) \text{vec}(\bar{U}^\top \Lambda_{\Sigma_N}^{-1/2} \Phi^\top Y \Psi \Lambda_{\Sigma_T}^{-1/2} \bar{V}) \\ &= (C \otimes B) \text{diag}(\text{vec}(J)) \text{vec}(\bar{Y}) \end{aligned}$$

Finally, applying identities 1 and 6 completes the proof.

**Theorem II.3.** *The cost function*

$$\begin{aligned} \xi(\theta) &= N(T-1) \ln(1-\theta^2) \\ &\quad + \text{tr}(E \Sigma_T^{-1}(\theta) E^\top \Sigma_N) + \frac{NT\alpha}{1-\theta^2} \end{aligned}$$

is minimised by a value of  $\theta$  that is the solution to the following cubic polynomial:

$$\frac{b}{2} + (N(T-1) - a - c - NT\alpha)\theta + \frac{b}{2}\theta^2 - N(T-1)\theta^3$$

where

$$\begin{aligned} a &= \text{tr}(E^\top \Sigma_N^{-1} E), \quad b = \text{tr}(E^\top \Sigma_N^{-1} E B_1), \\ \text{and } c &= \text{tr}(E^\top \Sigma_N^{-1} E B_2) \end{aligned}$$

*Proof.* First expand  $\Sigma_T^{-1}$  in terms of  $\theta$ .

$$\begin{aligned} \xi(\theta) &= N(T-1) \ln(1-\theta^2) + \frac{\text{tr}(E^\top \Sigma_N E) + NT\alpha}{1-\theta^2} \\ &\quad - \frac{\text{tr}(E^\top \Sigma_N^{-1} E B_1)}{1-\theta^2} \theta + \frac{\text{tr}(E^\top \Sigma_N^{-1} E B_2)}{1-\theta^2} \theta^2 \end{aligned}$$

Using the substitutions outlined previously, this simplifies.

$$\xi(\theta) = N(T-1) \ln(1-\theta^2) + \frac{a + NT\alpha - b\theta + c\theta^2}{1-\theta^2}$$

Taking the derivative with respect to  $\theta$  gives

$$\frac{\partial \xi}{\partial \theta} = -\frac{2\theta N(T-1)}{1-\theta^2} + \frac{2\theta(a + NT\alpha + c) - b(\theta^2 + 1)}{(1-\theta^2)^2}$$

Setting this equal to zero and rearranging completes the proof.

**Theorem II.4.** *Given a fitted GLS KGR model, a lower bound for the marginal variance of the prediction uncertainty for the latent signal  $F$  is given by*

$$\Omega_F = (\tilde{U}^{-\top} \circ (H^2 \tilde{U})) J (\tilde{V}^{-1} \circ (\tilde{V}^\top K))$$

where  $\circ$  is the Hadamard product,

$$S_T^\top \Sigma_T^{-1} S_T K = \tilde{V} \tilde{\Lambda}_K \tilde{V}^{-1}, \quad S_N^\top \Sigma_N^{-1} S_N H^2 = \tilde{U} \tilde{\Lambda}_H \tilde{U}^{-1}$$

and

$$J_{ij} = \frac{1}{\gamma + \tilde{\lambda}_i^K \tilde{\lambda}_j^H}$$

*Proof.* The Laplace approximation states that the posterior over the latent signal can be approximated by a Gaussian, with covariance given by the inverse Hessian (Fisher information matrix) of the posterior negative log likelihood.

$$\Sigma_F = \left[ - \frac{\partial^2 \ln \pi(F|Y, X)}{\partial \text{vec}(F)_i \partial \text{vec}(F)_j} \Big|_{F=F^*} \right]^{-1}$$

To calculate this, first take the derivative of  $-\ln \pi(F|Y, X)$  with respect to  $\text{vec}(F)$ .

$$- \frac{\partial \ln \pi(F|Y, X)}{\partial \text{vec}(F)} = (S_T^\top \Sigma_T^{-1} S_T \otimes S_N^\top \Sigma_N^{-1} S_N + \gamma K^{-1} \otimes H^{-2}) \text{vec}(F) - \text{vec}(S_N^\top \Sigma_N^{-1} Y \Sigma_T^{-1} S_T)$$

Taking the Jacobian of this vector function of  $F$  gives the Hessian. From this, we can see that  $\Sigma_F$  is given by

$$\Sigma_F = (S_T^\top \Sigma_T^{-1} S_T \otimes S_N^\top \Sigma_N^{-1} S_N + \gamma K^{-1} \otimes H^{-2})^{-1}$$

First, factorise  $\Sigma_F$  as follows

$$= (K \otimes H^2) (S_T^\top \Sigma_T^{-1} S_T K \otimes S_N^\top \Sigma_N^{-1} S_N H^2 + \gamma I_T \otimes I_N)^{-1}$$

Then substitute in the eigenvalue decompositions as previously defined.

$$= (K \otimes H^2) (\tilde{V} \tilde{\Lambda}_K \tilde{V}^{-1} \otimes \tilde{U} \tilde{\Lambda}_H \tilde{U}^{-1} + \gamma I_T \otimes I_N)^{-1}$$

Next, factorise this expression further

$$= (K \otimes H^2) (\tilde{V} \otimes \tilde{U}) \times (\tilde{\Lambda}_K \otimes \tilde{\Lambda}_H + \gamma I_T \otimes I_N)^{-1} (\tilde{V}^{-1} \otimes \tilde{U}^{-1})$$

Next, substitute in the definition of  $J$ .

$$= (K \tilde{V} \otimes H^2 \tilde{U}) \text{diag}(\text{vec}(J)) (\tilde{V}^{-1} \otimes \tilde{U}^{-1})$$

Since we are interested in the marginal variances, i.e. the diagonal of this matrix, we would like to avoid holding the

full  $(NT \times NT)$  matrix in memory. In order to achieve this, first consider the fact that the above matrix can be expressed as

$$\Sigma_F = \sum_{t=1}^T K \tilde{V} \delta_t \tilde{V}^{-1} \otimes H^2 \tilde{U} \text{diag}(J_t) \tilde{U}^{-1}$$

where  $\delta_t$  is an empty  $T \times T$  matrix, except element  $(t, t)$ , which is 1, and  $J_t$  is the  $t$ th column of  $J$ . This is effectively splitting each of the  $T$  operations involved in the Kronecker-product operation.

Now consider the diagonals of the above matrices. The diagonal of  $K \tilde{V} \delta_t \tilde{V}^{-1}$  is the Hadamard product between the  $t$ th row of  $\tilde{V}^{-1}$  and the  $t$ th column of  $K \tilde{V}$ . Denote this as  $\mathbf{a}_t$ .

$$\mathbf{a}_t = (\tilde{V}^{-\top})_t \circ (K \tilde{V})_t$$

The diagonal of the second matrix,  $\mathbf{b}_t$ , is

$$\mathbf{b}_t = \text{diag}^{-1}(H^2 \tilde{U} \text{diag}(J_t) \tilde{U}^{-1})$$

where  $\text{diag}^{-1}(\cdot)$  means to take the diagonal of a matrix and turn it into a vector. Using identity 9 from table 1, we can rewrite  $\mathbf{b}_t$  as

$$\mathbf{b}_t = (\tilde{U}^{-\top} \circ (H^2 \tilde{U})) J_t$$

Therefore,  $\Omega_F$  can be expressed as

$$\begin{aligned} \Omega_F &= \text{mat} \left( \sum_{t=1}^T \mathbf{a}_t \otimes \mathbf{b}_t \right) \\ &= \sum_{t=1}^T \mathbf{b}_t \mathbf{a}_t^\top \\ &= B A^\top \\ &= (\tilde{U}^{-\top} \circ (H^2 \tilde{U})) J (\tilde{V}^{-1} \circ (\tilde{V}^\top K)) \end{aligned}$$

□

## REFERENCES

- [1] K. B. Petersen and M. S. Pedersen, *The Matrix Cookbook*, 2012.
- [2] T. P. Minka, "Old and New Matrix Algebra Useful for Statistics," Tech. Rep., 2001. [Online]. Available: <http://citeseerx.ist.psu.edu/viewdoc/summary?doi=10.1.1.115.5808>