

SYSEN 5200: Systems Analysis Behavior and Optimization

Monte Carlo Simulation

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1. Suppose that you want to generate a random sample from a p.d.f. with the following distribution:

$$f(x) = \begin{cases} \frac{3x^2}{8} & \text{for } 0 \leq x \leq 2 \\ 0 & \text{otherwise} \end{cases} \quad (1)$$

the equivalent c.d.f can be derived as:

$$\begin{aligned} F(x) &= \int_{-\infty}^x f(t) dt \\ &= \int_0^x \frac{3t^2}{8} dt \\ &= \left. \frac{t^3}{8} \right|_0^x \\ &= \frac{x^3}{8} \end{aligned} \quad (2)$$

Therefore, we have:

$$F(x) = \begin{cases} 0 & \text{for } x < 0 \\ \frac{x^3}{8} & \text{for } 0 \leq x \leq 2 \\ 1 & \text{for } x > 2 \end{cases} \quad (3)$$

when inverted is:

$$F^{-1}(y) = \begin{cases} 0 & \text{for } y \leq 0 \\ \sqrt[3]{8y} & \text{for } 0 < y \leq 1 \\ 2 & \text{for } y > 1 \end{cases} \quad (4)$$

where $g(y) = F^{-1}(y)$ that recovers the r.v. X with the desired p.d.f. $f(x)$ from $Y \sim U[0, 1]$.

2. Suppose that the lifetime of each component i is independent and follows a distribution that is the maximum of zero and a normal random variable with mean μ_i and standard deviation σ_i (both in minutes), whose values are given by: $\mu_1 = 60, \sigma_1 = 15; \mu_2 = 50, \sigma_2 = 10; \mu_3 = 75, \sigma_3 = 15; \mu_4 = 75, \sigma_4 = 20; \mu_5 = 60, \sigma_5 = 15; \mu_6 = 70, \sigma_6 = 15$.

- (a) Considering the structure function of the following system:

$$\phi(x) = \min(\max(\min(x_1, x_2), x_3), x_4, \max(x_5, x_6)) \quad (5)$$

- (b) A simulation was conducted with over 1,000 replications to estimate the probability that the system will be functional for more than 60 minutes with a 95% confidence interval.

```

1  ## libraries
2  import numpy as np
3
4  ## structure func
5  def phi(x):
6      return min(max(min(x[0], x[1]), x[2]), x[3], max(x[4], x[5]))
7
8  ## simulation
9  def simulate():
10     x = np.maximum(0, np.random.normal(
11         [60, 50, 75, 75, 60, 70], ## mu
12         [15, 10, 15, 20, 15, 15] ## sigma
13     )
14 )
15     return phi(x) > 60
16
17 ## compute
18 n = 1001
19 results = [simulate() for i in range(n)]
20 prob = sum(results) / n
21 lower = prob - 1.96 * np.sqrt(prob * (1 - prob) / n)
22 upper = prob + 1.96 * np.sqrt(prob * (1 - prob) / n)
23
24 ## results
25 print(f"Probability: {prob:.4f}")
26 print(f"95% CI: ({lower:.4f}, {upper:.4f})")

```

Fig. Python 2(b)

Probability: **0.5834**

95% CI: **(0.5529, 0.6140)**

- (c) Suppose that the error of the estimate is to be bounded by 0.01 with 0.95 probability. Using Chebyshev's inequality, we estimate the number of replications is sufficient to provide this guarantee as:

$$\begin{aligned}
 P(|\hat{p} - p| \geq 0.01) &\leq \frac{\text{Var}(\hat{p})}{0.01^2} = \frac{p(1-p)/n}{0.01^2} \\
 n &\geq \frac{p(1-p)}{(0.01^2 \cdot 0.05)} \\
 n &\geq \frac{0.5 \cdot 0.5}{(0.01^2 \cdot 0.05)} = 50,000
 \end{aligned} \quad (6)$$

Assuming that the approximation from Central Limit Theorem holds, we have:

$$n \geq \left(\frac{z_{\alpha/2} \sigma}{E} \right)^2 = \left(\frac{1.96}{0.01} \right)^2 \cdot 0.5 \cdot 0.5 \approx 9,604. \quad (7)$$

- (d) Another simulation was conducted with over 1,000 replications to estimate the expected lifetime of the system in minutes with a 95% confidence interval.

```

1  ## libraries
2  import numpy as np
3
4  ## structure func
5  def phi(x):
6      return min(max(min(x[0], x[1]), x[2]), x[3], max(x[4], x[5]))
7
8  ## simulation
9  def simulate():
10     x = np.maximum(0, np.random.normal(
11         [60, 50, 75, 75, 60, 70], ## mu
12         [15, 10, 15, 20, 15, 15] ## sigma
13     )
14 )
15     return phi(x)
16
17 ## compute
18 n = 1001
19 results = [simulate() for i in range(n)]
20 mean = np.mean(results)
21 lower = mean - 1.96 * np.std(results, ddof=1) / np.sqrt(n)
22 upper = mean + 1.96 * np.std(results, ddof=1) / np.sqrt(n)
23
24 ## results
25 print(f"Expected L: {mean:.2f} minutes")
26 print(f"95% CI: ({lower:.2f}, {upper:.2f}) minutes")

```

Fig. Python 2(d)

Expected L: **61.80** min.
95% CI: (**61.05**, **62.56**) min.

- (e) Suppose that the error of the estimate is to be bounded by 1 minute with 0.95 probability. Using Chebyshev's inequality, we estimate the number of replications is sufficient to provide this guarantee as:

$$n \geq \frac{\sigma_1^2 + \sigma_2^2 + \sigma_3^2 + \sigma_4^2 + \sigma_5^2 + \sigma_6^2}{0.05} = \frac{15^2 + 10^2 + 15^2 + 20^2 + 15^2 + 15^2}{0.05} = 2,843 \quad (8)$$

Assuming that the approximation from Central Limit Theorem holds, we have:

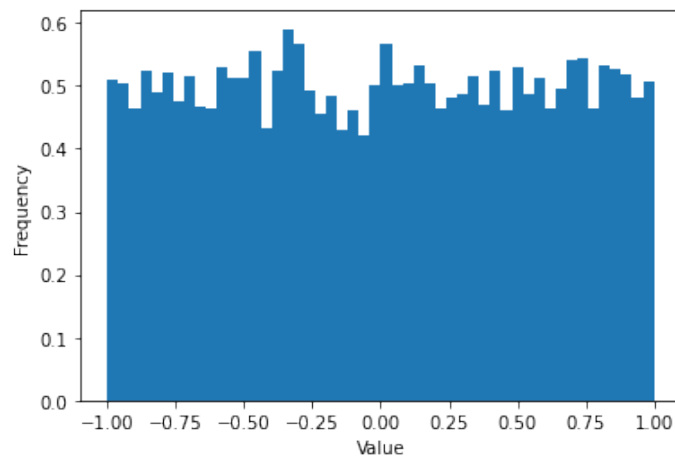
$$n \geq \left(\frac{z_{\alpha/2} \sigma}{E} \right)^2 = \left(\frac{1.96 \cdot 18.8}{1} \right)^2 \approx 546 \quad (9)$$

3. One application of Monte Carlo Simulation is estimating a difficult to evaluate integral. Here we will perform a simple simulation to estimate the area of a unit circle.

(a) First we will generate 10,000 points uniformly on $U[-1, 1]$ using 50 uniform-sized bins.

```
1  ## libraries
2  import numpy as np
3  import matplotlib.pyplot as plt
4
5  ## replicate
6  X = np.random.uniform(-1, 1, 10_000)
7
8  ## plot
9  plt.hist(
10     x = X,
11     bins = 50,
12     density = True
13 )
14
15 ## plot labels
16 plt.xlabel('Value')
17 plt.ylabel('Frequency')
18
19 ## plot preview
20 plt.show()
```

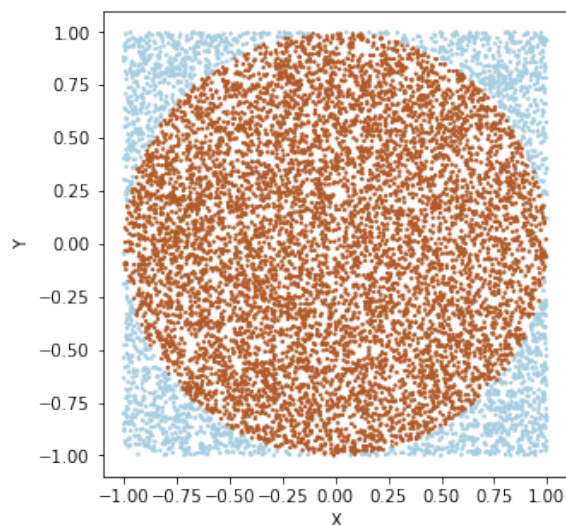
Fig. Python 3(a)



- (b) Then we generate 10,000 points uniformly distributed in a square around the origin with corners $(-1,-1)$, $(-1,1)$, $(1,-1)$, and $(1,1)$. We independently generate the X and Y -coordinates of a point to be $U[-1,1]$. Then we create an additional array Z of length 10,000, which indicates whether each random point lies within a circle of radius 1 centered at the origin.

```
1  ## libraries
2  import numpy as np
3  import matplotlib.pyplot as plt
4
5  ## replicate
6  X = np.random.uniform(-1, 1, 10_000)
7  Y = np.random.uniform(-1, 1, 10_000)
8
9  ## compute distance
10 R = np.sqrt(X**2 + Y**2)
11 Z = np.where(R <= 1, 1, 0)
12
13 ## plot
14 fig, ax = plt.subplots(figsize=(5, 5))
15 ax.scatter(
16     x = X,
17     y = Y,
18     c = Z,
19     s = 2.0,
20     cmap = plt.cm.Paired
21 )
22
23 ## plot labels
24 ax.set_xlabel('X')
25 ax.set_ylabel('Y')
26 ax.set_aspect('equal')
27
28 ## plot preview
29 plt.show()
```

Fig. Python 3(b)



- (c) Given n random points above, we can estimate π as $\hat{\pi}_n = 4N_{\text{circle}}/n$, where N_{circle} is the number of random points which lie within the circle. For n between 1 and 1000, we plot the relative error $|\pi_n - \pi|/\pi$, as well as the function $1/\sqrt{n}$.

```

1  ## libraries
2  import numpy as np
3  import matplotlib.pyplot as plt
4
5  ## pi
6  pi_true = np.pi
7
8  ## number of replications
9  ns = np.arange(1, 1001)
10
11 ## initialize estimations of pi and errors
12 pi_est = np.zeros_like(ns, dtype = float)
13 rel_err = np.zeros_like(ns, dtype = float)
14
15 ## replications and estimations
16 for i, n in enumerate(ns):
17     X = np.random.uniform(-1, 1, n)
18     Y = np.random.uniform(-1, 1, n)
19
20     ## compute distances
21     R = np.sqrt(X**2 + Y**2)
22     N_circle = np.count_nonzero(R <= 1)
23
24     ## estimate pi and compute error
25     pi_est[i] = 4 * N_circle / n
26     rel_err[i] = np.abs(pi_est[i] - pi_true) / pi_true
27
28 ## plot
29 fig, ax = plt.subplots()
30 ax.scatter(ns, rel_err, label = 'Relative Error', alpha = 0.5, s = 5)
31 ax.plot(ns, 1/np.sqrt(ns), label = '1 / sqrt(n)', color = 'orange')
32
33 ## plot labels
34 ax.set_xlabel('n')
35 ax.set_ylabel('Relative Error')
36 ax.legend()
37
38 ## plot preview
39 plt.show()

```

Fig. Python 3(c)

