

Standard deviation of the mean for finite samples

- When *measuring* the mean $\langle x \rangle = \frac{1}{N} \sum_{i=1}^N x_i$ of a random variable x governed by $P(x)$, we get an *estimate* of the real mean μ due to the *finite sample size* N .
- $\langle x \rangle \rightarrow \mu$ in the limit $N \rightarrow \infty$
- How good of an estimate is $\langle x \rangle$ for finite N ?
- Let's calculate this using the error propagation formula we derived in the last lecture.

Standard deviation of the mean

Let μ and σ_x be the true mean and standard deviation of $P(x)$.

We estimate μ by taking N samples and compute the mean

$$\langle x \rangle = \frac{1}{N} \sum_{i=1}^N x_i$$

What is the uncertainty in $\langle x \rangle$?

Let's repeat taking N measurements from $P(x)$ many times and compute $\langle x \rangle$ for each trial.

Standard deviation of the mean

- For every set of N measurements we will get a slightly different mean $\langle x \rangle_i$. The distribution of the means $P(\langle x \rangle)$ will be centred around the true mean μ .
- The standard deviation of $P(\langle x \rangle)$ will be **the standard error of the mean $\sigma_{\langle x \rangle}$** .
- We calculate $\sigma_{\langle x \rangle}$ using the error propagation formula for $f(x_1, x_2, \dots, x_N)$:

$$\sigma_f^2 = \left(\frac{\partial f}{\partial x_1} \right)^2 \sigma_{x_1}^2 + \left(\frac{\partial f}{\partial x_2} \right)^2 \sigma_{x_2}^2 + \dots + \left(\frac{\partial f}{\partial x_N} \right)^2 \sigma_{x_N}^2$$

Standard deviation of the mean

$$\langle x \rangle = \frac{1}{N} (x_1 + x_2 + \dots + x_N)$$

- Each x_i that appears in the sum is governed by $P(x)$ with a mean μ and standard deviation σ_x .
- $\langle x \rangle$ can be considered a function of N independent variables.
- Therefore,

$$\sigma_{\langle x \rangle} = \sqrt{\left(\frac{\partial \langle x \rangle}{\partial x_1}\right)^2 \sigma_{x_1}^2 + \left(\frac{\partial \langle x \rangle}{\partial x_2}\right)^2 \sigma_{x_2}^2 + \dots + \left(\frac{\partial \langle x \rangle}{\partial x_N}\right)^2 \sigma_{x_N}^2}$$

Standard deviation of the mean

- As the x_i are all governed by the same distribution $P(x)$ with standard deviation σ_x :

$$\sigma_{x_1} = \dots = \sigma_{x_N} = \sigma_x$$

- The partial derivatives in the sum are also the same

$$\frac{\partial \langle x \rangle}{\partial x_1} = \dots = \frac{\partial \langle x \rangle}{\partial x_N} = \frac{1}{N}$$

Standard deviation of the mean

Finally,

$$\sigma_{\langle x \rangle} = \sqrt{\left(\frac{\partial \langle x \rangle}{\partial x_1}\right)^2 \sigma_{x_1}^2 + \left(\frac{\partial \langle x \rangle}{\partial x_2}\right)^2 \sigma_{x_2}^2 + \dots + \left(\frac{\partial \langle x \rangle}{\partial x_N}\right)^2 \sigma_{x_N}^2}$$

$$\sigma_{\langle x \rangle} = \sqrt{\left(\frac{1}{N}\right)^2 \sigma_x^2 + \left(\frac{1}{N}\right)^2 \sigma_x^2 + \dots + \left(\frac{1}{N}\right)^2 \sigma_x^2} = \sqrt{N \cdot \left(\frac{1}{N}\right)^2 \sigma_x^2}$$

$$\sigma_{\langle x \rangle} = \frac{\sigma_x}{\sqrt{N}}$$

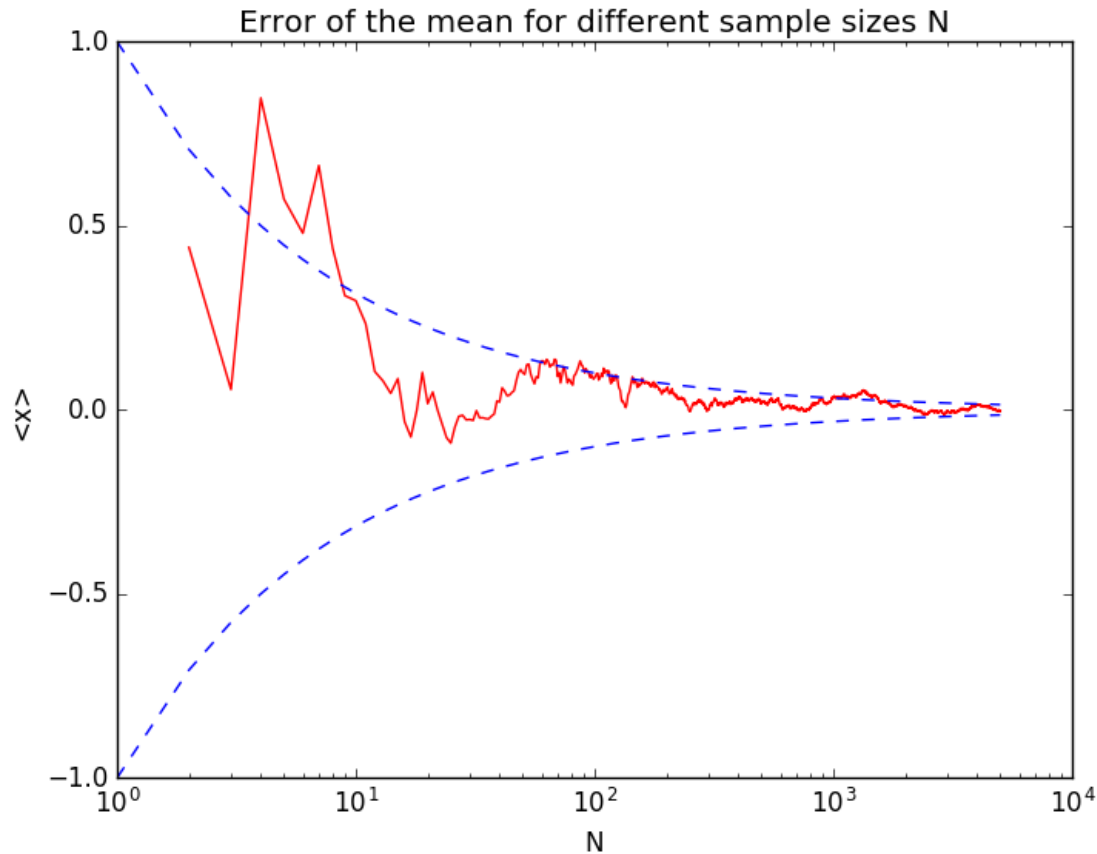
This is the **error in the mean** and reflects the uncertainty in the estimate of μ given by $\langle x \rangle$ measured from a finite sample size N

Remarks on the standard deviation of the mean

- $\sigma_{\langle x \rangle} = \frac{\sigma_x}{\sqrt{N}}$ reflects the uncertainty in the estimate of the mean.
- Error of the mean goes to zero as $N \rightarrow \infty$ as it should since $\langle x \rangle \rightarrow \mu$ in that limit.
- Don't confuse error of the mean $\sigma_{\langle x \rangle}$ with the standard deviation σ_x . The latter reflects the uncertainty in the measurements (i.e. error bars) which will not decrease as N increases, although larger N will give a better estimate of σ_x .
- Derivation does not assume any particular distribution $P(x)$.

Standard deviation of the mean

Plot shows $\langle x \rangle$ calculated from N x_i 's drawn from a Gaussian with $\mu = 0$ and $\sigma_x = 1$. Blue dotted line is $\pm\sigma_x/\sqrt{N}$.

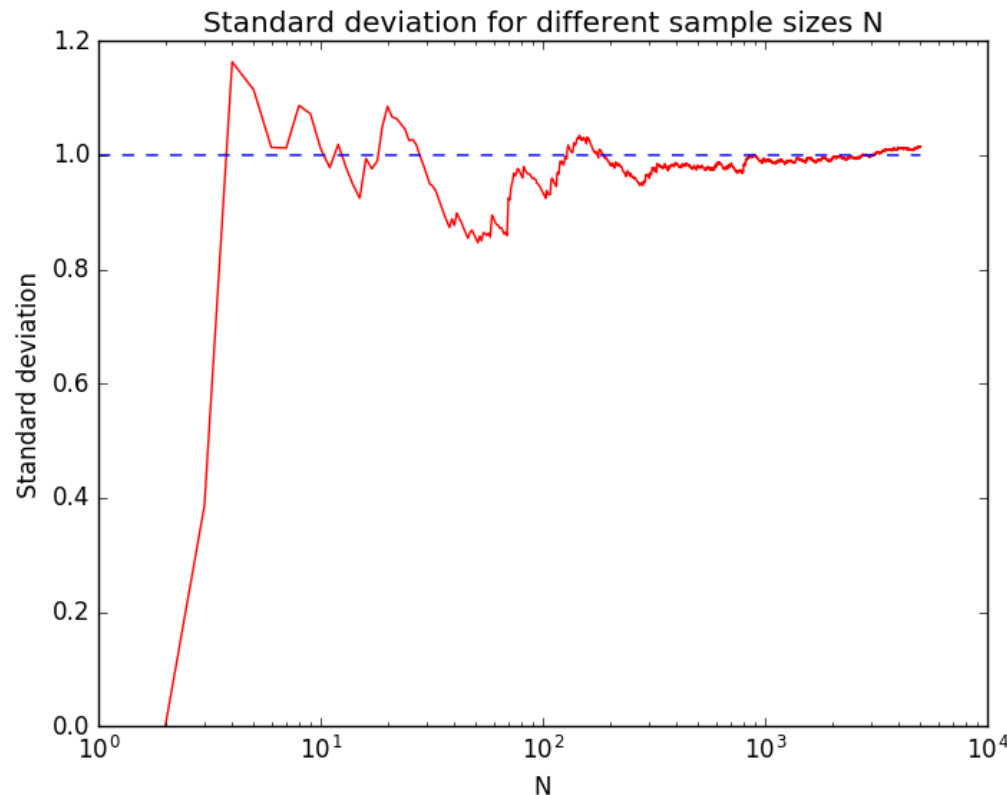


As N increases $\langle x \rangle$ gets closer to the true mean $\mu = 0$. $\langle x \rangle$ bounded by the error of the mean

Error of the standard deviation

Similar to the error of the mean, there is an error of the standard deviation ($= \sigma_x / \sqrt{2(N-1)}$).

$\sigma_{x,N}$ measured over a finite sample size N approaches true standard deviation (blue dotted line) in the limit $N \rightarrow \infty$



The Poisson process

Many counting experiments are governed by a Poisson process:

- the number of car accidents at a site or in an area
- the requests for individual documents on a web server
- Number of earthquakes in a certain area
- the number of deaths from horse kicks in the Prussian army.
- Incidence rate of rare diseases.
- Nuclear decay

What do all these processes have in common?

Conditions for a Poisson process

- Counting measurements.
- Events have low probability of occurring.
- Non-overlapping events.
- Constant long-time average rate.

Can be thought of as a binomial process in the limit of small success probability p and large N .

Recall binomial distribution: $B_{N,p}(x) = \binom{N}{x} p^x q^{N-x}$

With mean $\mu = Np$, and $\sigma = \sqrt{Np(1-p)}$.


Note $q = 1 - p$.

Poisson distribution - derivation

- Rewrite binomial distribution:

$$\begin{aligned} B_{N,p}(x) &= \binom{N}{x} p^x (1-p)^{N-x} \\ &= \frac{1}{x!} \frac{N!}{(N-x)!} p^x (1-p)^{-x} (1-p)^N \end{aligned}$$

Now take the limit of low success probability $p \ll 1$. This implies that the success probability $B_{N,p}(x)$ is close to zero for large x , the number of successful events. Therefore, $x \ll N$ in the region of interest which is the limit of low success rate.

$$\frac{N!}{(N-x)!} = N \cdot (N-1) \cdot \dots \cdot (N-x+1) \approx N^x$$


Poisson distribution - derivation

So the binomial distribution simplifies to

$$B_{N,p}(x) = \frac{1}{x!} N^x p^x (1-p)^{-x} (1-p)^N$$

$N^x p^x = (Np)^x = \mu^x$. Binomial expansion of fourth term yields

$$(1-p)^{-x} \approx 1 + px + \dots \approx 1 \text{ since } p \ll 1$$

Last term can be recast using $\mu = Np$:

$$(1-p)^N = \left[(1-p)^{\frac{1}{p}} \right]^\mu$$

Since $\lim_{p \rightarrow 0} (1+p)^{1/p} = \frac{1}{e}$, we finally obtain

$$P(x) = \frac{1}{x!} \mu^x e^{-\mu}$$

The Poisson distribution

$$P_{\mu}(x) = e^{-\mu} \frac{\mu^x}{x!}$$

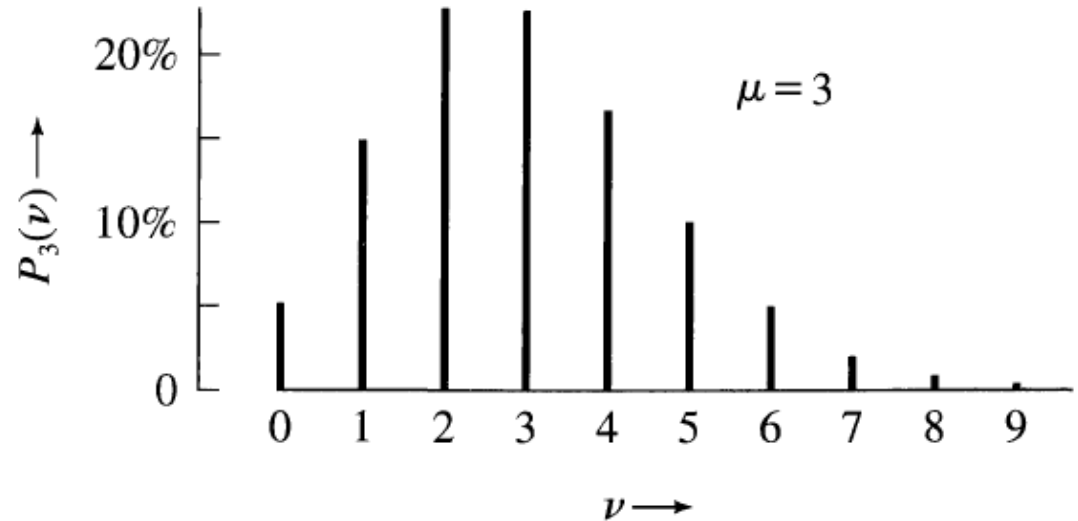
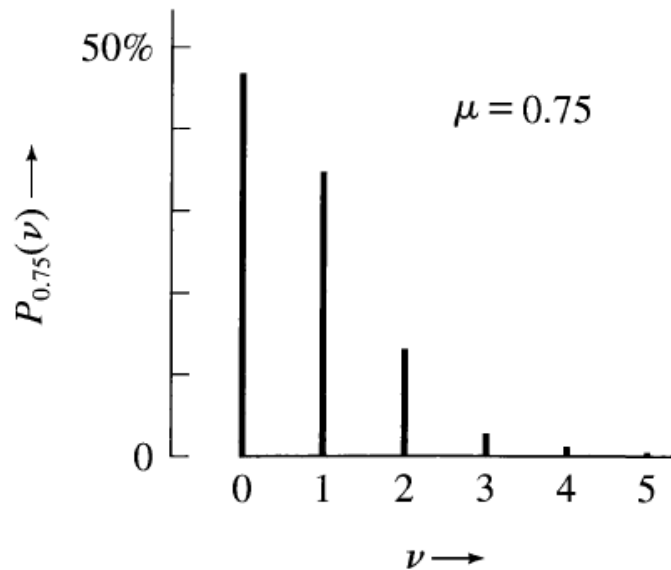
- Poisson distribution is the limit of a binomial distribution for small success probability p and large N .
- It is a **discrete** distribution.
- Poisson distribution only has *one* parameter – the mean μ .
- Standard deviation directly related to mean: $\sigma_x = \sqrt{\mu}$
- Similar to the binomial distribution, the Poisson distribution can be approximated by a Gaussian for large N . For the Poisson distribution this means large μ . (Recall binomial mean: $\mu = Np$)

$$\text{Gaussian: } G_{\mu,\sigma}(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\text{Binomial: } B_{N,p}(x) = \binom{N}{x} p^x q^{N-x} \approx \frac{1}{\sqrt{2\pi Npq}} e^{-\frac{(x-Np)^2}{2Npq}} \text{ for large } N$$

$$\text{Poisson: } P_{\mu}(x) = e^{-\mu} \frac{\mu^x}{x!} \approx \frac{1}{\sqrt{2\pi\mu}} e^{-\frac{(x-\mu)^2}{2\mu}} \text{ for large } \mu \text{ (good even for } \mu \gtrsim 10)$$

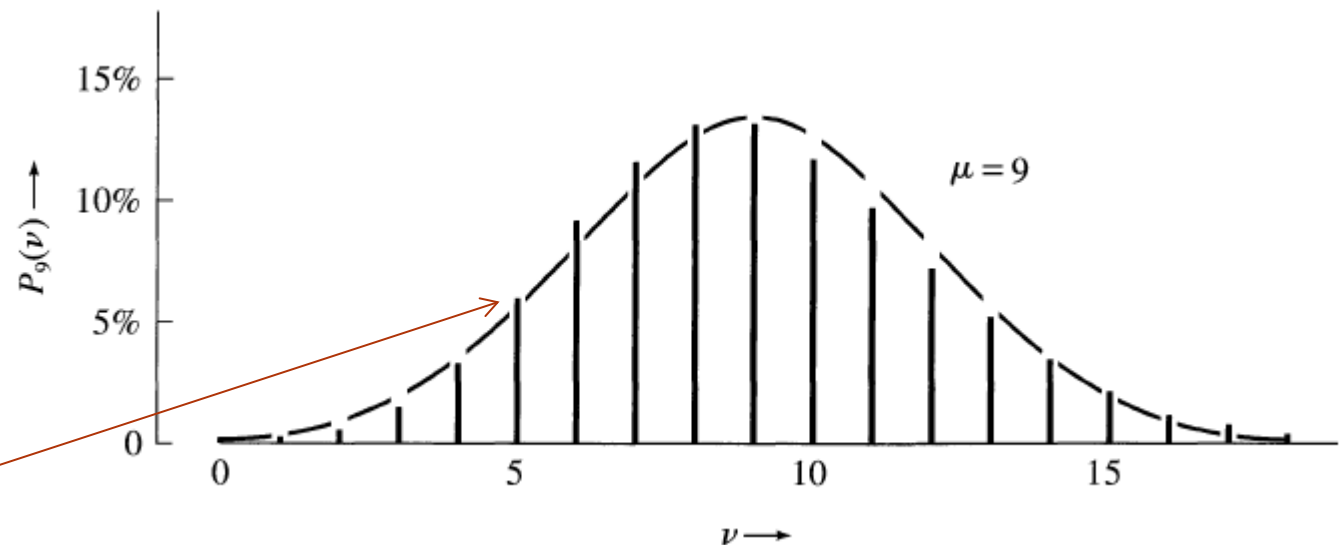
Poisson distributions for different μ 's



Asymmetric at low μ .

Becomes more bell shaped at larger μ

Gaussian approximation not bad even for $\mu = 9$



Poisson distribution - Normalisation

Check normalisation condition:

$$\begin{aligned}\sum_{x=0}^{\infty} P_{\mu}(x) &= \sum_{x=0}^{\infty} e^{-\mu} \frac{\mu^x}{x!} \\ &= e^{-\mu} \sum_{i=0}^{\infty} \frac{\mu^i}{i!} = e^{-\mu} \left(\frac{\mu^0}{0!} + \frac{\mu^1}{1!} + \frac{\mu^2}{2!} + \dots \right) \\ &= e^{-\mu} \left(\frac{\mu^0}{0!} + \frac{\mu^1}{1!} + \frac{\mu^2}{2!} + \dots \right) = e^{-\mu} \left(1 + \mu + \frac{\mu^2}{2!} + \dots \right)\end{aligned}$$

Note $x! = 0$. The term in the brackets is just a Taylor expansion of the exponential e^{μ} , so

$$\sum_{i=0}^{\infty} P_{\mu}(x) = e^{-\mu} e^{\mu} = 1$$

The mean of the Poisson distribution

Start from

$$\sum_{x=0}^{\infty} P_{\mu}(x) = \sum_{x=0}^{\infty} e^{-\mu} \frac{\mu^x}{x!} = 1$$

Differentiate w.r.t μ (use product rule):

$$\begin{aligned} \sum_{x=0}^{\infty} \left(-e^{-\mu} \frac{\mu^x}{x!} + x e^{-\mu} \frac{\mu^{x-1}}{x!} \right) &= 0 \\ - \sum_{x=0}^{\infty} e^{-\mu} \frac{\mu^x}{x!} + \sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^{x-1}}{x!} &= 0 \\ -1 + \sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^{x-1}}{x!} &= 0 \end{aligned}$$

In the last step we used the normalisation condition.

The mean of the Poisson distribution

$$\sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^{x-1}}{x!} = 1$$

Multiplying both sides with μ , we get

$$\sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^x}{x!} = \sum_{x=0}^{\infty} x P_{\mu}(x) = \langle x \rangle = \mu$$

So the mean of the Poisson distribution is just

$$\langle x \rangle = \mu$$

The standard deviation of the Poisson distribution

Use $\sigma_x = \sqrt{\langle x^2 \rangle - \langle x \rangle^2}$. Already derived $\langle x \rangle = \mu$.

To calculate second moment $\langle x^2 \rangle$, start with

$$\langle x \rangle = \sum_{x=0}^{\infty} x P_{\mu}(x) = \sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^x}{x!} = \mu$$

Differentiate w.r.t μ (product rule):

$$\sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^x}{x!} = \mu$$
$$\sum_{x=0}^{\infty} x^2 e^{-\mu} \frac{\mu^{x-1}}{x!} - \sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^x}{x!} = 1$$

Multiply both sides with μ

The standard deviation of the Poisson distribution

$$\begin{aligned}\sum_{x=0}^{\infty} x^2 e^{-\mu} \frac{\mu^x}{x!} - \mu \sum_{x=0}^{\infty} x e^{-\mu} \frac{\mu^x}{x!} &= \mu \\ \sum_{x=0}^{\infty} x^2 P_{\mu}(x) - \mu \sum_{x=0}^{\infty} x P_{\mu}(x) &= \mu \\ \langle x^2 \rangle - \mu \langle x \rangle &= \mu \\ \langle x^2 \rangle &= \mu^2 + \mu\end{aligned}$$

Therefore,

$$\sigma_x = \sqrt{\langle x^2 \rangle - \langle x \rangle^2} = \sqrt{\mu^2 + \mu - \mu^2} = \sqrt{\mu}$$

Note: Binomial $\sigma = \sqrt{Np(1-p)}$. For small p this reduces to $\sigma = \sqrt{Np} = \sqrt{\mu}$

$$\sigma_x = \sqrt{\mu}$$

Example: Nuclear decay

- The standard example in Physics for a Poisson process is nuclear decay. A radioactive sample may have 10^{20} nuclei (large N), each having a decay probability of down to $\sim 10^{-20}$ in one second (small p)
- Number of atoms dN decaying in time dt proportional to total number of atoms N , where the proportionality constant is the decay constant λ (units 1/s), which tells us the probability of an atom to decay per unit time.

$$\frac{dN}{dt} = -\lambda \cdot N(t)$$

Example: Nuclear decay

- Solution to this differential equation is

$$N(t) = N_0 e^{-\lambda t}$$

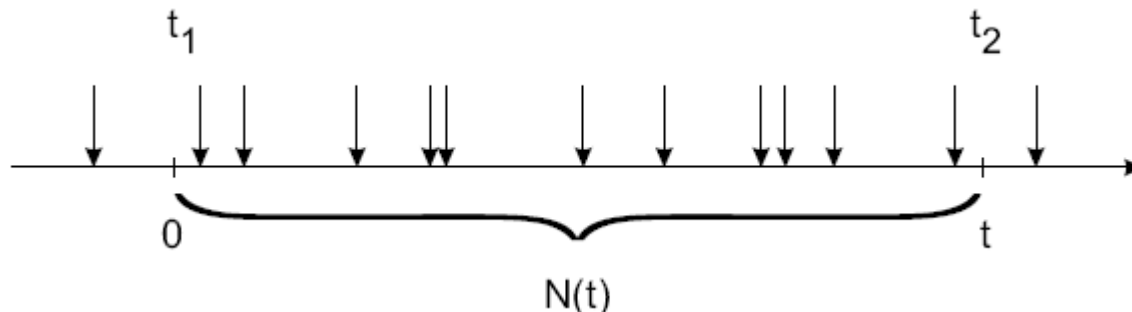
- At $t = \tau$, half of the atoms have decayed. τ is the half-life:

$$N(t = \tau) = \frac{1}{2} N_0 = N_0 e^{-\lambda \tau}$$
$$\tau = \frac{\ln 2}{\lambda}$$

- E.g. Cs^{137} has an half life of 30 years, which corresponds to $\lambda = 7.3 \cdot 10^{-10} \text{ s}^{-1}$. This is the probability for a Caesium atom to decay in 1 second.
- Since a typical sample contains many atoms (1 gram has $\sim 10^{21}$ atoms), get many decays per second.

Example: Nuclear decay

- For radioactive samples with long half lifes,
 $\frac{dN}{dt} = -\lambda N_0 e^{-\lambda t} \approx \text{constant}$ over experimental time scales ($\lambda \ll t \rightarrow e^{-\lambda t} \approx 1$).
- In this case, nuclear decay is a Poisson process with a constant rate of events λN_0 , even though decays occur very irregularly over any time interval.



Example: Nuclear decay

- Decays can be measured with a Geiger counter very precisely (assume no uncertainty in recording the number of counts).
- Let's say for a particular radioactive source we measure 1249 counts in 7 minutes. What is the statistical uncertainty in the number of counts?

$$N = 1249 \pm \sqrt{1249} = 1249 \pm 35$$

- What is the rate of decay (in 1/min) including the uncertainty?

$$R_{total} = \frac{1249 \pm 35}{7min} = 178 \pm 5$$

Example: Nuclear decay

- Since there are also counts from the natural background radiation, need to measure counts without the source and subtract it from the previously measured rate R_{total} .
- Without the source we measure 50 counts in 3 minutes. What is the rate and its uncertainty?

$$R_{bkgrnd} = \frac{50 \pm \sqrt{50}}{3min} = 17 \pm 2$$

- Therefore, the rate of the source is

$$R_{source} = R_{total} - R_{bkgrnd} = 161 \pm \sqrt{5^2 + 2^2} = 161 \pm 5$$

Remarks

- **Important:** In most applications of the Poisson distribution one is interested in the rate of events (number of events per unit time).
- **The square root error \sqrt{N} applies to the raw count of the events , not the rate!**
- Previous example: error in the rate is $\frac{\sqrt{1249}}{7\text{min}}$ **not** $\sqrt{\frac{1249}{7\text{min}}}$

Gaussian approximation

- Consider the Poisson distribution with $\mu = 64$. The probability of $x = 72$ is

$$P_{64}(72) = e^{-64} \frac{64^{72}}{72!} = 2.9\%$$

- This is tedious to calculate and can be problematic on a pocket calculator : $64^{72} \approx 10^{130}$
- Instead use Gaussian approximation

$$\begin{aligned} P_{64}(72) &\approx \int_{71.5}^{72.5} G_{64, \sqrt{64}}(72) dx = 3.0\% \\ &\approx G_{64, \sqrt{64}}(72) dx = G_{64, \sqrt{64}}(72) \cdot 1 \end{aligned}$$

Note: $dx = 72.5 - 71.5 = 1$

Gaussian approximation

- Calculating $Prob_{Poisson}(x \geq 72) = P_{64}(72) + P_{64}(73) + \dots$ is even more tedious.
- Use Gaussian approximation $Prob_{Poisson}(x \geq 72) \approx Prob_{Gaussian}(x \geq 71.5)$. **Since x is continuous variable in Gaussian distribution use lowest number that rounds up to 72 (so-called continuity correction). Likewise, for $Prob_{Poisson}(x \leq 72) \approx Prob_{Gaussian}(x \leq 72.5)$**
- 71.5 is 7.5 above the mean ($\mu = 64$), which is $\frac{7.5}{\sqrt{64}} = 0.94$ standard deviations above the mean.
- Use error function to calculate probability
 $Prob_{Gaussian}(x \geq 71.5, \mu = 64, \sigma = 8) = \frac{[1 - \text{erf}(t=0.94)]}{2} = 17.4\%$

Mastering physics issues

- Rounding tolerance for numerical answers is 2% by default.
- Default precision: Always state the result with 3 significant digits, e.g. $15.15789=15.2$.
- If units are required, a separate answer box for it will be there. Single box means numerical answer only. **In my questions I only ask you for the numerical answer (in units that are stated in the question)**