

# L03: Gaussians II

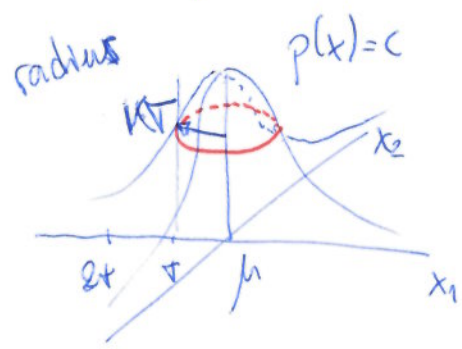
## \* Visualizing Gaussians (2D)

$$p(x) = \alpha \cdot e^{-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)}$$

find contours of constant  $p(x)$

$$K^2 = (x-\mu)^T \Sigma^{-1} (x-\mu) \quad , \quad K=1,2,\dots$$

$K=1 \rightarrow 1-\sigma$



isocontours are eqs  
of a quadratic curve.  
Ellipse.

## \* Mahalanobis distance.

$$\|x-y\|_{\Sigma} = \sqrt{(x-y)^T \Sigma^{-1} (x-y)}$$

\* Non-physical  $\Sigma$ : we use covariance projection and L.A.

(recall:) if  $y = Ax + b$  \  $x \sim \mathcal{N}(0, \mathbb{I})$

then  $\Sigma_y = A \Sigma_x A^T$

Problem: find  $A$  such that projects into  $\Sigma_y$ .

given  $\Sigma_x = \mathbb{I}$  (standard  $\mathcal{N}(0, \mathbb{I})$ ).

$$\Sigma_y = A \cdot \Sigma_x \cdot A^T = A \cdot \mathbb{I} \cdot A^T = AA^T$$

1) SVD decomposition:

$$\Sigma_y = U \cdot D \cdot V^T = U \cdot D \cdot U^T = \underbrace{U \cdot D^{1/2}}_A \cdot \underbrace{D^{1/2} \cdot U^T}_{\text{symmetric}}$$

2) Cholesky decomposition. (More efficient.)

$$\Sigma_y = L \cdot L^T \quad \begin{pmatrix} \Delta & 0 \\ 0 & \nabla \end{pmatrix}$$

Ex:

$$\Sigma = \begin{bmatrix} 4 & -2 \\ -2 & 10 \end{bmatrix} \quad \text{is contour for } K=1? \quad (1-\text{Sigma})$$

$$= \underbrace{\begin{bmatrix} a & 0 \\ b & c \end{bmatrix}}_L \underbrace{\begin{bmatrix} a & b \\ 0 & c \end{bmatrix}}_{L^T} = \begin{bmatrix} 4 & -2 \\ -2 & 10 \end{bmatrix}$$

$$a^2 = 4 \Rightarrow a = \pm 2$$

$$ba = ab = -2 \Rightarrow b = -1$$

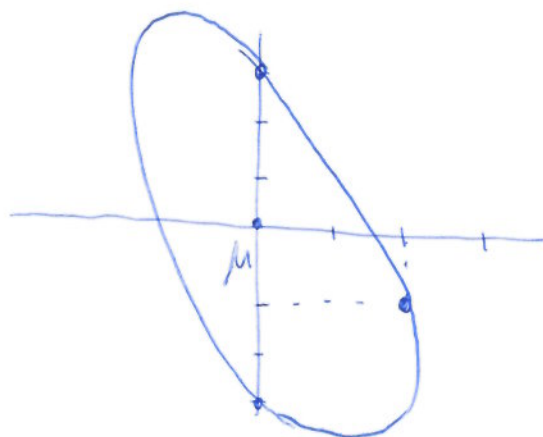
$$b^2 + c^2 = 10 \Rightarrow c = \sqrt{10 - b^2} = \pm 3$$

We choose positive values for diagonal  $\Rightarrow \exists! L$   
(unique solution)

Project points from the circumference  $r=1$

$$\begin{bmatrix} 2 & 0 \\ -1 & 3 \end{bmatrix} \begin{bmatrix} 1 \\ 0 \end{bmatrix} = \begin{bmatrix} 2 \\ -1 \end{bmatrix}$$

$$\begin{bmatrix} L & \end{bmatrix} \begin{bmatrix} 0 \\ 1 \end{bmatrix} = \begin{bmatrix} 0 \\ 3 \end{bmatrix}$$



note: Solution is centered at  $\mu$ , so we need to translate.

### \* Sampling from Gaussians

Most functions sample  $x \sim N(0, 1)$

Problem: how to sample  $y \sim N(\mu_y, \Sigma_y)$ ?

Solution: use covariance projection. Affine  $y = Ax + b$

$$1) \text{ Select } x \sim N(0, I) \quad x = \begin{bmatrix} N(0, 1) \\ N(0, 1) \\ \vdots \\ N(0, 1) \end{bmatrix} \left. \vphantom{\begin{bmatrix} N(0, 1) \\ N(0, 1) \\ \vdots \\ N(0, 1) \end{bmatrix}} \right\} \text{iid}$$

$$2) y \sim N(A\mu_x + b, A\Sigma_x A^T)$$

$$\Rightarrow \boxed{b = \mu_y}$$

$$3) \text{ Find } A \text{ such as } \boxed{A \cdot A^T = \Sigma_y} \quad (\text{Cholesky})$$

$$4) x \sim N(0, I) \quad \text{then} \quad y = Ax + b$$

equiv.  $y \sim N(\mu_y, \Sigma_y)$ .

# \* Conditioning a joint Gaussian distribution.

$$p(x_a, x_b) = \alpha \exp \left\{ -\frac{1}{2} \underbrace{\begin{bmatrix} x_a - \mu_a \\ x_b - \mu_b \end{bmatrix}^T \Sigma^{-1} \begin{bmatrix} x_a - \mu_a \\ x_b - \mu_b \end{bmatrix}}_{\Delta} \right\}$$

Problem:  $p(x_a | x_b)$   $x_a \in \mathbb{R}^n, x_b \in \mathbb{R}^m$

$$\Sigma = \begin{bmatrix} \Sigma_a & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_b \end{bmatrix} \quad \Sigma_{ab} = \Sigma_{ba}^T \quad (\Sigma \text{ symmetric})$$

Information matrix  $\Lambda = \Sigma^{-1}$ ,  $\Lambda = \begin{bmatrix} \Lambda_a & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_b \end{bmatrix}$

Expand the exponent  $\Delta$ :

$$\begin{aligned} \Delta = & -\frac{1}{2} (x_a - \mu_a)^T \Lambda_a (x_a - \mu_a) - \frac{1}{2} (x_a - \mu_a)^T \Lambda_{ab} (x_b - \mu_b) \\ & - \frac{1}{2} (x_b - \mu_b)^T \Lambda_{ba} (x_a - \mu_a) - \frac{1}{2} (x_b - \mu_b)^T \Lambda_b (x_b - \mu_b). \end{aligned}$$

Solution: Completing the square

intuition  $\rightarrow$  we want an exponent only dependent on  $x_a$   
since  $x_b$  is conditioning ("given").

$$\Delta = -\frac{1}{2} x_a^T \Sigma_{a|b}^{-1} x_a + x_a^T \Sigma_{a|b}^{-1} m - \frac{1}{2} m^T \Sigma_{a|b}^{-1} m + \text{const.}$$

Q: what happens to  $x_b$  and constant terms?



\*  $\boxed{\Sigma_{a|b}^{-1} = \Lambda_a}$  (I) (2nd order term)  $x_a^T \Sigma_{a|b}^{-1} x_a$

\* 1st order terms:

$$x_a^T \cdot \underbrace{\left( \Lambda_a \mu_a - \Lambda_{ab} (x_b - \mu_b) \right)}_{\Sigma_{a|b}^{-1} \mu_{a|b}}$$

$$\Rightarrow \mu_{a|b} = \Sigma_{a|b} \left( \Lambda_a \mu_a - \Lambda_{ab} (x_b - \mu_b) \right) \quad \text{(I)}$$

$$= \mu_a - \Sigma_{a|b} \Lambda_{ab} (x_b - \mu_b) \quad \text{(II)}$$

We'll use the foll. matrix equality:

$$\begin{bmatrix} A & B \\ C & D \end{bmatrix}^{-1} = \begin{bmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1}CMBD^{-1} \end{bmatrix}$$

$$M = (A - BD^{-1}C)^{-1} \quad (M^{-1} \text{ Schur complement})$$

$$\begin{bmatrix} \Sigma_a & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_b \end{bmatrix}^{-1} = \begin{bmatrix} \Lambda_a & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_b \end{bmatrix}$$

$$\rightarrow \Lambda_a = \left( \overbrace{\Sigma_a - \Sigma_{ab} \Sigma_b^{-1} \Sigma_{ba}}^{\Sigma_{ab}} \right)^{-1} \quad (\text{Schur complement})$$

$$\rightarrow \Lambda_{ab} = -\Lambda_a \cdot \Sigma_{ab} \Sigma_b^{-1}$$

$$\begin{aligned} \Rightarrow \mu_{a|b} &= \mu_a - \cancel{\Sigma_{ab}} \left( -\cancel{\Lambda_a} \Sigma_{ab} \Sigma_b^{-1} \right) (x_b - \mu_b) \\ &\quad \textcircled{\text{I}} \\ &= \mu_a + \Sigma_{ab} \Sigma_b^{-1} (x_b - \mu_b) \end{aligned}$$

$$p(x_a | x_b) = N(x_a; \mu_a + \Sigma_{ab} \Sigma_b^{-1} (x_b - \mu_b), \Sigma_a - \Sigma_{ab} \Sigma_b^{-1} \Sigma_{ba})$$

\* Marginalizing a joint Gaussian distribution.

problem:  $p(x_a) = \int p(x_a, x_b) dx_b$

(idea) solution: same as before, we will expand the exponent and complete the square, now twice

$$\int e^{\Delta} dx_b = e^{\Delta(x_a)} \underbrace{\int e^{\Delta(x_b)} dx_b}_{\eta} \cdot \underbrace{e^{\Delta(\text{constant})}}_{\eta'}$$

$$\Delta = -\frac{1}{2} (x_a - \mu_a)^T \Lambda_a (x_a - \mu_a) - \frac{1}{2} (x_a - \mu_a)^T \Lambda_{ab} (x_b - \mu_b) - \frac{1}{2} (x_b - \mu_b)^T \Lambda_{ba} (x_a - \mu_a) - \frac{1}{2} (x_b - \mu_b)^T \Lambda_b (x_b - \mu_b)$$

$$= -\frac{1}{2} x_b^T \Lambda_b x_b \quad (1)$$

$$\Delta(x_b) \left\{ \begin{aligned} &+ x_b^T (\Lambda_b \mu_b - \Lambda_{ba} (x_a - \mu_a)) \quad (2) \\ &- \frac{1}{2} m^T \Lambda_b \cdot m \end{aligned} \right.$$

$\hookrightarrow \Lambda_b \cdot m \Rightarrow m = \mu_b - \Lambda_b^{-1} \Lambda_{ba} (x_a - \mu_a)$

$$\Delta(x_a) \left\{ \begin{aligned} &+ \frac{1}{2} m^T \Lambda_b \cdot m \\ &- \frac{1}{2} x_a^T \Lambda_a x_a \quad (3) \\ &+ x_a^T (\Lambda_a \mu_a + \Lambda_{ab} \mu_b) \quad (4) \\ &+ \text{const.} \end{aligned} \right.$$

$$(\Lambda_b^{-1})^T = \Lambda_b^{-1}$$

$$\begin{aligned} \bullet \quad \frac{1}{2} m^T \Lambda_b m &= \frac{1}{2} (\mu_b - \Lambda_b^{-1} \Lambda_{ba} (x_a - \mu_a))^T \cdot \Lambda_b \cdot (\cdot) \stackrel{?}{=} \\ &= \frac{1}{2} \cancel{x_a^T \Lambda_{ab} \Lambda_b^{-1}} \cdot \cancel{\Lambda_b \cdot \Lambda_b^{-1} \Lambda_{ba} x_a} \\ &= x_a^T \Lambda_{ab} \Lambda_b^{-1} \cdot \Lambda_b (\mu_b + \Lambda_b^{-1} \Lambda_{ba} \mu_a) \\ &+ \text{const.} \end{aligned}$$

(3.8)

$$\begin{aligned}
\Delta(x_a) &= -\frac{1}{2} x_a^T \Lambda_a x_a + \frac{1}{2} x_a^T \Lambda_{ab} \Lambda_b^{-1} \Lambda_{ba} x_a \\
&\quad + x_a^T (\Lambda_a \mu_a + \cancel{\Lambda_{ab} \mu_b}) - x_a^T \cancel{\Lambda_{ab}} (\mu_b + \Lambda_b^{-1} \Lambda_{ba} \mu_a) \\
&\quad + \text{const} \\
&= -\frac{1}{2} x_a^T \underbrace{(\Lambda_a - \Lambda_{ab} \Lambda_b^{-1} \Lambda_{ba})}_{\Sigma_a} x_a \quad (\text{Schur complement!}) \\
&\quad + x_a^T (\Lambda_a - \Lambda_{ab} \Lambda_b^{-1} \Lambda_{ba}) \mu_a \\
&\quad + \text{const.}
\end{aligned}$$

$$p(x_a) = \int p(x_a, x_b) dx_b = \mathcal{N}(x_a; \mu_a, \Sigma_a)$$

Gaussian again!

Marginalizing a Gaussian is as simple as selecting the submatrix inside  $\Sigma$  and the corresponding mean!

Gaussians are their self conjugate priors.

Prob Rob 2.4, 3.1, 3.2