Pstat 115 HW2

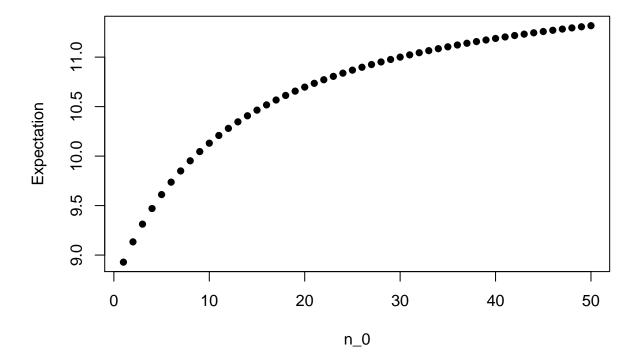
Kalvin Goode 2018/10/20

(1) (a) We see that $E(A) = \frac{120}{10} = 12 = \frac{12}{1} = E(B)$. So we expect to have same average of incidence of cancer. However, we see that $Var(A) = \frac{120}{10^2} = 1.2 < 12 = \frac{12}{1^2} = Var(B)$. Therefore, we are more certain on A because it has lower variance. (b) $y_A=c(12, 9, 12, 14, 13, 13, 15, 8, 15, 6)$ y_B=c(11, 11, 10, 9, 9, 8, 7, 10, 6, 8, 8, 9, 7) cat("Posterior distribution for A is gamma(",120+sum(y_A),",",10+length(y_A),")") ## Posterior distribution for A is gamma(237, 20) cat("Posterior distribution for B is gamma(",12+sum(y_B),",",1+length(y_B),")") ## Posterior distribution for B is gamma(125 , 14) cat("Posterior mean for A is",(120+sum(y_A))/(10+length(y_A))) ## Posterior mean for A is 11.85 cat("Posterior mean for B is",(12+sum(y_B))/(1+length(y_B))) ## Posterior mean for B is 8.928571 cat("Posterior variance for A is",(120+sum(y_A))/(10+length(y_A))^2) ## Posterior variance for A is 0.5925 cat("Posterior variance for B is",(12+sum(y_B))/(1+length(y_B))^2) ## Posterior variance for B is 0.6377551 cat("The 95% quantile-based credible intervals for theta_A is (",qgamma(0.025,120+sum(y_A),10+length(y_ ## The 95% quantile-based credible intervals for theta_A is (10.38924 , 13.40545) cat("The 95% quantile-based credible intervals for theta_B is (",qgamma(0.025,12+sum(y_B),1+length(y_B) ## The 95% quantile-based credible intervals for theta_B is (7.432064 , 10.56031) (c) $n_0=c(1:50)$ $a=12*n_0+sum(y_B)$ b=n_0+length(y_B) a/b

[1] 8.928571 9.133333 9.312500 9.470588 9.611111 9.736842 9.850000 [8] 9.952381 10.045455 10.130435 10.208333 10.280000 10.346154 10.407407 ## [15] 10.464286 10.517241 10.566667 10.612903 10.656250 10.696970 10.735294 ## [22] 10.771429 10.805556 10.837838 10.868421 10.897436 10.925000 10.951220

##

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## [29] 10.976190 11.000000 11.022727 11.044444 11.065217 11.085106 11.104167
## [36] 11.122449 11.140000 11.156863 11.173077 11.188679 11.203704 11.218182
## [43] 11.232143 11.245614 11.258621 11.271186 11.283333 11.295082 11.306452
## [50] 11.317460
plot(a/b,xlab="n_0",ylab="Expectation",pch=16)
```



(d)

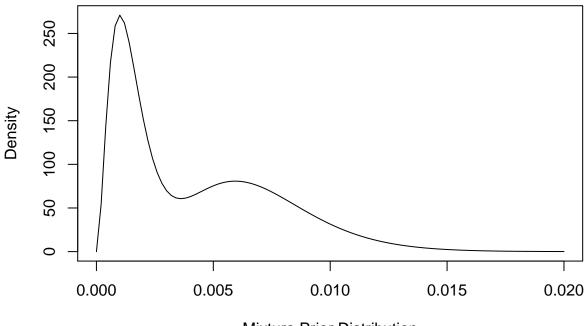
According to the 95% quantile-based credible intervals for θ_A and θ_B , we see that these two groups are not independent and are related. Therefore, it does not make sense to have $p(\theta_A), \theta_B) = p(\theta_A) \times p(\theta_B)$

(2)

(a)

The first expert has mean=3/2000 and variance= $3/2000^2$, and second expert has mean=7/1000 and variance= $7/1000^2$. The first expert has more confidence because the vaiance is smaller than that of second expert.

(b) curve((dgamma(x,3,2000)+dgamma(x,7,1000))/2,0,0.02,xlab="Mixture Prior Distribution", ylab="Density")



Mixture Prior Distribution

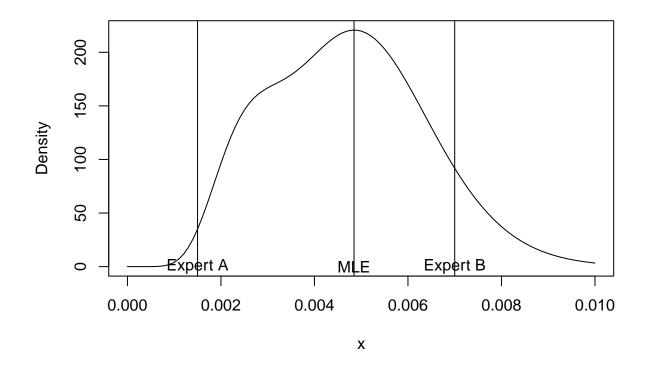
(c)

Let c_1 and c_2 be constants. By conjugation, we have $c_1gamma(x; (3+8), (2000+1767)) = Lgamma(x; 3, 2000)$ and $c_2gamma(x; (7+8), (1000+1767)) = Lgamma(x; 7, 1000)$. We know that $gamma(x; a, b) = \frac{x^{a-1}}{\Gamma(x)b^a}exp(\frac{-x}{b})$. Thus, $c_1 = \frac{11! \times 2000^3}{3767^{11} \times 3!}$ and $c_2 = \frac{15! \times 1000^7}{2767^{15} \times 7!}$ And the posterior distribution up to a proportionality is $\frac{1}{2}(c_1gamma(x; 11, 3767) + c_2gamma(x; 15, 2767))$

(d)

We have $p(\lambda|y) = \frac{L(\lambda;y)p(\lambda)}{K} = \frac{c_1gamma(x;11,3767) + c_2gamma(x;15,2767)}{c_1 + c_2}$. Notice that, from (c), we have $\frac{c_1}{c_1 + c_2} = 0.287$ and $\frac{c_1}{c_1 + c_2} = 0.712$. Thus, $p(\lambda|y) \sim 0.287gamma(11,3767) + 0.712gamma(15,2767)$

```
(e)
y<-function(x)
{
    fx<-0.28766*dgamma(x,11,3767)+0.71233*dgamma(x,15,2767)
}
curve(0.28766*dgamma(x,11,3767)+0.71233*dgamma(x,15,2767),0,0.01,xlab="x",ylab="Density")
abline(v=3/2000)
abline(v=7/1000)
MLE=optimize(y,c(0,0.01),maximum=TRUE)
text(3/2000, 0, labels='Expert A')
text(7/1000, 0, labels='Expert B')
text(MLE, 0, labels='MLE')
abline(v=MLE)</pre>
```



cat("The MLE is approximately")

The MLE is approximately

MLE

\$maximum

[1] 0.004847024

##

\$objective ## [1] 220.7259