# Holland's schema theorem

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## 1 The theorem

Let us fix the alphabet  $\Sigma = \{0, 1, \star\}$  with the special character  $\star$ . A symbol  $h = \star$  is called *wildcard*, otherwise  $h \neq \star$  is called *specification*. A *binary schema* is a string  $H = h_1 \dots h_n \in \Sigma^{n>0}$ .

A string  $S = s_1 \dots s_n \in \{0, 1\}^{n>0}$  fullfills a schema H if and only if there is a match for each specification:

$$s_i = h_i \forall i \in \{j | h_i \neq \star\}. \tag{1}$$

If a string S fullfills the schema H, we write  $S \in H$ .

The number of specifications in a schema is the order of that schema

$$o(H) = |\{i | h_i \neq \star, h_i \in H\}|. \tag{2}$$

The distance between the first and the last specification in a schema is the defining length of that schema

$$\delta(H) = \max_{h_i \neq \star, h_i \in H} i - \min_{h_i \neq \star, h_i \in H} i \tag{3}$$

At a given iteration t of our genetic algorithm, the number of binary strings that fullfill the schema H is  $r_{H,t}$ .

The average observed fitness at the t-th iteration is

$$\bar{f}(t) = \frac{1}{m} \sum_{i=0}^{m} f(b_{i,t})$$
 (4)

where  $b_{i,t}$  is the *i*-th individual in a population of m > 0 individuals (binary strings) at the *t*-th iteration.

The average observed fitness of the schema H at the t-th iteration is

$$\bar{f}(H,t) = \frac{1}{r_{H,t}} \sum_{i \in \{j | b_{j,t} \in H\}} f(b_{i,t}).$$
 (5)

**Theorem 1.1 (Holland)**  $\forall H \in \Sigma^{n>0}$ , given a genetic algorithm with:

- proportional selection
- one-point crossover with probability  $p_c$
- random mutation with probability  $p_m$

it holds

$$\mathbb{E}[r_{H,t+1}] \ge r_{H,t} \cdot \frac{\bar{f}(H,t)}{\bar{f}(t)} \cdot \left(1 - p_c \frac{\delta(H)}{n-1}\right) \cdot (1 - p_m)^{o(H)}. \tag{6}$$

#### 2 Proof

Since the probability of an individual  $b_{i,t}$  to be chosen for the (t+1)-th iteration

$$\mathbb{P}[\text{choose } b_{i,t}] = \frac{f(b_i)}{\sum_{i=0}^{m} f(b_{i,t})}$$

$$\tag{7}$$

the probability to choose an individual that fullfills the schema H is

$$\mathbb{P}[\text{choose a } b \in H] = \frac{\sum_{i \in \{j | b_{j,t} \in H\}} f(b_{i,t})}{\sum_{i=0}^{m} f(b_{i,t})} = p.$$
 (8)

Since each individual is chosen independently, the expected number of individuals that fullfill the schema H at iteration t+1 is

$$\mathbb{E}[r_{H,t+1}] = m \cdot p \tag{9}$$

$$= m \cdot \frac{\sum_{i \in \{j | b_{j,t} \in H\}} f(b_{i,t})}{\sum_{i=0}^{m} f(b_{i,t})}$$

$$= m \cdot \frac{r_{H,t}}{r_{H,t}} \cdot \frac{\sum_{i \in \{j | b_{j,t} \in H\}} f(b_{i,t})}{\sum_{i=0}^{m} f(b_{i,t})}$$

$$(10)$$

$$= m \cdot \frac{r_{H,t}}{r_{H,t}} \cdot \frac{\sum_{i \in \{j \mid b_{j,t} \in H\}} f(b_{i,t})}{\sum_{i=0}^{m} f(b_{i,t})}$$
(11)

$$= r_{H,t} \cdot \frac{\frac{\sum_{i \in \{j \mid b_{j,t} \in H\}} f(b_{i,t})}{r_{H,t}}}{\sum_{i=0}^{m} f(b_{i,t})}} = r_{H,t} \cdot \frac{\bar{f}(H,t)}{\bar{f}(t)}. \tag{12}$$

But it holds only if  $p_c = 0$  and  $p_m = 0$ . In the case in which these probability are not null, it is necessary to take them into account.

### 2.1 Crossover

**Property 2.1** If two individuals fullfill the schema, also their parents fullfill the schema.

**Property 2.2** If an individual fullfill the schema, and another individual does not, their children will not fullfill the same schema if and only if the cross site is between two specifications.

Let show the last property through an example. Let  $H = 01 \star 10 \star \star$  be the schema and  $b_i$  and  $b_j$  be two individuals, e.g.

$$b_i = 0111001 \in H$$
  
 $b_j = 1101101 \notin H$ .

If the cross site  $\gamma$ , i.e. the point in which the individuals are broken into two pieces in order to recombine their symbols, is before the last two symbols,

$$\operatorname{cross}_{\gamma}(0111001, 1101101) = 0111001, 1101101$$
 (13)

it is easy too see that only one of these two individuals fullfills the schema H.

If the cross site  $\bar{\gamma}$  is, for example, one position before the previous one, both the resulting individuals will not fullfill the schema H:

$$\operatorname{cross}_{\bar{\gamma}}(0111001, 1101101) = 0111101, 1101001.$$
 (14)

Thus, since the cross site is after an index in [1, n-1],

$$\mathbb{P}[\text{cross site is in the defining length}] = \frac{\delta(H)}{n-1}$$
 (15)

and the probability to be sure that the cross site is good is  $1 - \frac{\delta(H)}{n-1}$ . Actually, also a cross site inside the defining length could provide a good cross, but there is not certainty. Finally, the probability that an individual that fullfills the schema leads another individual that fullfills the same schema is  $p_s \geq 1 - p_c \frac{\delta(H)}{n-1}$ .

Since selection and crossover are independent operations, now

$$\mathbb{E}[r_{H,t+1}] = r_{H,t} \cdot \frac{\bar{f}(H,t)}{\bar{f}(t)} \cdot p_s \tag{16}$$

$$\geq r_{H,t} \cdot \frac{\bar{f}(H,t)}{\bar{f}(t)} \cdot \left(1 - p_c \frac{\delta(H)}{n-1}\right). \tag{17}$$

# 2.2 Mutation

After the crossover,  $r_{H,t+1}$  can only decrease if the mutation hits the wrong point. The mutation hit only one symbol and makes the logic not of the bit: m(0) = 1, m(1) = 0. The probability to choose a particular bit is  $p_m = 1/n$  and each bit is independent, so the mutation is not bad if it does not hit a specification. Since the number of specification of the schema H is o(H), the term  $(1 - p_m)^{o(H)}$  is added to the formula, and the theorem follows.