

# 1 Basic concept

## 1.1 Bias-variance decomposition

Define  $e(x) = f(x) - \hat{f}(x)$  and  $p(e)$  as its pdf.

$$\text{MSE} = E(e^2) = \int e^2 p de = (\int e p de)^2 + (\int e^2 p de - (\int e p de)^2) = E(e)^2 + \text{Var}(e) \quad (1)$$

## 1.2 A simple example

(A)  $Y$  follows binomial distribution  $B(n, \pi_h)$ .  $E(Y) = n\pi_h$ ,  $\text{Var}(Y) = n\pi_h(1 - \pi_h)$ . One can estimate  $f(0)$  by

$$f(0)_{\text{estimate}} = \frac{Y}{nh} \quad (2)$$

(B)

$$\begin{aligned} \pi_h &\approx hf(0) + \frac{f''(0)}{2} \int x^2 dx \\ &= hf(0) + \frac{f''(0)h^3}{24} \end{aligned} \quad (3)$$

Choosing  $h$  so that both  $h \ll 1$  and  $\pi_h \ll 1$  are true, we have

$$\begin{aligned} \text{MSE}(0) &= (E(\hat{f}(0)) - f(0))^2 + \text{Var}(\hat{f}(0)) \\ &\approx (\frac{\pi_h}{h} - f(0))^2 + \frac{\pi_h}{nh^2} \\ &= (\frac{f''(0)}{24})^2 h^4 + \frac{1}{nh} (f(0) + \frac{f''(0)h^2}{24}) \\ &\approx (\frac{f''(0)}{24})^2 h^4 + \frac{1}{nh} f(0) \end{aligned} \quad (4)$$

(C)

$$\frac{\partial \text{MSE}(0)}{\partial h} = 4Ah^3 - \frac{f(0)}{nh^2} \quad (5)$$

In order to minimize the MSE,  $h = (\frac{f(0)}{4An})^{\frac{1}{5}}$

# 2 Curve Fitting by linear smoothing

(A)

$$y_{\text{estimate}} = \beta_{\text{estimate}} x$$

$$\begin{aligned}
&= (\vec{x}^T \vec{x})^{-1} \vec{x}^T \vec{y} x \\
&= \frac{\sum_i x_i y_i}{\sum_i x_i^2} x \\
&= \sum_i \frac{x_i x}{\vec{x}^2} y_i
\end{aligned} \tag{6}$$

So we have  $\omega_i = \frac{x_i x}{\vec{x}^2}$

(B) See "linearsmoothing.R" and "weight.R"

### 3 Cross validation

(A) See "predictionerror.R"

(B) See "testmodel.R"

(C)

### 4 Local polynomial regression

(A) Define matrix  $R$  where  $R_{ij}(\vec{x}) = (x_i - x)^{j-1}$ . Then  $g(x_i, a) = (R\vec{a})_i$ . The cost function can be written as

$$\sum_{i=1}^n \omega_i (y_i - (R\vec{a})_i)^2 \tag{7}$$

Optimizing this with  $\vec{a}$  is equivalent to linear regression on  $f(\vec{x}) = \vec{x}^T \vec{a}$  given observed data  $X = R$  and  $\vec{y}$  and weighted cost function. The answer is

$$\vec{a} = (R^T \Omega R)^{-1} R^T \Omega \vec{y} \tag{8}$$

$$\Omega = \text{diag}(\omega_1, \omega_2, \dots) \tag{9}$$

The estimate of  $f(x)$  is just  $g(x, a) = a_0$ .

(B) Call the old weight function  $\omega$  and the new ones  $\gamma$

$$\begin{aligned}
(R^T \Omega R)^{-1} &= \left( \begin{array}{cc} \sum_i \omega_i & \sum_i R_{i2} \omega_i \\ \sum_i R_{i2} \omega_i & \sum_i R_{i2}^2 \omega_i \end{array} \right)^{-1} \\
&= \frac{1}{(\sum_i \omega_i)(\sum_i R_{i2}^2 \omega_i) - (\sum_i R_{i2} \omega_i)^2} \begin{pmatrix} \sum_i R_{i2}^2 \omega_i & -\sum_i R_{i2} \omega_i \\ -\sum_i R_{i2} \omega_i & \sum_i \omega_i \end{pmatrix} \\
R^T \Omega \vec{y} &= \begin{pmatrix} \sum_i \omega_i y_i \\ \sum_i R_{i2} \omega_i y_i \end{pmatrix}
\end{aligned} \tag{11}$$

$$\begin{aligned}
f(x) = a_0 &= \frac{(\sum_i R_{i2}^2 \omega_i)(\sum_i \omega_i y_i) - (\sum_i R_{i2} \omega_i)(\sum_i R_{i2} \omega_i y_i)}{(\sum_i \omega_i)(\sum_i R_{i2}^2 \omega_i) - (\sum_i R_{i2} \omega_i)^2} \\
&= \sum_i \frac{((\sum_j R_{j2}^2 \omega_j) - (\sum_j R_{j2} \omega_j) R_{i2}) \omega_i}{\sum_k ((\sum_j R_{j2}^2 \omega_j) - (\sum_j R_{j2} \omega_j) R_{k2}) \omega_k} y_i \quad (12)
\end{aligned}$$

Define weight function  $\gamma(x_i, x) = ((\sum_j R_{j2}^2 \omega_j) - (\sum_j R_{j2} \omega_j) R_{i2}) \omega_i$ , the above can be written as a weighted sum of  $y_i$ . Further evaluate  $\gamma$ :

$$\begin{aligned}
\gamma(x_i, x) &= ((\sum_j R_{j2}^2 \omega_j) - (\sum_j R_{j2} \omega_j) R_{i2}) \omega_i \\
&= \frac{1}{h^2} ((\sum_j (x_j - x)^2 K_j) - (\sum_j (x_j - x) K_j)(x_i - x)) K_i \\
&= \frac{1}{h^2} (s_2 - s_1(x_i - x)) K_i \quad (13)
\end{aligned}$$

Here we used  $K_i$  to denote  $K(\frac{x-x_i}{h})$ . Since the function will be normalized so we can ignore  $\frac{1}{h^2}$ .

(C) From  $y = f(x) + \epsilon$

$$\text{Var}(\vec{y}) = \sigma^2 \quad (14)$$

$$\text{Mean}(\vec{y}) = f(\vec{x}) \quad (15)$$

$$\begin{aligned}
\text{Mean}(a_0) &= \sum_i \gamma_i f(x_i) \\
\text{Var}(a_0) &= \text{TrVar}(H \vec{y}) \\
&= \sum_i |\gamma(x_i, x)|^2 \sigma^2 \quad (16)
\end{aligned}$$

(D) Define  $\vec{\mu} = \text{Mean}(\vec{y}) = f(\vec{x})$

$$\begin{aligned}
E(\sigma^2) &\propto E((\vec{y} - H \vec{y})^T (\vec{y} - H \vec{y})) \\
&= E(\vec{y}^T \vec{y}) - 2E(\vec{y}^T H \vec{y}) + E(\vec{y}^T H^2 \vec{y}) \\
&= E(\vec{y}^T \vec{y}) - 2 * (\vec{\mu}^T H \vec{\mu} + \sigma^2 \text{Tr}(H)) + \vec{\mu}^T (H^2) \vec{\mu} + \sigma^2 \text{Tr}(H^2) \quad (17)
\end{aligned}$$

Apply trace trick to the first terms:

$$E(\vec{y}^T \vec{y}) = \text{Tr}(E(\vec{y} \vec{y}^T)) = \text{Tr}(\sigma^2 + \vec{\mu} \vec{\mu}^T) = n\sigma^2 + \vec{\mu}^T \vec{\mu} \quad (18)$$

$$E(\sigma^2) = \sigma^2 + \frac{\vec{\mu}^T (1 - H)^2 \vec{\mu}}{n - 2\text{Tr}(H) + \text{Tr}(H^2)} \quad (19)$$

## 5 Gaussian Process

(A)

(B) Define  $C_{ij}$  to be element of the INVERSE of covariance matrix computed by covariance function, from the joint distribution

$$p(f(x_1), f(x_2), \dots, f(x_n), f(x^*)) \propto e^{-\frac{1}{2} \sum_{i,j} C_{ij} (f(x_i) - m(x_i))(f(x_j) - m(x_j))} \quad (20)$$

we have

$$\begin{aligned} p(f(x^*) | f(x_1), f(x_2), \dots, f(x_n)) &\propto e^{-\frac{1}{2} (C_{**} (f(x^*) - m(x^*))^2 + 2(f(x^*) - m(x^*)) \sum_i C_{*i} (f(x_i) - m(x_i)))} \\ &\propto e^{-\frac{C_{**}}{2} (f(x^*)^2 + 2f(x^*) (\sum_i \frac{C_{*i}}{C_{**}} (f(x_i) - m(x_i)) - m(x^*)))} \end{aligned} \quad (21)$$

So we have  $(x^* | x_1, \dots, x_n) \sim N(m(x^*) - \frac{1}{C_{**}} \sum_i C_{*i} (f(x_i) - m(x_i)), \frac{1}{C_{**}})$

(C) For any linear combination of  $(\vec{y}, \vec{\theta})$ :  $(\vec{a}^T \vec{y} + \vec{b}^T \vec{\theta})$ . Compute the moment generating function

$$E(e^{t\vec{a}^T \vec{y} + t\vec{b}^T \vec{\theta}}) = \int e^{t\vec{a}^T \vec{y} + t\vec{b}^T \vec{\theta}} p(\vec{y} | \vec{\theta}) p(\vec{\theta}) d\vec{y} d\vec{\theta} \quad (22)$$

Change the variable being integrated

$$\vec{y} \rightarrow \vec{y}' = \vec{y} - R\vec{\theta} \quad (23)$$

$$\vec{\theta} \rightarrow \vec{\theta}' = \vec{\theta} - \vec{m} \quad (24)$$

The Jacobian of this transformation is 1. The logarithmic pdf of  $e^{t\vec{a}^T \vec{y} + t\vec{b}^T \vec{\theta}}$  is:

$$t\vec{a}^T \vec{y}' + t(\vec{b}^T + \vec{a}^T R)\vec{\theta}' + t(\vec{b}^T + \vec{a}^T R)\vec{m} - \frac{\vec{y}'^T \Sigma^{-1} \vec{y}'}{2} - \frac{\vec{\theta}'^T V^{-1} \vec{\theta}'}{2} + \dots \quad (25)$$

Complete the square for  $\vec{y}'$ ,

$$-\frac{\vec{y}'^T \Sigma^{-1} \vec{y}'}{2} + t\vec{a}^T \vec{y}' = -\frac{(\vec{y}' - t\Sigma\vec{a})^T \Sigma^{-1} (\vec{y}' - t\Sigma\vec{a})}{2} + \frac{\vec{a}^T \Sigma \vec{a}}{2} t^2 \quad (26)$$

and similarly for  $\vec{\theta}'$ . The  $t$ -dependent part after the the integrating is

$$\exp \left( \frac{\vec{a}^T \Sigma \vec{a} + (\vec{b} + R^T \vec{a})^T V (\vec{b} + R^T \vec{a})}{2} t^2 + (\vec{b}^T + R^T \vec{a}) t \right) \quad (27)$$

The coefficient of  $t^2$  is positive definite. And the constant multiplied to this exponential must be 1 because  $M(0) = E(1) = 1$ .

## 6 In nonparametric regression and spatial smoothing

(A)

$$p(\vec{y}, f(\vec{x})) = p(\vec{y}|f(\vec{x}))p(f(\vec{x})) \propto e^{-\frac{|\vec{y}-f(\vec{x})|^2}{2\sigma^2}} e^{-\frac{f(\vec{x})^T C f(\vec{x})}{2}} \quad (28)$$

$$\begin{aligned} p(f(\vec{x})|\vec{y}) &\propto p(\vec{y}, f(\vec{x}))|_{f(\vec{x})} \\ &\propto e^{-\frac{f(\vec{x})^T (C + \frac{1}{\sigma^2}) f(\vec{x})}{2} + \frac{\vec{y}^T f(\vec{x})}{\sigma^2}} \\ &\propto e^{-\frac{(f(\vec{x}) - (C\sigma^2 + 1)^{-1}\vec{y})^T (C + \frac{1}{\sigma^2})(f(\vec{x}) - (C\sigma^2 + 1)^{-1}\vec{y})}{2}} \end{aligned} \quad (29)$$

Therefore  $f(\vec{x})|\vec{y} \sim N((C\sigma^2 + 1)^{-1}\vec{y}, \sigma^2(C\sigma^2 + 1)^{-1})$

(B) From previous question  $E(f(x^*)|f(\vec{x})) = -\vec{C}_*^T f(\vec{x})$  and  $E(f(\vec{x})|\vec{y}) = (C\sigma^2 + 1)^{-1}\vec{y}$ , where  $\vec{C}_* = \frac{1}{C_{**}}(C_{*1}, C_{*2} \dots)^T$ . we have

$$\begin{aligned} E(f(x^*)|\vec{y}) &= \int E(f(x^*)|f(\vec{x}))p(f(\vec{x})|\vec{y})df(\vec{x}) \\ &= - \int (\vec{C}_*^T f(\vec{x}))p(f(\vec{x})|\vec{y})df(\vec{x}) \\ &= -\vec{C}_*^T E(f(\vec{x})|\vec{y}) = \vec{W}^T \vec{y} \end{aligned} \quad (30)$$

Where  $\vec{W} = -\vec{C}_*^T (C\sigma + 1)^{-1}$ .

To compute variance, first compute  $E(f(x^*)^2)$ , using  $Var(x) = E(x^2) - E(x)^2$  and the fact  $x_i$  are independent of each other

$$\begin{aligned} E(f(x^*)^2|\vec{y}) &= \int E(f(x^*)^2|\vec{x})p(f(\vec{x})|\vec{y})df(\vec{x}) \\ &= \int \left(\frac{1}{C_{**}} + (\vec{C}_*^T f(\vec{x}))^2\right)p(f(\vec{x})|\vec{y})df(\vec{x}) \\ &= \frac{1}{C_{**}} + \frac{1}{C_{**}^2} \left(\sum_i C_{*i}^2 E(f(x_i)^2|\vec{y}) + \sum_{i \neq j} C_{*i} C_{*j} E(f(x_i)|\vec{y}) E(f(x_j)|\vec{y})\right) \\ &= \frac{1}{C_{**}} + \frac{1}{C_{**}^2} (\sigma^2 \sum_i C_{*i}^2 + (\vec{C}_*^T E(f(\vec{x})|\vec{y}))^2) \\ &= \frac{1}{C_{**}} + \sigma^2 |\vec{C}_*|^2 + E(f(x^*)|\vec{y})^2 \end{aligned} \quad (31)$$

$$Var(f(x^*)|\vec{y}) = E(f(x^*)^2|\vec{y}) - E(f(x^*)|\vec{y})^2 = \frac{1}{C_{**}} + \sigma^2 |\vec{C}_*|^2 \quad (32)$$