## 1 A simple Gaussian location model

(A) This is equivalent with determine the distribution x under a gaussian distribution with precision  $\omega$  and mean  $\mu$  where  $\omega \sim Gamma(\frac{d}{2}, \frac{\eta}{2\kappa})$ , the result should be

$$p(\theta) \sim (1 + \frac{\kappa(\theta - \mu)^2}{n})^{-\frac{d+1}{2}}$$
 (1)

Compare to the problem we have  $\nu=d,\,s^2=\frac{\eta}{d\kappa}$  and  $m=\mu$ 

(B)

$$\begin{split} p(\theta,\omega|\vec{y}) &= \frac{p(\vec{y}|\theta,\omega)p(\theta,\omega)}{\int p(\vec{y}|\theta',\omega')p(\theta',\omega')d\theta'd\omega'} \\ &\propto \omega^{\frac{n}{2}}e^{-\frac{\omega\sum_{i}(y_{i}-\theta)^{2}}{2}}\omega^{\frac{d+1}{2}-1}e^{-\omega\frac{\kappa(\theta-\mu)^{2}}{2}}e^{-\omega\frac{n}{2}} \\ &\propto \omega^{\frac{d+n+1}{2}-1}e^{-\frac{\omega(n+\kappa)(\theta-\frac{\mu\kappa+\sum_{i}y_{i}}{n+\kappa})^{2}}{2}}e^{-\frac{\omega(n+\sum_{i}y_{i}^{2}+\kappa\mu^{2}-\frac{(\mu\kappa+\sum_{i}y_{i})^{2}}{n+\kappa})^{2}}{2}} \end{split}$$

After rearanging parameters

$$\mu^* = \frac{\mu \kappa + \sum_i y_i}{n + \kappa} \tag{3}$$

$$\kappa^* = \kappa + n \tag{4}$$

$$d^* = d + n \tag{5}$$

$$\eta^* = \eta + \sum_{i} y_i^2 + \kappa \mu^2 - \frac{(\mu \kappa + \sum_{i} y_i)^2}{n + \kappa}$$
(6)

The distribution is of the same form as  $p(\theta, \omega)$ 

$$p(\theta, \omega | \vec{y}) \propto \omega^{\frac{d^*+1}{2} - 1} e^{-\frac{\omega \kappa^* (\theta - \mu^*)^2}{2}} e^{-\frac{\omega \eta^*}{2}}$$
 (7)

(C) Keeping only the part dependent on  $\theta$ 

$$p(\theta|\vec{y},\omega) = \frac{p(\vec{y},\theta,\omega)}{p(\vec{y},\omega)}$$

$$\propto e^{-\frac{\omega\kappa^*(\theta-\mu^*)^2}{2}} \sim N(\mu^*,(\omega\kappa^*)^{-1})$$
(8)

(D)

$$p(\omega|\vec{y}) = \frac{p(\omega, \vec{y})}{p(\vec{y})}$$

$$\propto \int p(\omega, \vec{y}, \theta') d\theta'$$

$$\propto \omega^{\frac{d^*+1}{2}-1} e^{-\frac{\omega\eta^*}{2}} \int e^{-\frac{\omega\kappa^*(\theta'-\mu^*)^2}{2}} d\theta'$$

$$\propto \omega^{\frac{d^*}{2}-1} e^{-\frac{\omega\eta^*}{2}} \sim \Gamma(\frac{d^*}{2}, \frac{\eta^*}{2})$$
(9)

$$p(\theta|\vec{y}) = \int p(\theta, \omega'|\vec{y}) d\omega'$$
(10)

Since  $p(\theta, \omega | \vec{y})$  is same to  $p(\theta, \omega)$  up to a parameter redefinition. So the result should be same to part A with parameter redefinition.

$$p(\theta|\vec{y}) \sim (1 + \frac{\kappa^*(\theta - \mu^*)^2}{\eta^*})^{-\frac{d^*+1}{2}}$$
 (11)

(F) False. In the limit  $\kappa$ , d and  $\eta$  goes to zero

$$p(\theta) \propto \frac{1}{\sqrt{1+\theta^2}} \tag{12}$$

When  $\theta \to \infty$ ,  $p(\theta) \to \frac{1}{\theta}$  which cannot be normalized.

(G) True. The starred paramters are still in healthy condition under this limit.

(H)

## $\mathbf{2}$ The conjugate Gaussian linear model

## **Basics** 2.1

(A)

$$p(\vec{\beta}|\vec{y},\omega) = \frac{p(\vec{\beta},\vec{y},\omega)}{p(\vec{y},\omega)}$$

$$\propto p(\vec{y}|\vec{\beta},\omega)p(\vec{\beta}|\omega)$$

$$\propto e^{-\frac{\omega(\vec{y}-X\vec{\beta})^T\Lambda(\vec{y}-X\vec{\beta})}{2}}e^{-\frac{\omega(\vec{\beta}-\vec{m})^TK(\vec{\beta}-\vec{m})}{2}}$$

$$\propto e^{-\frac{\omega(\vec{\beta}-\vec{m}^*)^TK^*(\vec{\beta}-\vec{m}^*)}{2}}$$

$$(13)$$

$$K^* = X^T\Lambda Y + K$$

$$K^* = X^T \Lambda X + K \tag{14}$$

$$\vec{m}^* = K^{*-1}(X^T \Lambda \vec{y} + K \vec{m})$$
 (15)

(B)

$$p(\omega|\vec{y}) \propto p(\omega, \vec{y})$$

$$\propto \int p(\vec{y}|\vec{\beta}', \omega) p(\vec{\beta}'|\omega) p(\omega) d\vec{\beta}'$$

$$\propto \omega^{\frac{d^*+1}{2}-1} e^{-\frac{\omega\eta^*}{2}} \int e^{-\frac{\omega(\vec{\beta}'-\vec{m}^*)^T K^*(\vec{\beta}'-\vec{m}^*)}{2}} d\vec{\beta}'$$

$$\propto \omega^{\frac{d^*}{2}-1} e^{-\frac{\omega\eta^*}{2}} \sim \Gamma(\frac{d^*}{2}, \frac{\eta^*}{2})$$
(16)

$$d^* = d + n \tag{17}$$

$$\eta^* = \eta + \vec{y}^T \Lambda \vec{y} + \vec{m}^T K \vec{m} - \vec{m}^{*T} K^* \vec{m}^*$$
 (18)

(C) Similar to previous problems

$$p(\omega, \vec{\beta}|\vec{y}) \propto \omega^{\frac{d^*+1}{2}-1} e^{-\frac{\omega\eta^*}{2}} e^{-\frac{\omega(\vec{\beta}-\vec{m}^*)^T K^*(\vec{\beta}-\vec{m}^*)}{2}}$$
 (19)

Again, reuse the derivation where we derive t-distribution from normal/gamma prior

$$p(\vec{\beta}|\vec{y}) = \int p(\omega', \vec{\beta}|\vec{y})d\omega'$$

$$\propto \left(1 + \frac{(\vec{\beta} - \vec{m}^*)^T K^* (\vec{\beta} - \vec{m}^*)}{\eta^*}\right)^{-\frac{d^*+1}{2}}$$
(20)

(D) The script "gdpgrowth.R" read data from "gdpgrowth.csv" and use the above linear model to fit the relation between column GR6096 and DENS60, using

$$\vec{\beta}_{\text{estimate}} = \vec{m}^* = K^{*-1} X^T \Lambda \vec{y} \tag{21}$$

## 2.2 A heavy tailed error model

(A) Let  $\vec{\alpha} = X\vec{\beta}$ 

$$p(\vec{y}|X, \vec{\beta}, \omega) = \int p(\vec{y}|X, \vec{\beta}, \omega \Lambda) \prod_{i} p(\lambda_{i}) d\lambda_{i}$$

$$\propto \prod_{i} \lambda_{i}^{\frac{1}{2}} e^{-\frac{\omega \lambda_{i}(y_{i} - \alpha_{i})^{2}}{2}} p(\lambda_{i}) d\lambda_{i}$$
(22)

This is the familiar normal/gamma prior, the result should be

$$p(\vec{y}|X,\vec{\beta},\omega) \propto \prod_{i} \left(1 + \frac{\lambda_i (y_i - \alpha_i)^2}{h}\right)^{-\frac{h+1}{2}}$$
 (23)

(B)

$$p(\lambda_{i}|\vec{y}, \vec{\beta}, \omega) \propto p(\vec{y}|\vec{\beta}, \omega, \Lambda)p(\lambda_{i})$$

$$\propto e^{-\lambda_{i}(\frac{h+\omega(y_{i}-\alpha_{i})^{2}}{2})}\lambda_{i}^{\frac{h+1}{2}-1}$$

$$\sim \Gamma(\frac{h+1}{2}, \frac{h+\omega(y_{i}-\alpha_{i})^{2}}{2}) \qquad (24)$$

(C) See script "MCMC on gdpgrowth.R" and "gibbs sampler.R". We generate the cycle

$$(\vec{\beta}_1|\vec{y},\omega_0,\Lambda_0),(\omega_1|\vec{y},\Lambda_0) \to (\Lambda_1|\vec{y},\vec{\beta}_1,\omega_1) \to (\vec{\beta}_2|\vec{y},\omega_1,\Lambda_1),(\omega_2|\vec{y},\Lambda_1) \to \cdots$$
(25)

where

$$(\vec{\beta}|\vec{y},\omega,\Lambda) \sim N(\vec{m}^*(\Lambda,\vec{y}),\omega K^*(\Lambda))$$

$$(\omega|\vec{y},\Lambda) \sim \Gamma(\frac{d^*}{2},\frac{\eta^*(\Lambda,\vec{y})}{2})$$

$$(\lambda_i|\vec{y},\vec{\beta},\omega) \sim \Gamma(\frac{h}{2},\frac{h+\omega(y_i-\alpha_i)^2}{2})$$
(26)