

## 1 Basic Model

$$y(t) = \sqrt{h(t)}u(t)$$

$$\ln(h(t)) = \alpha + \delta \ln(h(t-1)) + \sigma_\nu \nu(t)$$

$$t = 1, \dots, N$$

$$u(t), \nu(t) \sim N(0, 1)$$

## 2 Model Fitting

### 2.1 MCMC

From the model we have:

$$(y_t|h_t) \sim N(0, h_t)$$

$$(h_t|h_{t-1}, \omega) \sim N(\alpha + \delta \ln h_{t-1}, \sigma_\nu^2)$$

For other parameters, assume the prior distribution:

$$\alpha \sim N(\alpha_0, \sigma_\alpha^2)$$

$$\delta \sim N(\delta_0, \sigma_\delta^2)$$

$$\sigma_\nu^2 \sim IG\left(\frac{\nu_0}{2}, \frac{s_0}{2}\right)$$

We have the marginal distribution

$$\begin{aligned} p(y, h, \delta, \alpha, \sigma_\nu^2) &\propto \frac{1}{\sigma_\nu^{2+\nu_0}} \exp\left(-\frac{(\delta - \delta_0)^2}{2\sigma_\delta^2} - \frac{(\alpha - \alpha_0)^2}{2\sigma_\alpha^2} - \frac{s_0^2}{2\sigma_\nu^2}\right) \\ &\times \prod_{t=2}^N \frac{1}{h_t^{\frac{3}{2}} \sigma_\nu} \exp\left(-\frac{y_t^2}{2h_t} - \frac{(\ln h_t - \delta \ln h_{t-1} - \alpha)^2}{2\sigma_\nu^2}\right) \end{aligned} \quad (1)$$

From which we can derive the posterior distributions are

$$(\sigma_\nu^2|h, \alpha, \delta) \sim IG\left(\frac{\nu_0 + N - 1}{2}, \frac{s'}{2}\right)$$

$$s' = s_0 + (N-1)\alpha^2 + (1+\delta^2)S_2 - \delta^2(\ln h_N)^2 - (\ln h_1)^2 - 2\alpha((1-\delta)S_1 - \ln h_1 + \delta \ln h_N) - 2\delta S_3$$

$$(\delta|h, \alpha, \sigma_\nu^2) \sim N\left(\frac{\sigma_\nu^2 \delta_0 + \sigma_\delta^2(S_3 - \alpha(S_1 - \ln h_N))}{\sigma_\nu^2 + \sigma_\delta^2(S_2 - (\ln h_N)^2)}, \frac{\sigma_\nu^2 \sigma_\delta^2}{\sigma_\nu^2 + \sigma_\delta^2(S_2 - (\ln h_N)^2)}\right)$$

$$(\alpha|h, \sigma_\nu^2, \delta) \sim N\left(\frac{\sigma_\alpha^2((1-\delta)S_1 - \ln h_1 + \delta \ln h_N) + \sigma_\nu^2 \alpha_0}{\sigma_\nu^2 + (N-1)\sigma_\alpha^2}, \frac{\sigma_\nu^2 \sigma_\alpha^2}{\sigma_\nu^2 + (N-1)\sigma_\alpha^2}\right)$$

$$S_1 = \sum_{t=1}^N \ln h_t \quad S_2 = \sum_{t=1}^N (\ln h_t)^2 \quad S_3 = \sum_{t=2}^N \ln h_t \ln h_{t-1}$$

$$p(h_t|h_{t+1}, h_{t-1}, \delta, \alpha, \sigma_\nu^2) \propto \frac{1}{\sqrt{h_t}} \exp\left(-\frac{y_t^2}{2h_t}\right) \frac{1}{h_t} \exp\left(-\frac{(\ln h_t - \mu_t)^2}{2\sigma_t^2}\right) \quad (2)$$

$$\mu_t = \frac{\delta \ln h_{t+1} + \delta \ln h_{t-1} + (1 - \delta)\alpha}{1 + \delta^2}$$

$$\sigma^2 = \frac{\sigma_\nu^2}{1 + \delta^2}$$

In addition,  $h_1$  and  $h_N$  cannot be updated with 2, we will update them by directly drawing from autoregressive model of  $\ln h$ .

In summary, the outline of the algorithm is

1. Initialize  $h, \alpha, \delta, \sigma_\nu^2$
2. For  $t = 2, 3, \dots, N - 1$ , draw  $h_t$  from  $p(h_t|h_{t+1}, h_{t-1}, \delta, \alpha, \sigma_\nu^2)$
3. Draw  $\ln h_1$  from  $N(\alpha + \delta \ln h_2, \sigma_n u^2)$ ,  $\ln h_N$  from  $N(\alpha + \delta \ln h_{N-1}, \sigma_n u^2)$
4. Draw  $\sigma_\nu^2$  from  $(\sigma_\nu^2|h, \alpha, \delta)$
5. Draw  $\delta$  from  $(\delta|h, \alpha, \sigma_\nu^2)$
6. Draw  $\alpha$  from  $(\alpha|h, \delta, \sigma_\nu^2)$
7. Go to step 2

It is easy to simulate the posterior distribution of  $\sigma_\nu^2$ .  $\alpha$  and  $\delta$ , so the only nontrivial part of the MCMC is how to simulate  $p(h_t|h_{t+1}, h_{t-1}, \delta, \alpha, \sigma_\nu^2)$  which is a product of an inverse gamma and a lognormal distribution. Below we will give three sampling methods. The comparison of them will be the focus of this project.

## 2.2 Sampling Method 1: Metropolis-Hastings with Random Walk

Write (2) as a distribution of  $\ln h$  rather than  $h$

$$p(\ln h_t|h_{t+1}, h_{t-1}, \delta, \alpha, \sigma_\nu^2) \propto \frac{1}{\sqrt{h_t}} \exp\left(-\frac{y_t^2}{2h_t}\right) \exp\left(-\frac{(\ln h_t - \mu_t)^2}{2\sigma_t^2}\right)$$

Given  $\ln h_t^{i-1}$ . Each time we simply propose a  $\ln h_t^{i*}$  by drawing

$$N(\ln h_t^{i-1}, e^2)$$

where  $e^2$  is a preset parameter independent of other variables. And accept it with probability

$$\text{Min}\left(1, \frac{p(\ln h_t^{i*})}{p(\ln h_t^{i-1})}\right)$$

The algorithm is:

1. Draw  $\ln h_t^*$  from  $N(\ln h_t^{i-1}, e^2)$
2. Accept this value with probability  $\text{Min}(1, \frac{p(\ln h_t^{i*})}{p(\ln h_t^{i-1})})$
3. If accepted,  $h_t^i = h_t^*$ , else  $h_t^i = h_t^{i-1}$

### 2.3 Sampling Method 2: Metroplis-Hastings with Accept-Reject Sampling

This is the method proposed in . The idea is to refine the process of the proposing update in MH. We can "approximate" (2) by an inverse gamma distribution:

$$q(h_t) = \frac{\lambda^\phi}{\Gamma(\phi)} h^{-(\phi+1)} e^{-\frac{\lambda}{h_t}}$$

where

$$\lambda = \frac{1 - 2e^{\sigma^2}}{1 - e^{\sigma^2}} + \frac{1}{2}$$

$$\phi = (\lambda - 1)e^{\mu_t + \frac{\sigma^2}{2}} + \frac{y_t^2}{2}$$

and  $\sigma^2$  and  $\mu_t$  are defined under (2).

Define

$$c = 1.2 \left( \frac{p(h)}{q(h)} \right)_{h=\text{mode of } q}$$

We will propose candidate  $h_t^{i*}$  from  $IG(\lambda, \phi)$  and accept it with  $\text{Min}(1, \frac{p(h^*)}{cq(h^*)})$ , if rejected, repropose until accepted. The winner of accept-reject process will be the candidate of MH process with transition kernel  $f(h_t^*) = \text{Min}(p(h_t^*), cq(h_t^*))$ . The actual algorithm will be

1. Draw  $h_t^*$  from  $IG(\lambda, \phi)$ , note that both  $\lambda$  and  $\phi$  are functions of  $h_{t+1}, h_{t-1}$  and other parameters
2. Accept  $h_t^*$  with probability  $\text{Min}(1, \frac{p(h_t^*)}{cq(h_t^*)})$
3. If rejected, go to step 1.
4. If  $p(h_t^*) < cq(h_t^*)$ ,  $h_t^i = h_t^*$ . The algorithm ends.
5. Accept  $h_t^*$  with probability  $\text{Min}(1, \frac{p(h_t^*)/q(h_t^*)}{p(h_t^{i-1})/q(h_t^{i-1})})$
6. If accepted,  $h_t^i = h_t^*$ , else  $h_t^i = h_t^{i-1}$

## 2.4 Sampling Method 3: Pure Accept-Reject Sampling

$$\begin{aligned}
\ln p(\ln h_t | \dots) &= -\frac{1}{2} \ln h_t - \frac{y_t^2}{2h_t} - \frac{(\ln h_t - \mu_t)^2}{2\sigma^2} + \text{constants} \\
&\leq -\frac{1}{2} \ln h_t - \frac{y_t^2}{2} (\exp(-\mu_t)(1 + \mu_t - \ln h_t)) - \frac{(\ln h_t - \mu_t)^2}{2\sigma^2} + \text{constants} \\
&= -\frac{(\ln h_t - \mu'_t)^2}{2\sigma^2} + \text{constants}
\end{aligned} \tag{3}$$

Where

$$\mu'_t = \mu_t + \frac{\sigma^2}{2}(y_t^2 \exp(-\mu_t) - 1)$$

This observation leads to a standard reject-accept sampling:

1. Draw  $\ln h_t^*$  from  $N(\mu'_t, \sigma^2)$
2. Accept  $h_t^*$  with probability  $\text{Min}(1, g(h_t^*))$ , where
$$g(h_t) = \exp\left(\frac{y_t^2}{2}(\exp(-\mu_t))(1 + \mu_t - \ln h_t) - \frac{1}{h_t}\right)$$
3. If rejected, go to step 1.
4.  $h_t^i = h_t^*$

## 3 Test of Three Sampling Method

### 3.1 Data

### 3.2 Result

## 4 Comparison with GARCH

### 4.1 Data

## 5 Reference

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