

Package ‘HDNRA’

February 27, 2024

Type Package

Title High-Dimensional Location Testing with Normal-Reference Approaches

Version 1.0.0

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Description We provide a collection of various classical tests and latest normal-reference tests for comparing high-dimensional mean vectors including two-sample and general linear hypothesis testing (GLHT) problem. Some existing tests for two-sample problem [see Bai, Zhidong, and Hewa Saranadasa.(1996) <<https://www.jstor.org/stable/24306018>>; Chen, Song Xi, and Ying-Li Qin.(2010) <[doi:10.1214/09-aos716](https://doi.org/10.1214/09-aos716)>; Srivastava, Muni S., and Meng Du.(2008) <[doi:10.1016/j.jmva.2006.11.002](https://doi.org/10.1016/j.jmva.2006.11.002)>; Srivastava, Muni S., Shota Katayama, and Yutaka Kano.(2013)<[doi:10.1016/j.jmva.2012.08.014](https://doi.org/10.1016/j.jmva.2012.08.014)>]. Normal-reference tests for two-sample problem [see Zhang, Jin-Ting, Jia Guo, Bu Zhou, and Ming-Yen Cheng.(2020) <[doi:10.1080/01621459.2019.1604366](https://doi.org/10.1080/01621459.2019.1604366)>; Zhang, Jin-Ting, Bu Zhou, Jia Guo, and Tianming Zhu.(2021) <[doi:10.1016/j.jspi.2020.11.008](https://doi.org/10.1016/j.jspi.2020.11.008)>; Zhang, Liang, Tianming Zhu, and Jin-Ting Zhang.(2020) <[doi:10.1016/j.ecosta.2019.12.002](https://doi.org/10.1016/j.ecosta.2019.12.002)>; Zhang, Liang, Tianming Zhu, and Jin-Ting Zhang.(2023) <[doi:10.1080/02664763.2020.1834516](https://doi.org/10.1080/02664763.2020.1834516)>; Zhang, Jin-Ting, and Tianming Zhu.(2022) <[doi:10.1080/10485252.2021.2015768](https://doi.org/10.1080/10485252.2021.2015768)>; Zhang, Jin-Ting, and Tianming Zhu.(2022) <[doi:10.1007/s42519-021-00232-w](https://doi.org/10.1007/s42519-021-00232-w)>; Zhu, Tianming, Pengfei Wang, and Jin-Ting Zhang.(2023) <[doi:10.1007/s00180-023-01433-6](https://doi.org/10.1007/s00180-023-01433-6)>]. Some existing tests for GLHT problem [see Fujikoshi, Yasunori, Tetsuto Himeno, and Hirofumi Wakaki.(2004) <[doi:10.14490/jjss.34.19](https://doi.org/10.14490/jjss.34.19)>; Srivastava, Muni S., and Yasunori Fujikoshi.(2006) <[doi:10.1016/j.jmva.2005.08.010](https://doi.org/10.1016/j.jmva.2005.08.010)>; Yamada, Takayuki, and Muni S. Srivastava.(2012) <[doi:10.1080/03610926.2011.581786](https://doi.org/10.1080/03610926.2011.581786)>; Schott, James R.(2007) <[doi:10.1016/j.jmva.2006.11.007](https://doi.org/10.1016/j.jmva.2006.11.007)>; Zhou, Ting Zhang.(2017) <[doi:10.1016/j.jspi.2017.03.005](https://doi.org/10.1016/j.jspi.2017.03.005)>]. Normal-reference tests for GLHT problem [see Zhang, Jin-Ting, Jia Guo, and Bu Zhou.(2017) <[doi:10.1016/j.jmva.2017.01.002](https://doi.org/10.1016/j.jmva.2017.01.002)>; Zhang, Jin-Ting, Bu Zhou, and Jia Guo.(2022) <[doi:10.1016/j.jmva.2021.104816](https://doi.org/10.1016/j.jmva.2021.104816)>; Zhu, Tianming, Liang Zhang, and Jin-Ting Zhang.(2022) <[doi:10.5705/ss.202020.0362](https://doi.org/10.5705/ss.202020.0362)>; Zhu, Tianming, and Jin-Ting Zhang.(2022) <[doi:10.1007/s00180-021-01110-6](https://doi.org/10.1007/s00180-021-01110-6)>; Zhang, Jin-Ting, and Tianming Zhu.(2022) <[doi:10.1016/j.csda.2021.107385](https://doi.org/10.1016/j.csda.2021.107385)>].

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URL <https://nie23wp8738.github.io/HDNRA/>

BugReports <https://github.com/nie23wp8738/HDNRA/issues>

Encoding UTF-8

Roxygen list(markdown = TRUE)

RoxygenNote 7.3.1

LinkingTo Rcpp, RcppArmadillo

Imports expm, Rcpp, Rdpack, readr, stats

Suggests devtools, dplyr, knitr, rmarkdown, spelling, testthat (>= 3.0.0), tidyR

RdMacros Rdpack

Depends R (>= 2.10)

LazyData true

Language en-US

Config/testthat/edition 3

NeedsCompilation yes

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HDNRA-package

*HDNRA: High-Dimensional Location Testing with Normal-Reference Approaches***Description**

We provide a collection of various classical tests and latest normal-reference tests for comparing high-dimensional mean vectors including two-sample and general linear hypothesis testing (GLHT) problem. Some existing tests for two-sample problem [see Bai, Zhidong, and Hewa Saranadasa.(1996) <https://www.jstor.org/stable/24306018>; Chen, Song Xi, and Ying-Li Qin.(2010) [doi:10.1214/09aos716](https://doi.org/10.1214/09aos716); Srivastava, Muni S., and Meng Du.(2008) [doi:10.1016/j.jmva.2006.11.002](https://doi.org/10.1016/j.jmva.2006.11.002); Srivastava, Muni S., Shota Katayama, and Yutaka Kano.(2013) [doi:10.1016/j.jmva.2012.08.014](https://doi.org/10.1016/j.jmva.2012.08.014)]. Normal-reference tests for two-sample problem [see Zhang, Jin-Ting, Jia Guo, Bu Zhou, and Ming-Yen Cheng.(2020) [doi:10.1080/01621459.2019.1604366](https://doi.org/10.1080/01621459.2019.1604366); Zhang, Jin-Ting, Bu Zhou, Jia Guo, and Tianming Zhu.(2021) [doi:10.1016/j.jspi.2020.11.008](https://doi.org/10.1016/j.jspi.2020.11.008); Zhang, Liang, Tianming Zhu, and Jin-Ting Zhang.(2020) [doi:10.1016/j.ecosta.2019.12.002](https://doi.org/10.1016/j.ecosta.2019.12.002); Zhang, Liang, Tianming Zhu, and Jin-Ting Zhang.(2023) [doi:10.1080/02664763.2020.1834516](https://doi.org/10.1080/02664763.2020.1834516); Zhang, Jin-Ting, and Tianming Zhu.(2022) [doi:10.1080/10485252.2021.2015768](https://doi.org/10.1080/10485252.2021.2015768); Zhang, Jin-Ting, and Tianming Zhu.(2022) [doi:10.1007/s4251902100232w](https://doi.org/10.1007/s4251902100232w); Zhu, Tianming, Pengfei Wang, and Jin-Ting Zhang.(2023) [doi:10.1007/s00180023014336](https://doi.org/10.1007/s00180023014336)]. Some existing tests for GLHT problem [see Fujikoshi, Yasunori, Tetsuto Himeno, and Hirofumi Wakaki.(2004) [doi:10.14490/jjss.34.19](https://doi.org/10.14490/jjss.34.19); Srivastava, Muni S., and Yasunori Fujikoshi.(2006) [doi:10.1016/j.jmva.2005.08.010](https://doi.org/10.1016/j.jmva.2005.08.010); Yamada, Takayuki, and Muni S. Srivastava.(2012) [doi:10.1080/03610926.2011.581786](https://doi.org/10.1080/03610926.2011.581786); Schott, James R.(2007) [doi:10.1016/j.jmva.2006.11.007](https://doi.org/10.1016/j.jmva.2006.11.007); Zhou, Bu, Jia Guo, and Jin-Ting Zhang.(2017) [doi:10.1016/j.jspi.2017.03.005](https://doi.org/10.1016/j.jspi.2017.03.005)]. Normal-reference tests for GLHT problem [see Zhang, Jin-Ting, Jia Guo, and Bu Zhou.(2017) [doi:10.1016/j.jmva.2017.01.002](https://doi.org/10.1016/j.jmva.2017.01.002); Zhang, Jin-Ting, Bu Zhou, and Jia Guo.(2022) [doi:10.1016/j.jmva.2021.104816](https://doi.org/10.1016/j.jmva.2021.104816); Zhu, Tianming, Liang Zhang, and Jin-Ting Zhang.(2022) [doi:10.5705/ss.202020.0362](https://doi.org/10.5705/ss.202020.0362); Zhu, Tianming, and Jin-Ting Zhang.(2022) [doi:10.1007/s00180021-011106](https://doi.org/10.1007/s00180021-011106); Zhang, Jin-Ting, and Tianming Zhu.(2022) [doi:10.1016/j.csda.2021.107385](https://doi.org/10.1016/j.csda.2021.107385)].

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See Also

Useful links:

- <https://nie23wp8738.github.io/HDNRA/>
- Report bugs at <https://github.com/nie23wp8738/HDNRA/issues>

corneal	<i>HDNRA_data corneal</i>
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Description

This dataset was acquired during a keratoconus study, a collaborative project involving Ms.Nancy Tripoli and Dr.Kenneth L.Cohen of Department of Ophthalmology at the University of North Carolina, Chapel Hill. The fitted feature vectors for the complete corneal surface dataset collectively into a feature matrix with dimensions of 150×2000 .

Usage

```
data(corneal)
```

Format

'corneal':

A data frame with 150 observations on the following 4 groups.

normal group1 row 1 to row 43 in total 43 rows of the feature matrix correspond to observations from the normal group

unilateral suspect group2 row 44 to row 57 in total 14 rows of the feature matrix correspond to observations from the unilateral suspect group

suspect map group3 row 58 to row 78 in total 21 of the feature matrix correspond to observations from the suspect map group

clinical keratoconus group4 row 79 to row 150 in total 72 of the feature matrix correspond to observations from the clinical keratoconus group

References

Smaga Ł, Zhang J (2019). "Linear hypothesis testing with functional data." *Technometrics*, **61**(1), 99–110. doi:[10.1080/00401706.2018.1456976](https://doi.org/10.1080/00401706.2018.1456976).

Examples

```
library(HDNRA)
data(corneal)
dim(corneal)
group1 <- as.matrix(corneal[1:43, ])
dim(group1)
group2 <- as.matrix(corneal[44:57, ])
dim(group2)
group3 <- as.matrix(corneal[58:78, ])
dim(group3)
group4 <- as.matrix(corneal[79:150, ])
dim(group4)
```

COVID19

*HDNRA_data COVID19***Description**

A COVID19 data set from NCBI with ID GSE152641. The data set profiled peripheral blood from 24 healthy controls and 62 prospectively enrolled patients with community-acquired lower respiratory tract infection by SARS-COV-2 within the first 24 hours of hospital admission using RNA sequencing.

Usage

```
data(COVID19)
```

Format**'COVID19':**

A data frame with 86 observations on the following 2 groups.

healthy group1 row 2 to row 19, and row 82 to 87, in total 24 healthy controls

patients group2 row 20 to 81, in total 62 prospectively enrolled patients

References

Thair SA, He YD, Hasin-Brumshtein Y, Sakaram S, Pandya R, Toh J, Rawling D, Remmel M, Coyle S, Dalekos GN, others (2021). "Transcriptomic similarities and differences in host response between SARS-CoV-2 and other viral infections." *Iscience*, **24**(1). doi:10.1016/j.isci.2020.101947.

Examples

```
library(HDNRA)
data(COVID19)
dim(COVID19)
group1 <- as.matrix(COVID19[c(2:19, 82:87), ])
dim(group1)
group2 <- as.matrix(COVID19[-c(1:19, 82:87), ])
dim(group2)
```

glhtbf_zgz2017

*Test proposed by Zhou et al. (2017)***Description**

Zhou et al. (2017)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data under heteroscedasticity.

Usage

```
glhtbf_zgz2017(Y,G,n,p)
```

Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
G	A known full-rank coefficient matrix ($q \times k$) with $\text{rank}(\mathbf{G}) < k$.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, \dots, k.$$

It is of interest to test the following GLHT problem:

$$H_0 : \mathbf{G}\mathbf{M} = \mathbf{0}, \quad \text{vs. } H_1 : \mathbf{G}\mathbf{M} \neq \mathbf{0},$$

where $\mathbf{M} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_k)^\top$ is a $k \times p$ matrix collecting k mean vectors and $\mathbf{G} : q \times k$ is a known full-rank coefficient matrix with $\text{rank}(\mathbf{G}) < k$.

Let $\bar{\mathbf{y}}_i, i = 1, \dots, k$ be the sample mean vectors and $\hat{\boldsymbol{\Sigma}}_i, i = 1, \dots, k$ be the sample covariance matrices.

Zhou et al. (2017) proposed the following U-statistic based test statistic:

$$T_{ZGZ} = \|\mathbf{C}\hat{\boldsymbol{\mu}}\|^2 - \sum_{i=1}^k h_{ii} \text{tr}(\hat{\boldsymbol{\Sigma}}_i)/n_i,$$

where $\mathbf{C} = [(\mathbf{G}\mathbf{D}\mathbf{G}^\top)^{-1/2}\mathbf{G}] \otimes \mathbf{I}_p$, $\mathbf{D} = \text{diag}(1/n_1, \dots, 1/n_k)$, and h_{ij} is the (i, j) th entry of the $k \times k$ matrix $\mathbf{H} = \mathbf{G}^\top (\mathbf{G}\mathbf{D}\mathbf{G}^\top)^{-1} \mathbf{G}$.

They showed that under the null hypothesis, T_{ZGZ} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Zhou et al. (2017).

p.value the p -value of the test proposed by Zhou et al. (2017).

References

Zhou B, Guo J, Zhang J (2017). “High-dimensional general linear hypothesis testing under heteroscedasticity.” *Journal of Statistical Planning and Inference*, **188**, 36–54. doi:10.1016/j.jspi.2017.03.005.

Examples

```
set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
avec <- seq(1, k)
```

```

Y <- list()
for (g in 1:k) {
  a <- avec[g]
  y <- (-2 * sqrt(a * (1 - rho)) + sqrt(4 * a * (1 - rho) + 4 * p * a * rho)) / (2 * p)
  x <- y + sqrt(a * (1 - rho))
  Gamma <- matrix(rep(y, p * p), nrow = p)
  diag(Gamma) <- rep(x, p)
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
G <- cbind(diag(k - 1), rep(-1, k - 1))
glhtbf_zgz2017(Y, G, n, p)

```

glhtbf_zz2022

Test proposed by Zhang and Zhu (2022)

Description

Zhang and Zhu (2022)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data under heteroscedasticity.

Usage

```
glhtbf_zz2022(Y, G, n, p)
```

Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
G	A known full-rank coefficient matrix ($q \times k$) with $\text{rank}(\mathbf{G}) < k$.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i. i. d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, \dots, k.$$

It is of interest to test the following GLHT problem:

$$H_0 : \mathbf{G}\mathbf{M} = \mathbf{0}, \quad \text{vs. } H_1 : \mathbf{G}\mathbf{M} \neq \mathbf{0},$$

where $\mathbf{M} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_k)^\top$ is a $k \times p$ matrix collecting k mean vectors and $\mathbf{G} : q \times k$ is a known full-rank coefficient matrix with $\text{rank}(\mathbf{G}) < k$.

Let $\bar{\mathbf{y}}_i, i = 1, \dots, k$ be the sample mean vectors and $\hat{\boldsymbol{\Sigma}}_i, i = 1, \dots, k$ be the sample covariance matrices.

Zhang and Zhu (2022) proposed the following U-statistic based test statistic:

$$T_{ZZ} = \|\mathbf{C}\hat{\boldsymbol{\mu}}\|^2 - \sum_{i=1}^k h_{ii} \text{tr}(\hat{\boldsymbol{\Sigma}}_i)/n_i,$$

where $C = [(GDG^T)^{-1/2}G] \otimes I_p$, $D = \text{diag}(1/n_1, \dots, 1/n_k)$, and h_{ij} is the (i, j) th entry of the $k \times k$ matrix $H = G^T (GDG^T)^{-1} G$.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p -value of the test proposed by Zhang and Zhu (2022).

statistic the test statistic proposed by Zhang and Zhu (2022).

beta0 the parameter used in Zhang and Zhu (2022)'s test.

beta1 the parameter used in Zhang and Zhu (2022)'s test.

df estimated approximate degrees of freedom of Zhang and Zhu (2022)'s test.

References

Zhang J, Zhu T (2022). "A new normal reference test for linear hypothesis testing in high-dimensional heteroscedastic one-way MANOVA." *Computational Statistics & Data Analysis*, **168**, 107385. doi:10.1016/j.csda.2021.107385.

Examples

```
set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
avec <- seq(1, k)
Y <- list()
for (g in 1:k) {
  a <- avec[g]
  y <- (-2 * sqrt(a * (1 - rho)) + sqrt(4 * a * (1 - rho) + 4 * p * a * rho)) / (2 * p)
  x <- y + sqrt(a * (1 - rho))
  Gamma <- matrix(rep(y, p * p), nrow = p)
  diag(Gamma) <- rep(x, p)
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
G <- cbind(diag(k - 1), rep(-1, k - 1))
glhtbf_zzg2022(Y, G, n, p)
```

glhtbf_zzg2022

Test proposed by Zhang et al. (2022)

Description

Zhang et al. (2022)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data under heteroscedasticity.

Usage

```
glhtbf_zzg2022(Y, G, n, p)
```


Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
G	A known full-rank coefficient matrix ($q \times k$) with $\text{rank}(\mathbf{G}) < k$.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, \dots, k.$$

It is of interest to test the following GLHT problem:

$$H_0 : \mathbf{G}\mathbf{M} = \mathbf{0}, \quad \text{vs.} \quad H_1 : \mathbf{G}\mathbf{M} \neq \mathbf{0},$$

where $\mathbf{M} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_k)^\top$ is a $k \times p$ matrix collecting k mean vectors and $\mathbf{G} : q \times k$ is a known full-rank coefficient matrix with $\text{rank}(\mathbf{G}) < k$.

Zhang et al. (2022) proposed the following test statistic:

$$T_{ZZG} = \|\mathbf{C}\hat{\boldsymbol{\mu}}\|^2,$$

where $\mathbf{C} = [(\mathbf{G}\mathbf{D}\mathbf{G}^\top)^{-1/2}\mathbf{G}] \otimes \mathbf{I}_p$ with $\mathbf{D} = \text{diag}(1/n_1, \dots, 1/n_k)$, and $\hat{\boldsymbol{\mu}} = (\bar{\mathbf{y}}_1^\top, \dots, \bar{\mathbf{y}}_k^\top)^\top$ with $\bar{\mathbf{y}}_i, i = 1, \dots, k$ being the sample mean vectors.

They showed that under the null hypothesis, T_{ZZG} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p -value of the test proposed by Zhang et al. (2022)

statistic the test statistic proposed by Zhang et al. (2022).

beta the parameters used in Zhang et al. (2022)'s test.

df estimated approximate degrees of freedom of Zhang et al. (2022)'s test.

References

Zhang J, Zhou B, Guo J (2022). "Linear hypothesis testing in high-dimensional heteroscedastic one-way MANOVA: A normal reference L^2 -norm based test." *Journal of Multivariate Analysis*, **187**, 104816. doi:10.1016/j.jmva.2021.104816.

Examples

```
set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
avec <- seq(1, k)
```

```

Y <- list()
for (g in 1:k) {
  a <- avec[g]
  y <- (-2 * sqrt(a * (1 - rho)) + sqrt(4 * a * (1 - rho) + 4 * p * a * rho)) / (2 * p)
  x <- y + sqrt(a * (1 - rho))
  Gamma <- matrix(rep(y, p * p), nrow = p)
  diag(Gamma) <- rep(x, p)
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
G <- cbind(diag(k - 1), rep(-1, k - 1))
glhtbf_zzg2022(Y, G, n, p)

```

glht_fhw2004

Test proposed by Fujikoshi et al. (2004)

Description

Fujikoshi et al. (2004)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_fhw2004(Y, X, C)
```

Arguments

- | | |
|---|--|
| Y | An $n \times p$ response matrix obtained by independently observing a p -dimensional response variable for n subjects. |
| X | A known $n \times k$ full-rank design matrix with $\text{rank}(\mathbf{G}) = k < n$. |
| C | A known matrix of size $q \times k$ with $\text{rank}(\mathbf{C}) = q < k$. |

Details

A high-dimensional linear regression model can be expressed as

$$\mathbf{Y} = \mathbf{X}\mathbf{\Theta} + \boldsymbol{\epsilon},$$

where $\mathbf{\Theta}$ is a $k \times p$ unknown parameter matrix and $\boldsymbol{\epsilon}$ is an $n \times p$ error matrix.

It is of interest to test the following GLHT problem

$$H_0 : \mathbf{C}\mathbf{\Theta} = \mathbf{0}, \quad \text{vs.} \quad H_1 : \mathbf{C}\mathbf{\Theta} \neq \mathbf{0}.$$

Fujikoshi et al. (2004) proposed the following test statistic:

$$T_{FW} = \sqrt{p} \left[(n - k) \frac{\text{tr}(\mathbf{S}_h)}{\text{tr}(\mathbf{S}_e)} - q \right],$$

where \mathbf{S}_h and \mathbf{S}_e are the matrices of sums of squares and products due to the hypothesis and the error, respectively.

They showed that under the null hypothesis, T_{FW} is asymptotically normally distributed.

Value

A (list) object of S3 class `hstest` containing the following elements:

statistic the test statistic proposed by Fujikoshi et al. (2004).

p.value the p -value of the test proposed by Fujikoshi et al. (2004).

References

Fujikoshi Y, Himeno T, Wakaki H (2004). “Asymptotic results of a high dimensional MANOVA test and power comparison when the dimension is large compared to the sample size.” *Journal of the Japan Statistical Society*, **34**(1), 19–26. doi:[10.14490/jjss.34.19](https://doi.org/10.14490/jjss.34.19).

Examples

```
set.seed(1234)
k <- 3
q <- k-1
p <- 50
n <- c(25,30,40)
rho <- 0.01
Theta <- matrix(rep(0,k*p),nrow=k)
X <- matrix(c(rep(1,n[1]),rep(0,sum(n)),rep(1,n[2]),rep(0,sum(n)),rep(1,n[3])),ncol=k,nrow=sum(n))
y <- (-2*sqrt(1-rho)+sqrt(4*(1-rho)+4*p*rho))/(2*p)
x <- y+sqrt((1-rho))
Gamma <- matrix(rep(y,p*p),nrow=p)
diag(Gamma) <- rep(x,p)
U <- matrix(ncol = sum(n),nrow=p)
for(i in 1:sum(n)){
  U[,i] <- rnorm(p,0,1)
}
Y <- X%%Theta+t(U)%%Gamma
C <- cbind(diag(q),-rep(1,q))
glht_fhw2004(Y,X,C)
```

glht_sf2006

*Test proposed by Srivastava and Fujikoshi (2006)***Description**

Srivastava and Fujikoshi (2006)’s test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_sf2006(Y,X,C)
```

Arguments

- | | |
|---|--|
| Y | An $n \times p$ response matrix obtained by independently observing a p -dimensional response variable for n subjects. |
| X | A known $n \times k$ full-rank design matrix with $\text{rank}(\mathbf{G}) = k < n$. |
| C | A known matrix of size $q \times k$ with $\text{rank}(\mathbf{C}) = q < k$. |

Details

A high-dimensional linear regression model can be expressed as

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\Theta} + \boldsymbol{\epsilon},$$

where $\boldsymbol{\Theta}$ is a $k \times p$ unknown parameter matrix and $\boldsymbol{\epsilon}$ is an $n \times p$ error matrix.

It is of interest to test the following GLHT problem

$$H_0 : \mathbf{C}\boldsymbol{\Theta} = \mathbf{0}, \quad \text{vs.} \quad H_1 : \mathbf{C}\boldsymbol{\Theta} \neq \mathbf{0}.$$

Srivastava and Fujikoshi (2006) proposed the following test statistic:

$$T_{SF} = [2q\hat{a}_2(1 + (n - k)^{-1}q)]^{-1/2} \left[\frac{\text{tr}(\mathbf{B})}{\sqrt{p}} - \frac{q}{\sqrt{n - k}} \frac{\text{tr}(\mathbf{W})}{\sqrt{(n - k)p}} \right].$$

where \mathbf{W} and \mathbf{B} are the matrix of sum of squares and products due to error and the error, respectively, and $\hat{a}_2 = [\text{tr}(\mathbf{W}^2) - \text{tr}^2(\mathbf{W})/(n - k)]/[(n - k - 1)(n - k + 2)p]$. They showed that under the null hypothesis, T_{SF} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Srivastava and Fujikoshi (2006).

p.value the p -value of the test proposed by Srivastava and Fujikoshi (2006).

References

Srivastava MS, Fujikoshi Y (2006). “Multivariate analysis of variance with fewer observations than the dimension.” *Journal of Multivariate Analysis*, **97**(9), 1927–1940. doi:10.1016/j.jmva.2005.08.010.

Examples

```
set.seed(1234)
k <- 3
q <- k-1
p <- 50
n <- c(25,30,40)
rho <- 0.01
Theta <- matrix(rep(0,k*p),nrow=k)
X <- matrix(c(rep(1,n[1]),rep(0,sum(n)),rep(1,n[2]),rep(0,sum(n)),rep(1,n[3])),ncol=k,nrow=sum(n))
y <- (-2*sqrt(1-rho)+sqrt(4*(1-rho)+4*p*rho))/(2*p)
x <- y+sqrt((1-rho))
Gamma <- matrix(rep(y,p*p),nrow=p)
diag(Gamma) <- rep(x,p)
U <- matrix(ncol = sum(n),nrow=p)
for(i in 1:sum(n)){
  U[,i] <- rnorm(p,0,1)
}
Y <- X%%Theta+t(U)%%Gamma
C <- cbind(diag(q),-rep(1,q))
glht_sf2006(Y,X,C)
```

glht_ys2012

Test proposed by Yamada and Srivastava (2012)

Description

Yamada and Srivastava (2012)'test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_ys2012(Y,X,C)
```

Arguments

Y	An $n \times p$ response matrix obtained by independently observing a p -dimensional response variable for n subjects.
X	A known $n \times k$ full-rank design matrix with $\text{rank}(\mathbf{G}) = k < n$.
C	A known matrix of size $q \times k$ with $\text{rank}(\mathbf{C}) = q < k$.

Details

A high-dimensional linear regression model can be expressed as

$$\mathbf{Y} = \mathbf{X}\mathbf{\Theta} + \boldsymbol{\epsilon},$$

where $\mathbf{\Theta}$ is a $k \times p$ unknown parameter matrix and $\boldsymbol{\epsilon}$ is an $n \times p$ error matrix.

It is of interest to test the following GLHT problem

$$H_0 : \mathbf{C}\mathbf{\Theta} = \mathbf{0}, \quad \text{vs. } H_1 : \mathbf{C}\mathbf{\Theta} \neq \mathbf{0}.$$

Yamada and Srivastava (2012) proposed the following test statistic:

$$T_{YS} = \frac{(n-k) \text{tr}(\mathbf{S}_h \mathbf{D}_{\mathbf{S}_e}^{-1}) - (n-k)pq/(n-k-2)}{\sqrt{2q[\text{tr}(\mathbf{R}^2) - p^2/(n-k)]c_{p,n}}},$$

where \mathbf{S}_h and \mathbf{S}_e are the variation matrices due to the hypothesis and error, respectively, and $\mathbf{D}_{\mathbf{S}_e}$ and \mathbf{R} are diagonal matrix with the diagonal elements of \mathbf{S}_e and the sample correlation matrix, respectively. $c_{p,n}$ is the adjustment coefficient proposed by Yamada and Srivastava (2012). They showed that under the null hypothesis, T_{YS} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Yamada and Srivastava (2012).

p.value the p -value of the test proposed by Yamada and Srivastava (2012).

References

Yamada T, Srivastava MS (2012). "A test for multivariate analysis of variance in high dimension." *Communications in Statistics-Theory and Methods*, **41**(13-14), 2602–2615. doi:10.1080/03610926.2011.581786.

Examples

```

set.seed(1234)
k <- 3
q <- k-1
p <- 50
n <- c(25,30,40)
rho <- 0.01
Theta <- matrix(rep(0,k*p),nrow=k)
X <- matrix(c(rep(1,n[1]),rep(0,sum(n)),rep(1,n[2]),rep(0,sum(n)),rep(1,n[3])),ncol=k,nrow=sum(n))
y <- (-2*sqrt(1-rho)+sqrt(4*(1-rho)+4*p*rho))/(2*p)
x <- y+sqrt((1-rho))
Gamma <- matrix(rep(y,p*p),nrow=p)
diag(Gamma) <- rep(x,p)
U <- matrix(ncol = sum(n),nrow=p)
for(i in 1:sum(n)){
  U[,i] <- rnorm(p,0,1)
}
Y <- X%%Theta+t(U)%%Gamma
C <- cbind(diag(q),-rep(1,q))
glht_ys2012(Y,X,C)

```

glht_zgz2017

Test proposed by Zhang et al. (2017)

Description

Zhang et al. (2017)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_zgz2017(Y,G,n,p)
```

Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
G	A known full-rank coefficient matrix ($q \times k$) with $\text{rank}(\mathbf{G}) < k$.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, \dots, k.$$

It is of interest to test the following GLHT problem:

$$H_0 : \mathbf{GM} = \mathbf{0}, \quad \text{vs.} \quad H_1 : \mathbf{GM} \neq \mathbf{0},$$

where $\mathbf{M} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_k)^\top$ is a $k \times p$ matrix collecting k mean vectors and $\mathbf{G} : q \times k$ is a known full-rank coefficient matrix with $\text{rank}(\mathbf{G}) < k$.

Zhang et al. (2017) proposed the following test statistic:

$$T_{ZGZ} = \|\mathbf{C}\hat{\boldsymbol{\mu}}\|^2,$$

where $\mathbf{C} = [(\mathbf{G}\mathbf{D}\mathbf{G}^\top)^{-1/2}\mathbf{G}] \otimes \mathbf{I}_p$, and $\hat{\boldsymbol{\mu}} = (\bar{\mathbf{y}}_1^\top, \dots, \bar{\mathbf{y}}_k^\top)^\top$, with $\bar{\mathbf{y}}_i, i = 1, \dots, k$ being the sample mean vectors and $\mathbf{D} = \text{diag}(1/n_1, \dots, 1/n_k)$.

They showed that under the null hypothesis, T_{ZGZ} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Zhang et al. (2017)

p.value the p -value of the test proposed by Zhang et al. (2017).

beta the parameters used in Zhang et al. (2017)'s test.

df estimated approximate degrees of freedom of Zhang et al.(2017)'s test.

References

Zhang J, Guo J, Zhou B (2017). "Linear hypothesis testing in high-dimensional one-way MANOVA." *Journal of Multivariate Analysis*, **155**, 200–216. doi:10.1016/j.jmva.2017.01.002.

Examples

```
set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Y <- list()
for (g in 1:k) {
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
G <- cbind(diag(k - 1), rep(-1, k - 1))
glht_zgz2017(Y, G, n, p)
```

Description

Zhu and Zhang (2022)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_zz2022(Y, G, n, p)
```

Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
G	A known full-rank coefficient matrix ($q \times k$) with $\text{rank}(\mathbf{G}) < k$.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i. i. d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, \dots, k.$$

It is of interest to test the following GLHT problem:

$$H_0 : \mathbf{G}\mathbf{M} = \mathbf{0}, \quad \text{vs.} \quad H_1 : \mathbf{G}\mathbf{M} \neq \mathbf{0},$$

where $\mathbf{M} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_k)^\top$ is a $k \times p$ matrix collecting k mean vectors and $\mathbf{G} : q \times k$ is a known full-rank coefficient matrix with $\text{rank}(\mathbf{G}) < k$.

Zhu and Zhang (2022) proposed the following test statistic:

$$T_{ZZ} = \|\mathbf{C}\hat{\boldsymbol{\mu}}\|^2 - q \text{tr}(\hat{\boldsymbol{\Sigma}}),$$

where $\mathbf{C} = [(\mathbf{G}\mathbf{D}\mathbf{G}^\top)^{-1/2}\mathbf{G}] \otimes \mathbf{I}_p$, and $\hat{\boldsymbol{\mu}} = (\bar{\mathbf{y}}_1^\top, \dots, \bar{\mathbf{y}}_k^\top)^\top$, with $\bar{\mathbf{y}}_i, i = 1, \dots, k$ being the sample mean vectors and $\hat{\boldsymbol{\Sigma}}$ being the usual pooled sample covariance matrix of the k samples.

They showed that under the null hypothesis, T_{ZZ} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p -value of the test proposed by Zhu and Zhang (2022).

statistic the test statistic proposed by Zhu and Zhang (2022).

beta0 the parameter used in Zhu and Zhang (2022)'s test.

beta1 the parameter used in Zhu and Zhang (2022)'s test.

df estimated approximate degrees of freedom of Zhu and Zhang (2022)'s test.

References

Zhu T, Zhang J (2022). "Linear hypothesis testing in high-dimensional one-way MANOVA: a new normal reference approach." *Computational Statistics*, **37**(1), 1–27. doi:10.1007/s0018002101110-6.

Examples

```

set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Y <- list()
for (g in 1:k) {
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
G <- cbind(diag(k - 1), rep(-1, k - 1))
glht_zzz2022(Y, G, n, p)

```

glht_zzz2022

Test proposed by Zhu et al. (2022)

Description

Zhu et al. (2022)'s test for general linear hypothesis testing (GLHT) problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
glht_zzz2022(Y, X, C)
```

Arguments

Y	An $n \times p$ response matrix obtained by independently observing a p -dimensional response variable for n subjects.
X	A known $n \times k$ full-rank design matrix with $\text{rank}(\mathbf{G}) = k < n - 2$.
C	A known matrix of size $q \times k$ with $\text{rank}(\mathbf{C}) = q < k$.

Details

A high-dimensional linear regression model can be expressed as

$$\mathbf{Y} = \mathbf{X}\mathbf{\Theta} + \boldsymbol{\epsilon},$$

where $\mathbf{\Theta}$ is a $k \times p$ unknown parameter matrix and $\boldsymbol{\epsilon}$ is an $n \times p$ error matrix.

It is of interest to test the following GLHT problem

$$H_0 : \mathbf{C}\mathbf{\Theta} = \mathbf{0}, \quad \text{vs. } H_1 : \mathbf{C}\mathbf{\Theta} \neq \mathbf{0}.$$

Zhu et al. (2022) proposed the following test statistic:

$$T_{ZZZ} = \frac{(n - k - 2)}{(n - k)pq} \text{tr}(\mathbf{S}_h \mathbf{D}^{-1}),$$

where S_h and S_e are the variation matrices due to the hypothesis and error, respectively, and D is the diagonal matrix with the diagonal elements of $S_e/(n - k)$. They showed that under the null hypothesis, T_{ZZZ} and a chi-squared-type mixture have the same limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p -value of the test proposed by Zhu et al. (2022)

statistic the test statistic proposed by Zhu et al. (2022).

df estimated approximate degrees of freedom of Zhu et al. (2022)'s test.

References

Zhu T, Zhang L, Zhang J (2023). "Hypothesis Testing in High-Dimensional Linear Regression: A Normal Reference Scale-Invariant Test." *Statistica Sinica*. doi:10.5705/ss.202020.0362.

Examples

```
set.seed(1234)
k <- 3
q <- k - 1
p <- 50
n <- c(25, 30, 40)
rho <- 0.01
Theta <- matrix(rep(0, k * p), nrow = k)
X <- matrix(c(rep(1, n[1]), rep(0, sum(n)), rep(1, n[2]), rep(0, sum(n)), rep(1, n[3])),
  ncol = k, nrow = sum(n)
)
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
U <- matrix(ncol = sum(n), nrow = p)
for (i in 1:sum(n)) {
  U[, i] <- rnorm(p, 0, 1)
}
Y <- X %*% Theta + t(U) %*% Gamma
C <- cbind(diag(q), -rep(1, q))
glht_zzz2022(Y, X, C)
```

ks_s2007

Test proposed by Schott (2007)

Description

Schott, J. R. (2007)'s test for one-way MANOVA problem for high-dimensional data with assuming that underlying covariance matrices are the same.

Usage

```
ks_s2007(Y, n, p)
```

Arguments

Y	A list of k data matrices. The i th element represents the data matrix ($p \times n_i$) from the i th population with each column representing a p -dimensional observation.
n	A vector of k sample sizes. The i th element represents the sample size of group i , n_i .
p	The dimension of data.

Details

Suppose we have the following k independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, \dots, k.$$

It is of interest to test the following one-way MANOVA problem:

$$H_0 : \boldsymbol{\mu}_1 = \dots = \boldsymbol{\mu}_k, \quad \text{vs.} \quad H_1 : H_0 \text{ is not true.}$$

Schott (2007) proposed the following test statistic:

$$T_S = [\text{tr}(\mathbf{H})/h - \text{tr}(\mathbf{E})/e] / \sqrt{N-1},$$

where $\mathbf{H} = \sum_{i=1}^k n_i (\bar{\mathbf{y}}_i - \bar{\mathbf{y}})(\bar{\mathbf{y}}_i - \bar{\mathbf{y}})^\top$, $\mathbf{E} = \sum_{i=1}^k \sum_{j=1}^{n_i} (\mathbf{y}_{ij} - \bar{\mathbf{y}}_i)(\mathbf{y}_{ij} - \bar{\mathbf{y}}_i)^\top$, $h = k - 1$, and $e = N - k$, with $N = n_1 + \dots + n_k$. They showed that under the null hypothesis, T_S is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Schott (2007).

p.value the p -value of the test proposed by Schott (2007).

References

Schott JR (2007). “Some high-dimensional tests for a one-way MANOVA.” *Journal of Multivariate Analysis*, **98**(9), 1825–1839. doi:10.1016/j.jmva.2006.11.007.

Examples

```
set.seed(1234)
k <- 3
p <- 50
n <- c(25, 30, 40)
rho <- 0.1
M <- matrix(rep(0, k * p), nrow = k, ncol = p)
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Y <- list()
for (g in 1:k) {
  Z <- matrix(rnorm(n[g] * p, mean = 0, sd = 1), p, n[g])
  Y[[g]] <- Gamma %*% Z + t(t(M[g, ])) %*% (rep(1, n[g]))
}
ks_s2007(Y, n, p)
```

tsbf_cq2010

*Test proposed by Chen and Qin (2010)***Description**

Chen and Qin (2010)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_cq2010(y1, y2)
```

Arguments

- | | |
|----|---|
| y1 | The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation. |
| y2 | The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation. |

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Chen and Qin (2010) proposed the following test statistic:

$$T_{CQ} = \frac{\sum_{i \neq j}^{n_1} \mathbf{y}_{1i}^\top \mathbf{y}_{1j}}{n_1(n_1 - 1)} + \frac{\sum_{i \neq j}^{n_2} \mathbf{y}_{2i}^\top \mathbf{y}_{2j}}{n_2(n_2 - 1)} - 2 \frac{\sum_{i=1}^{n_1} \sum_{j=1}^{n_2} \mathbf{y}_{1i}^\top \mathbf{y}_{2j}}{n_1 n_2}.$$

They showed that under the null hypothesis, T_{CQ} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Chen and Qin (2010)

p.value the p -value of the test proposed by Chen and Qin (2010).

References

Chen SX, Qin Y (2010). "A two-sample test for high-dimensional data with applications to gene-set testing." *The Annals of Statistics*, **38**(2). doi:10.1214/09a0716.

Examples

```

set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1*p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2*p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %*% Z1 + mu1%*(rep(1,n1))
y2 <- Gamma2 %*% Z2 + mu2%*(rep(1,n2))
tsbf_cq2010(y1, y2)

```

tsbf_skk2013

Test proposed by Srivastava et al. (2013)

Description

Srivastava et al. (2013)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_skk2013(y1, y2)
```

Arguments

- | | |
|----|---|
| y1 | The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation. |
| y2 | The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation. |

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Srivastava et al. (2013) proposed the following test statistic:

$$T_{SKK} = \frac{(\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)^\top \hat{\mathbf{D}}^{-1} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2) - p}{\sqrt{2\widehat{\text{Var}}(\hat{q}_n)c_{p,n}}},$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors, $\hat{\mathbf{D}} = \hat{\mathbf{D}}_1/n_1 + \hat{\mathbf{D}}_2/n_2$ with $\hat{\mathbf{D}}_i, i = 1, 2$ being the diagonal matrices consisting of only the diagonal elements of the sample covariance matrices. $\widehat{\text{Var}}(\hat{q}_n)$ is given by equation (1.18) in Srivastava et al. (2013), and $c_{p,n}$ is the adjustment coefficient proposed by Srivastava et al. (2013). They showed that under the null hypothesis, T_{SKK} is asymptotically normally distributed.

Value

A (list) object of S3 class `hstest` containing the following elements:

statistic the test statistic proposed by Srivastava et al. (2013)

p.value the p -value of the test proposed by Srivastava et al. (2013)

cpn the adjustment coefficient proposed by Srivastava et al. (2013)

References

Srivastava MS, Katayama S, Kano Y (2013). “A two sample test in high dimensional data.” *Journal of Multivariate Analysis*, **114**, 349–358. doi:10.1016/j.jmva.2012.08.014.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(rep(0, p))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1*p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2*p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %*% Z1 + mu1%*(rep(1,n1))
y2 <- Gamma2 %*% Z2 + mu2%*(rep(1,n2))
tsbf_skk2013(y1, y2)
```

tsbf_zwz2023

*Test proposed by Zhu et al. (2023)***Description**

Zhu et al. (2023)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_zwz2023(y1, y2)
```

Arguments

y1 The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation.

y2 The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation.

Details

Suppose we have two independent high-dimensional samples:

$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}$, are i. i. d. with $E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i$, $\text{Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i$, $i = 1, 2$.

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhu et al. (2023) proposed the following test statistic:

$$T_{ZWZ} = \frac{n_1 n_2 n^{-1} \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2}{\text{tr}(\hat{\boldsymbol{\Omega}}_n)},$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors and $\hat{\boldsymbol{\Omega}}_n$ is the estimator of $\text{Cov}[(n_1 n_2 / n)^{1/2} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)]$. They showed that under the null hypothesis, T_{ZWZ} and an F-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p-value of the test proposed by Zhu et al. (2023).

statistic the test statistic proposed by Zhu et al. (2023).

df1 estimated approximate degrees of freedom d_1 of Zhu et al. (2023)'s test.

df2 estimated approximate degrees of freedom d_2 of Zhu et al. (2023)'s test.

References

Zhu T, Wang P, Zhang J (2023). "Two-sample Behrens–Fisher problems for high-dimensional data: a normal reference F-type test." *Computational Statistics*, 1–24. doi:10.1007/s00180023014336.

Examples

```

set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1*p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2*p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %*% Z1 + mu1%*(rep(1,n1))
y2 <- Gamma2 %*% Z2 + mu2%*(rep(1,n2))
tsbf_zwz2023(y1, y2)

```

tsbf_zz2022

Test proposed by Zhang and Zhu (2022)

Description

Zhang and Zhu (2022)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_zz2022(y1, y2)
```

Arguments

- | | |
|----|---|
| y1 | The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation. |
| y2 | The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation. |

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhang and Zhu (2022) proposed the following test statistic:

$$T_{ZZ} = \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2 - \text{tr}(\hat{\mathbf{\Omega}}_n),$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors and $\hat{\mathbf{\Omega}}_n$ is the estimator of $\text{Cov}(\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)$. They showed that under the null hypothesis, T_{ZZ} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p-value of the test proposed by Zhang and Zhu (2022).

statistic the test statistic proposed by Zhang and Zhu (2022).

beta0 parameter used in Zhang and Zhu (2022)'s test.

beta1 parameter used in Zhang and Zhu (2022)'s test.

df estimated approximate degrees of freedom of Zhang and Zhu (2022)'s test.

References

Zhang J, Zhu T (2022). "A further study on Chen-Qin's test for two-sample Behrens–Fisher problems for high-dimensional data." *Journal of Statistical Theory and Practice*, **16**(1), 1. doi:10.1007/s4251902100232w.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma2 %*% Z2 + mu2 %*% (rep(1, n2))
tsbf_zz2022(y1, y2)
```

tsbf_zzg2021

Test proposed by Zhang et al. (2021)

Description

Zhang et al. (2021)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_zzg2021(y1, y2)
```

Arguments

y1 The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation.

y2 The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation.

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhang et al.(2021) proposed the following test statistic:

$$T_{ZZGZ} = \frac{n_1 n_2}{n} \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2,$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors. They showed that under the null hypothesis, T_{ZZGZ} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p-value of the test proposed by Zhang et al. (2021).

statistic the test statistic proposed by Zhang et al. (2021).

beta parameter used in Zhang et al. (2021)'s test.

df estimated approximate degrees of freedom of Zhang et al. (2021)'s test.

References

Zhang J, Zhou B, Guo J, Zhu T (2021). "Two-sample Behrens-Fisher problems for high-dimensional data: A normal reference approach." *Journal of Statistical Planning and Inference*, **213**, 142–161. [doi:10.1016/j.jspi.2020.11.008](https://doi.org/10.1016/j.jspi.2020.11.008).

Examples

```

set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1*p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2*p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %%% Z1 + mu1%%(rep(1, n1))
y2 <- Gamma2 %%% Z2 + mu2%%(rep(1, n2))
tsbf_zzz2021(y1, y2)

```

tsbf_zzz2023

Test proposed by Zhang et al. (2023)

Description

Zhang et al. (2023)'s test for testing equality of two-sample high-dimensional mean vectors without assuming that two covariance matrices are the same.

Usage

```
tsbf_zzz2023(y1, y2, cutoff)
```

Arguments

y1	The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation.
y2	The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation.
cutoff	An empirical criterion for applying the adjustment coefficient

Details

Suppose we have two independent high-dimensional samples:

$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}$, are i.i.d. with $E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i$, $\text{Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}_i$, $i = 1, 2$.

The primary object is to test

$$H_0 : \mu_1 = \mu_2 \text{ versus } H_1 : \mu_1 \neq \mu_2.$$

Zhang et al.(2023) proposed the following test statistic:

$$T_{ZZZ} = \frac{n_1 n_2}{np} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)^\top \hat{\mathbf{D}}_n^{-1} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2),$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors, and $\hat{\mathbf{D}}_n = \text{diag}(\hat{\Sigma}_1/n + \hat{\Sigma}_2/n)$ with $n = n_1 + n_2$. They showed that under the null hypothesis, T_{ZZZ} and a chi-squared-type mixture have the same limiting distribution.

Value

A (list) object of S3 class `hstest` containing the following elements:

p.value the p-value of the test proposed by Zhang et al. (2023)’s test.

statistic the test statistic proposed by Zhang et al. (2023)’s test.

df estimated approximate degrees of freedom of Zhang et al. (2023)’s test.

cpn the adjustment coefficient used in Zhang et al. (2023)’s test.

References

Zhang L, Zhu T, Zhang J (2023). “Two-sample Behrens–Fisher problems for high-dimensional data: a normal reference scale-invariant test.” *Journal of Applied Statistics*, **50**(3), 456–476. doi:10.1080/02664763.2020.1834516.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho1 <- 0.1
rho2 <- 0.2
a1 <- 1
a2 <- 2
w1 <- (-2 * sqrt(a1 * (1 - rho1)) + sqrt(4 * a1 * (1 - rho1) + 4 * p * a1 * rho1)) / (2 * p)
x1 <- w1 + sqrt(a1 * (1 - rho1))
Gamma1 <- matrix(rep(w1, p * p), nrow = p)
diag(Gamma1) <- rep(x1, p)
w2 <- (-2 * sqrt(a2 * (1 - rho2)) + sqrt(4 * a2 * (1 - rho2) + 4 * p * a2 * rho2)) / (2 * p)
x2 <- w2 + sqrt(a2 * (1 - rho2))
Gamma2 <- matrix(rep(w2, p * p), nrow = p)
diag(Gamma2) <- rep(x2, p)
Z1 <- matrix(rnorm(n1*p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2*p, mean = 0, sd = 1), p, n2)
y1 <- Gamma1 %*% Z1 + mu1%*(rep(1,n1))
y2 <- Gamma2 %*% Z2 + mu2%*(rep(1,n2))
tsbf_zzz2023(y1,y2,cutoff=1.2)
```

ts_bs1996

*Test proposed by Bai and Saranadasa (1996)***Description**

Bai and Saranadasa (1996)'s test for testing equality of two-sample high-dimensional mean vectors with assuming that two covariance matrices are the same.

Usage

```
ts_bs1996(y1, y2)
```

Arguments

- y1** The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation.
- y2** The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation.

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i. i. d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Bai and Saranadasa (1996) proposed the following centralised L^2 -norm-based test statistic:

$$T_{BS} = \frac{n_1 n_2}{n} \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2 - \text{tr}(\hat{\boldsymbol{\Sigma}}),$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors and $\hat{\boldsymbol{\Sigma}}$ is the pooled sample covariance matrix. They showed that under the null hypothesis, T_{BS} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Bai and Saranadasa (1996)

p.value the p -value of the test proposed by Bai and Saranadasa (1996).

References

Bai Z, Saranadasa H (1996). "Effect of high dimension: by an example of a two sample problem." *Statistica Sinica*, 311–329. <https://www.jstor.org/stable/24306018>.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho <- 0.1
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma %*% Z2 + mu2 %*% (rep(1, n2))
ts_bs1996(y1, y2)
```

ts_sd2008

Test proposed by Srivastava and Du (2008)

Description

Srivastava and Du (2008)'s test for testing equality of two-sample high-dimensional mean vectors with assuming that two covariance matrices are the same.

Usage

```
ts_sd2008(y1, y2)
```

Arguments

- | | |
|----|---|
| y1 | The data matrix (p by n1) from the first population. Each column represents a p -dimensional observation. |
| y2 | The data matrix (p by n2) from the first population. Each column represents a p -dimensional observation. |

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Srivastava and Du (2008) proposed the following test statistic:

$$T_{SD} = \frac{n^{-1}n_1n_2(\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)^\top \mathbf{D}_S^{-1}(\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2) - \frac{(n-2)p}{n-4}}{\sqrt{2 \left[\text{tr}(\mathbf{R}^2) - \frac{p^2}{n-2} \right] c_{p,n}}},$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors, \mathbf{D}_S is the diagonal matrix of sample variance, \mathbf{R} is the sample correlation matrix and $c_{p,n}$ is the adjustment coefficient proposed by Srivastava and Du (2008). They showed that under the null hypothesis, T_{SD} is asymptotically normally distributed.

Value

A (list) object of S3 class `htest` containing the following elements:

statistic the test statistic proposed by Srivastava and Du (2008).

p.value the p -value of the test proposed by Srivastava and Du (2008).

cpn the adjustment coefficient proposed by Srivastava and Du (2008).

References

Srivastava MS, Du M (2008). “A test for the mean vector with fewer observations than the dimension.” *Journal of Multivariate Analysis*, **99**(3), 386–402. doi:[10.1016/j.jmva.2006.11.002](https://doi.org/10.1016/j.jmva.2006.11.002).

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho <- 0.1
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma %*% Z2 + mu2 %*% (rep(1, n2))
ts_sd2008(y1, y2)
```

ts_zgzc2020

Test proposed by Zhang et al. (2020)

Description

Zhang et al. (2020)’s test for testing equality of two-sample high-dimensional mean vectors with assuming that two covariance matrices are the same.

Usage

```
ts_zgzc2020(y1, y2)
```

Arguments

- | | |
|----|--|
| y1 | The data matrix (p by n_1) from the first population. Each column represents a p -dimensional observation. |
| y2 | The data matrix (p by n_2) from the first population. Each column represents a p -dimensional observation. |

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhang et al.(2020) proposed the following test statistic:

$$T_{ZGZC} = \frac{n_1 n_2}{n} \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2,$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors. They showed that under the null hypothesis, T_{ZGZC} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p-value of the test proposed by Zhang et al. (2020).

statistic the test statistic proposed by Zhang et al. (2020).

beta parameter used in Zhang et al. (2020)'s test.

df estimated approximate degrees of freedom of Zhang et al. (2020)'s test.

References

Zhang J, Guo J, Zhou B, Cheng M (2020). "A simple two-sample test in high dimensions based on L 2-norm." *Journal of the American Statistical Association*, **115**(530), 1011–1027. doi:10.1080/01621459.2019.1604366.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(rep(0, p))
mu2 <- mu1
rho <- 0.1
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma %*% Z2 + mu2 %*% (rep(1, n2))
ts_zgzc2020(y1, y2)
```


ts_zz2022

Test proposed by Zhang and Zhu (2022)

Description

Zhang and Zhu (2022)'s test for testing equality of two-sample high-dimensional mean vectors with assuming that two covariance matrices are the same.

Usage

```
ts_zz2022(y1, y2)
```

Arguments

y1	The data matrix (p by n_1) from the first population. Each column represents a p -dimensional observation.
y2	The data matrix (p by n_2) from the first population. Each column represents a p -dimensional observation.

Details

Suppose we have two independent high-dimensional samples:

$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}$, are i.i.d. with $E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i$, $\text{Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}$, $i = 1, 2$.

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhang et al.(2022) proposed the following test statistic:

$$T_{ZZ} = \frac{n_1 n_2}{n} \|\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2\|^2 - \text{tr}(\hat{\boldsymbol{\Sigma}}),$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors and $\hat{\boldsymbol{\Sigma}}$ is the pooled sample covariance matrix. They showed that under the null hypothesis, T_{ZZ} and a chi-squared-type mixture have the same normal or non-normal limiting distribution.

Value

A (list) object of S3 class htest containing the following elements:

p.value the p-value of the test proposed by Zhang and Zhu (2022).

statistic the test statistic proposed by Zhang and Zhu (2022).

beta0 parameter used in Zhang and Zhu (2022)'s test

beta1 parameter used in Zhang and Zhu (2022)'s test

df estimated approximate degrees of freedom of Zhang and Zhu (2022)'s test.

References

Zhang J, Zhu T (2022). "A revisit to Bai–Saranadasa's two-sample test." *Journal of Nonparametric Statistics*, **34**(1), 58–76. doi:10.1080/10485252.2021.2015768.

Examples

```

set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho <- 0.1
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma %*% Z2 + mu2 %*% (rep(1, n2))
ts_zzz2020(y1, y2)

```

ts_zzz2020

Test proposed by Zhang et al. (2020)

Description

Zhang et al. (2020)'s test for testing equality of two-sample high-dimensional mean vectors with assuming that two covariance matrices are the same.

Usage

```
ts_zzz2020(y1, y2)
```

Arguments

y1	The data matrix (p by n_1) from the first population. Each column represents a p -dimensional observation.
y2	The data matrix (p by n_2) from the first population. Each column represents a p -dimensional observation.

Details

Suppose we have two independent high-dimensional samples:

$$\mathbf{y}_{i1}, \dots, \mathbf{y}_{in_i}, \text{ are i.i.d. with } E(\mathbf{y}_{i1}) = \boldsymbol{\mu}_i, \text{ Cov}(\mathbf{y}_{i1}) = \boldsymbol{\Sigma}, i = 1, 2.$$

The primary object is to test

$$H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2 \text{ versus } H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2.$$

Zhang et al.(2020) proposed the following test statistic:

$$T_{ZZZ} = \frac{n_1 n_2}{np} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2)^\top \hat{\mathbf{D}}^{-1} (\bar{\mathbf{y}}_1 - \bar{\mathbf{y}}_2),$$

where $\bar{\mathbf{y}}_i, i = 1, 2$ are the sample mean vectors, $\hat{\mathbf{D}}$ is the diagonal matrix of sample covariance matrix. They showed that under the null hypothesis, T_{ZZZ} and a chi-squared-type mixture have the same limiting distribution.

Value

A (list) object of S3 class `htest` containing the following elements:

p.value the p-value of the test proposed by Zhang et al. (2020).

statistic the test statistic proposed by Zhang et al. (2020).

df estimated approximate degrees of freedom of Zhang et al. (2020)’s test.

References

Zhang L, Zhu T, Zhang J (2020). “A simple scale-invariant two-sample test for high-dimensional data.” *Econometrics and Statistics*, **14**, 131–144. doi:10.1016/j.ecosta.2019.12.002.

Examples

```
set.seed(1234)
n1 <- 20
n2 <- 30
p <- 50
mu1 <- t(t(rep(0, p)))
mu2 <- mu1
rho <- 0.1
y <- (-2 * sqrt(1 - rho) + sqrt(4 * (1 - rho) + 4 * p * rho)) / (2 * p)
x <- y + sqrt((1 - rho))
Gamma <- matrix(rep(y, p * p), nrow = p)
diag(Gamma) <- rep(x, p)
Z1 <- matrix(rnorm(n1 * p, mean = 0, sd = 1), p, n1)
Z2 <- matrix(rnorm(n2 * p, mean = 0, sd = 1), p, n2)
y1 <- Gamma %*% Z1 + mu1 %*% (rep(1, n1))
y2 <- Gamma %*% Z2 + mu2 %*% (rep(1, n2))
ts_zzz2020(y1, y2)
```

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