

# Day 12: Inference for a single proportion or difference of two (independent) proportions (Sections 8.1-8.2)

BSTA 511/611

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# MoRitz's tip of the day: code folding

- With code folding we can hide or show the code in the html output by clicking on the **Code** buttons in the html file.
- Note the `</> Code` button on the top right of the html output.

▼ Code

```
2*pnorm(-0.3607455)
```

[1] 0.7182897

► Code

```
[1] 0.7182897
```

```
1 ---  
2 title: "Day 12: Inference for a single proportion or difference of  
3 two (independent) proportions (Sections 8.1-8.2)"  
4 subtitle: "BSTA 511/611"  
5 author: "Meike Niederhausen, PhD"  
6 institute: "OHSU-PSU School of Public Health"  
7 date: "11/8/2023"  
8 categories: ["Week 7"]  
9 format:  
10   html:  
11     link-external-newwindow: true  
12     toc: true  
13     code-fold: show  
14     code-tools: true  
15     source: repo  
16   execute:  
17     echo: true  
18     freeze: auto # re-render only when source changes  
19   # editor: visual  
20   editor_options:  
21     chunk_output_type: console
```

Creates button at top right of html output that lets the user select: Hide All Code, Show All Code, or View Source

Can specify location of source file for View Source user option. Since this file is stored in a GitHub repository, I specified repo.

**Day 12: Inference for a single proportion or difference of two (independent) proportions (Sections 8.1-8.2)**

BSTA 511/611

WEEK 7

Code ▾

Show All Code

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View Source

See more information at <https://quarto.org/docs/output-formats/html-code.html#folding-code>

# Where are we?

CI's and hypothesis tests for different scenarios:

$$\text{point estimate} \pm z^*(\text{or } t^*) \cdot SE, \text{ test stat} = \frac{\text{point estimate} - \text{null value}}{SE}$$

| Day | Book | Population parameter    | Symbol              | Point estimate             | Symbol                  | SE   |
|-----|------|-------------------------|---------------------|----------------------------|-------------------------|--|
| 10  | 5.1  | Pop mean                | $\mu$               | Sample mean                | $\bar{x}$               | $\frac{s}{\sqrt{n}}$   |
| 10  | 5.2  | Pop mean of paired diff | $\mu_d$ or $\delta$ | Sample mean of paired diff | $\bar{x}_d$             | $\frac{s_d}{\sqrt{n}}$   |
| 11  | 5.3  | Diff in pop means       | $\mu_1 - \mu_2$     | Diff in sample means       | $\bar{x}_1 - \bar{x}_2$ | $\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$<br><b>or pooled</b> |
| 12  | 8.1  | Pop proportion          | $p$                 | Sample prop                | $\hat{p}$               | ???  |
| 12  | 8.2  | Diff in pop proportions | $p_1 - p_2$         | Diff in sample proportions | $\hat{p}_1 - \hat{p}_2$ | ???  |

# Goals for today (Sections 8.1-8.2)

- Statistical inference for a single proportion or the difference of two (independent) proportions
  1. Sampling distribution for a proportion or difference in proportions
  2. What are  $H_0$  and  $H_a$ ?
  3. What are the SE's for  $\hat{p}$  and  $\hat{p}_1 - \hat{p}_2$ ?
  4. Hypothesis test
  5. Confidence Interval
  6. How are the SE's different for a hypothesis test & CI?
  7. How to run proportions tests in R
  8. Power & sample size for proportions tests (extra material)

# Motivating example

## One proportion

- A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year.
  - What is the CI for the proportion?
  - The study also reported that 36% of noncollege young males had participated in sports betting. Is the proportion for male college students different from 0.36?

## Two proportions

- There were 214 men in the sample of noncollege young males (36% participated in sports betting in the previous year).
- Compare the difference in proportions between the college and noncollege young males.
  - CI & Hypothesis test

Barnes GM, Welte JW, Hoffman JH, Tidwell MC. Comparisons of gambling and alcohol use among college students and noncollege young people in the United States. J Am Coll Health. 2010 Mar-Apr;58(5):443-52. doi: 10.1080/07448480903540499. PMID: 20304756; PMCID: PMC4104810.

# Steps in a Hypothesis Test

1. Set the **level of significance**  $\alpha$
2. Specify the **null** ( $H_0$ ) and **alternative** ( $H_A$ ) **hypotheses**
  1. In symbols
  2. In words
  3. Alternative: one- or two-sided?
3. Calculate the **test statistic**.
4. Calculate the **p-value** based on the observed test statistic and its sampling distribution
5. Write a **conclusion** to the hypothesis test
  1. Do we reject or fail to reject  $H_0$ ?
  2. Write a conclusion in the context of the problem

# Step 2: Null & Alternative Hypotheses

Null and alternative hypotheses in **words** and in **symbols**.

## One sample test

- $H_0$ : The population proportion of young male college students that participated in sports betting in the previous year is 0.36.
- $H_A$ : The population proportion of young male college students that participated in sports betting in the previous year is not 0.36.

$$H_0 : p = 0.36$$

$$H_A : p \neq 0.36$$

## Two samples test

- $H_0$ : The difference in population proportions of young male college and noncollege students that participated in sports betting in the previous year is 0.
- $H_A$ : The difference in population proportions of young male college and noncollege students that participated in sports betting in the previous year is not 0.

$$H_0 : p_{coll} - p_{noncoll} = 0$$

$$H_A : p_{coll} - p_{noncoll} \neq 0$$

# Sampling distribution of $\hat{p}$

- $\hat{p} = \frac{X}{n}$  where  $X$  is the number of “successes” and  $n$  is the sample size.
- $X \sim Bin(n, p)$ , where  $p$  is the population proportion.
- For  $n$  “big enough”, the normal distribution can be used to approximate a binomial distribution:

$$Bin(n, p) \rightarrow N\left(\mu = np, \sigma = \sqrt{np(1 - p)}\right)$$

- Since  $\hat{p} = \frac{X}{n}$  is a linear transformation of  $X$ , we have for large n:

$$\hat{p} \sim N\left(\mu_{\hat{p}} = p, \sigma_{\hat{p}} = \sqrt{\frac{p(1 - p)}{n}}\right)$$

- *How we apply this result to CI's and test statistics is different!!!*

## Step 3: Test statistic

Sampling distribution of  $\hat{p}$  if we assume  $H_0 : p = p_0$  is true:

$$\hat{p} \sim N\left(\mu_{\hat{p}} = p, \sigma_{\hat{p}} = \sqrt{\frac{p(1-p)}{n}}\right) \sim N\left(\mu_{\hat{p}} = p_0, \sigma_{\hat{p}} = \sqrt{\frac{p_0 \cdot (1-p_0)}{n}}\right)$$

Test statistic for a one sample proportion test:

$$\text{test stat} = \frac{\text{point estimate} - \text{null value}}{SE} = z_{\hat{p}} = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0 \cdot (1-p_0)}{n}}}$$

---

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year.

$$z_{\hat{p}} = \frac{94/269 - 0.36}{\sqrt{\frac{0.36 \cdot (1-0.36)}{269}}} - 0.3607455$$

What is the test statistic when testing  $H_0 : p = 0.36$  vs.  $H_A : p \neq 0.36$ ?

## Step “3b”: Conditions satisfied?

### Conditions:

1. Independent observations?

- The observations were collected independently.

2. The number of **expected successes and expected failures is at least 10**.

- $n_1 p_0 \geq 10, n_1(1 - p_0) \geq 10$
- 

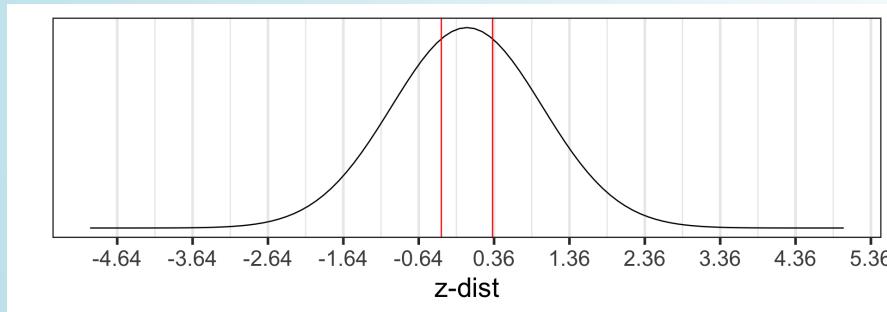
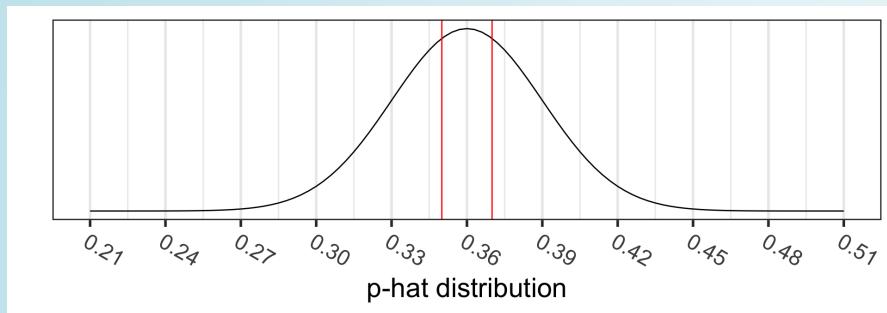
**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year.

Testing  $H_0 : p = 0.36$  vs.  $H_A : p \neq 0.36$ .

Are the conditions satisfied?

## Step 4: p-value

The **p-value** is the **probability** of obtaining a test statistic *just as extreme or more extreme* than the observed test statistic assuming the null hypothesis  $H_0$  is true.



Calculate the *p*-value:

$$2 \cdot P(\hat{p} < 0.35)$$

$$= 2 \cdot P\left(Z_{\hat{p}} < \frac{94/269 - 0.36}{\sqrt{\frac{0.36 \cdot (1-0.36)}{269}}}\right)$$

$$= 2 \cdot P(Z_{\hat{p}} < -0.3607455)$$

$$= 0.7182897$$

```
1  2*pnorm(-0.3607455)
```

```
[1] 0.7182897
```

# Step 5: Conclusion to hypothesis test

$$H_0 : p = 0.36$$

$$H_A : p \neq 0.36$$

- Recall the  $p$ -value = 0.7182897
- Use  $\alpha = 0.05$ .
- Do we reject or fail to reject  $H_0$ ?

## Conclusion statement:

- Stats class conclusion
  - There is insufficient evidence that the (population) proportion of young male college students that participated in sports betting in the previous year is different than 0.36 (  $p$ -value = 0.72).
- More realistic manuscript conclusion:
  - In a sample of 269 male college students, 35% had participated in sports betting in the previous year, which is not different from 36% (  $p$ -value = 0.72).

# 95% CI for population proportion

What to use for SE in CI formula?

$$\hat{p} \pm z^* \cdot SE_{\hat{p}}$$

Sampling distribution of  $\hat{p}$ :

$$\hat{p} \sim N\left(\mu_{\hat{p}} = p, \sigma_{\hat{p}} = \sqrt{\frac{p(1-p)}{n}}\right)$$

Problem: We don't know what  $p$  is - it's what we're estimating with the CI.

Solution: approximate  $p$  with  $\hat{p}$ :

$$SE_{\hat{p}} = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

---

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year. Find the 95% CI for the population proportion.

$$94/269 \pm 1.96 \cdot SE_{\hat{p}}$$

$$SE_{\hat{p}} = \sqrt{\frac{(94/269)(1 - 94/269)}{269}}$$
$$(0.293, 0.407)$$

## Interpretation:

We are 95% confident that the (population) proportion of young male college students that participated in sports betting in the previous year is in (0.29, 0.41).

# Conditions for one proportion: **test vs. CI**

## Hypothesis test conditions

### 1. *Independent observations*

- The observations were collected independently.

### 2. The number of **expected** successes and **expected** failures is at least 10.

$$n_1 p_0 \geq 10, \quad n_1(1 - p_0) \geq 10$$

## Confidence interval conditions

### 1. *Independent observations*

- The observations were collected independently.

### 2. The number of successes and failures is at least 10:

$$n_1 \hat{p}_1 \geq 10, \quad n_1(1 - \hat{p}_1) \geq 10$$

Inference for difference of two  
independent proportions

$$\hat{p}_1 - \hat{p}_2$$

# Sampling distribution of $\hat{p}_1 - \hat{p}_2$

- $\hat{p}_1 = \frac{X_1}{n_1}$  and  $\hat{p}_2 = \frac{X_2}{n_2}$ ,
  - $X_1$  &  $X_2$  are the number of “successes”
  - $n_1$  &  $n_2$  are the sample sizes of the 1st & 2nd samples
- Each  $\hat{p}$  can be approximated by a normal distribution, for “big enough”  $n$
- Since the difference of independent normal random variables is also normal, it follows that for “big enough”  $n_1$  and  $n_2$

$$\hat{p}_1 - \hat{p}_2 \sim N\left(\mu_{\hat{p}_1 - \hat{p}_2} = p_1 - p_2, \sigma_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{p_1 \cdot (1 - p_1)}{n_1} + \frac{p_2 \cdot (1 - p_2)}{n_2}}\right)$$

where  $p_1$  &  $p_2$  are the population proportions, respectively.

- *How we apply this result to CI's and test statistics is different!!!*

## Step 3: Test statistic (1/2)

Sampling distribution of  $\hat{p}_1 - \hat{p}_2$ :

$$\hat{p}_1 - \hat{p}_2 \sim N\left(\mu_{\hat{p}_1 - \hat{p}_2} = p_1 - p_2, \sigma_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{p_1 \cdot (1 - p_1)}{n_1} + \frac{p_2 \cdot (1 - p_2)}{n_2}}\right)$$

Since we assume  $H_0 : p_1 - p_2 = 0$  is true, we “pool” the proportions of the two samples to calculate the SE:

$$\text{pooled proportion} = \hat{p}_{pool} = \frac{\text{total number of successes}}{\text{total number of cases}} = \frac{x_1 + x_2}{n_1 + n_2}$$

Test statistic:

$$\text{test statistic} = z_{\hat{p}_1 - \hat{p}_2} = \frac{\hat{p}_1 - \hat{p}_2 - 0}{\sqrt{\frac{\hat{p}_{pool} \cdot (1 - \hat{p}_{pool})}{n_1} + \frac{\hat{p}_{pool} \cdot (1 - \hat{p}_{pool})}{n_2}}}$$

## Step 3: Test statistic (2/2)

$$\text{test statistic} = z_{\hat{p}_1 - \hat{p}_2} = \frac{\hat{p}_1 - \hat{p}_2 - 0}{\sqrt{\frac{\hat{p}_{pool} \cdot (1 - \hat{p}_{pool})}{n_1} + \frac{\hat{p}_{pool} \cdot (1 - \hat{p}_{pool})}{n_2}}}$$

$$\text{pooled proportion} = \hat{p}_{pool} = \frac{\text{total number of successes}}{\text{total number of cases}} = \frac{x_1 + x_2}{n_1 + n_2}$$

---

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year, and out of 214 noncollege young males 36% had.

What is the test statistic when testing  $H_0 : p_{coll} - p_{noncoll} = 0$  vs.  
 $H_A : p_{coll} - p_{noncoll} \neq 0$ ?

$$z_{\hat{p}_1 - \hat{p}_2} = \frac{94/269 - 77/214 - 0}{\sqrt{0.354 \cdot (1 - 0.354) \left( \frac{1}{269} + \frac{1}{214} \right)}} \\ = -0.2367497$$

# Step “3b”: Conditions satisfied?

## Conditions:

- *Independent observations & samples*
  - The observations were collected independently.
  - In particular, observations from the two groups weren't paired in any meaningful way.
- The number of expected successes and expected failures is at least 10 *for each group*
  - using the pooled proportion:
    - $n_1\hat{p}_{pool} \geq 10, n_1(1 - \hat{p}_{pool}) \geq 10$
    - $n_2\hat{p}_{pool} \geq 10, n_2(1 - \hat{p}_{pool}) \geq 10$

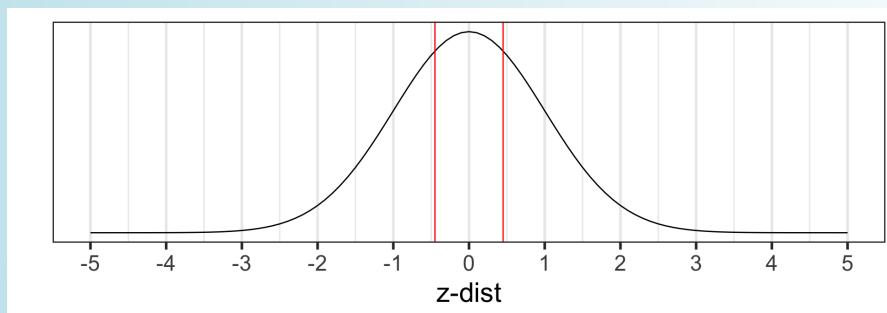
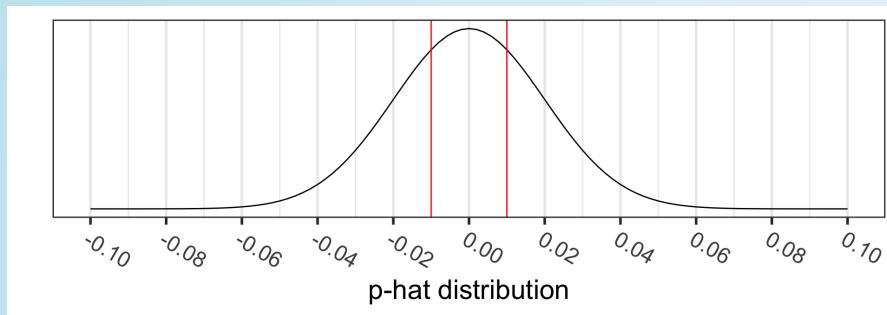
**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year, and out of 214 noncollege young males 36% had.

Testing  $H_0 : p_{coll} - p_{noncoll} = 0$  vs.  $H_A : p_{coll} - p_{noncoll} \neq 0$ .

Are the conditions satisfied?

## Step 4: p-value

The **p-value** is the **probability** of obtaining a test statistic *just as extreme or more extreme* than the observed test statistic assuming the null hypothesis  $H_0$  is true.



Calculate the *p*-value:

$$\begin{aligned} & 2 \cdot P(\hat{p}_1 - \hat{p}_2 < 0.35 - 0.36) \\ &= 2 \cdot P\left(Z_{\hat{p}_1 - \hat{p}_2} < \frac{94/269 - 77/214 - 0}{\sqrt{0.354 \cdot (1 - 0.354)(\frac{1}{269} + \frac{1}{214})}}\right) \\ &= 2 \cdot P(Z_{\hat{p}} < -0.2367497) \\ &= 0.812851 \end{aligned}$$

```
1 2*pnorm(-0.2367497)
```

```
[1] 0.812851
```

# Step 5: Conclusion to hypothesis test

$$H_0 : p_{coll} - p_{noncoll} = 0$$
$$H_A : p_{coll} - p_{noncoll} \neq 0$$

- Recall the  $p$ -value = 0.812851
- Use  $\alpha = 0.05$ .
- Do we reject or fail to reject  $H_0$ ?

## Conclusion statement:

- Stats class conclusion
  - There is insufficient evidence that the difference in (population) proportions of young male college and noncollege students that participated in sports betting in the previous year are different ( $p$ -value = 0.81).
- More realistic manuscript conclusion:
  - 35% of young male college students ( $n=269$ ) and 36% of noncollege young males ( $n=214$ ) participated in sports betting in the previous year ( $p$ -value = 0.81).

# 95% CI for population difference in proportions

What to use for SE in CI formula?

$$\hat{p}_1 - \hat{p}_2 \pm z^* \cdot SE_{\hat{p}_1 - \hat{p}_2}$$

SE in sampling distribution of  $\hat{p}_1 - \hat{p}_2$

$$\sigma_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{p_1 \cdot (1 - p_1)}{n_1} + \frac{p_2 \cdot (1 - p_2)}{n_2}}$$

Problem: We don't know what  $p$  is - it's what we're estimating with the CI.

Solution: approximate  $p_1, p_2$  with  $\hat{p}_1, \hat{p}_2$ :

$$SE_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{\hat{p}_1 \cdot (1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2 \cdot (1 - \hat{p}_2)}{n_2}}$$

---

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year, and out of 214 noncollege young males 36% had. Find the 95% CI for the difference in population proportions.

$$\frac{94}{269} - \frac{77}{214} \pm 1.96 \cdot SE_{\hat{p}_1 - \hat{p}_2}$$

$$SE_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{94/269 \cdot (1 - 94/269)}{269} + \frac{77/214 \cdot (1 - 77/214)}{214}}$$

**Interpretation:**

We are 95% confident that the difference in (population) proportions of young male college and noncollege students that participated in sports betting in the previous year is in (-0.127, 0.106).

# Conditions for difference in proportions: test vs. CI

## Hypothesis test conditions

### 1. Independent observations & samples

- The observations were collected independently.
- In particular, observations from the two groups weren't paired in any meaningful way.

### 2. The number of **expected** successes and **expected** failures is at least 10 *for each group* - using the pooled proportion:

- $n_1\hat{p}_{pool} \geq 10, n_1(1 - \hat{p}_{pool}) \geq 10$
- $n_2\hat{p}_{pool} \geq 10, n_2(1 - \hat{p}_{pool}) \geq 10$

## Confidence interval conditions

### 1. Independent observations & samples

- The observations were collected independently.
- In particular, observations from the two groups weren't paired in any meaningful way.

### 2. The number of successes and failures is at least 10 *for each group*.

- $n_1\hat{p}_1 \geq 10, n_1(1 - \hat{p}_1) \geq 10$
- $n_2\hat{p}_2 \geq 10, n_2(1 - \hat{p}_2) \geq 10$

# R: 1- and 2-sample proportions tests

```
prop.test(x, n, p = NULL,  
          alternative = c("two.sided", "less", "greater"),  
          conf.level = 0.95,  
          correct = TRUE)
```

- 2 options for data input
  1. Summary counts
    - `x` = vector with counts of “successes”
    - `n` = vector with sample size in each group
  2. Dataset
    - `x = table()` of dataset
    - Need to create a dataset based on the summary stats if do not already have one
- Continuity correction

# R: 1-sample proportion test

“1-prop z-test”

# Summary stats input for 1-sample proportion test

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year.

Test  $H_0 : p = 0.36$  vs.  $H_A : p \neq 0.36$ ?

```
1 .35*269 # number of "successes"; round this value
[1] 94.15
1 prop.test(x = 94, n = 269,
             p = 0.36,
             alternative = "two.sided",
             correct = FALSE) # x = # successes & n = sample size
# null value p0
# 2-sided alternative
# no continuity correction

1-sample proportions test without continuity correction

data: 94 out of 269, null probability 0.36
X-squared = 0.13014, df = 1, p-value = 0.7183
alternative hypothesis: true p is not equal to 0.36
95 percent confidence interval:
 0.2949476 0.4081767
sample estimates:
      p 
0.3494424
```

Can `tidy()` test output:

```
1 prop.test(x = 94, n = 269, p = 0.36, alternative = "two.sided", correct = FALSE) %>%
2   tidy() %>% gt()
```

| estimate  | statistic | p.value   | parameter | conf.low  | conf.high | method  | alternative |
|-----------|-----------|-----------|-----------|-----------|-----------|---|-------------|
| 0.3494424 | 0.1301373 | 0.7182897 | 1         | 0.2949476 | 0.4081767 | 1-sample proportions test without continuity correction | two.sided   |

# Dataset input for 1-sample proportion test (1/2)

Since we don't have a dataset, we first need to create a dataset based on the results:

"out of 269 male college students, 35% had participated in sports betting in the previous year"

```
1 .35*269 # number of "successes"; round  
[1] 94.15
```

```
1 SportsBet1 <- tibble(  
2   Coll = c(rep("Bet", 94),  
3             rep("NotBet", 269-94))  
4 )
```

```
1 glimpse(SportsBet1)
```

Rows: 269  
Columns: 1  
\$ Coll <chr> "Bet", "B...

```
1 SportsBet1 %>% tabyl(Coll)
```

| Coll   | n   | percent   |
|--------|-----|-----------|
| Bet    | 94  | 0.3494424 |
| NotBet | 175 | 0.6505576 |

R code for proportions test requires input as a base R **table**:

```
1 table(SportsBet1$Coll)
```

| Bet | NotBet |
|-----|--------|
| 94  | 175    |

# Dataset input for 1-sample proportion test (2/2)

- When using a dataset, `prop.test` requires the input `x` to be a table
- Note that we do not also specify `n` since the table already includes all needed information.

```
1 prop.test(x = table(SportsBet1$Coll),      # table() of data
2           p = 0.36,                      # null value p0
3           alternative = "two.sided",    # 2-sided alternative
4           correct = FALSE)            # no continuity correction
```

```
1-sample proportions test without continuity correction

data: table(SportsBet1$Coll), null probability 0.36
X-squared = 0.13014, df = 1, p-value = 0.7183
alternative hypothesis: true p is not equal to 0.36
95 percent confidence interval:
 0.2949476 0.4081767
sample estimates:
      p
0.3494424
```

Compare output with summary stats method:

```
1 prop.test(x = 94, n = 269,          # x = # successes & n = sample size
2           p = 0.36,                  # null value p0
3           alternative = "two.sided", # 2-sided alternative
4           correct = FALSE)       # no continuity correction
5 tidy() %>% gt()
```

| estimate  | statistic | p.value   | parameter | conf.low  | conf.high | method  | alternative |
|-----------|-----------|-----------|-----------|-----------|-----------|---|-------------|
| 0.3494424 | 0.1301373 | 0.7182897 | 1         | 0.2949476 | 0.4081767 | 1-sample proportions test without continuity correction | two.sided   |

# Continuity correction: 1-prop z-test with vs. without CC

- Recall that when we approximated the binomial distribution with a normal distribution to calculate a probability,
- that we included a continuity correction (CC)
- to account for approximating a discrete distribution with a continuous distribution.

```
1 prop.test(x = 94, n = 269, p = 0.36, alternative = "two.sided",
2           correct = FALSE) %>%
3   tidy() %>% gt()
```

| estimate  | statistic | p.value   | parameter | conf.low  | conf.high | method  | alternative |
|-----------|-----------|-----------|-----------|-----------|-----------|---|-------------|
| 0.3494424 | 0.1301373 | 0.7182897 | 1         | 0.2949476 | 0.4081767 | 1-sample proportions test without continuity correction | two.sided   |

```
1 prop.test(x = 94, n = 269, p = 0.36, alternative = "two.sided",
2           correct = TRUE) %>%
3   tidy() %>% gt()
```

| estimate  | statistic  | p.value   | parameter | conf.low  | conf.high | method   | alternative |
|-----------|------------|-----------|-----------|-----------|-----------|--|-------------|
| 0.3494424 | 0.08834805 | 0.7662879 | 1         | 0.2931841 | 0.4100774 | 1-sample proportions test with continuity correction | two.sided   |

Differences are small when sample sizes are large.

# R: 2-samples proportion test

“2-prop z-test”

# Summary stats input for 2-samples proportion test

**Example:** A 2010 study found that out of 269 male college students, 35% had participated in sports betting in the previous year, and out of 214 noncollege young males 36% had. Test  $H_0 : p_{coll} - p_{noncoll} = 0$  vs.  $H_A : p_{coll} - p_{noncoll} \neq 0$ .

```
1 # round the number of successes:  
2 .35*269 # number of "successes" in college students  
  
[1] 94.15  
  
1 .36*214 # number of "successes" in noncollege students  
  
[1] 77.04  
  
1 NmbrBet <- c(94, 77)                      # vector for # of successes in each group  
2 TotalNmbr <- c(269, 214)                    # vector for sample size in each group  
3  
4 prop.test(x = NmbrBet,                         # x is # of successes in each group  
5       n = TotalNmbr,                          # n is sample size in each group  
6       alternative = "two.sided",            # 2-sided alternative  
7       correct = FALSE)                      # no continuity correction
```

```
2-sample test for equality of proportions without continuity correction  
  
data: NmbrBet out of TotalNmbr  
X-squared = 0.05605, df = 1, p-value = 0.8129  
alternative hypothesis: two.sided  
95 percent confidence interval:  
-0.09628540  0.07554399  
sample estimates:  
prop 1    prop 2  
0.3494424 0.3598131
```

# Dataset input for 2-samples proportion test (1/2)

Since we don't have a dataset, we first need to create a dataset based on the results:

"out of 269 male college students, 35% had participated in sports betting in the previous year, and out of 214 noncollege young males 36% had"

```
1 # round the number of successes:  
2 .35*269 # college students
```

```
[1] 94.15
```

```
1 .36*214 # noncollege students
```

```
[1] 77.04
```

```
1 SportsBet2 <- tibble(  
2   Group = c(rep("College", 269),  
3               rep("NonCollege", 214)),  
4   Bet = c(rep("yes", 94),  
5           rep("no", 269-94),  
6           rep("yes", 77),  
7           rep("no", 214-77))  
8 )
```

```
1 glimpse(SportsBet2)
```

Rows: 483  
Columns: 2  
\$ Group <chr> "College", "College", "College",  
"College", "College", "College"...  
\$ Bet <chr> "yes", "yes", "yes", "yes", "yes",  
"yes", "yes", "yes", "yes", "...

```
1 SportsBet2 %>% tabyl(Group, Bet)
```

|   | Group      | no  | yes |
|---|------------|-----|-----|
| 1 | College    | 175 | 94  |
| 2 | NonCollege | 137 | 77  |

R code for proportions test requires input as a base R **table**:

```
1 table(SportsBet2$Group,  
2        SportsBet2$Bet)
```

|   |            | no  | yes |
|---|------------|-----|-----|
| 1 | College    | 175 | 94  |
| 2 | NonCollege | 137 | 77  |

# Dataset input for 2-samples proportion test (2/2)

- When using a dataset, `prop.test` requires the input `x` to be a table
- Note that we do not also specify `n` since the table already includes all needed information.

```
1 prop.test(x = table(SportsBet2$Group, SportsBet2$Bet),  
2           alternative = "two.sided",  
3           correct = FALSE)
```

```
2-sample test for equality of proportions without continuity correction  
  
data: table(SportsBet2$Group, SportsBet2$Bet)  
X-squared = 0.05605, df = 1, p-value = 0.8129  
alternative hypothesis: two.sided  
95 percent confidence interval:  
 -0.07554399  0.09628540  
sample estimates:  
 prop 1   prop 2  
0.6505576 0.6401869
```

Compare output with summary stats method:

```
1 prop.test(x = NmbrBet,  
2             n = TotalNmbr,  
3             alternative = "two.sided",  
4             correct = FALSE) %>%  
5 tidy() %>% gt()
```

| estimate1 | estimate2 | statistic  | p.value   | parameter | conf.low   | conf.high  | method  | alternative |
|-----------|-----------|------------|-----------|-----------|------------|------------|---|-------------|
| 0.3494424 | 0.3598131 | 0.05605044 | 0.8128509 | 1         | -0.0962854 | 0.07554399 | 2-sample test for equality of proportions without continuity correction | two.sided   |

# Continuity correction: 2-prop z-test with vs. without CC

- Recall that when we approximated the binomial distribution with a normal distribution to calculate a probability,
- that we included a continuity correction (CC)
- to account for approximating a discrete distribution with a continuous distribution.

```
1 prop.test(x = NmbrBet, n = TotalNmbr, alternative = "two.sided",
2           correct = FALSE) %>% tidy() %>% gt()
```

| estimate1 | estimate2 | statistic  | p.value   | parameter | conf.low   | conf.high  | method  | alternative |
|-----------|-----------|------------|-----------|-----------|------------|------------|---|-------------|
| 0.3494424 | 0.3598131 | 0.05605044 | 0.8128509 | 1         | -0.0962854 | 0.07554399 | 2-sample test for equality of proportions without continuity correction | two.sided   |

```
1 prop.test(x = NmbrBet, n = TotalNmbr, alternative = "two.sided",
2           correct = TRUE) %>% tidy() %>% gt()
```

| estimate1 | estimate2 | statistic  | p.value   | parameter | conf.low   | conf.high  | method   | alternative |
|-----------|-----------|------------|-----------|-----------|------------|------------|--|-------------|
| 0.3494424 | 0.3598131 | 0.01987511 | 0.8878864 | 1         | -0.1004806 | 0.07973918 | 2-sample test for equality of proportions with continuity correction | two.sided   |

Differences are small when sample sizes are large.

# Power & sample size for testing proportions

# Sample size calculation for testing one proportion

- Recall in our sports betting example that the null  $p_0 = 0.36$  and the observed proportion was  $\hat{p} = 0.35$ .
  - The  $p$ -value from the hypothesis test was not significant.
  - How big would the sample size  $n$  need to be in order for the  $p$ -value to be significant?
- **Calculate  $n$** 
  - given  $\alpha$ , power ( $1 - \beta$ ), “true” alternative proportion  $p$ , and null  $p_0$ :

$$n = p(1-p) \left( \frac{z_{1-\alpha/2} + z_{1-\beta}}{p - p_0} \right)^2$$

```
1 p <- 0.35
2 p0 <- 0.36
3 alpha <- 0.05
4 beta <- 0.20 #power=1-beta; want >=80% power
5 n <- p*(1-p)*((qnorm(1-alpha/2) + qnorm(1-beta)) /
6                         (p-p0))^2
7 n
[1] 17856.2
1 ceiling(n)
[1] 17857
```

We would need a sample size of at least 17,857!

# Power calculation for testing one proportion

Conversely, we can calculate how much power we had in our example given the sample size of 269.

- **Calculate power,**
  - given  $\alpha$ ,  $n$ , “true” alternative proportion  $p$ , and null  $p_0$

$$1 - \beta = \Phi(z - z_{1-\alpha/2}) + \Phi(-z - z_{1-\alpha/2}) \quad , \quad \text{where } z = \frac{p - p_0}{\sqrt{\frac{p(1-p)}{n}}}$$

$\Phi$  is the probability for a standard normal distribution

```
1 p <- 0.35; p0 <- 0.36; alpha <- 0.05; n <- 269
2 (z <- (p-p0)/sqrt(p*(1-p)/n))
[1] -0.343863
1 (Power <- pnorm(z - qnorm(1-alpha/2)) + pnorm(-z - qnorm(1-alpha/2)))
[1] 0.06365242
```

If the population proportion is 0.35 instead of 0.36, we only have a 6.4% chance of correctly rejecting  $H_0$  when the sample size is 269.

# R package **pwr** for power analyses

- Specify all parameters *except for* the one being solved for.
- One proportion

```
pwr.p.test(h = NULL, n = NULL, sig.level = 0.05, power = NULL, alternative = c("two.sided","less","greater"))
```

- Two proportions (same sample sizes)

```
pwr.2p.test(h = NULL, n = NULL, sig.level = 0.05, power = NULL, alternative = c("two.sided","less","greater"))
```

- Two proportions (different sample sizes)

```
pwr.2p2n.test(h = NULL, n1 = NULL, n2 = NULL, sig.level = 0.05, power = NULL, alternative = c("two.sided", "less","greater"))
```

---

*h* is the effect size, and calculated using an arcsine transformation:

$$h = \text{ES.h}(p_1, p_2) = 2 \arcsin(\sqrt{p_1}) - 2 \arcsin(\sqrt{p_2})$$

See PASS documentation for

- testing 1 proportion using effect size vs. other ways of powering a test of 1 proportion
- testing 2 proportions using effect size vs. other ways of powering a test of 2 proportions.

# pwr: sample size for one proportion test

```
pwr.p.test(h = NULL, n = NULL, sig.level = 0.05, power = NULL, alternative = c("two.sided","less","greater"))
```

- $h$  is the effect size:  $h = \text{ES.h}(p_1, p_2)$ 
  - $p_1$  and  $p_2$  are the two proportions being tested
  - one of them is the null proportion  $p_0$ , and the other is the alternative proportion

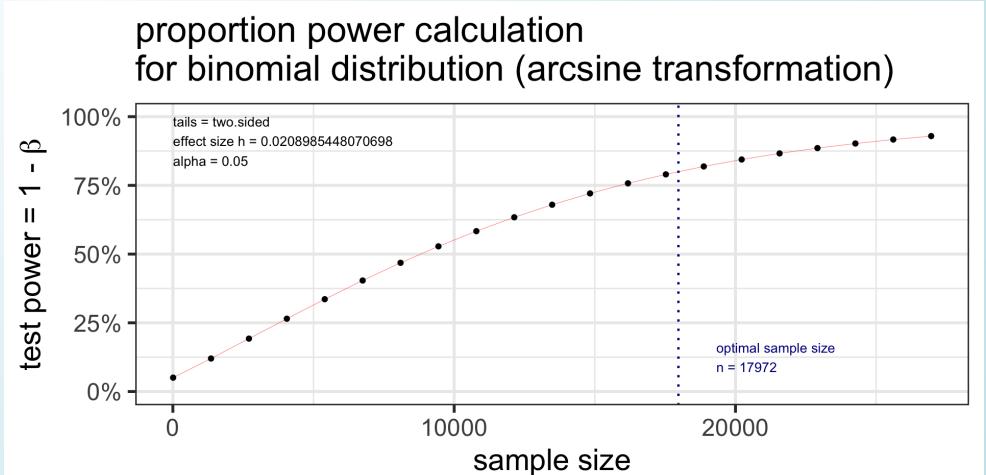
Specify all parameters *except for* the sample size:

```
1 library(pwr)
2
3 p.n <- pwr.p.test(
4   h = ES.h(p1 = 0.36, p2 = 0.35),
5   sig.level = 0.05,
6   power = 0.80,
7   alternative = "two.sided")
8 p.n
```

```
proportion power calculation for binomial distribution (arcsine transformation)

h = 0.02089854
n = 17971.09
sig.level = 0.05
power = 0.8
alternative = two.sided
```

```
1 plot(p.n)
```



# pwr: power for one proportion test

```
pwr.p.test(h = NULL, n = NULL, sig.level = 0.05, power = NULL, alternative = c("two.sided","less","greater"))
```

- $h$  is the effect size:  $h = \text{ES.h}(p_1, p_2)$ 
  - $p_1$  and  $p_2$  are the two proportions being tested
  - one of them is the null proportion  $p_0$ , and the other is the alternative proportion

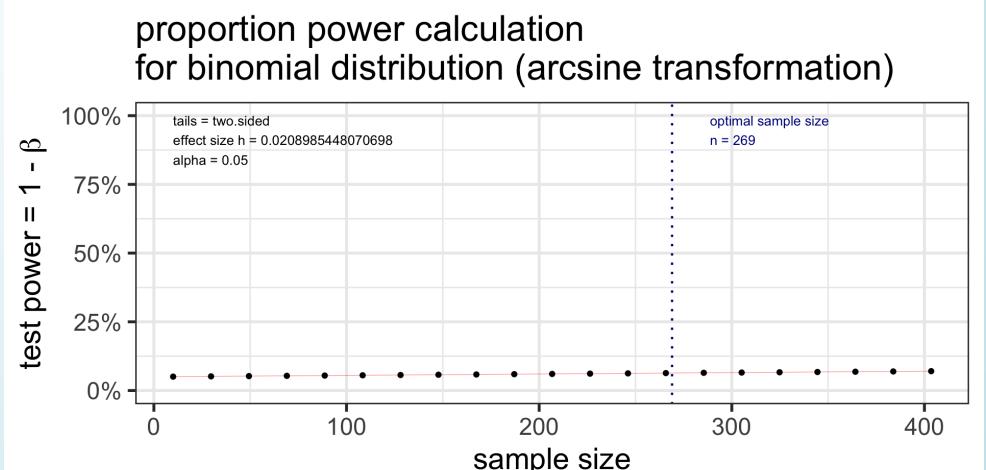
Specify all parameters except for the power:

```
1 library(pwr)
2
3 p.power <- pwr.p.test(
4   h = ES.h(p1 = 0.36, p2 = 0.35),
5   sig.level = 0.05,
6   # power = 0.80,
7   n = 269,
8   alternative = "two.sided")
9 p.power
```

```
proportion power calculation for binomial distribution (arcsine transformation)
```

```
h = 0.02089854
n = 269
sig.level = 0.05
power = 0.06356445
alternative = two.sided
```

```
1 plot(p.power)
```



# pwr: sample size for two proportions test

- Two proportions (same sample sizes)

```
pwr.2p.test(h = NULL, n = NULL, sig.level = 0.05, power = NULL,  
alternative = c("two.sided","less","greater"))
```

- $h$  is the effect size: `h = ES.h(p1, p2); p1` and `p2` are the two proportions being tested

Specify all parameters *except for* the sample size:

```
1 p2.n <- pwr.2p.test(  
2   h = ES.h(p1 = 0.36, p2 = 0.35),  
3   sig.level = 0.05,  
4   power = 0.80,  
5   alternative = "two.sided")  
6 p2.n
```

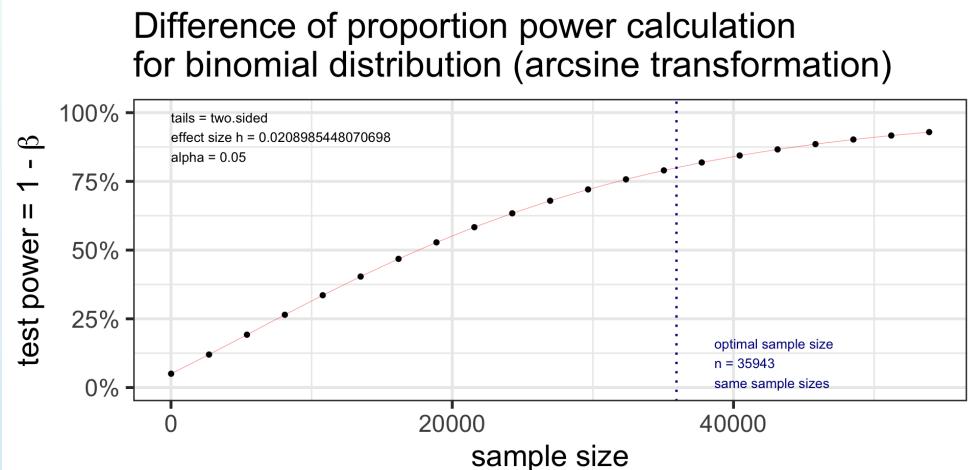
Difference of proportion power calculation  
for binomial distribution (arcsine transformation)

```
h = 0.02089854  
n = 35942.19  
sig.level = 0.05  
power = 0.8  
alternative = two.sided
```

NOTE: same sample sizes

Note:  $n$  in output is the **number per sample!**

```
1 plot(p2.n)
```



# pwr: power for two proportions test

- Two proportions (different sample sizes)

```
pwr.2p2n.test(h = NULL, n1 = NULL, n2 = NULL, sig.level = 0.05, power =  
NULL, alternative = c("two.sided", "less", "greater"))
```

- $h$  is the effect size: `h = ES.h(p1, p2); p1` and `p2` are the two proportions being tested

Specify all parameters *except for* the power:

```
1 p2.n2 <- pwr.2p2n.test(  
2   h = ES.h(p1 = 0.36, p2 = 0.35),  
3   n1 = 214,  
4   n2 = 269,  
5   sig.level = 0.05,  
6   # power = 0.80,  
7   alternative = "two.sided")  
8 p2.n2
```

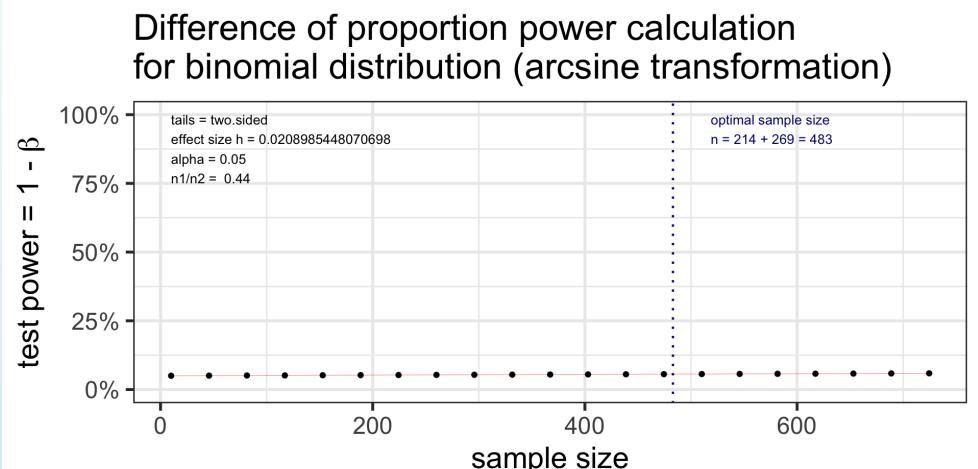
difference of proportion power calculation  
for binomial distribution (arcsine transformation)

```
h = 0.02089854  
n1 = 214  
n2 = 269  
sig.level = 0.05  
power = 0.05598413  
alternative = two.sided
```

NOTE: different sample sizes

Note:  $n$  in output is the **number per sample!**

```
1 plot(p2.n2)
```



# Where are we?

CI's and hypothesis tests for different scenarios:

$$\text{point estimate} \pm z^*(\text{or } t^*) \cdot SE, \text{ test stat} = \frac{\text{point estimate} - \text{null value}}{SE}$$

| Day | Book | Population parameter    | Symbol              | Point estimate             | Symbol                  | SE   |
|-----|------|-------------------------|---------------------|----------------------------|-------------------------|--|
| 10  | 5.1  | Pop mean                | $\mu$               | Sample mean                | $\bar{x}$               | $\frac{s}{\sqrt{n}}$   |
| 10  | 5.2  | Pop mean of paired diff | $\mu_d$ or $\delta$ | Sample mean of paired diff | $\bar{x}_d$             | $\frac{s_d}{\sqrt{n}}$   |
| 11  | 5.3  | Diff in pop means       | $\mu_1 - \mu_2$     | Diff in sample means       | $\bar{x}_1 - \bar{x}_2$ | $\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$ or pooled               |
| 12  | 8.1  | Pop proportion          | $p$                 | Sample prop                | $\hat{p}$               | $\sqrt{\frac{p(1-p)}{n}}$  |
| 12  | 8.2  | Diff in pop proportions | $p_1 - p_2$         | Diff in sample proportions | $\hat{p}_1 - \hat{p}_2$ | $\sqrt{\frac{p_1 \cdot (1-p_1)}{n_1} + \frac{p_2 \cdot (1-p_2)}{n_2}}$ |