# Lecture 9: $K \ge 2$ Classes and k-Nearest Neighbours

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STAT 447B: Methods for Statistical Learning

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# Today's Learning Goals

1 Review of Linear Discriminant Analysis

2  $K \ge 2$  Classes

3 k-Nearest Neighbours



### Normal Model for $\mathbf{X}|Y=c$

- Model the conditional distribution of X | Y instead of Y | X
- For example, we can assume that

$$\mathbf{X} \mid Y = c_{j} \sim \mathsf{MN}\left(\mu_{j}, \mathbf{\Sigma}\right)$$

(MN = multivariate normal, with dimension the number of variables in <math>X)

- The classes differ in their X mean vectors
- The class distributions are estimated by

$$\hat{\mathit{f}}_{j}\left(\mathbf{X}\right) \, \sim \, \mathsf{MN}\left(\hat{\mu}_{j},\widehat{\mathbf{\Sigma}}\right)$$

using the sample mean of each group and the pooled sample covariance matrix

• We can then find, for a given  $\mathbf{x}$ , the class  $\mathbf{j}$  that has the largest  $\hat{f}_j(\mathbf{x}) p_j$ 

# Fisher's Linear Discriminant Analysis for **NORMAL** Populations

Writing  $f_1$  for MN  $(\mu_1, \Sigma)$  and  $f_2$  for MN  $(\mu_2, \Sigma)$  then

$$f_1(\mathbf{x}) p_1 > f_2(\mathbf{x}) p_2 \quad \Leftrightarrow \quad \ln\left(\frac{f_1(\mathbf{x}) p_1}{f_2(\mathbf{x}) p_2}\right) > 0 \quad \Leftrightarrow \quad \mathbf{a}^T \mathbf{x} + \mathbf{b} > 0$$

for some vector  $\mathbf{a} \in \mathbb{R}^p$  and number  $\mathbf{b} \in \mathbb{R}$ . In other words, boundaries between classes are **linear**. Furthermore, we can estimate this linear boundary because

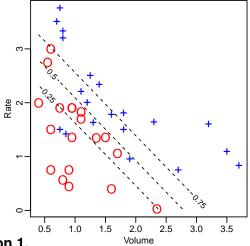
$$\mathbf{a} = \mathbf{\Sigma}^{-1} \; (\mu_1 - \mu_2)$$

and

$$oldsymbol{b} = -rac{1}{2} \, \left( \mu_1 - \mu_2 
ight)^T oldsymbol{\Sigma^{-1}} \left( \mu_1 + \mu_2 
ight) - \ln \left( rac{
ho_2}{
ho_1} 
ight)$$



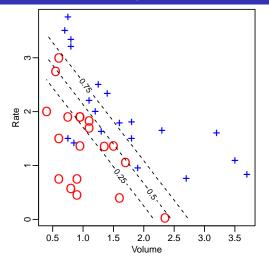
#### Vaso Constriction Data: LDA Fit



Clicker question 1.



#### Vaso Constriction Data: Logistic Fit





### LDA Classification Rule for NORMAL Populations

Note that if  $f_1$  is MN  $(\mu_1, \Sigma)$  and  $f_2$  is MN  $(\mu_2, \Sigma)$  then

$$f_1(\mathbf{x}) p_1 > f_2(\mathbf{x}) p_2 \Leftrightarrow \mathbf{a}^T \mathbf{x} + \mathbf{b} > 0$$

- Boundaries between classes are linear.
- Furthermore

$$\mathbf{a} = \mathbf{\Sigma}^{-1} \; (\mu_1 - \mu_2)$$

and

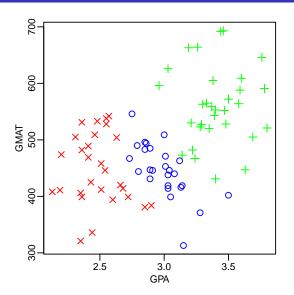
$$b = -rac{1}{2} (\mu_1 - \mu_2)^T \sum_{1}^{T} (\mu_1 + \mu_2) - \ln \left( rac{
ho_2}{
ho_1} 
ight)$$

 There is nothing special about having two classes. We can do the same with more than 2 classes

#### **Admissions Data**

- Admissions data from a Graduate School of Business
- From Applied Mulivariate Statistical Analysis by Richard A. Johnson and Dean W. Wichern, Prentice Hall, 2002
- The data are in the file T11-6.DAT
- 3 classes
  - y = 1 "admit student" (plotted as +)
  - y = 2 "do not admit student" (plotted as  $\times$ )
  - y = 3 "borderline" (plotted as ∘)
- 2 explanatory variables: GPA and GMAT, i.e.,  $\mathbf{x} = (\text{GPA}, \text{GMAT})$

#### **Admissions Data**





#### Admissions Data: LDA

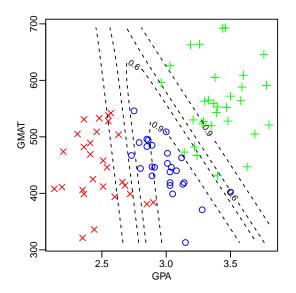
- 3 classes, so 3 multivariate normal distributions  $f_j$  for  $\mathbf{X}$ , conditional on the class
- As before we have prior class probabilities p<sub>i</sub>
- The 3 class codes are  $c_1 = 1, c_2 = 2, c_3 = 3$
- · Math is essentially the same as for 2 classes
  - $Pr(Y = c_i | \mathbf{X} = \mathbf{x}) \propto f_i(\mathbf{x}) p_i$
  - LDA therefore gives posterior probabilities

$$\Pr(Y = c_j \mid \mathbf{X} = \mathbf{x}) = \frac{f_j(\mathbf{x}) p_j}{\sum_{m=1}^{3} f_m(\mathbf{x}) p_m}$$

•  $\hat{y}(\mathbf{x})$  is the  $c_i$  with the largest estimated  $\Pr(Y = c_i | \mathbf{X} = \mathbf{x})$ 



### Admissions Data: LDA Maximum Probability





#### LDA: Drawbacks

 Assuming normality leads to a certain class of estimators for the boundaries and the conditional class probabilities

$$\hat{\mathbf{a}}^T \mathbf{x} + \hat{\mathbf{b}} > 0$$

where

$$\hat{\mathbf{a}} = \hat{\boldsymbol{\Sigma}}^{-1} \; (\bar{\boldsymbol{x}}_1 - \bar{\boldsymbol{x}}_2)$$

and

$$\hat{\boldsymbol{b}} = -\frac{1}{2} \; (\bar{\boldsymbol{x}}_1 - \bar{\boldsymbol{x}}_2)^T \; \hat{\boldsymbol{\Sigma}}^{-1} \; (\bar{\boldsymbol{x}}_1 + \bar{\boldsymbol{x}}_2) - \ln \left(\frac{\rho_2}{\rho_1}\right)$$

Works well for many applications, but ...
normality may not be reasonable, covariance matrix estimation
not robust to outliers, ...

# Logistic Classification

- When we have K > 2 classes we propose
  - For classes *j* = 1, . . . , *K* − 1

$$\eta_j(\mathbf{x}_i) = \beta_{0j} + \boldsymbol{\beta}_j^T \mathbf{x}_i$$

For class K

$$\eta_{\mathbf{K}}(\mathbf{x}_i) = 1$$

Furthermore we take

$$\Pr\left(Y = c_j \mid \mathbf{x}\right) = \frac{\exp\left(\eta_j(\mathbf{x}_i)\right)}{1 + \sum_{m=1}^{K-1} \exp\left(\eta_m(\mathbf{x}_i)\right)} \quad \text{for } j = 1, \dots, \frac{K}{\text{UBC}}$$



# Logistic Classification

 Under this model, again the boundary between any two classes is always linear

$$\ln \left( \frac{\Pr\left( \mathbf{Y} = \mathbf{c}_{j} \, | \, \mathbf{x} \right)}{\Pr\left( \mathbf{Y} = \mathbf{c}_{\ell} \, | \, \mathbf{x} \right)} \right) = \left( \beta_{0j} - \beta_{0\ell} \right) + \left( \boldsymbol{\beta}_{j} - \boldsymbol{\beta}_{\ell} \right)^{T} \mathbf{x}$$

Hence

$$\Pr(Y = c_j | \mathbf{x}) = \Pr(Y = c_\ell | \mathbf{x}) \Leftrightarrow (\beta_{0j} - \beta_{0\ell}) + (\beta_j - \beta_\ell)^T \mathbf{x} = 0$$

• How do we estimate the parameters  $(\beta_{0j}, \beta_j)$  for j = 1, ..., K-1?



# Logistic Classification

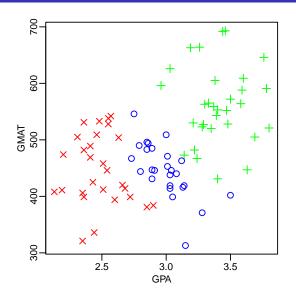
Maximum likelihood with a multinomial model

$$L(y_1, \dots, y_n, \mathbf{x}_1, \dots, \mathbf{x}_n, \beta_{0,1}, \dots, \beta_{0,K-1}, \boldsymbol{\beta}_1, \dots, \boldsymbol{\beta}_{K-1})$$

$$= \prod_{i=1}^n \Pr(Y_i = y_i \mid \mathbf{x}_i; \beta_{0,1}, \dots, \beta_{0,K-1}, \boldsymbol{\beta}_1, \dots, \boldsymbol{\beta}_{K-1})$$



#### **Admissions Data**





# Admissions Data: Logistic Regression in R

```
> # Logistic regression using multinom in nnet
> library(nnet)
> admiss.logistic <- multinom(y ~ GPA + GMAT, data = adr</pre>
  maxit = 10000)
# weights: 12 (6 variable)
initial value 93.382045
iter 10 value 15.844299
iter 20 value 7.434088
iter 30 value 7.245347
iter 620 value 5.395291
```



converged

iter 630 value 5.393841 final value 5.392175

# Admissions Data: Logistic Regression in R

```
> print(summary(admiss.logistic))
Call:
multinom(formula = y ~ GPA + GMAT, data = admiss, maxit = 10000)
Coefficients:
  (Intercept) GPA GMAT
 485.9823 -117.37076 -0.3227173
    167.3553 -31.06165 -0.1458875
Std. Errors:
  (Intercept) GPA GMAT
2 0.7093202 2.648911 0.0182661
3 0.2417068 1.733160 0.0123214
```

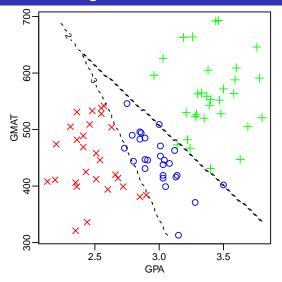


#### Clicker question 2.

AIC: 22.78435

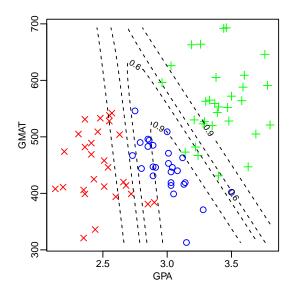
Residual Deviance: 10.78435

#### Admissions Data: Logistic Classification Boundaries





### Admissions Data: LDA Maximum Probability





### Logistic Classification Versus LDA

- Both use MLE estimates for the parameters
- Both estimate  $Pr(Y = c_i | \mathbf{x})$
- However, the results might be different
- Which class decision boundaries do you prefer for the Admissions Data? Clicker question 3.



# Nearest Neighbours

• If we knew the class probabilities  $Pr(Y = y | \mathbf{x})$  the optimal classification rule is

$$\hat{y}(\mathbf{x}) = \arg \max_{h} \Pr(Y = h | \mathbf{x})$$

- So we need to estimate  $\Pr(Y = y \mid \mathbf{x})$ . That's what we have been doing. Another way is to estimate it locally for each  $\mathbf{x}$
- Nearest neighbours is a natural way to do so. For each x let

$$\widehat{Pr}(Y = y | \mathbf{x}) = \text{proportion of points of class } y \text{"near" } \mathbf{x}$$



# k-Nearest Neighbours

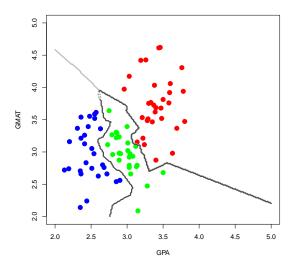
#### Algorithm:

- 1 Choose a fixed predetermined number of neighbours, k (usually k = 1, 3, 5, ...)
- For each x
  - 1 Find the *k* nearest neighbours of **x** in the training data (nearest defined by Euclidean distance, say)
  - $2 n_i$  is the number of neighbours of class i
  - 3  $\hat{y}(\mathbf{x}) = \text{the class } j \text{ with the largest } n_j$

No parametric model for  $Pr(Y | \mathbf{x})!$  Very flexible!



#### Admissions Data: 1-NN





#### Admissions Data: 1-NN

